Quantum Algorithm for Exact Monte Carlo Sampling

Nicolas Destainville, ^{1,2} Bertrand Georgeot, ^{1,2} and Olivier Giraud ^{1,2,3}

¹Université de Toulouse; UPS; Laboratoire de Physique Théorique (IRSAMC); F-31062 Toulouse, France

²CNRS; LPT (IRSAMC); F-31062 Toulouse, France

³LPTMS, CNRS and Université Paris-Sud, UMR 8626, Bâtiment 100, 91405 Orsay, France

(Received 9 March 2010; published 23 June 2010)

We build a quantum algorithm which uses the Grover quantum search procedure in order to sample the exact equilibrium distribution of a wide range of classical statistical mechanics systems. The algorithm is based on recently developed exact Monte Carlo sampling methods, and yields a polynomial gain compared to classical procedures.

DOI: 10.1103/PhysRevLett.104.250502 PACS numbers: 03.67.Ac, 05.10.Ln

The possibility of using quantum mechanics to treat information and perform computation has attracted great interest in the recent past (see, e.g., [1] for a review). Quantum algorithms have been devised, which solve computational problems faster than their classical counterpart, such as the factorization algorithm of Shor [2]. However, relatively few problems have been identified which are amenable to quantum speedup. While many works have proposed methods to simulate quantum systems using a quantum processor (see, e.g., [1] and references therein), fewer have tried to build quantum algorithms to speed up classical physical problems [3]. In particular, statistical physics is the source of many computational problems which have led to great efforts of several communities to develop efficient classical algorithms. For example, the goal of many Monte Carlo algorithms is to sample a configuration set Ω from an equilibrium probability distribution π [4]. It is therefore important to explore the possibilities to use quantum computers in order to speed up such problems. Some quantum algorithms have been proposed to approximate the partition functions of certain statistical physics models [5], or even obtain it exactly in very specific cases [6]. In the very recent past, many works have focused on quantum algorithms implementing classical Markov Chain Monte Carlo (MCMC) methods through quantum walks [7–10]. In general, these methods give a quadratic gain compared to classical simulations.

Here we consider another type of MCMC algorithm recently developed, the "coupling from the past" (CFTP) procedure of Propp and Wilson, which leads in finite time to the exact equilibrium distribution [11]. We propose a quantum algorithm combining this CFTP procedure and the quantum search procedure of Grover [12], enabling a quadratic speedup over the classical algorithm, without using quantum walks. Our method enables to sample the exact equilibrium distribution in finite time for a wide class of systems, while previous algorithms either provide an approximate version whose error has to be controlled [5,7–10], or are restricted to specific models [6]. Our algorithm is also rather simple compared to these other methods, while yielding a comparable polynomial gain.

One of the key issues in classical MCMC algorithms is that they must be iterated sufficiently many times so that the final state is a "typical" configuration, in other words has a probability distribution close to the stationary one, π , independently of the initial state. In order to get close to the correct distribution π , one should be able to know when sufficient convergence is achieved. In a few particular cases, it is possible to calculate analytically the relaxation time of the algorithm, i.e., the typical time needed to reach stationarity. But in practice, estimating or bounding relaxation times is a notoriously difficult mathematical problem [13–18] and one has to rely on heuristic arguments to infer that stationarity has approximately been reached.

An elegant alternative way to circumvent this issue has been proposed by Propp and Wilson in 1996 [11,19]. As detailed below, the CFTP technique is a reformulation of the MCMC procedure that generates exact samples, in the sense that they are exactly distributed according to the stationary distribution π . Thus successive calls of the algorithm generate totally uncorrelated samples (see below). The basic idea is to run the Markov chain "from the past," from a time -t up to t = 0. Now suppose that there exists a time -T such that at t = 0 all states have coalesced (or "coupled"); i.e., their evolution through the algorithm has led to the same state x_c of the configuration set Ω . Then any initial configuration at $t = -\infty$ would lead to the same state x_c , which can thus be seen as the result of an infinite time simulation. The state x_c is consequently distributed exactly according to the stationary distribution. The difficulty of the procedure dwells in the necessity to track the evolution of the whole set Ω , whereas in standard Monte Carlo sampling only one state of Ω is tracked.

When stored in a computer memory, the configuration set is always finite. Thus we will consider a discrete time Markov chain [20] on a finite configuration set Ω of cardinality N. Our quantum algorithm consists in replacing the classical evolution of the N states of Ω by a quantum evolution of a single quantum state, namely, the superposition of the N ones, $\frac{1}{\sqrt{N}}\sum_{x}|x\rangle$; then the Grover quantum search procedure is applied on top of this quantum evolution to find efficiently if the system has coalesced.

Let us now detail the classical CFTP algorithm. Without loss of generality, a state $x \in \Omega$ (with Ω of cardinal N) can be coded by *n* classical bits $b_i = 0$, 1 as $x = b_1, ..., b_n$ and $N \leq 2^n$. Let P be the transition matrix of the Markov chain. Its elements are the transition probabilities P(x, y), with $P(x, y) \ge 0$ being the conditional probability that the chain is in the state x at time t + 1 given that it was in the state y at time t. The chain is supposed to be reversible, which means that it satisfies the detailed balance condition [4,20]: there exists a probability distribution on Ω , denoted by π , such that $\pi(x)P(y,x) = \pi(y)P(x,y)$ for all states x and y. This condition ensures that π is a stationary distribution. We assume that π exists and is unique, in which case it coincides with the equilibrium distribution (see [20] for further details). If $P(x, T|x_0, 0)$ is the probability that the chain is in the state x at time T given that it was in the state x_0 at t = 0, then [20]

$$\lim_{T \to \infty} P(x, T | x_0, 0) = \pi(x). \tag{1}$$

This result is central in traditional Monte Carlo sampling: if the algorithm (the Markov chain) is iterated long enough, then the probability distribution of its final state is close to the stationary one.

A Monte Carlo step at time t can be seen as a map f_t : $\Omega \to \Omega$, determined by a randomly generated parameter α_t as $f_t(\cdot) = \phi(\cdot, \alpha_t)$. Thus once the random numbers α_t are set, each step of the algorithm is completely deterministic. If T is the duration (number of steps), then the algorithm is entirely coded by the map $F_T = f_T \circ \dots \circ$ $f_2 \circ f_1$. A standard way of performing Monte Carlo sampling consists in following the dynamics of a single initial state during a sufficiently large time and averaging a physical observable \mathcal{O} over time iterates. In this case, the statistical error on the numerical measure of $\langle \mathcal{O} \rangle$ can be estimated using the relaxation time $\tau_{\mathcal{O}}$ of \mathcal{O} . Indeed, the algorithm behaves as if roughly $T/ au_{\mathcal{O}}$ independent realizations were measured, leading to an error $Err_{\mathcal{O}} =$ $\sqrt{2\tau_{\mathcal{O}}/T}\Delta\mathcal{O}$, where T is the total simulation duration and $\Delta \mathcal{O}$ the standard deviation of \mathcal{O} [4]. In principle, $\tau_{\mathcal{O}}$ can itself be numerically measured through the correlation function of \mathcal{O} , $C_{\mathcal{O}}(s) = \langle \mathcal{O}(t+s)\mathcal{O}(t)\rangle_t - \langle \mathcal{O}\rangle^2 \propto$ $\exp(-s/\tau_{\mathcal{O}})$. However, $C_{\mathcal{O}}(s)$ is itself an equilibrium quantity that can be measured only if the system has reached stationarity, and the measured $\tau_{\mathcal{O}}$ may be only representative of a long transient regime instead of equilibrium. This is particularly critical in disordered, glassy systems where it is impossible to ascertain whether equilibrium has indeed been reached, because the system is likely to get trapped for a long time in the many metastable states [21-23].

Instead of iterating the states in the future as in the traditional method explained just above, the CFTP procedure goes from the past: we suppose that we have at our disposal a sequence of random numbers $\alpha_{-1}, \alpha_{-2}, \ldots$ before starting the algorithm. The CFTP algorithm constructs the iterates of all the states $x \in \Omega$ until they have all

reached the same state ("coalescence"). The essential subroutine of the algorithm [let us call it $\Pi(T)$] calculates the N computational paths from time -T up to time 0 through the map $G_T = f_{-1} \circ \ldots \circ f_{-T}$ and tests at the end whether all histories have coalesced to the same state. If the coalescence test fails, the same procedure is started again from an earlier time. The algorithm reads:

T=0 repeat $T=T+\Delta T$ (go ΔT steps back in time) $\Pi(T)$ (follow all N paths and test coalescence) until coalescence is achieved

At coalescence, T is such that one has $G_T(x) = x_c$ for any $x \in \Omega$. Thus for any $t' \leq -T$, $G_{-t'}(x) = x_c$. In particular, $\lim_{t' \to -\infty} G_{-t'}(x) = x_c$. Therefore x_c can be seen as the result of a Monte Carlo algorithm of infinite duration and is exactly distributed according to π . It is proven that with probability 1 the algorithm returns a value in finite time [11]. Successive calls of the algorithm return totally uncorrelated samples, and $\langle \mathcal{O} \rangle$ is now calculated by averaging over realizations instead of time. The statistical error is now perfectly controlled: $\operatorname{Err}_{\mathcal{O}} = \frac{\Delta \mathcal{O}}{\sqrt{R}}$ where R is the number of realizations (independent calls of the algorithm).

The CFTP can be applied to any MCMC problem [24]. In specific instances, it is sufficient to follow the history of a small subset of the N states. This is the case, e.g., if a partial order on Ω makes it sufficient to follow only extremal configurations. Unfortunately, in general the N states should be followed in parallel, which represents an often prohibitive computational cost. We will propose below a quantum algorithm reducing this cost.

We denote the average coalescence time of the algorithm by $\hat{\tau}$. To what extent is it related to relaxation times $\tau_{\mathcal{O}}$ as discussed above? In principle, $\tau_{\mathcal{O}}$ depends on the observable \mathcal{O} . But $\tau_{\mathcal{O}}$ is bounded above by (and in general on the same order of magnitude as) the relaxation time of the Markov chain, denoted by τ [25]. This latter time measures the speed of convergence of the probability distribution to π and is equal to the inverse of the first gap of the transition matrix P [17,18]. Furthermore, τ is itself bounded above by (and in general on the same order of magnitude as) $\hat{\tau}$ [17], which makes CFTP-type techniques so useful to estimate convergence rates, even on theoretical grounds. All in all, generally speaking, $\hat{\tau} \sim \tau \sim \tau_{\mathcal{O}}$. Running the algorithm yields an accurate estimate of $\hat{\tau}$ and thus of relaxation times.

Let us now turn to our quantum algorithm. The essential subroutine $\Pi(T)$ of the classical CFTP algorithm follows the history of the N states $x \in \Omega$ and tests coalescence of all states. The quantum algorithm will start from a register holding the N initial states $|x\rangle$, $0 \le x \le N-1$, coded on n qubits, and follow each history in parallel. A second register holds the results of the successive application of the maps f_t . Calculation of the iterates is done with the help of ancilla registers. To illustrate the computational scheme, we will first specialize our presentation to the case of the

Ising model with Glauber dynamics. In this case the observable \mathcal{O} could be, e.g., the magnetization. Starting from the totally separable state, Hadamard gates are applied to each qubit to put the register into an equal superposition of all states. Suppose that after t iterations the quantum state reads

$$\sum_{x=0}^{N-1} |x\rangle |H_{t,T}(x)\rangle |0\rangle |0\rangle |0\rangle, \tag{2}$$

where $H_{t,T} = f_{-T+t-1} \circ \dots \circ f_{-T}$. The next Monte Carlo step consists in flipping a certain spin i with some probability function of the energy difference ΔE between configuration $H_{t,T}(x)$ and the same configuration with spin i flipped. This spin i may be chosen at random; alternatively one can consider each spin one after another (sequential sweep). The spin is flipped with probability p = $[1 + \exp(\beta \Delta E)]^{-1}$ or left unchanged with probability 1 – p. Here β is the inverse temperature. In terms of quantum registers, one has to evaluate the energies associated with configurations $|y\rangle = |H_{tT}(x)\rangle$ and $|y^{(i)}\rangle = X^{(i)}|y\rangle$ (where the Pauli matrix $X^{(i)}$ flips spin i), by arithmetic operations controlled by the second register. The probability p is then calculated on the third register and the random number α_t , uniformly distributed on [0,1], is put on the fourth register. The sign of $\alpha_t - p$ is computed, and the one-qubit fifth register $|s\rangle$ is set to $|0\rangle$ if $\alpha - p \ge 0$ and $|1\rangle$ if $\alpha - p < 0$. The bit-flip matrix $X^{(i)}$ is applied, controlled by $|s\rangle$. The last three registers are then reset to $|0\rangle$ in the usual way by running the operations backwards. The circuit in Fig. 1 shows one step of the iteration algorithm before reset of these registers.

After T steps the quantum state reads

$$\sum_{x=0}^{N-1} |x\rangle |G_T(x)\rangle. \tag{3}$$

If T is such that all $G_T(x)$ are equal then the histories have coalesced and a measure of the second register yields an element in Ω distributed exactly according to the stationary distribution π . Since all iterations are performed in parallel this step requires an average time $\hat{\tau}$. The crucial point in the CFTP algorithm is that the exact distribution is obtained if and only if all $G_T(x)$ are equal. Consider, for instance, the case where the $G_T(x)$ take two different

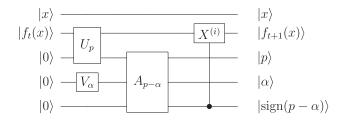


FIG. 1. Circuit for one step of the Monte Carlo algorithm. The unitary operator U_p calculates probability p, V_{α} is the operator implementing the random numbers α , and $A_{p-\alpha}$ is a modified adder circuit that gives the sign of $p-\alpha$.

values, say y_1 and y_2 . Then one might detect that the states have not coalesced onto a unique value as soon as one obtains different results after measuring the second register upon repeated runs of the procedure (with the same random numbers). If the states have not coalesced then the process has to be restarted from an earlier time. Obviously, only in the case where the probabilities of measuring y_1 and y_2 are both high will different outcomes be obtained quickly upon measurement. In the extreme case where there is a unique $x_0 \in \Omega$ such that $G_T(x_0) = y_2$ while all other x verify $G_T(x) = y_1$, almost all measurements of the second register will give y_1 , and the state will be almost indistinguishable from the state where all $G_T(x)$ are equal. Since the CFTP algorithm requires to distinguish these cases, the idea is to use the Grover algorithm to amplify the probability amplitude of the unknown noncoalesced state $|x_0\rangle$, so that it can be detected. If we first measure the second register and consistently get the value y_1 then our aim is to detect whether all x verify $G_T(x) = y_1$ or not. Since we know the value of y_1 , we can attach to our quantum state a one-qubit register in the state $|z\rangle = \frac{1}{\sqrt{2}}(|0\rangle - |1\rangle)$. We then perform the operation $|x\rangle|G_T(x)\rangle|z\rangle \mapsto |x\rangle|G_T(x)\rangle|z+$ $\varphi(x)$, where $\varphi(x) = 0$ if $G_T(x) = y_1$ and $\varphi(x) = 1$ otherwise, and erase all the registers but the first one. This gives a phase $e^{i\pi}$ to states which do not verify $G_T(x) = y_1$. One can thus apply Grover iterations to magnify the amplitude of the noncoalesced states. The whole sequence above corresponds to one "oracle" step of the Grover iteration, and has to be performed using the same random numbers α_t .

What is the speedup on $\Pi(T)$ obtained by this procedure? Suppose that after T time steps M states x are not coalesced. We consider the case $M \ll N$, since otherwise noncoalescence is easily detected by a few measurements or even classically. Then $O(\sqrt{N/M})$ Grover iterations are required (even though M is unknown [26]), each using O(T) operations. The total number of quantum gates in this case is $\sim T\sqrt{N/M}$. Note that, as the random numbers can be generated classically once and for all beforehand, their computational cost is $\sim T$, thus negligible.

As explained above, the complete algorithm proceeds by performing a certain number of calls of $\Pi(T)$ until coalescence. Let us evaluate the speedup of the quantum algorithm for $\Pi(T)$ compared to the classical procedure. We consider that preliminary runs enable quickly to estimate a suitable $\Delta T \lesssim \hat{\tau}$ which is gradually improved. To iterate classically the Monte Carlo steps on one computational path will cost a certain number of computational operations g(N). Let us first suppose that the dynamics is rapid, with a short coalescence time $\hat{\tau}$ with $g(N) \sim \ln^a N$. This means that the dynamics is polynomial in the physical system size n. The classical algorithm will need to follow $\sim N/M$ computational paths to detect the ones which did not coalesce. The total cost is $\sim g(N)N/M \sim$ $(N\ln^a N)/M$. In contrast, in the case of the quantum algorithm, $O(\sqrt{N/M})$ calls to the oracle are required, and the oracle performs the N evolutions in parallel in also

 $\sim g(N) \sim \ln^a N$ operations. The total speedup of the quantum algorithm will be $O(\sqrt{N/M})$, a quadratic gain. If the dynamics is torpid, with a long coalescence time $\hat{\tau}$ with $g(N) \sim N^c$, then the classical algorithm requires $\sim N^{c+1}/M$ operations whereas the quantum one costs only $\sim N^{c+1/2}/\sqrt{M}$ steps. The relative gain gets smaller for increasing c, going from quadratic for small c to almost zero for $c \to \infty$.

The estimates above assume that all initial states, stored in a *n*-bit register and coded by all numbers x between 0 and N-1, are physically admissible. This is the case, e.g., for spin problems. However, in many interesting cases one must inspect a large number of states $(N = 2^n)$, of which only a small subset $(N_a = N^b, b < 1)$ are admissible. This is the case, e.g., for dimer, spanning tree, or hard core lattice gas problems. In this situation, one can use a modified version of the above quantum algorithm. Indeed, starting from an equal superposition of the N states, one can build in the Grover oracle [before each $\Pi(T)$ step] a subroutine which recognizes the nonadmissible states, and overwrites in this case the second register with a known admissible state x_a , so that the latter is used as an initial state in the dynamics. In this case, the quantum algorithm for M = O(1) still requires $\sim g(N)\sqrt{N}$ steps, since the Grover search is applied on the whole Hilbert space of dimension N. To obtain both the equilibrium distribution and the relaxation time, the classical algorithm needs $\sim N$ operations to identify admissible states, and $\sim g(N)N^b$ operations to run the CFTP on the admissible states. If $g(N) \sim \ln^a N$, the gain is unchanged. If $g(N) \sim$ N^c , the gain is $\max(c+b,1)/(c+1/2)$. The gain is polynomial in all cases except for b < 1/2 and c > 1/2 (very long relaxation time and very few admissible states).

The algorithm proposed here presents several advantages compared to the recently proposed method for simulating Markov Chain systems [7–10]. These procedures use quantum walks to approximate the stationary distribution, in a time typically quadratically faster than the classical convergence time. The speedup over classical computation is therefore comparable but, in our case, we obtain a sampling of the exact stationary distribution rather than an approximate version of it with errors that have to be controlled. Our quantum algorithm is also simpler. Another advantage of our method is that the relaxation time is directly related to the average coalescence time [17] and thus can be accurately measured.

It has been proven that the Grover algorithm is optimal, in the sense that the number of calls to the oracle cannot be lower (see [1] and references therein). Our quantum algorithm can therefore be improved only by speeding up the oracle part. This may be possible by combining our algorithm with techniques used in the other approaches of [5–10]. As the algorithm developed here is very general, applying to a wide class of systems without any structure taken into account, tailored algorithms may achieve a larger gain in specific cases.

- [1] M. A. Nielsen and I. L. Chuang, *Quantum Computation* and *Quantum Information* (Cambridge Univ. Press, Cambridge, 2000).
- [2] P. W. Shor, in *Proceedings of the 35th Annual Symposium* on the Foundations of Computer Science (IEEE Computer Society, Los Alamitos, CA, 1994), p. 124.
- [3] D. A. Meyer, Phil. Trans. R. Soc. A 360, 395 (2002); B. Georgeot, Phys. Rev. A 69, 032301 (2004); B. Georgeot and O. Giraud, Phys. Rev. E 77, 046218 (2008).
- [4] M. E. J. Newman and G. T. Barkema, Monte Carlo Methods in Statistical Mechanics (Clarendon Press, Oxford, 1999).
- [5] D. Aharonov et al., arXiv:quant-ph/0702008; M. Van den Nest et al., Phys. Rev. A 80, 052334 (2009).
- [6] D. A. Lidar and O. Biham, Phys. Rev. E 56, 3661 (1997);
 J. Geraci and D. A. Lidar, Commun. Math. Phys. 279, 735 (2008).
- [7] M. Szegedy, in *Proceedings of the 45th Annual IEEE Symposium on Foundations of Computer Science* (IEEE, New York, 2004), p. 32.
- [8] F. Magniez et al., in Proceedings of the 39th Annual ACM Symposium on Theory of Computing (ACM, Washington, DC, 2007), p. 575.
- [9] R.D. Somma *et al.*, Phys. Rev. Lett. **101**, 130504 (2008).
- [10] P. Wocjan and A. Abeyesinghe, Phys. Rev. A 78, 042336 (2008); P. Wocjan *et al.*, Phys. Rev. A 80, 022340 (2009);
 D. Poulin and P. Wocjan, Phys. Rev. Lett. 103, 220502 (2009).
- [11] J. G. Propp and D. B. Wilson, Random Struct. Algorithms 9, 223 (1996); J. G. Propp and D. B. Wilson, Discrete Math Theor. Comput. Sci. 41, 181 (1998).
- [12] L. K. Grover, Phys. Rev. Lett. 79, 325 (1997).
- [13] D. Aldous, L. Lovasz, and P. Winkler, Stoch. Proc. Appl. 71, 165 (1997).
- [14] N. Destainville, Phys. Rev. Lett. 88, 030601 (2002).
- [15] D. B. Wilson, Ann. Appl. Probab. 14, 274 (2004).
- [16] C. Moore and S. Mertens, *The Nature of Computation* (Oxford University Press, Oxford, UK, 2009).
- [17] D. Aldous, Séminaire de probabilités XVII, Lecture Notes in Mathematics Vol. 986 (Springer, New York, 1983), p. 243.
- [18] P. Diaconis and D. Strook, Ann. Appl. Probab. 1, 36 (1991).
- [19] M. A. Novotny, in *Computer Simulation Studies in Condensed Matter Physics XII*, edited by D. P. Landau *et al.*, Springer Proceedings in Physics Vol. 85 (Springer, Berlin, 2000).
- [20] G. Grimmett and D. Stirzaker, Probability and Random Processes (Clarendon Press, Oxford, 1982).
- [21] J.-P. Bouchaud *et al.*, in *Spin Glasses and Random Fields*, edited by A. P. Young (World Scientific, Singapore, 1997).
- [22] F. Ritort and P. Sollich, Adv. Phys. **52**, 219 (2003).
- [23] C. Chanal and W. Krauth, Phys. Rev. Lett. 100, 060601 (2008).
- [24] J. G. Propp and D. B. Wilson, J. Algorithms 27, 170 (1998).
- [25] $\tau_{\mathcal{O}} = \tau$ if the average value of \mathcal{O} on the first eigenmode of P is nonvanishing; otherwise $\tau_{\mathcal{O}} < \tau$.
- [26] M. Boyer et al., Fortschr. Phys. 46, 493 (1998).