

Non-Markovian persistence and nonequilibrium critical dynamics

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The persistence exponent θ for the global order parameter $M(t)$ of a system quenched from the disordered phase to its critical point describes the probability, $p(t) \sim t^{-\theta}$, that $M(t)$ does not change sign in the time interval t following the quench. We calculate θ to $O(\epsilon^2)$ for model A of Hohenberg and Halperin [Rev. Mod. Phys. **49**, 435 (1977)] (and to order ϵ for model C) and show that at this order $M(t)$ is a non-Markov process. Consequently, to our knowledge, θ is a new exponent. The calculation is performed by expanding around a Markov process, using a simplified version of the perturbation theory recently introduced by Majumdar and Sire [Phys. Rev. Lett. **77**, 1420 (1996)]. [S1063-651X(97)50707-0]

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The ‘‘persistence exponent’’ θ , which characterizes the decay of the probability that a stochastic variable exceeds a threshold value (typically its mean value) throughout a time interval, has attracted a great deal of recent interest [1–11]. One of the most surprising properties of this exponent is that its value is highly nontrivial even in simple systems. For example, θ is irrational for the $q > 2$ Potts model in one dimension [6] (where the fraction of spins that have not changed their state in the time t after a quench to $T=0$ decays as $t^{-\theta}$) and is apparently not a simple fraction for the diffusion equation [9,10] (where the fraction of space where the diffusion field has always exceeded its mean decays as $t^{-\theta}$).

A recent study of nonequilibrium model A critical dynamics [12], where a system coarsens at its critical point starting from a disordered initial condition, looked at the probability $P(t_1, t_2)$ that the global magnetization does not change sign during the interval $t_1 < t < t_2$ [11]. The persistence exponent for this system is defined by $P(t_1, t_2) \sim (t_1/t_2)^\theta$ in the limit $t_2/t_1 \rightarrow \infty$. Explicit results were obtained for the one-dimensional (1D) Ising model, the $n \rightarrow \infty$ limit of the $O(n)$ model, and to order $\epsilon = 4 - d$ near dimension $d = 4$. For these systems it was found that the value of θ was related to the dynamic critical exponent z , the static critical exponent η , and ‘‘nonequilibrium’’ exponent λ [which describes the decay of the autocorrelation with the initial condition, $\langle \phi(\mathbf{x}, t) \phi(\mathbf{x}, 0) \rangle \sim t^{-\lambda/z}$] by the scaling relation $\theta z = \lambda - d + 1 - \eta/2$. This relation may be derived from the assumption that the dynamics is Markovian, which is indeed the case for all of the cases considered in that paper.

From a consideration of the structure of the diagrams that appear at order ϵ^2 (and higher order), however, it was argued that the dynamics of the global order parameter should not be Markovian to all orders, implying that the exponent θ does not obey exactly that ‘‘Markovian scaling relation’’ [11]. Thus, to our knowledge, θ is a new exponent. Monte Carlo simulations in two dimensions indeed suggest weak violation of the Markov scaling relation [11].

In this paper we present an explicit calculation of the non-Markovian properties of the global order parameter. The nonequilibrium magnetization-magnetization correlation

function is calculated to order ϵ^2 , and this is then used to calculate θ to the same order, using a perturbative method proposed by Majumdar and Sire (MS) [8], valid in the vicinity of a Markov process. The Markov scaling relation is shown explicitly to be violated at order ϵ^2 , supporting our claim that θ is a new independent exponent.

Before discussing the calculation of θ , however, we provide first a simpler, and more transparent, formulation of the perturbation theory than that given in MS. In particular the final result, Eq. (14), does not appear explicitly in MS [13].

Let $y(t)$ be a Gaussian stochastic process with zero mean, whose correlation function obeys dynamical scaling, i.e., $\langle y(t_1)y(t_2) \rangle = t_1^\alpha \Phi(t_1/t_2)$. Let $T = \ln t$ and $x(T) = y(t)/\langle y^2(t) \rangle^{1/2}$. Then $x(t)$ is a Gaussian stationary process with zero mean, i.e., its correlation function is translationally invariant, $\langle x(T_1)x(T_2) \rangle = A(T_2 - T_1)$. Notice that $A(0) = 1$ by construction, a convention that we shall adopt throughout this paper (in contrast to that of Ref. [8]). If the persistence probability of y decays algebraically in t , then the persistence probability of $x(T)$ decays as $\sim \exp(-\theta T)$ for $T \rightarrow \infty$.

The persistence probability may be expressed as the ratio of two path integrals, as follows [8]:

$$P(x(T') > 0; 0 < T' < T) = \frac{\int_{x>0} D x(T) \exp(-S)}{\int D x(T) \exp(-S)}, \quad (1)$$

where

$$S = \frac{1}{2} \int_0^T dT_1 \int_0^T dT_2 x(T_1) G(T_1, T_2) x(T_2). \quad (2)$$

Here $G(T_1, T_2)$ is the matrix inverse of the correlation matrix $\langle x(T_1)x(T_2) \rangle \equiv A(T_2 - T_1)$. Notice that G is not simply a function of $T_2 - T_1$ (unless we impose periodic boundary conditions).

In MS this path-integral formalism was used to map the Markov process onto a quantum harmonic oscillator in imaginary time, developing the perturbation theory in the formalism of quantum mechanics. We shall merely use path integrals as a convenient notation, performing all our calculations within the natural framework of stochastic processes.

Let $x^0(T)$ be a stationary Gaussian Markov process, i.e., one defined by

$$\frac{dx^0}{dT} = -\mu x^0 + \xi(T), \quad (3)$$

where ξ is a Gaussian white noise, with $\langle \xi(T)\xi(T') \rangle = 2\mu\delta(T-T')$. The noise strength has been chosen so that the autocorrelation function is $A^0(T) = \exp(-\mu T)$.

Suppose the process $x(T)$ is perturbatively close to a Markov process, in the sense that $G = G^0 + \epsilon g$. Then we can expand the exponentials in the path integrals in Eq. (1) and reexponentiate, so that to $O(\epsilon)$ the numerator becomes

$$\int_{\mathcal{C}} Dx(T) e^{-S} = \int_{\mathcal{C}} Dx(T) \exp\left(-S^0 - \frac{\epsilon}{2} \int_0^T dT_1 \int_0^T dT_2 \times g(T_1, T_2) A_{\mathcal{C}}^0(T_1, T_2) + O(\epsilon^2)\right), \quad (4)$$

where the subscript \mathcal{C} represents the constraint $x(T') > 0$ ($0 < T' < T$) on the paths in the integral in the numerator of Eq. (1), and

$$A_{\mathcal{C}}^0(T_1, T_2) \equiv \frac{\int_{\mathcal{C}} Dx(T) x(T_1) x(T_2) e^{-S^0}}{\int_{\mathcal{C}} Dx(T) e^{-S^0}} \quad (5)$$

is the correlation function for the Markov process, averaged (and normalized) only over the paths consistent with the constraint \mathcal{C} . The denominator in Eq. (1) is given by an identical expression, except that $A_{\mathcal{C}}^0$ is replaced by A^0 , the unconstrained correlation function.

By virtue of the constraint, $A_{\mathcal{C}}^0$ will not be strictly translationally invariant for finite T . In the limit $T \rightarrow \infty$, however, the double time integral in Eq. (4) reduces to T times an infinite integral over the relative time $T_2 - T_1$, with $A_{\mathcal{C}}^0(T_1, T_2)$ replaced by its stationary limit $A_{\mathcal{C}}^0(T_2 - T_1)$. Similarly, g will be translationally invariant in this regime, giving

$$\int_0^T dT_1 \int_0^T dT_2 g(T_1, T_2) A_{\mathcal{C}}^0(T_1, T_2) \rightarrow T \int_{-\infty}^{\infty} (d\omega/2\pi) \tilde{g}(\omega) \tilde{A}_{\mathcal{C}}^0(\omega), \quad (6)$$

where we have used the translational invariance to write the final result in Fourier space [15]. Note that the zeroth-order result $\int_{x>0} Dx(T) \exp(-S^0) / \int Dx(T) \exp(-S^0)$ is just the persistence probability of the stationary Gaussian Markov process $x^0(T)$, which decays as $\exp(-\mu T)$ as $T \rightarrow \infty$.

Using (1), (4) and (6), we find that the persistence exponent may be written in the form

$$\begin{aligned} \theta &\equiv \lim_{T \rightarrow \infty} -\frac{1}{T} \ln[P(x(T') > 0; 0 < T' < T)] \\ &= \mu + \epsilon \int_0^{\infty} \frac{d\omega}{2\pi} \tilde{g}(\omega) [\tilde{A}_{\mathcal{C}}^0(\omega) - \tilde{A}^0(\omega)] + O(\epsilon^2). \end{aligned} \quad (7)$$

where the term in $\tilde{A}^0(\omega)$ is the $O(\epsilon)$ contribution from the denominator in Eq. (1), and we have exploited the $\omega \rightarrow -\omega$ symmetry of the integrand.

We now calculate $A_{\mathcal{C}}^0(T)$. The conditional probability $Q(x, T | x_0, 0)$ for the stationary Markov process may be obtained directly from Eq. (3):

$$Q(x, T | x_0, 0) = \left[\frac{1}{2\pi(1 - e^{-2\mu T})} \right]^{1/2} \exp\left[-\frac{(x - x_0 e^{-\mu T})^2}{2(1 - e^{-2\mu T})} \right]. \quad (8)$$

The conditional probability $Q^+(x_2, T_2 | x_1, T_1)$ that the process goes to (x_2, T_2) , given that it started from (x_1, T_1) , without x ever being negative is given by the method of images:

$$Q^+(2|1) = Q(x_2, T_2 | x_1, T_1) - Q(x_2, T_2 | -x_1, T_1), \quad (9)$$

where we have adopted an obvious shorthand notation for the arguments of Q^+ .

To calculate the joint probability $P^+(x_1, T_1; x_2, T_2)$ that the process passes through x_1 at T_1 and x_2 at T_2 , averaged only over paths where $x(T)$ is always positive, we consider a path starting at (x_i, T_i) and finishing at (x_f, T_f) , passing through (x_1, T_1) and (x_2, T_2) without ever crossing the origin. Then the required stationary limit is

$$P^+(x_1, T_1; x_2, T_2) = \lim_{T_i \rightarrow -\infty, T_f \rightarrow \infty} \frac{Q^+(f; 2; 1|i)}{Q^+(f|i)}. \quad (10)$$

The Markov property means that we can decompose $Q^+(f; 2; 1|i) = Q^+(f|2)Q^+(2|1)Q^+(1|i)$. Using Eqs. (8) and (9) in Eq. (10), we find

$$\begin{aligned} P^+(x_1, 0; x_2, T) &= \frac{2}{\pi} (1 - e^{-2\mu T})^{-1/2} x_1 x_2 e^{\mu T} \\ &\quad \times \exp\left[-\frac{(x_1^2 + x_2^2)}{2(1 - e^{-2\mu T})} \right] \sinh\left(\frac{x_1 x_2}{2 \sinh \mu T} \right). \end{aligned} \quad (11)$$

It is now straightforward to evaluate the autocorrelation function:

$$\begin{aligned} A_{\mathcal{C}}^0(T) &= \int_0^{\infty} dx_1 \int_0^{\infty} dx_2 x_1 x_2 P^+(x_1, 0; x_2, T) \\ &= \frac{2}{\pi} [3(1 - e^{-2\mu T})^{1/2} + (e^{\mu T} + 2e^{-\mu T}) \sin^{-1} e^{-\mu T}]. \end{aligned} \quad (12)$$

Equation (7) for θ can now be expressed as a real-time integral as follows. We first write $A(T) = A^0(T) + \epsilon a(T)$, and we note that in Fourier space $[\tilde{A}(\omega)]^{-1} = \tilde{G}(\omega) = \tilde{G}^0(\omega) + \epsilon \tilde{g}(\omega)$. Using $A^0 = \exp(-\mu T)$ gives $\tilde{g}(\omega) = -\tilde{a}(\omega)(\omega^2 + \mu^2)^2 / 4\mu^2$. Inserting this in (7), and transforming to real time, gives

$$\begin{aligned} \theta &= \mu - \frac{\epsilon}{4\mu^2} \int_0^{\infty} dT a(T) \left(\mu^2 - \frac{d^2}{dT^2} \right)^2 [A_{\mathcal{C}}^0(T) - A^0(T)] \\ &= \mu \left\{ 1 - \epsilon \frac{2\mu}{\pi} \int_0^{\infty} a(T) [1 - \exp(-2\mu T)]^{-3/2} dT \right\}. \end{aligned} \quad (14)$$

The final result is remarkably compact. Since $\epsilon a(T)$ is just the perturbation to the Markov correlator $A^0(T) = e^{-\mu T}$, the normalization $A(T) = 1$ forces $a(0) = 0$. This is sufficient to converge the integral in Eq. (14) provided $a(T)$ vanishes more rapidly than $T^{1/2}$. Equation (14) has recently been used to calculate persistence exponents for interface growth in a class of generalized Edwards-Wilkinson models [14].

As was remarked earlier, the problem of nonequilibrium critical dynamics is Markovian to first order in $\epsilon = 4 - d$. In the thermodynamic limit the global order parameter is Gaussian because, at time t , it is the sum of $[L/\xi(t)]^d$ (essentially) statistically independent contributions, where L is the system size and $\xi \sim t^{1/z}$ is the length scale over which critical correlations have been established. Corrections to the Gaussian distribution can be expressed in terms of higher cumulants of the normalized total magnetization $M(t)/\langle M^2(t) \rangle^{1/2}$. Using the translational invariance of the system with respect to space it is easy to show that for large L the $2N$ -point cumulant is smaller by a factor $(t^{1/z}/L)^{(N-1)d}$ than the Gaussian part of the $2N$ -point correlation function. The perturbative approach discussed in the first part of this paper can therefore be applied. To calculate the lowest non-Markovian term in θ , we need to calculate the autocorrelation function of the total magnetization $M(t)$ to order ϵ^2 , i.e., we need to calculate the autocorrelation function $A(t_1, t_2) = \langle M(t_1)M(t_2) \rangle / \langle M^2(t_1) \rangle^{1/2} \langle M^2(t_2) \rangle^{1/2}$, which in the scaling regime depends only on the ratio t_2/t_1 . The necessary techniques of dynamical field theory, incorporating the extra renormalization associated with the random initial condition (and responsible for the nonequilibrium exponent λ), have been developed by Janssen *et al.* [16,17].

Models A and C of Hohenberg and Halperin [12] are defined by Langevin equations for a nonconserved n -component vector order-parameter field $\vec{s}(\mathbf{r}, t)$ and (for model C) a noncritical conserved density $m(\mathbf{r}, t)$:

$$\partial_t \vec{s} = - \frac{\delta H}{\delta \vec{s}} + \vec{\zeta}, \quad (15)$$

$$\partial_t m = \rho \nabla^2 \frac{\delta H}{\delta m} + \eta \quad (16)$$

with the Hamiltonian

$$H[\vec{s}, m] = \int d^d r \left[\frac{\tau}{2} \vec{s}^2 + \frac{1}{2} (\nabla \vec{s})^2 + \frac{g}{4!} (\vec{s}^2)^2 + \frac{1}{2} m^2 + \frac{\gamma}{2} m \vec{s}^2 - h_m m \right]. \quad (17)$$

The external field h_m is to be adjusted such that $\langle m \rangle = 0$. The Langevin noises $\vec{\zeta}$ and η are Gaussian random forces with zero mean and correlators $\langle \zeta_i(\mathbf{r}, t) \zeta_j(\mathbf{r}', t') \rangle = 2 \delta_{ij} \delta(\mathbf{r} - \mathbf{r}') \delta(t - t')$, $\langle \eta(\mathbf{r}, t) \eta(\mathbf{r}', t') \rangle = -2\rho \nabla^2 \delta(\mathbf{r} - \mathbf{r}') \delta(t - t')$, while the initial conditions $\vec{s}_0(\mathbf{r})$, $m_0(\mathbf{r})$ are Gaussian random variables with distribution $P[\vec{s}_0, m_0] \propto \exp(-H_0[\vec{s}_0, m_0])$, where $H_0[\vec{s}_0, m_0] = \int d^d r [\tau_0 \vec{s}_0^2 / 2 + m_0^2 / 2c_0]$.

We first consider ‘‘model A’’ dynamics (where for $n > 1$ ‘‘persistence’’ is associated with a given component of the order parameter). For model A, Eq. (16) is discarded, and the terms in m are omitted from Eq. (17). The calculation of the autocorrelation function is straightforward in principle [16,17], but algebraically tedious, and the final result is [with $T = \ln(t_2/t_1)$]

$$A(T) = e^{-\mu T} \left[1 - \frac{3(n+2)}{4(n+8)^2} \epsilon^2 F_A(e^T) + O(\epsilon^3) \right], \quad (18)$$

where $\mu = (\lambda - d + 1 - \eta/2)/z$ from the one-loop calculation [11] (equivalent to the ‘‘Markov scaling relation’’), and

$$\begin{aligned} F_A(x) = & -\ln \frac{4}{3} [2 \ln(2x) + (x-1) \ln(x-1) - (x+1) \ln(x+1)] - 2(\ln 2)^2 - \frac{\pi^2}{6} + 4 \ln 2 - (x-1) \ln \left(\frac{x-1}{2x} \right) \\ & + (x+1) \ln \left(\frac{x+1}{2x} \right) + (x-1) \ln \left(\frac{x-1}{2x} \right) \ln \left(\frac{3x-1}{2x} \right) - (x+1) \ln \left(\frac{x+1}{2x} \right) \ln \left(\frac{3x+1}{2x} \right) - (3x+1) \ln \left(\frac{3x+1}{2x} \right) \\ & + (3x-1) \ln \left(\frac{3x-1}{2x} \right) - \frac{(x-1)}{2} \left[\ln \left(\frac{3x-1}{2x} \right) \right]^2 + \frac{(x+1)}{2} \left[\ln \left(\frac{3x+1}{2x} \right) \right]^2 - (x-1) \text{Li}_2 \left(\frac{x-1}{2x} \right) + (x+1) \text{Li}_2 \left(\frac{x+1}{2x} \right) \\ & - 2(x+1) \text{Li}_2 \left(\frac{x+1}{4x} \right) + 2(x-1) \text{Li}_2 \left(\frac{x-1}{4x} \right) + (x+1) \text{Li}_2 \left(\frac{2x}{3x+1} \right) - (x-1) \text{Li}_2 \left(\frac{2x}{3x-1} \right), \end{aligned}$$

and $\text{Li}_2(x) \equiv -\int_0^x dt \ln(1-t)/t$ is the dilogarithm function. The function $F_A(e^T)$ is a bounded, monotonically increasing, function of T in $(0, \infty)$. It vanishes as $T \ln T$ for $T \rightarrow 0$ [satisfying the requirement for convergence at $T=0$ of the integral in Eq. (14)], while $F(\infty) = 0.057622 \dots$

The non-Markov nature of the process $M(t)$ at order ϵ^2 follows from the fact that, at this order, $A(T)$ is no longer a

simple exponential. Substituting $a(T) = A(T) - e^{-\mu T}$ from Eq. (18) into Eq. (14), using $\mu = (1/2) + O(\epsilon)$, we find (after some algebra)

$$\theta = \mu \left\{ 1 + \frac{3(n+2)}{4(n+8)^2} \epsilon^2 \alpha \right\}, \quad (19)$$

where

$$\begin{aligned} \alpha = & 4(\sqrt{2}-2\sqrt{3}+\sqrt{6})+8\sqrt{2}\ln 2-4(\sqrt{2}-1)\ln 3 \\ & -2(1+2\sqrt{2})\ln(3+2\sqrt{2})-14\ln(5+2\sqrt{6}) \\ & +10\ln(7+4\sqrt{3})+8\sqrt{2}\ln[(4+\sqrt{2}-\sqrt{6})/(4-\sqrt{2} \\ & -\sqrt{6})]-4\sqrt{2}\ln[(2\sqrt{3}-2+\sqrt{2})/(2\sqrt{3}-2-\sqrt{2})] \\ = & 0.271577604975\dots \end{aligned}$$

This result can be compared with recent simulation data for the Ising model in two [11,19,20] and three [19] dimensions. For $d=2$, using $\lambda=1.585\pm 0.006$ [21], and $\eta=1/4$ (exact) gives $\mu z=0.460\pm 0.006$. Ignoring non-Markov corrections, one would obtain $\theta z=\mu z$, smaller than the measured value $\theta z=0.505\pm 0.020$ (the finite-size scaling method used in [11] naturally determines the combination θz [20]). The non-Markov correction factor in Eq. (19) is, for $n=1$, $(1+0.0075438\dots, \epsilon^2)\approx 1.030$ for $\epsilon=2$. The ‘‘improved’’ estimate for θz becomes 0.474 ± 0.006 , closer to, but still somewhat smaller than, the numerical estimate.

For $d=3$, one has $z=2.032\pm 0.004$, $\lambda=2.789\pm 0.006$ [21], and $\eta=0.032\pm 0.003$, giving $\mu=0.380\pm 0.003$. Multiplying by the non-Markov correction factor for $\epsilon=1$, i.e., 1.0075, gives $\theta=0.383\pm 0.003$, compared to the numerical result $\theta\approx 0.41$ [19]. A direct expansion to order ϵ^2 , using the known expansions for z , λ , and η , gives (specializing to $n=1$) $\theta=1/2-\epsilon/12+(\alpha-2\ln 3)\epsilon^2/72-2\epsilon^2/81+O(\epsilon^3)$, i.e., $\theta\approx 0.365$ for $d=3$, slightly lower than that obtained using the best numerical estimates of z , λ , and η and only using the ϵ expansion for the non-Markov correction.

A similar approach can be applied to ‘‘model C,’’ defined by the full set of equations (15)–(17). In this case, one ob-

tains non-Markovian corrections already at order ϵ . The autocorrelation function is given by (for $n=1$)

$$A(T)=\exp(-\mu T)\left[1-\frac{\epsilon}{6}F_C(e^T)+O(\epsilon^2)\right], \quad (20)$$

$$\begin{aligned} F_C(x) = & \ln 2 - \frac{x-1}{2}\ln(x-1) - \frac{x+1}{2}\ln(x+1) \\ & + x\ln x - \frac{x-1}{2x}. \end{aligned} \quad (21)$$

Again, $F_C(e^T)$ vanishes like $T\ln T$ for $T\rightarrow 0$, while $F_C(\infty)=\ln 2-1/2$. Inserting $a(T)=A(T)-\exp(-\mu T)$ from Eq. (20) into Eq. (14) gives

$$\theta = \mu \left[1 + \frac{7-4\sqrt{2}}{12}\epsilon + O(\epsilon^2) \right], \quad (22)$$

where $\mu=(\lambda-d+1-\eta/2)/z$ as before, but now the dynamical exponents z and λ take their model-C values [12,18].

In summary, we have computed to order ϵ^2 the persistence exponent θ for the global order parameter $M(t)$ of models A and C. At this order, the dynamics of $M(t)$ are non-Markovian, and θ is a new exponent, not related to the usual static and dynamic exponents. The calculation was performed by expanding around a Markov process, using a simplified form of the perturbation theory introduced by Majumdar and Sire.

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- [1] B. Derrida, A. J. Bray, and C. Godrèche, *J. Phys. A* **27**, L357 (1994).
 - [2] A. J. Bray, B. Derrida, and C. Godrèche, *Europhys. Lett.* **27**, 175 (1994).
 - [3] D. Stauffer, *J. Phys. A* **27**, 5029 (1994).
 - [4] E. Ben-Naim, P. L. Krapivsky, and S. Redner, *Phys. Rev. E* **50**, 2474 (1994).
 - [5] B. Derrida, P. M. C. de Oliveira, and D. Stauffer, *Physica A* **224**, 604 (1996).
 - [6] B. Derrida, V. Hakim, and V. Pasquier, *Phys. Rev. Lett.* **75**, 751 (1995).
 - [7] E. Ben-Naim, L. Frachebourg, and P. L. Krapivsky, *Phys. Rev. E* **53**, 3078 (1996).
 - [8] S. N. Majumdar and C. Sire, *Phys. Rev. Lett.* **77**, 1420 (1996).
 - [9] S. N. Majumdar, C. Sire, A. J. Bray, and S. Cornell, *Phys. Rev. Lett.* **77**, 2867 (1996).
 - [10] B. Derrida, V. Hakim, and R. Zeitak, *Phys. Rev. Lett.* **77**, 2871 (1996).
 - [11] S. N. Majumdar, A. J. Bray, S. J. Cornell, and C. Sire, *Phys. Rev. Lett.* **77**, 3704 (1996).
 - [12] P. C. Hohenberg and B. I. Halperin, *Rev. Mod. Phys.* **49**, 435 (1977).
 - [13] The result (14) can be deduced after some algebra from Eqs. (4), (5), and (7) of Ref. [8].
 - [14] J. Krug, H. Kallabis, S. N. Majumdar, S. Cornell, A. J. Bray, and C. Sire (unpublished).
 - [15] Note that $A_C^0(\infty)>0$ [see Eq. (13)], so $\tilde{A}_C^0(\omega)$ implicitly contains a term proportional to $\delta(\omega)$.
 - [16] H. K. Janssen, B. Schaub, and B. Schmittman, *Z. Phys.* **73**, 539 (1989).
 - [17] H. K. Janssen, in *From Phase Transitions to Chaos, Topics in Modern Statistical Physics*, edited by G. Györgyi, I. Kondor, L. Sasvári, and T. Tél (World Scientific, Singapore, 1992), pp. 68–91.
 - [18] K. Oerding and H. K. Janssen, *J. Phys. A* **26**, 3369 (1993).
 - [19] D. Stauffer, *Int. J. Mod. Phys. C* **7**, 753 (1996).
 - [20] L. Schülke and B. Zheng (unpublished) have recently measured θ directly for the Ising model in $d=2$. Their result $\theta=0.237\pm 0.003$ is consistent with $\theta z=0.505\pm 0.020$ [11] for z in the accepted range 2.13–2.17.
 - [21] P. Grassberger, *Physica A* **214**, 547 (1995).