

**Crossover from Linear to Square-Root Market Impact**Frédéric Bucci,<sup>1,2</sup> Michael Benzaquen,<sup>3,2,4</sup> Fabrizio Lillo,<sup>5</sup> and Jean-Philippe Bouchaud<sup>4,6,2</sup><sup>1</sup>*Scuola Normale Superiore di Pisa, Piazza dei Cavalieri 7, 56126 Pisa, Italy*<sup>2</sup>*Chair of Econophysics and Complex Systems, Ecole polytechnique, 91128 Palaiseau Cedex, France*<sup>3</sup>*Ladhyx, UMR CNRS 7646 & Department of Economics, Ecole polytechnique, 91128 Palaiseau Cedex, France*<sup>4</sup>*Capital Fund Management, 23-25, Rue de l'Université 75007 Paris, France*<sup>5</sup>*Department of Mathematics, University of Bologna, Piazza di Porta San Donato 5, 40126 Bologna, Italy*<sup>6</sup>*CFM-Imperial Institute of Quantitative Finance, Department of Mathematics, Imperial College, 180 Queen's Gate, London SW7 2RH, United Kingdom*

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Using a large database of 8 million institutional trades executed in the U.S. equity market, we establish a clear crossover between a linear market impact regime and a square-root regime as a function of the volume of the order. Our empirical results are remarkably well explained by a recently proposed dynamical theory of liquidity that makes specific predictions about the scaling function describing this crossover. Allowing at least two characteristic timescales for the liquidity (“fast” and “slow”) enables one to reach quantitative agreement with the data.

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Financial markets sputter enormous amounts of data that can now be used to test scientific theories at levels of precision comparable to those achieved in physical sciences (see, e.g., Ref. [1] for a recent example). Among the most remarkable empirical findings in the last decades is the “square-root impact law,” which quantifies how much prices are affected, *on average*, by large buy or sell orders, usually executed as a succession of smaller trades. Such a succession of small trades, all executed in the same direction (either buys or sells) and originating from the same market participant, is called a *metaorder*. A metaorder of total size  $Q$  impacts the price as  $\sim\sqrt{Q}$  and not proportionally to  $Q$  as naively expected and actually predicted by classical economics arguments [2]. The square-root law is surprisingly universal: it is found to be to a large degree independent of details such as the asset class, time period, execution style, and market venues [3–14]. In particular, the advent of electronic markets and high frequency trading has *not* altered the square-root behavior, in spite of radical changes in the microstructure of markets.

The universality of this square-root law, together with its insensitivity to the high frequency dynamics of prices, suggests that its interpretation should lie in some general properties of the low frequency, large scale dynamics of liquidity [15]. In fact, the publicly displayed liquidity at any given time is usually very small—typically on the order of  $10^{-2}$  of the total daily transaction volume in stock markets. Financial markets are the arena of a collective hide-and-seek game between buyers and sellers, resulting in a somewhat paradoxical situation where the total quantity that markets participants intend to trade is very large (0.5% of the total market capitalization changes hands every day

in stock markets) while most of this liquidity remains hidden, or “latent.” These observations have led to the development of a physics inspired, “locally linear order book” (LLOB) model for the coarse-grained dynamics of latent liquidity [7,15–17], which naturally explains why the impact of metaorders grows like the square root of its size in a certain regime of parameters [17]. In a nutshell, the LLOB model predicts a linearly growing equilibrium liquidity profile, which implies a square-root impact law even in the absence of any reaction of the liquidity providers to the incoming metaorder. But this LLOB model also suggests that for a given execution time  $T$ , the very small  $Q$  regime should revert to a linear behavior. The model, in fact, predicts the detailed shape of the crossover between linear and square-root impact. Deviations from a pure square root were observed in Ref. [9], where the authors fitted the data with a logarithmic function  $\ln(a + bQ)$ , which indeed behaves linearly for small arguments [18].

The aim of the present Letter is to test for the first time the detailed theoretical predictions of a crossover from linear to square root impact using the very large ANcerno [21] database of metaorders, executed on the U.S. equity market and issued by a diversified set of institutional investors. We find that the crossover between linear and square-root impact is well described by the theory, albeit the transaction volume at the crossover point is much smaller than the predicted one.

We argue that this can be accounted for by the coexistence of “slow” and “fast” agents in financial markets. Fast agents contribute to the total transaction volume but are unable to offer resistance against the execution of large

metaorders. Therefore, only slow agents are able to dampen market impact and only their contribution is relevant for shaping up the square-root law. We recall how the LLOB model can be augmented to account for multiple agent frequencies [23], and compute the impact crossover function within this extended framework, resulting in a remarkably good fit of the data.

Let us first briefly recall the basic ingredients of the LLOB model, as well as its main predictions. The fundamental quantity of interest is the density  $\varphi(x, t)$  of latent orders around price  $x$  at time  $t$ . Conventionally, one can choose  $\varphi$  to be positive for buy latent orders [corresponding to  $x < p(t)$ , where  $p(t)$  is the current transaction price] and negative for sell latent orders [corresponding to  $x > p(t)$ ]. As argued in Refs. [7,16,17], the coarse-grained dynamics of the latent liquidity close to the current price is well described by the following equation:

$$\partial_t \varphi = D \partial_{xx} \varphi - \nu \varphi + \lambda \text{sgn}(y) + m \delta(y), \quad (1)$$

where  $y := p(t) - x$ , and  $\nu$  describes order cancellation,  $\lambda$  new order deposition, and  $D \partial_{xx}$  limit price reassessments. The final ‘‘source’’ term corresponds to a metaorder of size  $Q$  executed at a constant rate  $m = Q/T$ , corresponding to a flux of orders localized at the transaction price  $x = p(t)$ . In the absence of a metaorder ( $m = 0$ ), Eq. (1) admits a stationary solution in the price reference frame, which is *linear* when  $y$  is small (hence the name LLOB):

$$\varphi_{st}(y) = \mathcal{L}y, \quad (2)$$

where  $\mathcal{L} = \lambda/\sqrt{D\nu}$  is a measure of liquidity. The linear behavior of the latent liquidity close to the transaction price is in fact a generic result, which holds much beyond the simple setting defined by Eq. (1), while being at the origin of the square-root impact law [7,15–17]. The total transaction rate  $J$  is simply given by the flux of orders through the origin, i.e.,  $J := D \partial_y \varphi_{st}|_{y=0} = D\mathcal{L}$ .

In the limit of a slow latent order book (i.e.,  $\nu T \ll 1$ ), the price trajectory  $p_m(t)$  during the execution of the metaorder [obtained as the solution of  $\varphi(p_m, t) = 0$ ] is given by the following self-consistent expression [17]:

$$p_m(t) = p_0(t) + y(t), \quad (3)$$

$$y(t) = \frac{m}{\mathcal{L}} \int_0^t \frac{ds}{\sqrt{4\pi D(t-s)}} \exp\left(-\frac{[y(t) - y(s)]^2}{4D(t-s)}\right), \quad (4)$$

where  $p_0(t)$  is the price trajectory in the absence of the metaorder that starts at  $t = 0$  and ends at  $t = T$ . Price impact is then defined as  $I := y(T)$ , and is found to be given by

$$I(Q) = \sqrt{\frac{DQ}{J}} \mathcal{F}(\eta), \quad \text{with } \eta := \frac{Q}{JT}, \quad (5)$$

where  $\eta$  is the *participation rate* and the scaling function  $\mathcal{F}(\eta) \approx \sqrt{\eta/\pi}$  for  $\eta \ll 1$  and  $\approx \sqrt{2}$  for  $\eta \gg 1$ . Hence,  $I(Q)$  is linear in  $Q$  for small  $Q$  at fixed  $T$ , and crosses over to a square-root for large  $Q$ . Note that in the square-root regime, impact is predicted to be *independent* of the execution time  $T$ . Many other results, such as the decay of impact for  $t > T$ , have been derived and discussed in Refs. [17,23].

We now turn to the ANcerno database to see how well Eq. (5) is supported empirically. Our sample covers for a total of 880 trading days, from January 2007 to June 2010 and we follow the cleaning procedure introduced in Ref. [9] to remove possible spurious effects. The sample is represented by around 8 million metaorders uniformly distributed in time and market capitalization [24]. Each metaorder in the database is characterized by a broker label, a stock symbol, the total number of traded shares  $Q$ , the sign  $\epsilon = \pm 1$  (buy or sell), the start time  $t_s$ , and the end time  $t_e$  of its execution. In line with the definition given above, and following Ref. [9], the participation rate is given by  $\eta = Q/V_T$  where  $V_T = V(t_e) - V(t_s)$  is the total volume traded in the market during the metaorder execution. In order to compare different stocks with very different daily volumes, we shall measure  $Q$  in units of the corresponding daily volume  $V_d$ , and introduce the volume fraction  $\phi := Q/V_d$ , which in the model notation is equal to  $Q/JT_d$ , where  $T_d = 1$  day. We will also measure execution in relative volume time and redefine the execution time  $T$  as  $T := [V(t_e) - V(t_s)]/V_d$ . Finally, we introduce rescaled log prices as  $p(t) := [\log P(t)]/\sigma_d$ , where  $\sigma_d = (P_{\text{high}} - P_{\text{low}})/P_{\text{open}}$  is the daily volatility estimated from the daily high, low and open prices  $P(t)$ .

The average price impact  $I(Q)$  for a given executed volume  $Q$ , as studied in most previous studies, is defined as

$$I(Q) = \mathbb{E}[\epsilon(p_e - p_s)|Q], \quad (6)$$

where  $p_s, p_e$  are, respectively, the midprice at the start and at the end of the metaorder. As shown in Refs. [9,13], the ANcerno data confirm that  $I(Q)$  is close to a square-root in an intermediate regime of volume fraction  $10^{-3} \lesssim \phi \lesssim 10^{-1}$ , but shows an approximate linear behavior for smaller volume fractions  $\phi \lesssim 10^{-3}$ . Note that data for large volume fractions  $\phi \gtrsim 10^{-1}$  are difficult to interpret, as they are prone to strong conditioning effects.

In order to test directly Eq. (5), we estimate the scaling function  $\mathcal{F}(\eta)$  by dividing the data into evenly populated bins of constant participation rate  $\eta$  and compute the conditional expectation of  $\epsilon(p_e - p_s)/\sqrt{\phi}$  for each bin. According to the LLOB model this expectation is equal to  $\sqrt{D/\sigma_d^2} \mathcal{F}(\eta)$ . Here and in the following, error bars are determined as standard errors.

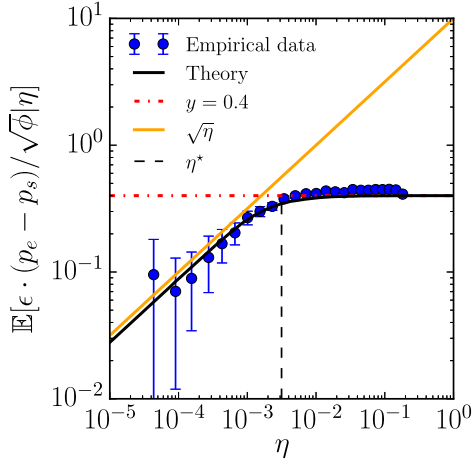


FIG. 1. Empirically determined scaling function  $\mathcal{F}(\eta)$  vs participation rate  $\eta$ . The data (blue points) interpolate between a  $\sqrt{\eta}$  behavior observed at small participation rates and an asymptotically constant regime  $\approx 0.4$  for large  $\eta$ , i.e., for  $\eta \gtrsim \eta^*$  with  $\eta^* \approx 3.15 \times 10^{-3}$ . Black line: prediction of LLOB, with an adjusted crossover  $\eta^* := J_s/J_f$  allowing for the existence of two categories of agents (fast and slow). The data points are obtained by restricting to metaorders with sufficient large order size, i.e.,  $\phi \gtrsim 10^{-5}$ .

The results are shown in Fig. 1 and are, up to a rescaling of both the  $x$  and  $y$  axis, remarkably well accounted for by the LLOB function  $\mathcal{F}(\eta)$ , that describes the crossover between a linear-in- $Q$  regime for small participation rates  $\eta$ , and a  $T$  independent,  $\sqrt{Q}$  regime at large  $\eta$ . Whereas a linear regime for small  $Q$ 's was already reported in Refs. [9,13], the scaling analysis provided here has not been attempted before. The fact that impact in the  $\sqrt{Q}$  regime chiefly depends on  $Q$  but not on  $T$  is compatible with the results of Refs. [9,25], but contradicts theories that assign the  $\sqrt{Q}$  dependence to *duration* of the metaorder, as in Refs. [3,26–28].

However, whereas the crossover between the two regimes should occur around  $\eta^* = 1$  within the original LLOB model, empirical data points towards a much smaller value  $\eta^* \sim 10^{-3}$ . This is actually consistent with the fact that all the empirical evidence for the square-root law reported in the literature concern moderate participation rates (typically in the range  $10^{-3} - 10^{-1}$ , see, e.g., Refs. [4,8,13,14]), but never in a regime where the volume of the metaorder becomes larger than that of the rest of the market, as would be requested within the LLOB specification. Note that in our sample, 70% of the metaorders are such that  $\eta > \eta^*$ .

In order to account for this large discrepancy in the value of  $\eta^*$ , we shall consider the extended LLOB model recently proposed by two of us to include agents with different time horizons, as is clearly the case in financial markets [23]. In the simplest case of a bi-modal distribution of agents (fast and slow), the LLOB formalism can be generalized

to describe two latent order book densities,  $\varphi_s(x, t)$  for the slow liquidity and  $\varphi_f(x, t)$  for the fast liquidity—for example, provided by high frequency traders. The corresponding dynamical equations then read [23]

$$\partial_t \varphi_o = D_o \partial_{xx} \varphi_o - \nu_o \varphi_o + \lambda_o \text{sign}(y) + m_o(t) \delta(y), \quad (7)$$

where  $y = p(t) - x$  and  $o = s, f$ , and where  $m_s(t)$  [ $m_f(t)$ ] is the fraction of the metaorder absorbed by the slow (fast) traders, with  $m_s(t) + m_f(t) = m$ . We allow the activity rate of the two categories of agents to be different through the coefficients  $D_o$ ,  $\nu_o$ , and  $\lambda_o$ . The interesting limit for our purpose is (i)  $J_s \ll J_f$  and  $m \ll J_f$ , where  $J_o = \lambda_o \sqrt{D_o/\nu_o}$ . These inequalities mean that (i) the total transaction rate  $J = J_s + J_f \approx J_f$  is dominated by fast traders and (ii) the flux corresponding to the metaorder is small compared to the total transaction rate of the market, as with most metaorders executed in liquid markets. (ii)  $\nu_s T \ll 1$  and  $\nu_f T \gg 1$ . As shown in Ref. [23] this implies that slow, persistent agents are able to resist the impact of the metaorder, whereas fast agents are playing the role of “transparent” intermediaries, only lubricating the high-frequency activity of markets. This double-frequency model can be solved exactly in some limits [23]. One should distinguish two cases, depending on whether the execution time  $T$  is larger or smaller than a certain  $T^\dagger := \nu_f^{-1} \eta^{*2} D_s/D_f$ , where  $\eta^* := J_s/J_f$ . For  $T > T^\dagger$ , the scaling result Eq. (5) is simply modified as

$$I(Q) = \sqrt{\frac{D_s Q}{J_s}} \mathcal{F}\left(\frac{\eta}{\eta^*}\right). \quad (8)$$

For  $T < T^\dagger$ , this result is further multiplied by  $\sqrt{T/T^\dagger}$ , with a shifted crossover point  $\eta^* \rightarrow \eta^* T^\dagger/T$ .

If we assume that  $T^\dagger$  is small enough for all data points (which needs to be checked *a posteriori*), then the prediction of the double-frequency model, Eq. (8), is precisely the same as the one of the standard LLOB model, up to a rescaling of the  $x$  axis by  $\eta^*$ , and of the  $y$  axis by a ratio  $\sqrt{D_s J/D J_s}$ . Figure 1 shows that the LLOB scaling prediction indeed reproduces the data very well, which allows a direct determination of  $\eta^* \approx J_s/J \approx 3.15 \times 10^{-3}$ . In other words, we find that most of the daily liquidity is provided by fast agents, whereas the resistance to a metaorder relies on a small fraction of slow agents. We have checked that neither the quality of the fit nor the value of  $\eta^*$  are significantly different in the period 2007–2008 and 2009–2010. We have also investigated the dependence of  $\eta^*$  on market capitalization and volatility. We find that low volatility-large cap. stocks are characterized by a larger value of  $\eta^*$  than high volatility-small cap. stocks, suggesting, perhaps counterintuitively, that the low

frequency activity is comparatively more important in low volatility-large cap. stocks.

The plateau value for  $\eta > \eta^*$ , on the other hand, leads to  $\sqrt{D_s J / D J_s} = 0.4 / \sqrt{2}$ , leading to  $\sqrt{D_s / D} \simeq 10^{-2}$ . Since  $\sqrt{D}$  should be close to the price volatility [17], we find that, consistently with its interpretation, the slow liquidity moves much more slowly than the price itself. These estimates in turn lead to  $T^\dagger \approx 45 \nu_f^{-1}$ , or  $\sim 45$  sec for  $\nu_f^{-1} = 1$  second. Since the median execution time of the metaorders in our sample is 35 min, we conclude that most metaorders in our sample are indeed longer than  $T^\dagger$ .

Still, a bi-modal distribution of trading frequencies is certainly an oversimplification. One should consider instead, as in Ref. [23], a continuous distribution of frequencies. Several empirical facts about the dynamics of financial markets (see, e.g., Refs. [15,29–31]) actually suggest that such a distribution is a power law. The numerical solution and the fitting procedure of such a general model is beyond the scope of the present Letter, but the simplified analysis of Ref. [23] suggests that the LLOB scaling function should be approximately valid, with a crossover value  $\eta^*$  that decreases as a power law of  $T$ . Intuitively, the critical participation rate  $\eta^*$  should be larger for small metaorder duration  $T$ , since there are less traders that can be considered fast on such short timescales and more traders that are “slower” than  $T$ . This intuition is indeed confirmed by Fig. 2, where we show the rescaled data as a function of  $\eta$ , for metaorders longer and shorter

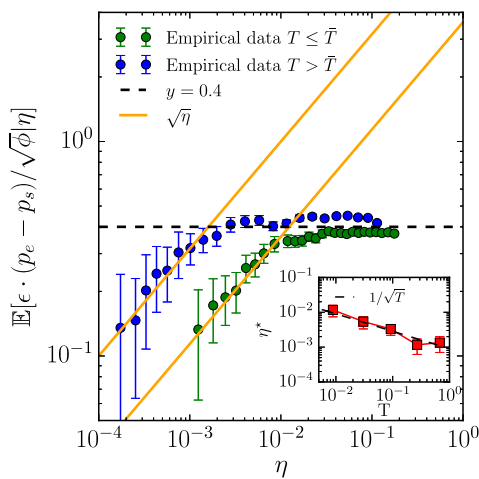


FIG. 2. (Main panel) Empirically determined scaling function  $\mathcal{F}(\eta)$  vs participation rate  $\eta$  for metaorders with duration  $T$  larger (blue dots) or smaller (green dots) than the median sample duration  $\bar{T} \approx 0.09$ . Note that the crossover value  $\eta^*$  is  $\sim 10$  times larger in the latter case. Note that the asymptotic value of  $\mathcal{F}$  (and hence the impact for a given  $Q$ ) is approximately independent of  $T$ , as predicted by the theory. Both empirical curves are obtained using metaorders with sufficient large order size, i.e.,  $\phi \gtrsim 10^{-5}$ . (Inset) Plot of the crossover participation rate  $\eta^*$  as a function of the execution time  $T$ , revealing an approximate  $T^{-1/2}$  behavior.

than the median execution time  $\bar{T} \approx 0.09$ . The crossover participation rate  $\eta^*$  for small durations is found to be 10 times larger for large durations. In the inset of Fig. 2 we show the  $T$  dependence of  $\eta^*$ , obtained by fitting the rescaled data by  $\mathcal{F}(\eta/\eta^*)$  using five bins of  $T$  containing the same number of data points ( $\sim 1.4 \times 10^6$ ), suggesting  $\eta^* \sim T^{-1/2}$  [32]. It would be very interesting to use this result to map out the frequency distribution of the hidden liquidity, but this requires going beyond the approximate solution of Ref. [23]. We leave this for a subsequent investigation [33].

In this Letter, we have used a very large dataset of orders executed in the U.S. equity market to quantitatively test a market impact model, which predicts a crossover from a linear (in volume) behavior for small volumes to a square-root behavior for intermediate volumes. The data unambiguously suggest the existence of such a crossover, and once again confirms the square-root law which, as emphasized on several previous occasions, is remarkably independent of the execution time—contradicting many early theories [3,26,27]. We have shown how the data points towards the existence of multiple timescales in the dynamics of liquidity, with its high-frequency component dominating the total market activity and its low-frequency component contributing to the concavity of the impact function. (For a complementary viewpoint, see Ref. [34]). Our results are interesting from two rather different points of views. One is that they represent a significant improvement in our understanding of the determinants of market impact which is both the main component of trading costs for institutional investors and an important aspect of the stability of financial markets. The second aspect is that we are entering an era where economic and financial data become of such quality that theoretical ideas can be tested with standards comparable to those of natural sciences.

The data were purchased by Imperial College from the company ANCerno Ltd. (formerly the Abel Noser Corporation), which is a widely recognized consulting firm that works with institutional investors to monitor their equity trading costs. See Ref. [22] for details.

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