Multiplicative Shot-Noise: A New Route to Stability of Plastic Networks

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Fluctuations of synaptic weights, among many other physical, biological, and ecological quantities, are driven by coincident events of two "parent" processes. We propose a multiplicative shot-noise model that can capture the behaviors of a broad range of such natural phenomena, and analytically derive an approximation that accurately predicts its statistics. We apply our results to study the effects of a multiplicative synaptic plasticity rule that was recently extracted from measurements in physiological conditions. Using mean-field theory analysis and network simulations, we investigate how this rule shapes the connectivity and dynamics of recurrent spiking neural networks. The multiplicative plasticity rule is shown to support efficient learning of input stimuli, and it gives a stable, unimodal synaptic-weight distribution with a large fraction of strong synapses. The strong synapses remain stable over long times but do not "run away." Our results suggest that the multiplicative shot-noise offers a new route to understand the tradeoff between flexibility and stability in neural circuits and other dynamic networks.

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Introduction.—Many natural processes are triggered by coincidences of two "parent" events. Examples include firefly flash synchronization [1], effects of simultaneous environmental stressors [2], applications of two-photon microscopy [3], and stimulus-reward associations in reinforcement learning [4,5]. In neuroscience, coactivation of pre- and postsynaptic neurons plays a crucial role in inducing synaptic plasticity [6,7], a primary mechanism underlying learning and memory.

The parent processes are often described by event-based models [8], among which the Poisson process is an appealing starting point owing to its memoryless property. Experimental studies show that the aforementioned coincidence-based phenomena often cannot be accurately described as sums over the parent shot-noise (Poisson) processes [2,7]. Specifically, induction of long-term plasticity was shown to depend strongly on the calcium flux into the postsynaptic neuron [9,10]. This flux, in turn, depends on coincident spiking activity of pre- and postsynaptic neurons, and is well described by the product of two shot-noise processes [7,11,12]. In contrast, most network-level studies of spike-timing-dependent plasticity (STDP) typically assume that the synaptic strength change is the sum over contributions of spike pairs, ignoring cooperative effects between spikes [13–17]. These models often cannot reproduce realistic spiking activity observed in vivo [18].

Motivated by the converging theoretical and experimental evidence, we propose a stochastic process whose fluctuations are triggered by multiplicative interactions between two parent shot-noise processes $[c_{\alpha}(t), \alpha = 1, 2]$. The rates of parent events are λ_{α} , and their amplitudes $a_{\alpha,i}$ are exponentially distributed with mean A_{α} . Events are referred to as "spikes," adopting the neuroscience terminology, but they may correspond to events in other domains. Spike times are denoted $\{t_{\alpha,i}\}$ and may be temporally correlated. We then have

$$\frac{\mathrm{d}c_{\alpha}(t)}{\mathrm{d}t} = -\frac{c_{\alpha}(t)}{\tau_{\alpha}} + \sum_{i} a_{\alpha,i} \delta(t - t_{\alpha,i}). \tag{1}$$

The decay timescales τ_{α} define a window during which coincidences can occur.

Our primary interest is a *multiplicative shot-noise* process [denoted C(t) and henceforth referred to as the *coincidence detector*], whose transient deviations from baseline are driven by the product $c_1 \times c_2$, with decay timescale τ_C [Fig. 1(a)],

$$\frac{\mathrm{d}\mathcal{C}(t)}{\mathrm{d}t} = -\frac{\mathcal{C}(t)}{\tau_{\mathcal{C}}} + \eta c_1(t)c_2(t). \tag{2}$$

The stochastic calculus of Poisson processes makes it difficult to analyze their products [19]. In contrast, existing nonlinear shot-noise models [20–23] are equivalent to transformations of a single Poisson process and are not suitable for studying statistics of coincidences.

We analyze the statistics of the coincidence detector and apply these results to gain insights to a longstanding problem in neuroscience: the stability of recurrent neuronal



FIG. 1. Analytical approximation of multiplicative shot-noise statistics. (a) An illustration of the process C, driven by the product $c_1 \times c_2$. (b) The cumulative distribution function (CDF) of C calculated through the method of moments (MM) [Eq. (4)] matches simulations well at high firing rates. (c) Relationship between the log expectation, mean, and variance of C [Eq. (5)]. Data collapse on a line for a range of firing rates and spike-time correlations (ρ). (d) At low firing rates, the CDF given by MM matches simulations well for $C \gtrsim \eta \tau_C A_1 A_2$, but poorly for $C \rightarrow 0$. The heuristic MM matches the simulation over the entire range of C. (e) Performance of heuristic MM for uncorrelated spike trains. The errors, measured by KS distance, are ≤ 0.2 for all firing rates. See Ref. [24] for definition of ρ and KS distances for other parameters.

networks subject to STDP [41]. Here, c_1 , c_2 are calcium transients induced by pre- and postsynaptic spikes, and C is the total calcium flux, which triggers plasticity [7]. Based on spiking network simulations and theoretical analysis of a reduced model, we show that when *individual synapses* in a recurrent network are subject to a nonlinear calciumbased plasticity rule, the empirical *macroscopic* network properties are reproduced (e.g., stable activity patterns, unimodal heavy-tailed synaptic-weight distributions [42]). Further, our results suggest that STDP in itself can support representations that remain stable over a timescale of hours, marking an important step toward understanding the prolonged retention of spatial memories in the face of plasticity and noise [43].

Statistics of multiplicative shot noise.—For simplicity, we assume that the decay timescales of the parent processes are identical, $\tau_1 = \tau_2 \equiv \tau$. This is consistent with calcium-induced plasticity [7,44] and cases where the parent processes are generated by similar agents (e.g., firefly flashes [1]).

Solving Eq. (1) gives filtered spike trains [Fig. 1(a)]. Using these solutions, we evaluate C(t) [Eq. (2)] at steady state. When the parent processes are uncorrelated,

$$\mathcal{C} \stackrel{d}{=} \eta \tau_{\mathcal{C}} \sum_{i,j} a_{1,i} a_{2,j} \times e^{-\frac{|t_{1,i} - t_{2,j}|}{\tau}} \times \frac{\mathcal{R}[\min(t_{1,i}, t_{2,j})]}{2\delta - 1}.$$
 (3)

Here, $\delta = \tau_C / \tau$ is the ratio between timescales [Eqs. (1) and (2)], and $\stackrel{d}{=}$ means "equal in distribution." The natural interpretations of the three factors in Eq. (3) are the stochastic amplitudes of synaptic transmission, the temporal window for coincident spikes, and $\mathcal{R}(x) = e^{-x/\tau_C} - e^{-2x/\tau}$ describes the firing-rate-dependent accumulation of multiple coincidences. Notably, \mathcal{R} represents a departure from summation over spike pairs [13–16]. Similar expressions for temporally correlated spike trains appear in Ref. [24].

We begin by formulating an analytical approximation of P_C , the distribution of C. At steady state, the shot-noise process [e.g., c_{α} in Eq. (1)] follows a gamma distribution [45,46] with shape and scale parameters $\lambda_{\alpha}\tau$, A_{α} . We think of the coincidences of the parent processes as events which drive fluctuations of C. Therefore, we assume that P_C can be approximated by a gamma distribution whose shape (k) and scale (σ) parameters measure the effective rate and amplitude of the coincident spikes.

We use the method of moments (MM), i.e., matching the first and second moments of $P_{\mathcal{C}}$ to a gamma distribution, to analytically estimate k, σ . We find, for uncorrelated spike trains,

$$k = \frac{\langle \mathcal{C} \rangle^2}{\operatorname{Var}(\mathcal{C})} = \frac{\lambda_1 \lambda_2 \tau^2}{\frac{\lambda_1 + \lambda_2}{\delta + 1} \tau + \frac{1}{2\delta + 1}},$$

$$\sigma = \frac{\operatorname{Var}(\mathcal{C})}{\langle \mathcal{C} \rangle} = \eta \tau_{\mathcal{C}} A_1 A_2 \left(\frac{\lambda_1 + \lambda_2}{\delta + 1} \tau + \frac{1}{2\delta + 1} \right). \quad (4)$$

See Ref. [24] for expressions including correlations.

We show that MM [Eq. (4)] is highly accurate in the high-firing-rate regime ($\langle C \rangle > \eta \tau_C A_1 A_2$) by comparing it to numerical simulations [Fig. 1(e), above the blue line]. However, for low firing rates (i.e., $\lambda_1 \tau, \lambda_2 \tau < 1$, which implies $\langle C \rangle < \eta \tau_C A_1 A_2$) and particularly at $C \rightarrow 0$, the MM is inaccurate [Fig. 1(d)]. The reason is that for low firing rates (and k < 1), the gamma probability density diverges at 0. Such a singularity cannot be captured by the mean and variance of P_C . Notably, either λ_1 or λ_2 being high suffices for the MM to be accurate, because the high-rate process provides a "background" on top of which the low-rate process can trigger coincidences.

In some applications of our theory, it may be important to accurately estimate $P_{\mathcal{C}}$ at $\mathcal{C} \to 0$, in the low-firing-rate regime. We obtain such an estimate by first noticing that in this regime, the maximum-likelihood estimate of the gamma distribution parameters (k, σ) yields a good approximation of $P_{\mathcal{C}}$. This estimate relies on the logexpectation variable $s = \ln \langle \mathcal{C} \rangle - \langle \ln \mathcal{C} \rangle$, which is indeed sensitive to the singularity at $\mathcal{C} \to 0$. Next, we show that the log expectation, the mean, and the variance of \mathcal{C} obey a simple relationship, irrespective of the firing rates and correlation [Fig. 1(c)],

$$\ln\left(\frac{s}{\operatorname{Var}(\mathcal{C})}\right) = -\ln 2 - 2\ln\langle\mathcal{C}\rangle + b(\delta)[-\ln\langle\mathcal{C}\rangle]_{+}^{2}, \quad (5)$$

where $[x]_+ = \max(x, 0)$, and $b(\delta)$ is fit directly to simulations. We find corrected shape and scale parameters $\tilde{k}, \tilde{\sigma}$ by computing *s* via Eq. (5) and using standard maximum-likelihood formulas [24]. We call this the "heuristic MM," and we use it when coincidences are rare (k < 1) and both firing rates are small $[\max(\lambda_1 \tau, \lambda_2 \tau) < 1]$. This approximation is accurate for all values of C, in the entire parameter space we explored [Fig. 1(e)].

Network stabilization by multiplicative synaptic plasticity.—We now leverage our results to study effects of a multiplicative plasticity on network structure. In this context, C represents the calcium influx into a neuron, triggered by coincident pre- and postsynaptic spikes. Large influx induces long-term potentiation (LTP, when $C > \theta_p$), while intermediate influx induces long-term depression (LTD, $\theta_p > C > \theta_d$) [7,9,10]. Given the potentiation and depression rates $\gamma_{p,d}$ and thresholds $\theta_{p,d}$, the synaptic-weight dynamics are

$$\tau_w \frac{\mathrm{d}w}{\mathrm{d}t} = \gamma_p \Theta(\mathcal{C}(t) - \theta_p) - \gamma_d \Theta(\mathcal{C}(t) - \theta_d) \equiv \Gamma_p - \Gamma_d. \quad (6)$$

Note that *w* has its lower bound at 0, and typically $\tau_w \gg \tau_C$. Based on our analytical approximation of P_C , we compute the total potentiation/depression rates $\Gamma_{p,d} \simeq [\gamma_{p,d}/\Gamma(\tilde{k})] \int_{\theta_{p,d}}^{\infty} C^{\tilde{k}-1} e^{-C/\tilde{\sigma}} dC$, which depend on

the spike-train properties and the plasticity rule [Eqs. (4), (5)].

Neuronal activity *in vivo* undergoes substantial firingrate fluctuations, generated by external input variability or intrinsic dynamics [47]. Below, we formulate a mean-field approximation, reducing the joint dynamics of neurons and synapses and accounting for network structure, to the effective dynamics of a pair of pre- and postsynaptic neurons and the synapse connecting them. The statistical properties of the reduced system recapitulate the network behavior.

In the reduced model, we assume that the neurons' firing rates (λ_1, λ_2) are sampled from P_{λ} and have correlation time $T_{\lambda} \sim 0.1-1$ s. P_{λ} may depend on the synaptic strength and will be determined self-consistently, accounting for network interactions. During an interval T_{2} , the rates (λ_1, λ_2) are approximately constant, so the weight change is $\Delta w = T_{\lambda}/\tau_w \times (\Gamma_p - \Gamma_d)$. Its distribution $P_{\text{step}}(\Delta w)$ is calculated from P_{λ} through Eq. (6) [Figs. 2(a) and 2(b)]. Thus, the synaptic dynamics are reduced to a 1D random walk on $w \ge 0$ [Fig. 2(c)] with weight-dependent step-size distribution P_{step} . We identify its steady-state distribution P_w with the synaptic-weight distribution of the network. Importantly, for $T_{\lambda} \sim 1$ s, Δw is not infinitesimal, so the small step-size approximation [13,14,48] is invalid. Next, we use the mean-field approach to study representative network architectures.

Weak synapses.—In this limit, the pre- and postsynaptic firing rates (λ_1, λ_2) are sampled *independently* of the



FIG. 2. Reduced model of synaptic-weight dynamics. (a),(b) Synaptic-weight change [(dw/dt), Eq. (6)] arising from independent (a) and equal (b) pre- and postsynaptic firing rates. In (a), the temporal correlation $\rho = 0$. The green line indicates $\lambda_1 + \lambda_2 = 50$ Hz. (c) Example synaptic-weight trajectory, illustrating the memory time *T* and the maximum weight w_{max} . (d) Synaptic-weight distribution, including analytical result for the tail behavior in the weak-synapse scenario. (Inset) Memory-time distribution of a synapse. (e),(f) In the feed-forward (strong-synapse) scenario, the synaptic-weight (e) and memory-time (f) distributions have heavier tails. Also shown in (e) is the potentiation-depression ratio $\alpha(w)$. (g) Joint distribution of (w_{max}, T) . Strong synapses are preferentially protected from forgetting. (h) Average memory time increases linearly with $\langle L_p \rangle / \langle L_d \rangle$ in the independent case, and nonlinearly in the feedforward case. See Ref. [24] for parameter values.

synaptic weight w. P_{step} is further assumed to be discrete, such that a synapse can be unchanged/potentiated/ depressed by fixed amounts $\Delta w = 0, L_p, -L_d$ with probabilities $\alpha_0, \alpha_p, \alpha_d$. Equivalently, $P_{\text{step}}(\Delta w) = \alpha_0 \delta(\Delta w) +$ $\alpha_d \delta(\Delta w + L_d) + \alpha_p \delta(\Delta w - L_p)$. Such a scenario would be expected in a network switching between "high" and "low" states. For illustration, we assume that L_p/L_d is an integer (see Ref. [24] for more general analysis). We define the potentiation-depression ratio $\alpha \equiv \alpha_p L_p / \alpha_d L_d$. When depression dominates ($\alpha < 1$), using results for random walks [49,50], P_w is unimodal, and its tail follows $P_w \sim$ $h(\beta)e^{-(\beta w/L_d)}, w \gg L_p$. The factors $\beta, h(\beta)$ were determined by analyzing the moment-generating function of w [24]. We find that the tail becomes heavier as overall potentiation and depression are more closely balanced ($\alpha_d L_d \gtrsim \alpha_p L_p$), or as L_p/L_d is larger with fixed α . Generally, when weight changes (Δw) are non-negligible relative to the mean, the distribution P_w is unimodal, and its tail behavior is sensitive to the high-order statistics of the step-size distribution P_{step} , in contrast to the case of infinitesimal Δw [14].

To study the process of forgetting in the reduced model, we envision a potentiated synapse with initial weight w_0 representing a certain memory. The memory time $T(w_0)$ is that synapse's first passage time to 0. Analysis of the random walk statistics [24] gives the average and tail behavior of $T(w_0)$,

$$\langle T(w_0) \rangle = \frac{w_0}{\alpha_d L_d - \alpha_p L_p} \quad \text{and} \quad P_T(t) \sim w_0 t^{-\frac{3}{2}} e^{-\frac{t}{\kappa}}.$$
 (7)

See Ref. [24] for expressions of κ . Similarly to P_w , the average memory time becomes longer and the tail becomes heavier as L_p/L_d increases, with fixed α [Fig. 2(e)].

Strong synapses.—When synapses are strong [51,52], P_{λ} becomes weight dependent. The postsynaptic neuron receives *feed-forward* weighted presynaptic input (firing rate λ_1 , weight w) and background input from the rest of the network. Both inputs switch between high and low firing rates. Here, using the heuristic MM to compute $P_{\text{step}}(\Delta w)$ requires knowing how spike-time correlations depend on w and the background input. This relationship is determined by matching the postsynaptic neuron with a leaky-integrate-and-fire neuron driven by presynaptic shot noise and background Gaussian noise [24]. We then numerically evaluate P_w , P_T at steady state, showing a substantially heavier tail when compared to the independent case [while fixing the overall potentiation-depression ratio $\alpha_0 = \langle \alpha(w) \rangle_{P_w}$, Figs. 2(e) and 2(f)].

Inspection of the joint distribution of the synaptic-weight running maximum and the memory time (w_{max}, T) suggests that strong synapses are specifically resistant to forgetting [Fig. 2(g)]. Moreover, the average memory time increases nonlinearly with the LTP-LTD amplitude ratio [Fig. 2(h)], compared to an approximately linear increase in the independent case [Eq. (7), again matching $\langle \alpha(w) \rangle_{P_w}$]. Similar results were observed in a reduced clustered recurrent network model [24].

Taken together, in the regime where a small number of inputs is sufficient to trigger a postsynaptic response, the multiplicative plasticity rule supports a unimodal synapticweight distribution in which strong synapses are preferentially protected from turnover.

Spiking network simulations.—We tested our results by simulating a network of leaky-integrate-and-fire neurons. The network consists of two excitatory (E) clusters which mutually inhibit each other indirectly via inhibitory clusters [Fig. 3(a), see Ref. [24] for networks with > 2 clusters]. Initially, intracluster $E \rightarrow E$ connections are strong, while intercluster connections are weak. Crucially, the probability that an $E \rightarrow E$ connection exists is independent of the cluster assignment. The initial structure may represent two mutually exclusive memories stored in the network that spontaneously switch on a timescale of ~ 0.2 s [Fig. 3(a)]. Structured inhibition is consistent with experiments showing inhibitory stimulus-specific ensembles, and it may arise from inhibitory plasticity [53–55]. We investigate memory retention when $E \rightarrow E$ synapses undergo multiplicative plasticity by examining the steady-state statistics of C, network structure, and dynamics.

In this network, potentiation is more likely in intracluster relative to intercluster synapses, so the tail of P_w for intracluster synapses is heavier [Fig. 3(b)]. Yet, notably, P_w is unimodal with only minimal saturation to the upper bound. To examine the stability of network structure, we plot the steady-state $E \rightarrow E$ weight matrix spectrum [Fig. 3(c)]. The spectral distribution's bulk follows the circular law for a network with independent, random weights, sampled from cluster-specific distributions [24,56]. Additionally, there are two outlying eigenvalues. The fact that the larger eigenvalue (corresponding to the "DC" eigenvector) does not saturate to its maximum possible value, together with the stability of the switching dynamics, suggests that there is no runaway potentiation of either cluster. The smaller eigenvalue corresponds to an eigenvector that follows from the clustered connectivity. Angles between the plastic network's eigenvectors and those of a network with perfect cluster structure are stable and much smaller than angles computed for a network with shuffled connections, indicating that network structure is preserved despite ongoing plasticity. As predicted by the mean-field analysis, strong synapses are protected from rapid turnover. The dynamics of the plastic network also retains the cluster properties, exhibiting larger intracluster spike-time correlations and larger avalanches than a shuffled network. Intriguingly, avalanche statistics are closely related to the synaptic-weight distribution [57].

To understand the implications of the multiplicative rule beyond stability, we extended the results in Ref. [58] and analytically computed the memory capacity of a Hopfield-like network, defined such that the variance of the synaptic-weight distribution is independent of the memory load [24]. Figure 3(h) shows that a heavier tail of the distribution, similarly to P_w in Fig. 2(e), leads to a



FIG. 3. Recurrent spiking neural network simulations with synapses subject to the multiplicative plasticity rule. (a) Initial network structure, and steady-state spiking activity. (b) Calcium influx and synaptic-weight (inset) distributions at steady state, for inter- and intracluster connections. (c) Eigenvalues (top) and eigenvectors (bottom) of the $E \rightarrow E$ steady-state synaptic-weight matrix. The second outlier suggests that the network's cluster structure is preserved at steady state. (d) The first two subspace angles between the plastic network and a network with perfect cluster structure as a function of time. The second angle in the plastic network remains far from $\pi/2$ compared to the shuffled network. (e) Joint distribution of (w_{max}, T) for intracluster $E \rightarrow E$ connections at steady state. Similarly to the reduced model [Fig. 2(c)], strong synapses are preferentially protected. (f),(g) The plastic network exhibits larger avalanches and stronger intracluster spike-timing correlations, compared to the shuffled network. (h) Memory capacity of a Hopfield-like network versus synaptic-weight variability. Synaptic weights are normalized such that the variance is independent of the memory load α . See Ref. [24] for simulation and calculation details.

marked increase in capacity. Furthermore, we demonstrate in Ref. [24] that the multiplicative plasticity rule supports efficient learning of structured connectivity [akin to Fig. 3(a)], reflecting the structure of an external input.

In Ref. [24] we explored the sensitivity of the spiking network stability results to changes of two key parameters, the potentiation/depression thresholds [$\theta_{p,d}$ in Eq. (6)] and the structure of inhibition. We additionally showed that networks with intrinsically bistable synapses [10] are also stable, but they do not exhibit realistic synaptic-weight distributions or activity-dependent protection of strong synapses.

Highlighting the importance of the multiplicative rule's statistics, we found that an additive plasticity rule with $C = c_1 + c_2$ [instead of Eq. (2)] rapidly leads to instability, and it is unable to efficiently learn the structure of an external input [24].

Our analysis offers insights to the two-timescale problem, where synaptic interactions determine network dynamics on short timescales and undergo neural-activity-dependent modifications on longer timescales. Importantly, we analyze the network in a regime where the strict separation of timescales does not hold. Previous studies utilizing plasticity rules where modifications depend (possibly nonlinearly) on sums over pre- and postsynaptic activity typically resulted in unrealistic synaptic-weight or firing-rate distributions, or required fast homeostatic mechanisms for stability [13–16,59–62]. The multiplicative structure of the plasticity rule analyzed here effectively eliminates modifications due to "spurious" activity.

while specific patterns of activity are responsible for potentiation and learning. The general structure of the multiplicative process introduced here suggests that our results could be applied to understand nonlinear and adaptive interacting systems in a broad range of scientific fields.

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