Extinction dynamics of spiral defect chaos

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(Received 30 April 2018; revised manuscript received 19 December 2018; published 9 January 2019)

Spatially extended excitable systems can exhibit spiral defect chaos (SDC) during which spiral waves continuously form and disappear. To address how this dynamical state terminates using simulations can be computationally challenging, especially for large systems. To circumvent this limitation, we treat the number of spiral waves as a stochastic population with a corresponding birth-death equation and use techniques from statistical physics to determine the mean episode duration of SDC. Motivated by cardiac fibrillation, during which the heart's electrical activity becomes disorganized and shows fragmenting spiral waves, we use generic models of cardiac electrophysiology. We show that the duration can be computed in minimal computational time and that it depends exponentially on domain size. Therefore, the approach can result in efficient and accurate predictions of mean episode duration which may be extended to more complex geometries and models.

DOI: 10.1103/PhysRevE.99.012407

I. INTRODUCTION

Spiral waves are generic solutions of spatially extended excitable systems. Under certain conditions, these spiral waves are unstable and break up, creating multiple, drifting spiral waves. The resulting dynamical state can be described as spiral defect chaos (SDC), present in a variety of different pattern-forming systems [1–9]. During SDC, spiral waves continuously break down to form new ones, and are removed through collisions with other spiral waves or with nonconducting boundaries, as has been shown by many computational studies. This stochastic competition between creation and annihilation persists until the last spiral wave is terminated, with its duration representing a stochastic event. Often, the mean episode duration τ of SDC is of interest, which is a statistical measure of the average until termination. Determining τ through direct simulations of spatially extended models of SDC can be challenging because a statistically significant quantification of this stochastic quantity requires the timeconsuming task of simulating a multitude of episodes [10,11]. This becomes even more problematic for large geometry sizes since τ typically increases as a function of the system size.

In this study, we will use generic excitable systems as an example and show how statistical physics techniques can be used to determine τ . Specifically, we are motivated by cardiac fibrillation and use two cardiac electrophysiological models. During fibrillation, spiral waves underly the irregular conduction patterns and many computational studies have reported that the number of spiral waves increase due to wave break or decrease due to collisions with other spiral waves or with nonconducting boundaries [12–16]. In other words, this fibrillatory state can be described by SDC and lasts until the last spiral wave is terminated.

In this study we are interested in the mean episode duration τ which is a statistical measure of the average time of annihilation of all spiral waves. Determining τ through direct simulations of spatially extended cardiac models is challenging because previous simulation studies have shown

that τ increases sharply as a function of the system size [17,18]. This increase is related to the so-called critical mass hypothesis which posits that fibrillation requires hearts with a minimal size [19,20].

Our method to compute τ treats number of spiral tips n as a stochastic quantity and casts its birth-death process into a master equation, a commonly used approach in the field of population dynamics [21–23]. This approach was also used previously for understanding spiral wave dynamics in spatially extended fluid dynamical systems [1–9]. Furthermore, it was used in recent studies that examined filament turbulence in phenomenological models and that described the dynamics of surface defects in terms of a master equation [24–26]. Contrary to these studies, we focus here on tips migrating in two dimensions and on termination events and the associated mean episode duration. In our case, the master equation describes the probability P(n,t) of having n spiral tips at time t as

$$\frac{dP(n,t)}{dt} = \sum_{r} [W_r(n-r)P(n-r,t) - W_r(n)P(n,t)],$$
(1)

where W_r are transition rates for the number of spiral tips to change by r tips and can be computed directly from spatially extended simulations of cardiac models. Since tips are created and annihilated either as pairs or as singlets, we only need to consider $r=\pm 1, \pm 2$. As a boundary condition we take n=0 to be absorbing. This means that there is no escape from the no-tip state and that all birth rates for n=0 vanish: $W_r(0)=0$. Furthermore, an additional boundary condition stems from the fact that for n=1 the pairwise death rate equals 0: $W_{-2}(1)=0$. Once the rates are known, we can construct a transition matrix which can be used to compute τ at minimal computational cost [27].

For large n, the death rate will exceed the birth rate since tips will have a high probability of colliding. As a result, the

number of spiral tips does not grow to very large numbers. If for small n the birth rate is larger than the death rate, then a long-lived (quasistationary) metastable state exists with a mean number of tips \bar{n} . The distribution associated with this metastable state is called the quasistationary distribution $P_{qs}(n)$ [28,29]. Note that for systems with an absorbing state at n=0, the stationary distribution trivially corresponds to P(n)=0 for all $n\neq 0$ and P(0)=1. In the quasistationary state, the number of tips fluctuates around the average value for prolonged periods of time and the mean episode duration can be computed using

$$\frac{1}{\tau} = \sum_{r < 0} W_r(-r) P_{qs}(-r). \tag{2}$$

Termination only occurs during rare escape events, corresponding to a large fluctuation away from the mean number of tips. As a consequence, standard equilibrium statistical physics approaches based on small fluctuations do not apply [29,30]. Instead, techniques from nonequilibrium statistical physics must be invoked to determine statistical quantities corresponding to extinction, including τ .

To illustrate our stochastic approach to quantifying termination dynamics, we carry out simulations of SDC using spatially extended electrophysiological models. We should stress, however, that the approach should also work for other systems that exhibit spiral wave dynamics, including the complex Ginzburg-Landau equation [31] or simple phenomenological models [32]. The "direct" simulations use the standard reaction-diffusion equation:

$$\partial_t V = D\nabla^2 V - I_{\text{ion}}/C_m,\tag{3}$$

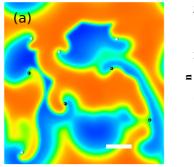
where V is the transmembrane potential, C_m ($\mu F cm^{-2}$) is the membrane capacitance, and $D\nabla^2$ expresses the intercellular coupling via gap junctions and diffusion constant D. The membrane currents in the electrophysiological model are denoted by $I_{\rm ion}$ which are governed by nonlinear evolution equations coupled to V. For our purposes, the precise form of $I_{\rm ion}$ is not important and we present results using the detailed Luo-Rudy (LR) model [33], modified to obtain spiral wave breakup as described in Qu *et al.* [34]. To stress the

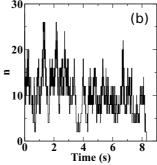
generality of our approach we also carry out simulations using the simplified Fenton-Karma (FK) model (parameter set 8) [35]. We perform the simulations in square two-dimensional computational domains although our approach can be equally well applied in more complex geometries. As boundary conditions, we consider both nonconducting and periodic boundary conditions, and we vary the area of the computational domain, which is equivalent to varying D while keeping the area constant. For both models, we use $C_m = 1 \mu F/cm^2$ while the diffusion constant is chosen to be $D = 0.0005 \text{ cm}^2/\text{ms}$ for the LR model and $D = 0.001 \text{ cm}^2/\text{ms}$ for the FK model. Simulations are carried out with a discretization of 0.025 cm, using a five-point stencil, and a time step of 0.025 ms, using explicit Euler integration. For both models, the conduction velocity along a cable is within the electrophysiological range: 33 cm/s for the LR model and 51 cm/s for the FK model. Errors in direct simulation results are reported as standard deviations.

II. RESULTS USING DIRECT SIMULATIONS

Starting with a random initial condition that contains multiple spiral waves, we solve the reaction-diffusion equation and keep track of the number of spiral tips using a standard algorithm [Fig. 1(a)] [35]. The number of tips fluctuates and the simulation ends after time T_e when the number of spiral tips reaches 0 [Fig. 1(b)]. We can compute the distribution of these termination times by repeating the simulations many times, starting with different and independent initial conditions. These conditions are created by perturbing multispiral states with randomly placed current stimuli in the form of a current stimulus of duration 2 ms and strength several times the excitation threshold. After perturbation, the system is allowed to evolve for another 100 ms before measurements are started. Our simulations reveal that this distribution is exponentially distributed, indicating that spiral wave termination can be well described as a Poisson process [Fig. 1(c)].

Next, we compute the birth and death rates as a function of the number of tips n using different domain sizes with nonconducting boundaries by quantifying the number of





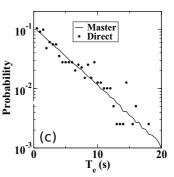


FIG. 1. Direct numerical simulations provide statistics of spiral tip dynamics. (a) Snapshot of a simulation of the LR model in a 7.5×7.5 cm computational domain with periodic boundary conditions. The voltage is represented using a color code with red (blue) corresponding to depolarized (repolarized) tissue. The location of the tips of counter- and clockwise rotation spiral waves are shown in black and white, respectively (scale bar: 1 cm). (b) Typical time trace of the number of spiral tip pairs. For this particular simulation, spiral tips spontaneously extinguished after 8.3 s. (c) Distribution of termination times for the direct simulations (symbols, computed using 400 termination events) and the master equation (solid line, computed using 10 000 termination events).

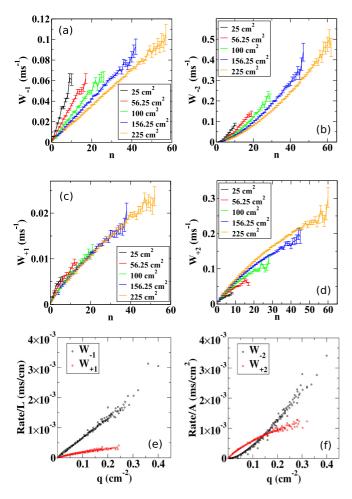


FIG. 2. Transition rates for the LR model computed using direct simulations and nonconducting boundaries. (a)–(d) The birth and death rates for $n \to n-1$ (a), $n \to n-2$ (b), $n \to n+1$ (c), and $n \to n+2$ (d) computed in a square geometry of various sizes. Error bars represent standard deviation. (e) The $W_{\pm 1}$ rates, normalized by the perimeter of the domain, as a function of the density of tips, q = n/A. (f) The $W_{\pm 2}$ rates, normalized by the area of the domain, as a function of the density of tips.

transitions per time interval. The results are presented for the LR model in Figs. 2(a)–2(d). The number of transitions observed in the simulations depends on the domain size and on n and no transitions are recorded above some critical value of n. Here, to increase accuracy, we only consider rates that are determined using at least 100 transitions in the simulation. As a consequence, rates are computed up to a certain maximum value of n. In addition, for increasing domain sizes, transitions for small n become increasingly rare. As a result, in large domains, the number of recorded transitions for small values of n may not reach 100. The rates corresponding to these values of n are therefore not included.

Examining the computed rates, we see that W_{-1} depends linearly on the number of spiral tips for all domain sizes [Fig. 2(a)]. The remaining rates, however, show a more complex dependence on the number of tips, indicating the existence of nontrivial long-range interactions between spiral tips [Figs. 2(b) and 2(d)]. As a result, the rate curves are not easily fit by simple power laws. Therefore, we employ

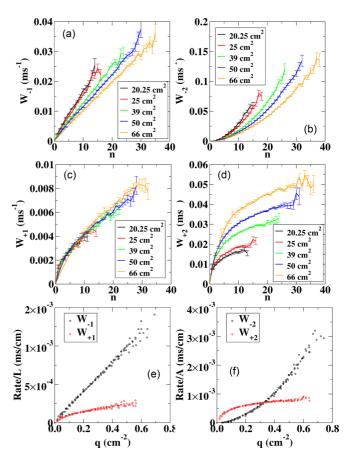


FIG. 3. Transition rates for the FK model computed using direct simulations and nonconducting boundaries. (a)–(d) The birth and death rates for $n \to n-1$ (a), $n \to n-2$ (b), $n \to n+1$ (c), and $n \to n+2$ (d) computed in a square geometry of various sizes. Error bars represent standard deviation. (e) The $W_{\pm 1}$ rates, normalized by the perimeter of the domain, as a function of the density of tips, q = n/A. (f) The $W_{\pm 2}$ rates, normalized by the area of the domain, as a function of the density of tips.

a smoothing spline fit to the data to determine rates corresponding to transition events with less than the minimum number. Note that this interpolation takes into account the zero rate for either n=1 (periodic boundary conditions) or n=0 (nonconducting boundaries).

In addition, we compute the rates for the FK model. The results, presented in Fig. 3, show the same linear dependence of W_{-1} on the number of spiral tips, along with a more complex dependence of the other rates. We can also compute the $W_{\pm 2}$ rates for domains that contain periodic boundary conditions. The results of these simulations are shown in Fig. 4 for both models and are qualitatively similar to the results presented in Figs. 2 and 3. As a consistency check, we can use these rates to compute the distribution of termination times. As expected, this distribution is exponential and agrees well with the one computed using direct simulations [Fig. 1(c)].

Importantly, we find that at large A all rates collapse onto a single curve when plotted as a function of the density q = n/A. Specifically, the $W_{\pm 2}$ rates are found to scale with the area as $W_{\pm 2}(n) \sim Aw_{\pm 2}(q)$ [Figs. 2(f), 3(f), and 4(c) and 4(f)], indicating that the birth and death rates only depend on the density and that tips are well mixed. Furthermore, the

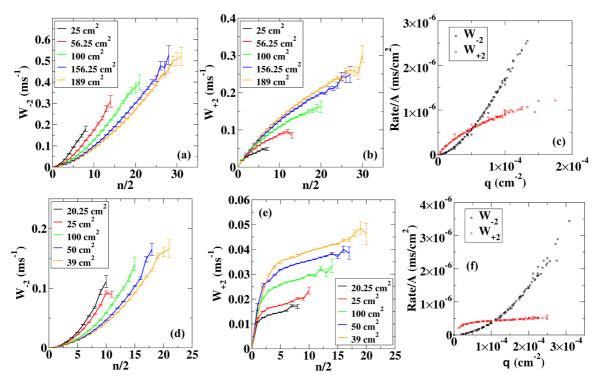


FIG. 4. Transition rates for the LR [(a)–(c)] and FK model [(d)–(f)] computed using direct simulations with periodic boundary conditions. (a),(b) and (d),(e) The birth and death rates for $n \to n-2$ [(a),(d)], $n \to n+2$ [(b),(e)] computed in a square geometry of various sizes with absorbing boundaries. Error bars represent standard deviation. (c),(f) The $W_{\pm 2}$ rates, normalized by the area of the domain, as a function of the density of tips.

 $W_{\pm 1}$ rates scale with the perimeter L as $W_{\pm 1}(n) \sim Lw_{\pm 1}(q)$ [Figs. 2(e) and 3(e)]. Here, and in the following, we will take the continuum limit such that q and functions that depend on this variable are considered to be continuous. Note that this observed linear scaling of W_{-1} with L implies that the death rate is proportional with the length of the nonconducting boundary and that creating ablation lesions will increase this rate. Furthermore, such scaling is expected if single tips annihilate through simple collision processes and get created near the boundaries.

III. RESULTS USING TRANSITION RATES

Once the transition rates are determined, it is straightforward to compute the quasistationary distribution $P_{qs}(n)$ using the transition matrix at minimal computational cost [Figs. 5(a), 5(b), 5(d), and 5(e)] [27]. For small domains, this can be carried out using the rates obtained in the simulations while for larger domains, where the rates for small n cannot be computed accurately, we can use the interpolated rates. As the domain size increases, the distribution shifts to larger values of n, and becomes more symmetric around its peak. Of course, the quasistationary distribution computed using the transition matrix agrees very well with the one determined using direction simulations. This agreement is shown for the largest domain size in Figs. 5(a), 5(b), 5(d), and 5(e) (symbols) but is also valid for other domain sizes. The average number of tips, \bar{n} , increases with system size and our simulations reveal that it depends linearly on the area of the computational domain for both boundary conditions [Fig. 5(c) and 5(f)].

For geometries that do not contain any nonconducting boundaries it is possible to derive closed-form solutions for the quasistationary distribution. In this case, n is always even and tips will be created and annihilated in pairs such that $W_{\pm 1}=0$. The quasistationary distribution can be obtained by setting the left-hand side in Eq. (1) to zero, resulting in the recursion relationship

$$P_{qs}(n) = P_{qs}(0) \prod_{j=2}^{n} W_{+2}(2j-2) / W_{-2}(2j), \tag{4}$$

where $P_{\rm qs}(0)$ can be determined by the normalization condition $\sum_{n=0}^{\infty} P_{\rm qs}(n)=1$ [23].

The deterministic equation corresponding to the master equation can be found in a straightforward manner [23]:

$$\frac{dn}{dt} = 2W_{+2}(n) - 2W_{-2}(n). (5)$$

As a consequence, the deterministic stationary state is determined by $W_{+2}(n^*) = W_{-2}(n^*)$. The maximum value of the quasistationary distribution occurs for $P_{\rm qs}(n-2)/P_{\rm qs}(n) \approx 1$, corresponding to $W_{+2}(\bar{n}-2) = W_{-2}(\bar{n})$. Therefore, for large values of A the stochastic average number can be well approximated by the deterministic average number, $\bar{n} \approx n^*$. Furthermore, using our numerically found scaling, we obtain $w_{+2}(q^*) = w_{-2}(q^*)$, where $q^* = n^*/A$. Hence, the average density is independent of the area and n^* , and thus \bar{n} , scale with A, consistent with the scaling found in the simulations [Figs. 5(c) and 5(f)]. For domains that contain nonconducting boundaries, the ± 1 rates are no longer zero and the

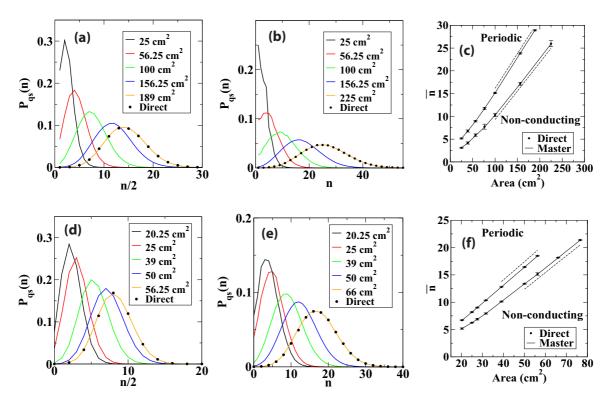


FIG. 5. Dependence of spiral tip dynamics on the domain size for the LR model [(a)-(c)] and the FK model [(d)-(f)]. (a),(b) (d),(e) The quasistationary distribution for periodic [(a),(d)] and nonconducting boundary conditions [(b),(e)] using different domain sizes as computed using the transition matrix. The symbols show the quasistationary distribution as computed using the direct simulations. (c),(f) The average number of tips as a function of the area of the computational domain, computed using direct simulations (symbols) and using the master equation approach (line). The dashed curves are straight lines.

corresponding deterministic equation reads

$$\frac{dn}{dt} = 2W_{+2}(n) - 2W_{-2}(n) + W_{+1}(n) - W_{-1}(n).$$
 (6)

Using our obtained scaling, we have for the deterministic stationary state:

$$2w_{+2}(q^*) - 2w_{-2}(q^*) + w_{+1}(q^*)/\sqrt{A} - w_{-1}(q^*)/\sqrt{A} = 0.$$
(7)

For large areas, the last two terms can be neglected and the average number of tips will again scale linearly with the area.

IV. MEAN EPISODE DURATION

To find the mean episode duration τ in the direct simulations, we average the termination times T_e obtained from each independent simulation. This computation becomes more and more time-consuming as A increases since termination becomes less and less likely. As a consequence, the number of determined termination events we consider vary from 400 for small domains to less than 10 for the largest areas still amenable to direct simulations. Our results reveal that τ increases sharply as the domain size becomes larger, consistent with earlier computational studies [18]. More specifically, τ displays an exponential dependence on the size of the domain, both for periodic and nonconducting boundary conditions (red symbols, Fig. 6), a result that agrees with the earlier study by Qu [17].

Rather than using direct simulations to determine an average value for T_e , it is straightforward to use the interpolated

transition rates and the resulting transition matrix to compute τ using simple matrix operations [27]. For this, we construct a transition matrix Q for all transient states n > 0, with elements Q_{ij} representing the probability of transitioning from state i to state j. The probability of reaching state j from state i in t steps is then given by the ijth entry of Q^t . Summing this over all time results in the so-called fundamental matrix $N = I + Q + Q^2 + \cdots = (I - Q)^{-1}$, where I is the identity matrix. Each element of the fundamental matrix N_{ii} represents the mean duration our system will spend in state j given an initial state i, which can be used to determine the quasistationary distribution. Moreover, the mean time to extinction τ is given by $N\vec{e}$, where \vec{e} is a column vector of ones. The confidence intervals for τ are computed through bootstrapping as follows. First we resample each transition rate by drawing a value from a binomial distribution with probability equal to the original transition rate and using the number of recorded transitions from the direct simulation. We then proceed by interpolating these resampled transition rates and computing τ from these interpolated transition rates. This is computed for 1000 trials and the confidence interval is determined from the 5th to the 95th percentile of the resulting τ across all trials.

The resulting values for τ agree well with the direct numerical simulations (black symbols, Fig. 6). Importantly, using the transition matrix allows us to estimate the mean episode duration for system sizes where determination of mean episode duration with direct simulations is impossible. For example, directly simulating a single extinction event

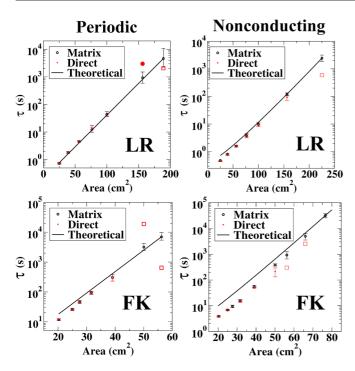


FIG. 6. Termination times as a function of the system size for the LR model (upper row) and the FK model (lower row). Red symbols show τ as a function of the area of the computational domain from direct simulations using domains with periodic boundary conditions (left column) and nonconducting boundaries (right column). Also shown are the results from the master equation approach (black symbols) and from the closed-form expression obtained using the WKB analysis (solid line). The red square represents the result of a single termination event computed using direct simulations while the solid red circle represents the computed time for a single computation that did not result in a termination.

on a domain with area $A = 225 \mu m^2$ and nonconducting boundaries was found to take approximately 100 hours of CPU time. Estimating τ from this single event is not useful as the error is large and generating a sufficient amount of termination events is not practical. Furthermore, for other larger domain sizes our direct simulations failed to produce a single termination event, even after 7 days of CPU time. Using the interpolated transition rates computed from this single, nonterminating event, however, we are still able to use the transition matrix (Fig. 6) to predict the mean episode duration. Moreover, τ can already be estimated using only a fraction of the data, and thus simulation time, further demonstrating the power of the approach. This is shown in Fig. 7 where we plot τ as a function of the fraction of computational data from a direct simulation of the LR model. Obviously, for larger fractions, the errors in the transition rates become smaller, resulting in smaller confidence intervals. Furthermore, the mean termination time converges as the fraction increases and can be reasonably well estimated from a small fraction of the entire dataset.

V. SCALING RESULTS

We can also use our stochastic analysis of termination to determine the scaling of τ with the area. For periodic

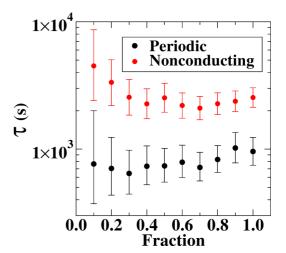


FIG. 7. Accuracy of the master equation approach. The mean episode duration computed using the master equation as a function of the fraction of computational data obtained using direct simulations of the LR model. Data segments of size indicated by fraction were started at random positions. Results are shown for $A = 225 \text{ cm}^2$ for the nonconducting case and for $A = 189.0625 \text{ cm}^2$ for the periodic case. The error bars are determined using bootstrapping and represent the 5% and 95% confidence interval.

boundary conditions, it is possible to obtain an analytical expression for τ [23]:

$$\tau(n_0) = \sum_{k=1}^{n_0/2} \phi[2(k-1)] \sum_{j=k}^{\infty} \frac{1}{\phi(2j)W_{+2}(2j)},$$
 (8)

where n_0 is the initial number of spiral tips, $\phi(k) = \prod_{i=1}^{k/2} W_{-2}(2i)/W_{+2}(2i)$, and $\phi(0) \equiv 1$. Using the numerically determined rates we find that τ quickly converges as n_0 becomes large and that $\tau(\bar{n})$ agrees well with the values obtained using the numerical methods. This is shown explicitly in Fig. 8 which plots the mean episode duration as a function of the initial number of tips for one particular domain size. We have verified that qualitatively similar results hold for other domain sizes.

To determine the scaling with the area we focus on the first term of this expression, $\tau(2) = \sum_{j=1}^{\infty} [\phi(2j)W_{+2}(2j)]^{-1}$. We can write $\phi(2j)$ as

$$\ln[\phi(2j)] = \sum_{z=1}^{j} \ln\left[\frac{W_{-2}(2z)}{W_{+2}(2z)}\right] \approx -\frac{A}{2} \int_{2/A}^{x} \ln\frac{w_{+2}(s)}{w_{-2}(s)} \, ds,$$

where we have used the fact that the transition rates scale with the area A and have defined s = 2z/A and x = 2j/A. As a result, the mean episode duration becomes

$$\tau \approx \int_0^\infty \frac{\exp\left[A \int_{2/A}^x \ln \sqrt{\frac{w_{+2}(s)}{w_{-2}(s)}} ds\right]}{2w_{+2}(x)} dx. \tag{10}$$

For large A, this integral will be sharply peaked around q^* . Thus, τ has the following scaling behavior [36]:

$$\tau \sim \exp\left[A \int_{2/A}^{q^*} \ln\sqrt{\frac{w_{+2}(s)}{w_{-2}(s)}} ds\right]$$
 (11)

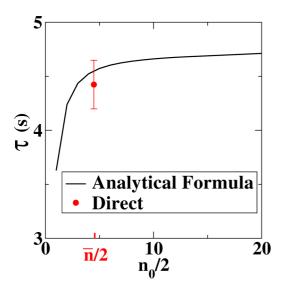


FIG. 8. Analytical formula for periodic boundary conditions. Mean episode duration τ as computed using the analytical formula in the main text as a function of the initial number of tips, n_0 . Symbol corresponds to the result from direct simulations ($\bar{n} \approx 9$). Results are shown for the LR, using a domain of size 7.5 cm \times 7.5 cm.

and, as an immediate consequence of the observed scaling of the transition rates, we find that τ scales exponentially with the area, consistent with our direct numerical results (Fig. 6).

We can also use approximation methods to determine the scaling of the mean episode duration by viewing the number of spiral tips as a stochastic population in a metastable state. This approach is particularly useful for domains containing nonconducting boundaries, for which it is no longer possible to derive an exact expression for τ . As long as A, equivalent to the total population size in models of population biology, is sufficiently large, we can use a dissipative WKB approximation, pioneered by Kubo $et\ al.\ [37]$. In this approximation, the quasistationary distribution is assumed to obey $P_{qs}(q) \sim e^{-AS(q)}$ where S(q) is a function called the action. We can now use our obtained scaling $W_r(n) = A^{r/2}w_r(q)$, together with the assumed form of $P_{qs}(q)$, and substitute them into the stationary form of Eq. (1). This equation is written in terms of the continuous rescaled variable q = n/A so that

 $n-r \rightarrow q-r/A$ [29]. For the periodic case, we take $S(q)=S_0(q)+O(A^{-1})$ while for absorbing boundaries, since the scaling of the ± 1 rates goes as \sqrt{A} while the ± 2 rates go as A, we use $S(q)=S_0(q)+A^{-1/2}S_1(q)+O(A^{-1})$. The resulting equation can then be expanded in terms of 1/A which yields, to O(1), a Hamilton-Jacobi equation

$$H(q, p) = \sum_{r} A^{r/2} w_r(q) (e^{rp} - 1) = 0,$$
 (12)

where $p(q) = \partial S/\partial q$ is the fluctuation momentum [28,29].

From the Hamiltonian H we can define the dynamics of p and q using $\frac{dp}{dt} = -\frac{\partial H}{\partial q}$ and $\frac{dq}{dt} = \frac{\partial H}{\partial p}$. The nontrivial solution of $H(q, p_a(q)) = 0$ corresponds to the activation trajectory in the q, p phase space [28–30,37,38]. This trajectory describes the most probable path along which the system evolves from the metastable state $(q^*, 0)$ to a point q in phase space. Since we are interested in extinction, we will consider the trajectory that connects $(q^*, 0)$ with [0, p(0)], the so-called "optimal" path to extinction [28]. This q, p phase space, along with the activation trajectory, is shown in Fig. 9 for periodic boundaries for both the LR and the FK model. The optimal path can be determined numerically but can also be determined using approximate closed-form relations. Specifically, for periodic boundary conditions, we find

$$S_0 = \int_{a^*}^{2/A} \ln \gamma_0 \, dq,\tag{13}$$

and $\gamma_0 = \sqrt{\frac{w_{-2}}{w_{+2}}}$. In Fig. 9, this corresponds to the area between the activation trajectory and the q axis, represented by the shaded part. Thus, we find that the mean episode duration scales as $\tau \sim e^{AS_0}$, consistent with Eq. (11). For absorbing boundaries, we can solve for S_1 perturbatively, yielding

$$S_1 = \int_{q^*}^{2/A} \frac{(\gamma_0 w_{+1} - w_{-1})(\gamma_0 - 1)}{2\gamma_0 (w_{+2} - w_{-2})} dq.$$
 (14)

Using our formulation, we can now compute the mean episode duration for any system size once the rates for a single domain \hat{A} are determined. Specifically, after computing for this particular domain both \hat{S}_0 and \hat{S}_1 and the corresponding mean episode duration $\hat{\tau}$, we can find τ as a function of the system size using

$$\tau(A) \approx \hat{\tau} e^{(A-\hat{A})\hat{S}_0 + (\sqrt{A} - \sqrt{\hat{A}})\hat{S}_1}.$$
 (15)

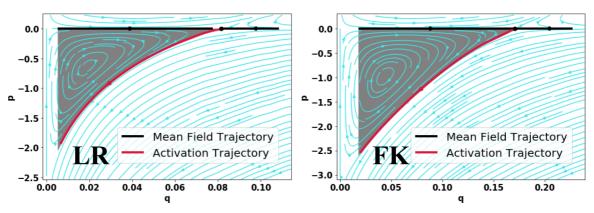


FIG. 9. WKB approach to spiral tip dynamics. Phase portrait of the Hamiltonian dynamics in q, p space for periodic boundary conditions in the LR and the FK model, showing the activation trajectory of the WKB Hamiltonian (red line). The shaded area represents the exponential factor S_0 .

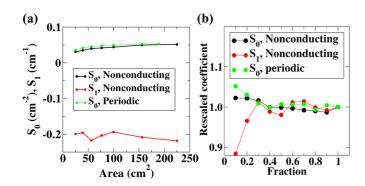


FIG. 10. WKB parameters as a function of area size for the LR model. (a) The exponential coefficients S_0 and S_1 as a function of domain size for periodic and nonconducting boundaries. (b) The exponential factors S_0 and S_1 , rescaled by their value computed at the largest computational data set, as a function of the fraction of computational data obtained using direct simulations. Data segments of size indicated by fraction were started at random positions. Shown are the results for $A = 225 \text{ cm}^2$ for the nonconducting case and for $A = 189.0625 \text{ cm}^2$ for the periodic case.

This scaling law for τ agrees well with the values of τ computed from the master equation, especially for larger values of A (Fig. 6), justifying the WKB approximation. For these larger domain sizes, the factors S_0 and S_1 converge, making the estimate from Eq. (15) to be more accurate, as shown in Fig. 10(a). This is consistent with the obtained quasistationary distributions which become more symmetric around their peak value for larger domain size (Fig. 5), rendering the WKB approximation more accurate. Furthermore, as is the case for τ (Fig. 7), both S_0 and S_1 can be estimated using the interpolated rates and only a fraction of the direct simulation data [Fig. 10(b)]. Thus, accurate estimates for arbitrary domain sizes do not require simulating actual termination events. Finally, the exponential scaling of τ with system size A reveals that, even though spiral wave driven fibrillation will always terminate, its mean episode duration depends critically on the size of the heart. Of course, this result is valid as long as the specifics of the model do not change. Other factors, including changes in electrophysiological parameters, can have an effect of mean termination duration. For large values of A, τ can be large while for very small values of A as found, for example, in rodents, the mean episode duration will be well below 1 s. These findings are fully consistent with the well-established critical mass hypothesis which posits that fibrillation only occurs in hearts of a minimum size [17,19,20].

VI. SUMMARY

In summary, we present a statistical approach to quantify spiral wave dynamics in spatially extended domains. This approach recasts the problem into a master equation, after which statistical physics methods can be employed. Our approach is valid for any model exhibiting SDC. Key in this approach are the transition rates, which were computed numerically from a limited set of direct simulations. Using a dissipative WKB approach, we find that the mean episode duration of SDC can be computed in minimal computational time. In addition, we show that this duration depends exponentially on domain size.

Our results should be generally applicable to any system exhibiting SDC. Here we have used electrophysiological models, motivated by the spiral wave dynamics observed during fibrillation when the heart's electrical activity becomes disorganized [12,39-41]. Clearly, our fibrillation model is an idealized and simplified version of the clinical reality since the heart is a heterogeneous three-dimensional object. Nevertheless, it is intriguing to note that the exponential dependence of the mean episode duration is consistent with the critical mass hypothesis which states that fibrillation requires a minimal organ size [17,19]. Furthermore, some of the simplifications of the current model might be overcome by future extensions. First, our approach should also be applicable to more complex, realistic geometries. Geometry data are routinely obtained in patients and electrophysiological models can be readily implemented using computational tools. This should allow us to compute rate equations in realistic geometries, after which we can use the same approach as detailed here. This extension might also be used to study the effect of different surgically created lesion sets and pharmacological interventions on the termination time and has the potential to be an important step toward determining optimal therapeutic interventions aimed at minimizing the duration of fibrillation episodes. Second, we may be able to extend the model to include tissue inhomogeneities. As long as the spiral wave is not trapped, one should be able to compute the creation and annihilation rates as carried out in this study. Third, the approach can be extended to include tissue with a nonzero thickness, appropriate for the ventricles and possibly for atrial tissue. To extend our approach to this type of problem we will need to track spiral wave tips on both surfaces. In addition, it would be interesting to further study the dependence of the rates on the number of tips. If, for example, rational functions for these rates can be derived, it should be possible to obtain analytical expressions for τ . Finally, it would be interesting to compare scaling of fibrillation in healthy and diseased hearts by simulating appropriate electrophysiological models.

ACKNOWLEDGMENTS

We thank Brian A. Camley and David A. Kessler for valuable suggestions. We gratefully acknowledge support from the National Institutes of Health (Grant No. R01 HL122384) and the American Heart Association (Grant No. 16PRE30930015).

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