Beyond the Death of Linear Response: 1/f Optimal Information Transport

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(Received 13 July 2009; revised manuscript received 26 May 2010; published 21 July 2010; publisher error corrected 26 July 2010)

Nonergodic renewal processes have recently been shown by several authors to be insensitive to periodic perturbations, thereby apparently sanctioning the death of linear response, a building block of nonequilibrium statistical physics. We show that it is possible to go beyond the "death of linear response" and establish a permanent correlation between an external stimulus and the response of a complex network generating nonergodic renewal processes, by taking as stimulus a similar nonergodic process. The ideal condition of 1/f noise corresponds to a singularity that is expected to be relevant in several experimental conditions.

DOI: 10.1103/PhysRevLett.105.040601 PACS numbers: 05.40.—a, 87.18.Tt, 87.19.lm, 89.70.Hj

There has been a surge of interest in understanding the dynamics of complex networks over the past decade with studies ranging from the topology of transportation webs to the connectivity of communication meshes to the dynamics of neuron networks. Most recently, the importance of the nascent theory of information exchange between complex networks has become evident.

In living neural networks, the connection between function and information transport is studied with experimental techniques of increasing efficiency [1] from which an attractive perspective is emerging; i.e., these complex networks live in a state of phase transition (collective, cooperative behavior), a critical condition that has the effect of optimizing information transmission [2]. From the studies of complex networks, it is evident that the statistical distributions for network properties are inverse power laws and that the power-law index is a measure of the degree of complexity. Intimate connections exist between neural organization and information theory, the empirical laws of perception [3], and the production of 1/f noise [4], with the surprising property that 1/f signals are encoded and transmitted by sensory neurons with higher efficiency than white noise signals [5]. Although 1/f noise production is interpreted by psychologists as a manifestation of human cognition [6], and by neurophysiologists [7] as a sign of neural activity, a theory explaining why this form of noise is important for communication purposes does not exist yet.

The well known stochastic resonance phenomenon [8] describes the transport of information through a random medium, obeying the prescriptions of Kubo linear response theory (LRT) [9], being consequently limited [10] to the stationary equilibrium condition. There are many complex networks that generate 1/f noise and violate this condition: Two relevant examples are blinking quantum dots [11] and liquid crystals [12]. The non-Poisson nature of

the renewal processes generated in these examples [13] is accompanied by ergodicity breakdown and nonstationary behavior [14].

The response of nonstationary networks to harmonic perturbation has recently been found by many authors [15–20] to fade away with time, an effect called death of linear response [15]. This result seems to call into question one of the fundamental theories of statistical physics, that being the fluctuation-dissipation theorem and the resulting LRT of Kubo [9]. The intuitive explanation of this effect is as follows. The complex network is prepared and perturbed at time t = 0. Experimental preparation generates a cascade of events, whose rate R(t) is a decreasing function of time, thereby making the response, which is proportional to R(t), fade away. Is this a general property, independent of whether the stimulus is periodic [15-20]? If this was a general result, it would be difficult, if not impossible, to explain the communication properties revealed by the recent neurophysiological literature.

The purpose of this Letter is to prove that *the death of LRT* actually rests on an extension of Kubo LRT to the nonstationary condition (NSLRT) and that consequently a nonergodic system, insensitive to perturbations with a fixed time scale, does respond to perturbations sharing the same nonergodic behavior. We shall argue that this important phenomenon explains why 1/f noise is an efficient stimulus for complex systems.

The NSLRT rests on the general LRT form

$$\sigma(t) = \langle \xi_S(t) \rangle = \varepsilon \int_0^t \chi(t, t') \xi_P(t') dt', \tag{1}$$

where the subscripts S and P denote the "system" network and the perturbing network, respectively. Note that $\langle \xi_S(t) \rangle$ is the Gibbs ensemble average over infinitely many responses $\xi_S(t)$ to $\xi_P(t)$ and $\epsilon \ll 1$ is the stimulus strength. We make the simplifying but realistic assumption that the

preparation of S [12] does not set a bias on S, so that $\langle \xi_S(0) \rangle = 0$. The function $\chi(t, s)$ is given by [21,22]

$$\chi(t, t') = \frac{d\Psi_S(t, t')}{dt'} = R_S(t')\Psi_S(t - t').$$
 (2)

The function $\Psi_S(t, t')$ is the autocorrelation function of $\xi_S(t)$, namely, the survival probability of age t', and $R_S(t)$ for the case of discrete signals considered here is the rate at which events are produced by the network S prepared at t=0, i.e., the bits per second encoded in $\xi_S(t)$. This rate is time-independent only in the Poisson case. In the non-Poisson case, this rate depends on time, thereby making $\Psi_S(t, t')$ nonstationary. The brand new survival probability $\Psi_S(t) = \Psi_S(t, t' = 0)$ is given by [16,21,22]

$$\Psi_S(t) = (1 + t/T_S)^{1-\mu_S},\tag{3}$$

from which the corresponding waiting-times probability density $\psi_S(t) = -d\Psi_S(t)/dt$ is derived. In the range of parameters $1 < \mu_S < 3$ considered here, it is known [23] that

$$R_S(t) \approx -\frac{\sin \pi \mu_S}{T_S} (T_S/t)^{2-\mu_S}$$
 for $1 < \mu_S < 2$, (4)

$$R_S(t) \approx \frac{1}{\tau_S} [1 + (T_S/t)^{\mu_S - 2}]$$
 for $2 < \mu_S < 3$, (5)

with $\tau_S = T_S/(\mu_S - 2)$ the mean value of $\psi_S(t)$.

When μ_S < 2, the experimental preparation of S induces a sequence of events, whose rate R_S tends to vanish for $t \to \infty$, yielding a perennial out-of-equilibrium condition and an explanation of the death of linear response [15–20] as well. In fact, the response to a harmonic perturbation of frequency f is proportional to $1/(ft)^{2-\mu_s}$ [16]. In the case $2 < \mu_S < 3$, on the contrary, the preparation-induced cascade of events, in the limit $t \to \infty$, becomes stationary and virtually identical to that of a Poisson process. The theoretical analysis of this Letter is done in the asymptotic time regime. Thus, we refer to the case $2 < \mu < 3$ as stationary, in contrast to the *nonstationary* case $\mu \leq 2$ of perennial transition. Before proceeding with the use of the NSLRT of Eq. (1), we point out some important properties of both signals $\xi_S(t)$ and $\xi_P(t)$. If necessary, the signal $\xi_P(t)$ must share the same properties as $\xi_S(t)$, and for simplicity they are both assumed to be dichotomous signals with random renewal fluctuations between the values +1 and -1. The survival probability in each state is given by Eq. (3) with parameters carrying the appropriate index: T_S , μ_S for S and T_P , μ_P for P. The spectrum of this type of fluctuating signal, as calculated in Refs. [24,25], is

$$S(f) \propto L^{\mu-2} f^{\mu-3},\tag{6}$$

valid for $\mu < 2$, remarkably, even though a stationary autocorrelation function cannot be defined in this case. In the case $\mu > 2$, $S(f) = A/f^{3-\mu}$, with A independent of L, the length of the sequence under study. Similarly to the rate of events $R_S(t)$, the spectral intensity per unit time tends to vanish for $\mu < 2$ as an effect of increasing L. The ideal 1/f

noise condition, corresponding to $\mu = 2$, generates instead a logarithmic decrease of the spectral intensity with time and consequently a spectrum virtually independent of L.

Averaging Eq. (1) over the external fluctuations $\xi_P(t)$, we obtain

$$\langle \sigma(t) \rangle = \langle \langle \xi_S(t) \rangle \rangle = \varepsilon \int_0^t \chi(t, t') \langle \xi_P(t') \rangle dt'.$$

As previously mentioned, the NSLRT of Eq. (1) rests on the preparation of S at time t=0. We apply the same preparation condition to P, thereby generating the cascades $R_S(t)$ and $R_P(t)$ described by Eqs. (4) and (5), with the appropriate indexing. Under this condition the relaxation of $\langle \xi_P(t) \rangle$ becomes identical to the survival probability $\Psi_P(t)$. Assuming the condition of Eq. (2), we have the following expression for the average response:

$$\langle \sigma(t) \rangle = \varepsilon \int_0^t R_S(t') \Psi_S(t-t') \Psi_P(t') dt'.$$
 (7)

The preparation of both S and P makes the average over many realizations of the response $\sigma(t)$ to a given stimulus P vanish for $t \to \infty$. While we refer the readers to Ref. [26] for details, hereby we prove that the intensity of the response of S with $\mu_S < 2$ to $\xi_P(t)$ does not decay if $\mu_P < 2$. This is what we mean by going beyond the LRT death, claimed by many researchers [15,17–20]. To prove this important fact, we study the cross-correlation (or input-output correlation) function between the system S and the stimulus P: $C(t) \equiv \langle\langle \xi_S(t)\xi_P(t)\rangle\rangle$ and the mutual information, which are used as indicators of aperiodic stochastic resonance [27]. Multiplying both sides of Eq. (1) by $\xi_P(t)$ and averaging over the fluctuations of the perturbation P, we obtain

$$\Phi(t) \equiv C(t)/\varepsilon = \int_0^t dt' R_S(t') \Psi_S(t-t') \Psi_P(t,t'). \tag{8}$$

Note that both Eqs. (7) and (8) depend on the survival probability of network P, but the former depends on the single time t', whereas the latter depends on both t' and t and is nonstationary. We limit ourselves to report the results for the asymptotic value Φ_{∞} of $\Phi(t)$. When $\xi_S(t)$ and $\xi_P(t)$ are not stationary, i.e., when $1 < \mu_S \le 2$ and $1 < \mu_P \le 2$, Eq. (8), in the limit $t \to \infty$, gives

$$\begin{split} \Phi_{\infty} &= \zeta(\mu_{S}, \mu_{P}) \\ &\equiv \Gamma(\mu_{S} + \mu_{P} - 2) \\ &\times \frac{{}_{3}F_{2}[\{\mu_{P} - 1, \mu_{P} - 1, \mu_{P} + \mu_{S} - 2\}, \{\mu_{P}, \mu_{P}\}, 1]}{\Gamma(2 - \mu_{P})\Gamma(\mu_{P})^{2}\Gamma(\mu_{S} - 1)}, \end{split}$$
(9)

where $_3F_2$ is the generalized hypergeometric function. In the case $2 < \mu_P < 3$, Φ_{∞} is simply zero.

In the case $2 < \mu_S < 3$, inserting into Eq. (8) expression (5) for $R_S(t)$ leads to

$$\Phi(t) \simeq \int_0^t dt' \frac{1}{\tau_S} \Psi_S(t - t') \Psi_P(t, t'). \tag{10}$$

TABLE I. Summary of the asymptotic values of the cross-correlation function $\Phi(t)$.

$\mu_{S\downarrow} \mu_{P\rightarrow}$	$1 < \mu_P \le 2$		$2 < \mu_P < 3$	
$1 < \mu_S \le 2$	$\Phi_{\infty} = \zeta(\mu_S, \mu_P)^{\mathrm{a}}$	I	$\Phi_{\infty} = 0$	II
$2 < \mu_S < 3$	$\Phi_{\infty} = 1$	III	$\Phi_{\infty} = \frac{\mu_S - 2}{\mu_S + \mu_P - 4}$	IV

^aSee Eq. (9).

Equation (10) is exact for $t \gg \tau_S$, and for $1 < \mu_P \le 2$ it leads [26] to $\Phi_{\infty} = 1$, while for $2 < \mu_P < 3$ it yields

$$\Phi_{\infty} = (\mu_{S} - 2)/(\mu_{P} + \mu_{S} - 4). \tag{11}$$

Results are summarized in Table I. For illustrative purposes, we supplement Table I with Fig. 1, showing the 3D plot of the cross-correlation function Φ_{∞} in the same parameter range: Square II and square III correspond to the condition of minimal and maximal correlation, respectively. Intuitively, it is so because of the difference of time scales between S and P in such regions. In III, fluctuations $\xi_S(t)$ and $\xi_P(t)$ have a finite and an infinite time scale, respectively, thereby allowing $\xi_S(t)$ to adapt to the stimulus-induced bias so as to yield maximal correlation. In II, the role of the time scales is inverted, and the bias induced by P on the longer (diverging) time scale of the process $\xi_S(t)$ is asymptotically averaged out due to the many intervening switching events of $\xi_P(t)$, producing no correlation. The vertex $\mu_S = \mu_P = 2$, representing a 1/f-noise system under the stimulus of a 1/f-noise perturbation, marks the abrupt transition from vanishing (square II) to maximal correlation (III).

Now let us proceed to the demonstration that the intensity of the response $\sigma(t)$ to a single realization of the stimulus does not decay, if $\Phi_{\infty} \neq 0$. We note that by definition the nonvanishing Φ_{∞} yields

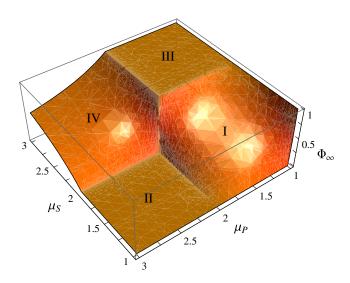


FIG. 1 (color online). The asymptotic limit of $\Phi(t)$ is displayed for μ_S , $\mu_P \in]1, 3[$. The vertex $\mu_S = \mu_P = 2$ marks the transition to a condition of maximal input-output cross correlation.

$$C(t) \equiv \sum_{i,j} ij p \big[\xi_S(t) = i \big| \xi_P(t) = j \big] p \big[\xi_P(t) = j \big] \rightarrow \varepsilon \Phi_{\infty},$$

where the conditional probability for the occurrence of a value of $\xi_S = i = \pm 1$, given the occurrence of a value of $\xi_P = j = \pm 1$, has been introduced. We note that for $t \to \infty$, on a time scale such that $\langle \xi_P(t) \rangle \sim \langle \xi_P(0) \rangle t^{1-\mu_P}$ is a second-order quantity $O(\varepsilon^2)$, we have $P[\xi_P(t) = j] = 1/2 + O(\varepsilon^2)$ and $\Phi(t) = \Phi_\infty + O(\varepsilon^2)$. Thus, due to the symmetry of the considered dichotomous processes:

$$p_t(\xi_S^i|\xi_P^j) \equiv p[\xi_S(t) = i|\xi_P(t) = j] \rightarrow \frac{1}{2} + ij\varepsilon \frac{\Phi_\infty}{2}. \quad (12)$$

In the same long-time scale, Eq. (12) yields

$$\langle \sigma(t) \rangle_{\pm} \equiv \sum_{i} p_{t}(\xi_{S}^{i} | \xi_{P}^{\pm 1}) i \simeq \pm \varepsilon \Phi_{\infty}, \tag{13}$$

where the subscript \pm indicates the value of ξ_P at time t. Summing Eq. (13) over the two values of ξ_P gives a total average null response, as expected. But if the magnitude $|\sigma(t)|$ of the response to a single instance of the input $\xi_P(t)$ is considered instead, its total average is

$$\langle |\sigma(t)| \rangle = \frac{1}{2} \sum_{\pm} \langle |\sigma(t)| \rangle_{\pm} \gtrsim \frac{1}{2} \sum_{\pm} |\langle \sigma(t) \rangle_{\pm}| \simeq \varepsilon \Phi_{\infty}, \quad (14)$$

where an equality holds if terms of order $O(\varepsilon^2)$ are neglected. Thus, when $\Phi_\infty > 0$, the response $\sigma(t)$ to a single instance of the input $\xi_P(t)$ does not die out and remains proportional to the stimulus, no matter how large t becomes. Square III is the plateau region of maximal cross correlation and response, together with the degenerate limit case $\mu_P = 1$ of square I. The term death of linear response is appropriate for the vanishing correlation of square II. The total average response $\langle \sigma(t) \rangle$ always tends to vanish for $t \to \infty$ for reasons that do not imply a lack of response except in the case of square II. The everlasting response to a single complex stimulus is confirmed numerically by Fig. 2, whose inset shows the correlation emerging from averaging over many realizations of the stimulus.

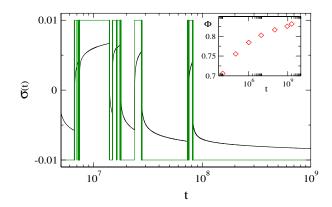


FIG. 2 (color online). Response $\sigma(t)$ (black line) to input $\xi_P(t)$ (green square line) rescaled by ε , for $\mu_S=1.9$, $T_S=10$, $\mu_P=1.55$, and $T_P=9$. Inset: Average of $\sigma(t)\xi_P(t)$ over $N=10^4$ inputs, converging to $\Phi_\infty=0.85$ as predicted by Eq. (9).

The reason for the striking difference between the response to a harmonic perturbation and the response to a nonergodic stimulus is intimately related to the emergence of 1/f noise and to its spectrum described by Eq. (6), which assigns the weight $S(f)/L = 1/(fL)^{3-\mu_P}$ to the spectral component of frequency f of a nonergodic stimulus. As a consequence, the stimulus generates, in time, smaller and smaller frequencies f, so as to keep $1/(fL)^{2-\mu_S}$ (i.e., the response intensity to frequency f [16]) finite, thereby yielding Eq. (14). The death of linear response [15,17–20] is caused by the fact that stimuli with fixed frequencies cannot cope with the decreasing frequency of the cascade of events of Eq. (4).

We have afforded a compelling proof that the intensity of the single realizations of $\sigma(t)$, with $\mu_S < 2$, does not decay if the perturbation $\xi_P(t)$ falls in the same complexity basin ($\mu_P < 2$). We refer to this phenomenon more appropriately as *complexity management* since, differently from the *complexity matching* conjecture advanced in Ref. [23], it does not imply maximum response when $\mu_S = \mu_P$.

Now we argue that 1/f stimuli generate the maximum information transport by looking at the mutual information

$$I(t) = \sum_{i,j} p_t(\xi_P^j) p_t(\xi_S^i | \xi_P^j) \log[p_t(\xi_S^i | \xi_P^j) / p_t(\xi_S^i)]. \quad (15)$$

By using Eq. (12), in fact, it follows that $I(t \to \infty) \simeq \varepsilon^2 \Phi_\infty^2$ and the information transmission rate is obtained by multiplying I(t) by the input rate [28], given by $R_P(t)$. If $\mu_P > 2$, Fig. 1 shows that $\Phi_\infty < 1$. Although square III in Fig. 1 indicates that all stimuli with $\mu_P \le 2$ induce maximal correlation, $\mu_P < 2$ corresponds to a stimulus with decaying events rate (input bits/sec) $R_P(t)$. So even if a response is produced in this regime, the rate of information transfer vanishes in time. Only at the crucial condition $\mu_P = 2$, of ideal 1/f noise, does this algebraic decay become logarithmic, and, consequently, a steady and maximal information transmission rate is achieved.

Experimental verification either on liquid crystals [12] or on ion channels, whose open/close dynamics has been reported to have 1/f properties [29], is desirable. In the latter case, by using a patch-clamp technique, a 1/f stimulus can be used as a stimulus and the correlation with the current output analyzed.

In conclusion, the NSLRT proposed in this Letter [30] explains not only the mystery of the efficient transport of information emerging from the latest theoretical and experimental results in neurophysiology [2,5], but many other forms of 1/f noise propagations, e.g., why ecological time series tend to exhibit 1/f noise if the underlying abiotic perturbations are 1/f noise [31].

M.B. and P.G. acknowledge support from ARO and Welch through Grants No. W911NU-05-1-0205 and No. B-1577.

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$$\chi(t, t') = d\Psi_S(t, t')/dt = \psi_S(t, t'),$$

which is called dynamic NSLRT and fits, e.g., the experimental results of Ref. [12]. Nonetheless, we adopt here Eq. (2), which yields a response to periodic stimuli identical to Refs. [15,17–20]. This choice makes more transparent our arguments on the role of $R_S(t)$, without altering the nature of the new NSLRT. The two prescriptions, as reported in Ref. [26], generate qualitatively equivalent results.

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