# Percolation on correlated random networks

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We consider a class of random, weighted networks, obtained through a redefinition of patterns in an Hopfieldlike model, and, by performing percolation processes, we get information about topology and resilience properties of the networks themselves. Given the weighted nature of the graphs, different kinds of bond percolation can be studied: stochastic (deleting links randomly) and deterministic (deleting links based on rank weights), each mimicking a different physical process. The evolution of the network is accordingly different, as evidenced by the behavior of the largest component size and of the distribution of cluster sizes. In particular, we can derive that weak ties are crucial in order to maintain the graph connected and that, when they are the most prone to failure, the giant component typically shrinks without abruptly breaking apart; these results have been recently evidenced in several kinds of social networks.

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## I. INTRODUCTION

Network theory is a fundamental tool for the modern understanding of complex systems: By a simple graph representation, where the elementary units of a system become nodes and their mutual interactions become links, many properties about the structure and dynamics of the system itself can be inferred [1].

Recently, the characterization of network dynamics has become a central issue: Networks are intrinsically dynamic and continuously accommodate novel members, lose their original elements, and build, erase, and rearrange their links [2]. The structural reorganization of networks may arise due, e.g., to a change in the resources providing the energy to maintain their links, or to a large stress [3,4]. In this context percolation [5,6] constitutes a very interesting process able to mimic a failure of or a damage to links and nodes. Moreover, percolation represents one of the simplest example of dynamical process on a graph, exhibiting a phase transition [7], and, indeed, it has been mapped into several other critical phenomena; as well, applications in epidemiology, traffic models, and the analysis of technological networks resilience have been rigorously studied [8–11].

Here we apply percolation processes as a means in order to probe the topology and the resilience of a network itself. We especially focus on a class of stochastic, weighted networks G recently introduced in Refs. [12,13]. Such networks are generated by assigning to each node a set of attributes and by linking two nodes whenever the pertaining attributes are similar enough: the larger the similarity, the stronger the link. As shown in Refs. [12,13], the resulting class of (weighted) networks G exhibits interesting properties such as imitative interactions (by construction), degree-degree correlation, high transitivity (i.e., a large clustering coefficient), and a properly tunable topology through a parameter  $\theta$ , which controls the distribution of attributes. Therefore, such networks constitute an efficient tool to describe several different systems that belong to disparate contexts, ranging from biological networks [14,15], to technological structures [16], and to social organizations [17–19].

Now, since the graphs under investigation are weighted, we can perform different kinds of percolation processes: random (where links are deleted in a purely random fashion) and deterministic (where links are deleted in rank order from the weakest to strongest, or vice versa). We especially focus on graphs  $\mathcal{G}$  obtained for  $\theta = 0$  and 0.25, corresponding (in the limit of large size) to fully connected weighted networks and accounting for an "unbiased" and "biased" pattern distribution, respectively. First, we consider the relative size of the largest connected component S as a function of the fraction f of links left: Numerical data suggest that a "giant component" emerges when the fraction f approaches a "critical" value  $f_c$ , which is found to scale with the system size according to  $V^{-\nu}$ , where  $\nu$  depends on the kind of dilution. The latter also controls the sharpness of the percolation, as well as the distribution of cluster sizes, showing that, for biased pattern distributions, weak ties play a crucial role as they can be used to build up a spanning tree; conversely, strongest links are typically redundant, and as a result, if weak ties are the most prone to failure, the system will exhibit a poor resilience.

The paper is organized as follows: In Sec. II we describe the correlated random networks we are focusing on, as well as the percolation processes we perform. Then in Sec. III we present the basic probability relations concerning the coupling distribution, from which we can infer qualitative information on the properties of the percolation transition; these properties are confirmed by the results of the numerical analysis that is reported in the following sections. In particular, the behavior of the giant component is studied in Sec. IV, the distribution of cluster sizes is described in Sec. V, and the behavior of the clustering coefficient is examined in Sec. VI. An overall discussion about the results and the perspectives of our work is contained in Sec. VII. In Appendix A we show that the class of graphs that we consider displays dissortative mixing in a wide region of the parameter set, while in Appendix B we present some analytic results that are valid for the random percolation process on  $\mathcal{G}$ .

## **II. THE MODEL**

We first introduce the class of networks on which we focus our analysis, and later we describe the percolative processes we will perform on such networks.

#### A. Network generation

Recently, a new approach to generate correlated random networks has been introduced [12,13]; the approach is based on a simple shift  $[-1, +1] \rightarrow [0, +1]$  in the definition of patterns in an Hopfield-like model, and it allows us to generate a broad variety of different topologies ranging from fully connected to small world to extremely diluted.

More precisely, we consider a set of V nodes, each endowed with a set of L attributes encoded by a binary string  $\xi$ ; the ensemble of strings is extracted according to the probabilities  $P(\xi^{\mu} = 0) = (1 - a)/2$  and  $P(\xi^{\mu} = 1) = (1 + a)/2$ , where the fixed parameter a belongs to the interval [-1,1]. Then the coupling between two generic nodes i and j is given by the rule

$$J_{ij} = \sum_{\mu=1}^{L} \xi_i^{\mu} \xi_j^{\mu}.$$
 (1)

Therefore, the wider the overlap between nonnull entries, the larger the weight associated to the link, with  $J_{ij} \in [0, L]$ ; the extreme case  $J_{ij} = 0$  means that there exists no link between nodes *i* and *j*. The values taken by the strings components admit the following interpretation:  $\xi_i^{\mu} = 1$  means that agent *i* is endowed with the particular feature  $\mu$ , this feature can represent a biological trait or an individual attitude according to the considered system (the absence of this particular feature corresponds to  $\xi_i^{\mu} = 0$ ). Then Eq. (1) states that agents show homophily. For example, in social networks, people interact with others of similar age, income, race, etc.

As shown in Refs. [12,13], the way a node is connected to the network is sensitively affected by the number  $\rho$  of nonnull entries present in the pertaining string, that is, for the *i*th node,  $\rho_i = \sum_{\mu} \xi_i^{\mu}$  (notice that since  $\rho$  is Poissonian, its average is given by  $\bar{\rho} = L(1 + a)/2$ ). In fact, one finds that the average probability  $\bar{P}_{\text{link}}(\rho_i; a)$  that *i* is connected to another generic node reads as

$$\bar{P}_{\text{link}}(\rho_i;a) = 1 - \left(\frac{1-a}{2}\right)^{\rho_i}.$$

Moreover, by averaging over all possible string arrangements, one finds for the average link probability p between two generic nodes

$$p = 1 - \left[1 - \left(\frac{1+a}{2}\right)^2\right]^L$$

The class of networks that are generated in this way exhibit different levels of correlation. For instance, it is easy to see [12,13] that two neighbors of a given node are more likely to be connected than they would be if the graph was purely random

generated; this kind of transitivity also affects the weights associated with the links [12,13]. Such networks also display a dissortative behavior. Indeed, the nodes having strings with small  $\rho$  typically possess a small coordination number, and they are more likely to be linked with nodes with large  $\rho$ . The mathematical aspects of the degree correlations are elaborated in Appendix A.

Finally, we introduce the parameter  $\alpha = L/V$ , which turns out to crucially control not only the topology but also the thermodynamic of the system [12,13]. Here we assume  $\alpha$  to be constant and finite, which means that, as the volume of the system grows, the length of the string increases proportionally; this corresponds the the so-called high-storage regime in neural networks [20]. Interestingly, as  $V \rightarrow \infty$  there exists a vanishingly small range of values for *a* giving rise to a nontrivial graph; such a range can be recognized by the following scaling:

$$a = -1 + \frac{\gamma}{V^{\theta}},\tag{2}$$

where  $\theta \ge 0$  and  $\gamma$  is a finite parameter. As explained in Ref. [13],  $\theta$  controls the connectivity regime of the network, ranging from fully connected (FC;  $0 \le \theta < 1/2$ ) to extremely diluted  $(1/2 < \theta < 1)$  to completely disconnected  $(\theta > 1)$ , while  $\gamma$  allows a fine tuning. In particular, here we focus on  $\theta < 1/2$  and  $\gamma < 2$ , corresponding to a FC regime: In this case topological disorder is lost, while disorder on couplings is still present; however, notice that for  $\theta = 0$  and  $\gamma = 2$ , the coupling distribution gets peaked at J = L, and disorder on couplings is relaxed as well.

In the following, we will refer to the weighted random graph, generated as explained above, as  $\mathcal{G}(\alpha, \theta, \gamma, L)$ , hence highlighting the dependence on the set of parameters that control its size and its topology. We also anticipate that we will focus only on the cases  $\theta = 0$  and 0.25 corresponding to weighted, complete graphs. Of course, for these cases there is no topological correlation among links (the clustering coefficient is equal to 1 and assortativity is neutral), though correlation among link couplings is retained.

#### **B.** Percolation processes

Given an arbitrary graph, bond percolation consists in deleting the existing links with some probability 1 - for, in other terms, in occupying links with probability f; nodes connected together form clusters. When f exceeds a given system-dependent threshold (or critical) value  $f_c$ , a macroscopic cluster, i.e., a cluster occupying a finite fraction of all available sites, also called a giant component, is formed. For various network architectures and space dimensionalities this transition is typically continuous, or second order, as the system properties change continuously at the critical point [5,21]. For instance, on random networks à la Erdös-Rényi (ER) [22], one starts from a set of V nodes and adds links such that the probability f that two nodes are joined by a link is the same for all pairs of nodes. When f < 1/V, the largest component remains minuscule, its number of vertices scaling as log V; in contrast, if f > 1/V, there is a component of size linear in V. Thus, the fraction of vertices in the largest component undergoes a continuous phase transition at f = 1/V.

As explained before, in the graph under study quenched weights are assigned to the edges, and this allows us to think of different kinds of processes, each corresponding to different physical situations: The deletion of a link may mimic the failure of the link itself due to overload [4] or, rather, to error or attack that may affect randomly any link [23]. In the former case links with higher weight are the first to be deleted, while in the latter case deletion occurs randomly. In other kinds of situations we can think that nodes transfer a signal to neighbors, and the passage of information is effective only when the tie strength is larger than some noise level [24]. Therefore, as the level of noise grows, more and more links starting from the weakest ones get ineffective. To summarize, we deal with the following processes:

(1) Random percolation (RP): Starting from the original graph  $\mathcal{G}(\alpha, \theta, \gamma, L)$  we consider each link, and we remove it with probability 1 - f, *independently* of the couple of adjacent nodes, in such a way that f is the fraction of links left; as f is tuned from 0 to 1 we range from a completely disconnected graph to the original graph.

(2) Deterministic-Weak percolation (WP): Starting from  $\mathcal{G}(\alpha, \theta, \gamma, L)$ , we remove all links with weight *smaller* than a given threshold  $\iota$ ; that is to say, as  $\iota$  is tuned from 0 to L, we remove links in rank order from the weakest to strongest ties.

(3) Deterministic-Strong percolation (SP): Starting from  $\mathcal{G}(\alpha, \theta, \gamma, L)$ , we remove all links with weight *larger* than a given threshold  $\iota$ ; analogously to the previous case, this corresponds to remove links in rank order from the strongest to the weakest ties.

In order to evaluate the impact of removing ties, we measure the relative size of the largest connected component S, providing the fraction of nodes that can all reach each other through connected paths, as a function of the fraction f of links left f. We also measure the average squared size  $\bar{S} = \sum_{s=1}^{V} n_s s^2 / V$ , where  $n_s$  is the number of clusters containing s nodes. According to percolation theory, if the (infinite) network collapses because of a phase transition at  $f_c$ , then  $\bar{S}$  diverges as f approaches  $f_c^-$  [5,25].

### **III. COUPLING DISTRIBUTION**

The coupling distribution  $P_{\text{coupl}}(J; a, L)$  plays an important role as for deterministic processes, so that it is worth recalling some previous results [12] and deepening its dependence on the system parameters.

The probability for two strings  $\xi_i$  and  $\xi_j$  (with  $\rho_i$  and  $\rho_j$  nonnull entries) to be connected by a link with weight *J* is just the probability that the strings display *J* effective matchings; this has been found to be [12]

$$P_{\text{match}}(J;\rho_i,\rho_j,L) = \frac{\binom{L}{J}\binom{L-J}{\rho_i-J}\binom{L-\rho_i}{\rho_j-J}}{\binom{L}{\rho_i}\binom{L}{\rho_j}},$$
(3)

from which we can write that, in the average, the coupling distribution reads off as

$$P_{\text{coupl}}(J; a, L) = \sum_{\rho_i=0}^{L} \sum_{\rho_j=0}^{L} P_{\text{match}}(J; \rho_i, \rho_j, L) \\ \times P_1(\rho_i; a, L) P_1(\rho_j; a, L),$$
(4)

being  $P_1(\rho; a, L) = {L \choose \rho} [(1 + a)/2]^{\rho} [(1 - a)/2]^{L-\rho}$  the probability that a given string displays  $\rho$  nonnull entries. Therefore, we get

$$P_{\text{coupl}}(J; a, L) = {\binom{L}{J}} (\tilde{a} + 1)^{-2L} \sum_{\rho_i=0}^{L} \sum_{\rho_j=0}^{L} {\binom{L-J}{\rho_i-J}} \times {\binom{L-\rho_i}{\rho_2-J}} \tilde{a}^{\rho_i+\rho_j},$$
(5)

where we called  $\tilde{a} = (1 + a)/(1 - a)$ . Since we are focusing on the case  $L = \alpha V$ , with a string bias *a* given by Eq. (2), it is convenient to rewrite the coupling distribution as a function of the effective parameters, namely,

$$P_{\text{coupl}}(J; \alpha, \theta, \gamma, L)$$

$$= \binom{L}{J} \left[ 1 - \frac{\gamma}{2(L/\alpha)^{\theta}} \right]^{2L} \sum_{\rho_i=0}^{L} \sum_{\rho_j=0}^{L} \binom{L-J}{\rho_i - J} \binom{L-\rho_i}{\rho_j - J} \times \left[ \frac{\gamma}{2(L/\alpha)^{\theta} - \gamma} \right]^{\rho_i + \rho_j}.$$
(6)

The previous expression shows that for systems large enough, and  $\alpha$  and L fixed, the distribution gets peaked at smaller Jas  $\theta$  is increased (when  $\theta > 0.5$  only couplings with value 0 or 1 display nonvanishing probability), and the same holds for fluctuations. Moreover, a link is absent with probability  $P_{\text{coupl}}(0; \alpha, \theta, \gamma, L) = [1 - \gamma^2/4(\alpha/L)^{2\theta}]^L$ , which decays to zero for  $\theta < 0.5$ .

In particular, in the following analysis we assume  $\gamma = 1$ ,  $\alpha = 0.1$ , and  $\theta = 0$  or 0.25; for  $\theta = 0$  we can write explicitly  $P_{\text{court}}(J; 0, 1, 0, 1, L)$ 

$$= \frac{L!}{J!} 2^{-2L} \sum_{\rho_i=0}^{L} \sum_{\rho_j=0}^{L} \frac{1}{(\rho_i - J)!(\rho_j - J)!(L - \rho_i - \rho_j + J)!},$$
(7)

and similarly for the latter. In Fig. 1 we show a comparison of the two cases where numerical data are fitted with curves given by Eq. (6). Data corroborate that, even at relatively small sizes,



FIG. 1. (Color online) Coupling distributions  $P_{\text{coupl}}(J; \alpha, \theta, \gamma, L)$  from different values of *L* (depicted in different colors, as shown by the legend) and for  $\gamma = 1$ ,  $\alpha = 0.1$ ,  $\theta = 0$  (left panel), or  $\theta = 0.25$  (right panel). Curves represent Eq. (6).



FIG. 2. (Color online) Log-log scale plot of the average coupling  $\overline{J}(\alpha, \theta, \gamma, L)$  (main figure) and its fluctuations  $\Delta J(\alpha, \theta, \gamma, L)$  (inset), for  $\theta = 0$  ( $\bigcirc$ ) and  $\theta = 0.25$  ( $\Box$ ), as shown in the legend. Symbols represent numerical data, while curves represent analytical estimates from Eqs. (8) and (9), respectively.

the analytical formula above provides a good approximation and that the distribution gets broader for larger *L* and smaller  $\theta$ . More precisely, we calculate the average coupling  $\bar{J}(\alpha, \theta, \gamma, L)$ and its fluctuations  $\Delta J[(\alpha, \theta, \gamma, L)]$  as

$$\bar{J}(\alpha,\theta,\gamma,L) = \sum_{J=0}^{L} J P_{\text{coupl}}(J;\alpha,\theta,\gamma,L) = \frac{\gamma^2}{4} \frac{L}{(L/\alpha)^{2\theta}},$$
 (8)

$$\Delta J(\alpha, \theta, \gamma, L) = \sum_{I=0}^{L} (J - \bar{J})^2 P_{\text{coupl}}(J; \alpha, \theta, \gamma, L), \quad (9)$$

where for the closed-form expression in Eq. (8) we used Eq. (6). Relevant results are shown in Fig. 2, where, again, the comparison between analytical estimates and numerical data is successful.

Interestingly, from the width of the distribution one can infer information about the sharpness of the deterministic percolation: A broader distribution is expected to give rise to a less sharp transition. Moreover, we notice that the case  $\gamma = 1$  and  $\theta = 0$  corresponds to a = 0; namely, it corresponds to an unbiased distribution for strings, and this yields a rather symmetric coupling distribution: As a consequence, SP and WP are expected to behave similarly. Conversely, when the coupling distribution is not symmetric, as for  $\theta = 0.25$ , different behaviors emerge. All these points are deepened in the next section.

Finally, we notice that for  $\theta = 0.25$  relatively small sizes give rise to nonfully connected structures; that is, the coupling probability is nonnull for J = 0. As we derived from Eq. (6), the probability that a link is absent decreases slowly with the size, and such a finite-size effect gets negligible only for  $V \sim 10^5$ . Indeed, we find that finite-size effects enhance the skewness positivity of the distribution.

## **IV. PERCOLATION TRANSITIONS**

In the following analysis we generate the graph  $\mathcal{G}(\alpha, \theta, \gamma, L)$ , and, while performing a dilution process (either



FIG. 3. (Color online) Main figure: Relative size of the largest connected component *S* vs the fraction of links left *f* for a system of size V = 5700 and  $\theta = 0$ . Fluctuations on these data, obtained by averaging over several realizations of the structure, are approximately 4%. Inset:  $f_c$  vs system size; symbols represent numerical data, while curves represent the best fit given by a power law with exponents  $\nu_{\text{RP}} \approx 1$  and  $\nu_{\text{WP}} \approx \nu_{\text{SP}} \approx 0.5$ . Different percolation processes are compared as shown by the legend.

deterministic or random), we measure the number of clusters and their size; such results are then averaged over  $10^2$  realizations of  $\mathcal{G}(\alpha, \theta, \gamma, L)$  in order to account for the stochasticity of the graph itself. As explained in the previous section, in the thermodynamic limit both  $\theta = 0$  and 0.25 give rise to fully connected structures, so that, for large enough sizes, the random percolation process recovers the well-known results holding for ER graphs [26].

In order to evaluate the impact of removing ties, we measure the relative size of the largest connected component S as a function of the fraction of links left f. Results obtained for  $\theta = 0$  and 0.25 are shown in Figs. 3 and 4, respectively.

Let us comment on the results of Fig. 3. First, we notice that when weak links are deleted first the graph starts to be disconnected (S < 1) at a value of f that is rather large; that is, weak ties are crucial to maintain the overall connection of the graph, and, in this sense, they work as bridges. Moreover, the WP transition is smoother than the one obtained from a random deletion of edges. This suggests that the deletion of weak ties vields the disconnection, from the giant component, of single nodes (indeed, those displaying small  $\rho$ ) or of small clusters. Otherwise stated, as f is increased from 0 to 1, we first connect nodes displaying large overlap, hence forming a strong main component, while nodes with small  $\rho$  are likely to remain isolated or to form small clusters, which are successively annexed to the giant component: The process is therefore quite gentle. On the other hand, when links are introduced randomly, clusters grow up in a more uniform way, so that links merging disjoint components can give rise to a faster increase in the size of the giant component. As for the SP process, when  $\theta = 0$  and  $\gamma = 1$ , strings are homogeneously distributed (a = 0), so that, as mentioned above, no qualitative differences are expected between SP and WP; in particular, in this peculiar case (a = 0) strong links also turn out to be



FIG. 4. (Color online) Main figure: Relative size of the largest connected component *S* vs the fraction of links left *f* for a system of size V = 5700 and  $\theta = 0.25$ . Fluctuations on these data, obtained by averaging over several realizations of the structure, are approximately 4%. Different percolation processes are compared (the legend is the same as in Fig. 3). The inset on the right highlights the case of SP, where a jump in *S* occurs around  $f \sim 0.1$ ; notice that this jump results from the fact that *f* cannot be continuously tuned (see also Fig. 1). Inset on the left:  $f_c$  vs system size; symbols represent numerical data, while curves represent the best fit given by a power law with, in particular,  $v_{RP} \approx 1$ .

crucial in maintaining the graph connected: Being  $\tilde{f}$  the largest fraction of links for which S < 1, we get  $\tilde{f}_{RP} < \tilde{f}_{SP} < \tilde{f}_{WP}$ .

Let us now consider results for  $\theta = 0.25$  shown in Fig. 4. As for the RP, slight quantitative changes in S(f) with respect to the previous case are due to finite-size effects, while deterministic processes (DPs) are also affected by the positive skewness of the coupling distribution. More precisely, strings now display only rare nonnull entries so that small components (typically made up of very close or even identical strings) can arise during a WP (this explains the sharper transition); also, strong ties are rather unlikely, and their deletion does not modify the connection of the giant component so that we can derive that they are redundant; that is, they typically do not participate to the spanning tree. For the sizes considered here we now have  $\tilde{f}_{SP} < \tilde{f}_{RP} < \tilde{f}_{WP}$ .

As anticipated, the occurrence of a percolation transition is envisaged by a singularity in the average squared size  $\bar{S}$  at  $f_c$ , due to a network collapse as f approaches  $f_c$ . We get consistent estimates for  $f_c$  by evaluating the value of f corresponding to a maximum in the derivative of S and to the singularity in  $\overline{S}$ . Results for  $\theta = 0$  and 0.25 are shown in the insets of Figs. 3 and 4, respectively; the sets of values for  $f_c$  have also been fitted with power-law functions. Of course, for the RP the expected exponent  $\nu = -1$  is recovered [5], while for WP and SP when  $\theta = 0$  we find comparable exponents  $\approx 0.5$ ; for  $\theta = 0.25$  finite-size effects prevent to get sound estimates for  $\nu$ , although a different behavior of SP with respect to the other cases is apparent. In any case, the RP corresponds to smaller values of  $f_c$ , meaning that a smaller fraction of links is necessary to carry a giant component; that is, a smaller degree of redundancy is retained.

Finally, some analytical insights for the case of arbitrary  $\theta$  and random dilution are presented in the Appendix B, where we show consistency with known results about noncorrelated networks.

### V. CLUSTER DISTRIBUTIONS

While previous results offer a global description of the network dynamics, in this section we focus on the evolution of the internal organization of clusters by measuring the distribution  $N(\rho, s)$ , representing the number of nodes corresponding to a string with  $\rho$  nonnull entries and belonging to a cluster of size s.

In Fig. 5 we show three sets of snapshots of  $N(\rho,s)$ for the case  $\theta = 0$ ; each row represents a different dilution process (from top to bottom: RP, WP, SP), while each column represents a different regime (from left to right:  $f < f_c$ ,  $f \approx$  $f_c$ ,  $f > f_c$ ). While for the random dilution intermediate- and high-dilution regimes (panels b and c) are characterized by the existence of several clusters of different sizes, for deterministic dilution (panel e, f and h, i) a node basically either belongs to the largest component or is isolated. Moreover, in the former case due to the homogeneity underlying the process, curves are all peaked at around  $\bar{\rho}$ ; namely, the set of attributes characterizing a given node does not affect the cluster size the node belongs to. Conversely, for WP (SP), larger (smaller) values of s yield distributions  $N(\rho, s)$  peaked at larger values for  $\rho$ . In particular, when weak ties are the most prone to failure, nodes displaying strings with large  $\rho$  are the most likely to belong to the giant component.

Analogous results for the case  $\theta = 0.25$  are depicted in Fig. 6. Now, in the intermediate regime, also for the WP a few small clusters emerge, while the abrupt jump in S(f) evidenced for the SP is recovered here by the fact that *s* assumes only two values: either 1 or *V*.

We conclude this section with a remark. By focusing only on the topology of the graphs  $\mathcal{G}(\alpha, \theta, \gamma, L)_{f, P}$ , for a fixed parameter set and fixed f, we can compare the level of organization of the graph resulting from a different percolative process P. This can be attained by means of entropy measures [27], which, given a particular ensemble, provide the normalized logarithm of the number of networks in that ensemble, hence estimating how effective the features characterizing the ensemble are. Here we can fix the parameter set  $(a, \theta, \gamma, L)$  and f (this somehow fixes the "energy" of the system) and measure the entropy within a configuration approach, namely, working out the degree sequences; this approach is also related to a hidden variable model [13,27], consistently with the assignation of attributes. For instance, for  $\theta = 0$ , when the dilution is low (most links still present) the entropy of the ensemble WP is expected to be larger due to the presence of a few isolated nodes that yield a larger number of configurations; vice versa, when the dilution is increased the entropy of the ensemble RP is expected to prevail. Further analysis on this point may lead to speculate that a failure of a limited number of nodes is likely to involve only weak ties, while when the failure is wider it is more likely to involve any generic node. Similar reasoning can be extended in order to account also from an energy contribution due to the coupling.



FIG. 5. (Color online) Number of nodes  $N(\rho, s)$  corresponding to a string with  $\rho$  nonnull entries and belonging to a cluster of size *s*, plotted as a function of  $\rho$ , while different values of *s* and shown in different colors. The system considered has size V = 17500,  $\theta = 0$ ,  $\alpha = 0.1$ , and  $\gamma = 1$ . The nine panels are arranged in such a way that each row represents a different dilution process (from top to bottom: RP, WP, SP), while each column represents a different regime (from left to right:  $f < f_c$ ,  $f \approx f_c$ ,  $f > f_c$ ).

### VI. CLUSTERING AND CORRELATIONS

In this section we want to focus the attention on the properties of clustering and of correlation among links as we dilute them.

Before proceeding it is worth recalling that the clustering coefficient C provides a measure of the transitivity of the graph, and it can be calculated as the average over nodes i of the local clustering coefficient  $C_i$ , defined as the actual number of links between the vertices within the neighborhood of i, divided by the maximum number of links that can exist between them, that is,

$$C = \frac{1}{V} \sum_{i=1}^{V} C_i = \frac{1}{V} \sum_{i=1}^{V} \frac{2E_i}{z_i(z_i - 1)},$$
 (10)

where  $E_i$  is the number of links among nodes that are connected to *i* (node *i* is not included),  $z_i$  is the number of neighbors (also called degree) of *i*, and one conventionally sets  $C_i = 0$  for  $z_i = 0, 1$ . A graph is often referred to as small-world, if its diameter is small (scaling as log *V*, which is is verified by  $\mathcal{G}$ ), and its average clustering coefficient is significantly higher than the one relevant to a random graph constructed on the same vertex set, meaning C = f. As evidenced in Refs. [12,13], the graph under study can be defined as small-world.

As shown in Fig. 7 (upper panel), when  $\theta = 0$ , the clustering coefficient relevant to the graph obtained with a random dilution just corresponds, as expected, to the coefficient pertaining to an analogous ER graph; conversely, the clustering coefficients for deterministic dilution are larger and display a less trivial profile. In fact, for the RP, starting from small f we first build up a set of uncorrelated small components having zero or very small coefficient, so that their contribution to C is negligible; a significant and regular increase in C is only set up from the percolation threshold (dashed line in the figures).

As for the WP, starting from small f we first connect nodes having strings with large  $\rho$ , and these form a highly clustered component which already contributes to C; as f is increased the largest component gradually expands and C consistently grows; when  $f \gtrsim f_c$  most nodes are connected and new links



FIG. 6. (Color online) Number of nodes  $N(\rho, s)$  corresponding to a string with  $\rho$  nonnull entries and belonging to a cluster of size *s*, plotted as a function of  $\rho$ , while different values of *s* and shown in different colors. The system considered has size V = 17500,  $\theta = 0.25$ ,  $\alpha = 0.1$  and  $\gamma = 1$ . The arrangement of panels is the same as in the previous figure.

serve to connect low-degree nodes so that the rate of growth of C is reduced; finally, when the network is connected (no isolated nodes) all new links determine an improvement in the clustering so that there is an acceleration in the growth of C.

Similar arguments apply for the SP percolation: For small f only very weak ties are introduced, and these, due to the relatively homogeneity of the graph (a = 0), are sufficient to build up a structures component that progressively grows determining a larger and larger C.

In Fig. 8 we show results for the case  $\theta = 0.25$ ; due to the finite-size effects affecting SP, we focus just on the cases RP and WP. Of course, for the RP no qualitative changes are evidenced with respect to the case  $\theta = 0$ , while for the WP, as f is increased, detached small clusters are now more likely to occur due to the sparsity of nonnull entries in strings, and this explains the fact that C is now qualitatively comparable with the ER case.

Another interesting coefficient that we introduce, in order to monitor the evolving topology as links are removed, is a slightly modified version of the clustering coefficient, which we denote as  $\tilde{C}$  and define as the average of the local dilution  $\tilde{C}_i$ , given by the fraction of links within the subgraph containing the node *i* and all its neighbors, that is,

$$\tilde{C} = \frac{1}{V} \sum_{i=1}^{V} \tilde{C}_i = \frac{1}{V} \sum_{i=1}^{V} \frac{2\tilde{E}_i}{z_i(z_i+1)},$$
(11)

where  $\tilde{E}_i$  is the number of links connecting any couple of nodes belonging to the subgraph, including *i* itself, and one conventionally sets  $\tilde{C} = 0$  for  $z_i = 0$ . We remark that  $\tilde{C}_i$  differs from  $C_i$  by the fact that here we count all links within the neighborhood of *i*, including those stemming from *i*, i.e.,  $\tilde{E}_i =$  $E_i + z_i$ , from which  $\tilde{C}_i = C_i + \sum_i 2/(z_i + 1)$ . Hence, once *f* fixed, by comparing *C* and  $\tilde{C}$ , one can derive information about the arrangement of existing links, either highly clusterized, so to form a small-sized connected component with nodes having relatively large degree (comparable *C* and  $\tilde{C}$ ) or highly scattered, so to eventually form a large connected component with nodes having relatively small degree (large  $\tilde{C}$ , small *C*).

As shown in Fig. 7, again for  $\theta = 0$  qualitative differences emerge between deterministic and random dilution. For deterministic percolations  $\tilde{C}$  follows a behavior similar to *C*: This results from the fact that, basically, we have only  $-C_{\rm RP}$  $-C_{\rm WP}$ 

 $C_{SP}$ 

O 0.5





FIG. 7. (Color online) Clustering coefficients *C* (top panel) and  $\tilde{C}$  (bottom panel) as a function of *f* and for  $\theta = 0$ , V = 5700; different percolative processes are compared as explained by the legend. The vertical dashed and dotted lines highlight the percolation threshold for RP and WP, respectively.

one component, which keeps on growing as f is increased, the remaining nodes being mainly isolated. Vice versa, for random percolation we can distinguish three different regimes, demarcated by two extremal points: In the first regime we have the emergence of several small components (e.g., dimers, trimers), each with small but nonnull contributions to  $\tilde{C}$ ; in the second regime such small components start to merge, and this yields a reduction in  $\tilde{C}$ ; finally, when the graph has reached a connected status, increasing the number of links can just produce an increase in  $\tilde{C}$ .

For  $\theta = 0.25$ , also the WP transition display the same multiregime behavior, which confirms the picture above.



FIG. 8. (Color online) Clustering coefficients *C* (top panel) and  $\tilde{C}$  (bottom panel) as a function of *f* and for  $\theta = 0.25$ , V = 5700; different percolative processes are compared as explained by the legend. The vertical dashed and dotted lines highlight the percolation threshold for RP and WP, respectively.

#### VII. CONCLUSIONS AND PERSPECTIVES

In this work we analyzed the evolution, under percolation process, of a class of weighted graphs  $\mathcal{G}$  introduced in Refs. [12,13], whose topological properties arise from imitative interaction among nodes and can be properly varied by tuning the parameters  $(\alpha, \theta, \gamma, L)$ . In particular, here we fixed  $\alpha = 0.1$ ,  $\gamma = 1$ , and  $\theta = 0$  or 0.25, while the size  $(V = L/\alpha)$ is the number of nodes) is varied; such a situation corresponds, for large enough volumes, to fully connected graphs, still retaining a nontrivial distribution  $P_{\text{coupl}}(J; \alpha, \theta, \gamma, L)$  for the coupling strength J associated to any link. This allows to perform and compare different percolation processes: random (RP, where links are randomly extracted for deletion), deterministic-weak (WP, where links are deleted starting from the weakest ones), and deterministic-strong (SP, where links are deleted starting from the strongest ones).

Our results highlight that weak ties are the most crucial in order to ensure the overall connection of the system; that is, the size of the largest component starts to be smaller than V when only few (weak) links are deleted. When  $\theta$  approaches value 0.5 from below (and in the presence of finite-size effects, which affect the skewness of the coupling distribution) one can see that the spanning tree underlying  $\mathcal{G}$  is mostly made up of weak links, while strong linkes are unlikely and mainly redundant. Hence the robustness of  $\mathcal{G}$  sensitively depends on which ties are the most prone to failure. The fact that weak ties are fundamental to maintain the whole graph connected is consistent with the so-called theory of weak ties [28,29], according to which, in social systems, weak ties work as bridges between different subcommunities.

Moreover, we showed that removing in rank order, from the weakest to the strongest ties, shrinks the network, but does not precipitously break it apart, in such a way that the percolation is rather smooth. A similar phenomenon has been evidenced in the context of social networks where, when all "declared friendships" are considered, the graph is highly connected, but when only "strong" links are retained, selecting, first, "maintained" relationships and, second, "mutual" relationships, nodes get gradually disconnected forming only small subclusters [19]. Conversely, as we underlined, RP gives rise to more structures subclusters while diluting.

A possible extension of this work could consider noncomplete graphs ( $\theta > 0.5$ ) with random deletion of nodes so to evaluate whether also for such correlated networks, degree-degree correlation yields qualitative changes in the percolation behavior as expected from Ref. [30].

Analysis similar to those performed here can involve different connecting rules [see Eq. (1) and Ref. [24]] in order to figure out a possible relation between the kind of interaction (e.g., imitative or anti-imitative) and the dynamic behavior. Also, a possible mapping between the dilution obtained via cutting a fraction 1 - f of links and via a progressive reduction of the parameter *a* may be figured out.

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### **APPENDIX A: DEGREE-DEGREE CORRELATION**

In this Appendix we aim to show that overpercolated networks  $\mathcal{G}$  display negative assortativity by calculating how the average degree of nodes belonging to the neighborhood of *i* depends on the degree of *i* itself. In fact, we can write that in a graph  $\mathcal{G}(\alpha, \theta, \gamma, L)$  the probability for a node *i* to have neighbors that display on the average *z* neighbors is

$$P_{\text{deg-deg}}(z;\rho_i,a,L) = \frac{1}{N} \sum_{\rho_j=1}^{L} P_{\text{link}}(\rho_i,\rho_j,L) P_{\text{deg}}(z;\rho_j,a,V) P_1(\rho_j;a,L),$$

where  $\mathcal{N}$  is the normalization factor,  $P_{\text{link}}(\rho_i, \rho_j, L) = 1 - P_{\text{match}}(0; \rho_i, \rho_j, L)$  is the probability that there exists a link connecting *i* and *j*, and  $P_{\text{deg}}(z; \rho_j, a, V) = {V \choose z} [\bar{P}_{\text{link}}(\rho_j; a)]^z [1 - \bar{P}_{\text{link}}(\rho_j; a)]^{V-z}$  is the probability that node *j* has *z* neighbors. Hence one finds that the average degree for *i*'s neighbors is

$$\begin{split} \tilde{z}(\rho_i; a, L) &= \sum_{z=0}^{V} P_{\text{deg-deg}}(z; \rho_i, a, L) z \\ &= \frac{1}{\mathcal{N}} \sum_{\rho_j=1}^{L} P_{\text{link}}(\rho_i, \rho_j, L) P_1(\rho_j; a, L) \bar{z}(\rho_j; a, L, V), \end{split}$$

being  $\bar{z}(\rho_j; a, V) = V\{1 - [(1 - a)/2]^{\rho_j}\}$  the average degree for node *j*. With some algebra one gets to

$$\tilde{z}(\rho_i; a, L) = 1 - \left(\frac{1-a}{2}\right)^L \left(\frac{3+a}{2}\right)^L \times \left[1 - \left(\frac{2}{3+a}\right)^{\rho_i}\right] \left[1 - \left(\frac{1-a}{2}\right)^{\rho_i}\right]^{-1}.$$
 (A1)

Now, noticing that 0 < (1 - a)/2 < 2/(3 + a) < 1, we can deduce that  $\tilde{z}(\rho_i; a, L, V)$  is decreasing with  $\rho_i$ , namely, with  $\bar{z}(\rho_i; a, L, V)$ , so that, as long as the mean-field approach developed here is valid [12,13], the graph displays dissortativity.

#### APPENDIX B: ANALYTICAL RESULTS ON RP

The percolation problem has been studied over different kinds of structure, both analytically and numerically [31,32]; in particular, within the so-called configuration model approach [33,34], we can exploit the generating function formalism to get some insights into the problem. First, being  $\bar{P}_{degree}(k)$  the average degree distribution for the generic graph  $\mathcal{G}$  (here we drop the dependence on the parameter set to lighten the notation), we define

$$G_0(x) = \sum_{k=0}^{\infty} \bar{P}_{degree}(k) x^k, \ G_1(x) = \sum_{k=0}^{\infty} Q(k) x^k,$$
 (B1)

where  $Q(k) = (k + 1)\overline{P}_{degree}(k + 1)/\overline{z}$ . Now, assuming V large and the clustering not significant, Q(k) is the so-called excess degree [34], representing the degree distribution of the vertex at the end of a randomly chosen edge; notice that  $G'_0(1) = \overline{z}$  and  $G_1(x) = G'_0(x)/\overline{z}$ . Recalling that  $\overline{P}_{degree}(k) =$ 

$$\sum_{\rho=0}^{L} {\binom{V}{k}} [1 - (\frac{1-a}{2})^{\rho}]^{k} (\frac{1-a}{2})^{\rho(V-k)} P_{1}(\rho; a, L)$$
 (see Refs. [12,13]), we can write

$$G_0(x) = \sum_{\rho=0}^{L} P_1(\rho; a, L) \left[ x + (1-x) \left( \frac{1-a}{2} \right)^{\rho} \right]^V.$$

Moreover, for uniform link deletion probability, the mean cluster is [35]

$$\bar{s} = 1 + fG'_0(1) + \frac{f^2G'_0(1)}{1 - fG'_1(1)},$$
 (B2)

which diverges when  $1 - f G'_1(1) = 0$ ; this point marks the percolation threshold of the system: for  $f > f_c = 1/G'_1(1)$  a giant component of connected vertices is established. Therefore, consistently with the Molloy-Reed criterion [36], when  $G'_1(1) < 1$  the graph consists of many small components, while when  $G'_1(1) > 1$  a giant component can emerge. Here we find

$$G_1'(1) = \frac{G_0''(1)}{\bar{z}} = \frac{\bar{z}}{[1 - h(a)]^2} [1 - 2h(a) + g(a)], \quad (B3)$$

where  $h(a) = [(3 - 2a - a^2)/4]^L = p$  and  $g(a) = [(1 - a)(5 - a^2)/8]^L$ . Assuming  $a = -1 + \gamma(\alpha/L)^{\theta}$  and posing  $\tilde{\gamma} = \gamma(\alpha/L)^{\theta}/2$ , we can write

$$G'_{1}(1) = \frac{\bar{z}}{[1 - (1 - \tilde{\gamma}^{2})^{L}]^{2}} [1 - 2(1 - \tilde{\gamma}^{2})^{L} + (1 - 2\tilde{\gamma}^{2} + \tilde{\gamma}^{3})^{L}] \underset{L \to \infty}{\to} \bar{z},$$
(B4)

in analogy with the percolation threshold expected for the ER graph. In Fig. 9 we show  $G'_1(1)$  as a function of  $\gamma$  and *a* and for a finite value of *L*.



FIG. 9. (Color online) Natural logarithm of  $G'_1(1)$  for L = 20,  $\theta = 0.5$  and different values of  $\gamma$  and  $\alpha$  as given by Eq. (B4). Whenever G'(1) > 1 a giant component emerges.

Finally, the formalism developed in the first part of this appendix has also been used to find an expression for the global clustering coefficient or transitivity of the configuration model [37]:

$$c \equiv \frac{3N_{\triangle}}{N_3} = \frac{\bar{z}}{V} \left(\frac{\bar{z}^2 - \bar{z}}{\bar{z}^2}\right)^2 = \frac{1}{Vz} \left[\frac{G_0'(1)G_1'(1)}{z}\right]^2, \quad (B5)$$

where  $N_{\Delta}$  is the number of triangles in the network and  $N_3$  is the number of connected triples of vertices [34]. Notice that from Eq. (10), *c* is given by the coefficient expected

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for the ER graph, namely,  $\bar{z}/V$ , times an extra factor such that when the degree distribution is highly skewed, given that the factor  $\bar{z}^2/\bar{z}^2$  can be rather large, *c* is not necessarily negligible for the graph sizes relatively large. Interestingly, we find

$$c = \frac{[1 - 2h(a) + g(a)^2]^2}{[1 - h(a)]^3},$$
 (B6)

which, for  $a \in [-1,1]$  is always larger than p = 1 - h(a), hence confirming the large degree of cliquishness of  $\mathcal{G}$ .

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