# Time decay of excitations at quasi-one-dimensional trapping

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The survival probability of a particle that performs a quasi-one-dimensional random walk on a lattice with randomly distributed traps is considered. The derivations are based on the approximation of weak transitions between one-dimensional chains where detailed results on the survival probability are available, neglecting returns to the same chain. It is shown that calculated curves of the survival probability coincide rather well with data of Monte Carlo simulation at times  $t \sim 1/\delta_2$ , where  $\delta_2$  is the degree of quasi-one-dimensionality.

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### I. INTRODUCTION

The problem of random walks of electronic excitations in low-dimensional systems with randomly distributed traps is rather actual. There are a number of theoretical papers studying the migration and trapping of excitations in one-dimensional (1D) systems [1-8], and far less in quasi-one-dimensional (Q1D) objects [9]. Wieting, Fayer, and Dlott [9] studied in detail the time dependence of the mean number of new sites visited by a walker in the case of Q1D energy migration in two- and three-dimensional topology. However, their expression for the trapping rate and survival probability is based on the Rosenstock approximation [10], which is rather rough at long times in 1D and Q1D systems.

At the same time in a number of experimental papers the energy migration and trapping are studied by timeresolved luminescence spectroscopy in Q1D crystals (antiferromagnetic [11-15] and molecular [16,17]). These investigations have shown that even the insignificant probability of an excitation off-chain hop results in essential changes in the luminescence decay curves. In all likelihood, while studying the energy migration in multichain biopolymers [18], the Q1D character of transport should be taken into account.

The present paper is concerned with a simple model describing the decay law of survival probability due to Q1D random walks on a lattice with randomly distributed traps. This approach is based on the exact [2,8] or approximated [5,10] solutions for the survival probability at strictly 1D random walks at high and low trap concentrations, respectively. In spite of some simplifications, the calculated survival probability at times corresponding to the time of an off-chain hop is in good agreement with the results of Monte Carlo (MC) simulations.

#### **II. THE DECAY OF SURVIVAL PROBABILITY**

Consider the following problem: a walker performs Q1D random walks in three-dimensional topology on the periodic lattice with randomly distributed quenched traps. When a walker hops to a trap site, it is captured instantaneously (efficient trapping). The discrete time and nearest-neighbor transitions are considered. At each step with relative probabilities  $W_{\text{on}}$ ,  $W_{\text{off }1}$ ,  $W_{\text{off }2}$  a walker is able to transit to one of six neighbor sites. These probabilities are determined as

on chain (1D direction) 
$$W_{on} = \frac{F_{on}}{F_{tot}}$$
, (1)  
off chain 
$$\begin{cases} W_{off 1} = \frac{F_{off 1}}{F_{tot}}, \\ W_{off 2} = \frac{F_{off 2}}{F_{tot}}, \\ F_{tot} = F_{on} + F_{off 1} + F_{off 2}, \end{cases}$$

where  $F_{on}$  and  $F_{off 1}$ ,  $F_{off 1}$  are hopping rates on the chain and off chain, respectively. It follows from the quasione-dimensionality condition that  $F_{off 1}$ ,  $F_{off 2} \ll F_{on}$ . For the sake of simplicity we assume  $W_{off 1} = W_{off 2} = W_{off}$  and indicate  $W_{on} = \delta_1$ ,  $2W_{off} = \delta_2$ , then  $\delta_1 + \delta_2 = 1$ .

Let S(n) be the number of new sites visited by a walker in *n* steps. Then the survival probability (the probability that a walker will be not trapped) is

$$f(n) = p^{S(n)} , \qquad (2)$$

where p = 1 - C and C is the trap concentration. Further, consider Q1D random walks for a certain motion route (or configuration). It is assumed that at the  $l_1+1, l_2+1, \ldots, l_k+1$  step a walker makes an off-chain hop  $(l_k < n)$ . Thus, after n steps, a walker makes k offchain hops. If we assume that after leaving the given chain a walker never hops back again, then the mean survival probability for the given configuration is

$$\langle f(n) \rangle = p^k \langle p^{S(l_1)} \rangle \langle p^{S[l_2 - (l_1 + 1)]} \rangle \cdots \langle p^{S[n - (l_k + 1)]} \rangle .$$
(3)

Such averaging is possible due to the above assumption of the factors in Eq. (3) being of statistically independent values. By averaging all the configurations  $\{l_1, l_2, \ldots, l_k\}$  and all off-chain hops we get the mean survival probability (hereafter the survival probability) for Q1D random walks

$$\Phi(n) = \delta_1^n \langle p^{S(n)} \rangle + \delta_2 \delta_1^{n-1} p \sum_{l_1=0}^{n-1} \langle p^{S(l_1)} \rangle \langle p^{S[n-(l_1+1)]} \rangle + \delta_2^2 \delta_1^{n-2} p^2 \sum_{l_1=0}^{n-2} \sum_{l_2=l_1+1}^{n-1} \langle p^{S(l_1)} \rangle \langle p^{S[l_2-(l_1+1)]} \rangle \langle p^{S[n-(l_2+1)]} \rangle + o(\delta_2^3) .$$
(4)

The first term of this expression determines the survival probability for a walker that has not gone to a neighboring chain after *n* steps; the second term corresponds to that for a walker that has made one off-chain hop, the third term to that for a walker that has made two off-chain hops, and so on. It is essential that in Eq. (4)  $\Phi(n)$  is expressed by survival probabilities  $\langle p^{S[l_{i+1}-(l_i+1)]} \rangle$  for strictly 1D walks. For this case the well-known solutions are exact for high concentrations [2,8] and approximated for low concentrations [2,5-8].

If we consider the times when a walker makes one offchain hop (the walker's lifetime on a chain) that corresponds to  $n \sim 1/\delta_2$ , it is possible to leave only the two first terms, without a large error in the series (4). Since the parameter  $\delta_2$  is  $\sim 10^{-6} - 10^{-9}$  for a number of crystals [11-17], a decrease of  $\Phi(n)$  at times  $\sim 1/\delta_2$  and concentrations  $C \sim 10^{-2} - 10^{-3}$  will be less than  $10^{-2}$  from the initial value (a similar decrease of  $\Phi(n)$  is characteristic of 1D systems at the same concentrations and times [1,2,5-8,11-15]). The latter allows us to use this approximation for comparison with the experimental luminescence decay of an exciton population in a Q1D system with traps.

Proceeding to integration at  $n \gg 1$ , and taking into account that the integrand is symmetric with respect to the permutation of the  $l_1$  and  $n - l_1$  indices,

$$\Phi(n) = \delta_1 \langle p^{S(n)} \rangle + 2\delta_2 \delta_1^{n-1} p \int_0^{n/2} \langle p^{S(l_1)} \rangle \langle p^{S(n-l_1)} \rangle dl_1 .$$
(5)

At low concentrations (C = 0.01) the integrand can be calculated using the Rosenstock approximation [10] for the short-time limit and the approximation obtained in [5] for the long-time limit. The survival probability in the 1D system as described by the Rosenstock approximation [10] is

$$\langle f(m) \rangle^{\text{Ros}} = \exp(-C\sqrt{8m/\pi})$$
 (6)

According to [6], at C = 0.01, the expression (6) correlates fairly well with MC data for  $m \le 2 \times 10^4$ . In the long-time limit, when  $n \ge 2 \times 10^4$  an approximation obtained by Anlauf [5] is used:

$$\langle f(m) \rangle^{An} = \frac{8}{\pi} \left[ \frac{2}{3\pi} \right]^{1/2} x^{3/2} \left[ 1 + \frac{17}{18} \frac{1}{x} + \frac{205}{645} \frac{1}{x^2} \right]$$
  
  $\times \exp(-\frac{3}{2}x) , \qquad (7)$ 

$$x = (-\pi \ln p)^{2/3} m^{1/3}$$
.

The substitution of Eqs. (6) and (7) for Eq. (5) and a change of *m* for  $l_1, n - l_1, n$  makes it possible to calculate a number of curves for some  $\delta_2$  values by the numerical integration.



FIG. 1. Decay of  $\Phi(n)$  for the trap concentration C=0.01 at various values of the quasi-one-dimensionality degree  $\delta_2$ . The circles denote the MC simulation, the solid lines the results from Eq. (5), and the dashed lines the results from Eq. (8). (a) From top to bottom:  $\delta_2 \le 10^{-7}$ ;  $\delta_2 = 5 \times 10^{-6}$ ,  $10^{-5}$ . The thin line is the exact decay form of  $\Phi(n)$  for strictly 1D migration as obtained in [8]. (b) From top to bottom:  $\delta_2 = 2 \times 10^{-5}$ ,  $4 \times 10^{-5}$ .

## III. MC SIMULATION AND DISCUSSION

In order to check Eq. (5), the curves of the survival probability were simulated for Q1D random walks by the MC method. The expression (2) was used as the basis of the algorithm, making it possible to reduce computation time relative to the direct computation of the survival probability when it was necessary to store the trap configuration. For introducing randomness in the stepping a fast random number generator with a sequence period of  $2^{49}-1$  was written. Depending on the  $\delta_2$  value, realizations from  $2 \times 10^3$  to  $8 \times 10^3$  were simulated. The time intervals were  $n_{\text{max}} = 2 \times 10^5$  steps for C = 0.01 and  $n_{\text{max}} = 4 \times 10^3$  for C = 0.1. We used periodic boundary conditions with  $L = 3600 \times 10 \times 10$  sites (L is a site of the lattice) for simulation. The walker started in the center of the lattice with S(n) = 1 at n = 0.

Figure 1 presents the dependences of  $\Phi(n)$  [Eq. (5), solid lines] and MC simulation (circles) at various  $\delta_2$  values. For  $\delta_2 \le 10^{-7}$ , MC data coincide with the exact solution [thin line in Fig. 1(a)] for strictly 1D random walks [8]. With an increase of  $\delta_2$  the calculated curves coincide well with MC data for  $\delta_2 = 5 \times 10^{-6}$ ,  $10^{-5}$ , and a difference between the MC simulation and Eq. (5) is observed only beginning from  $\delta_2 \ge 2 \times 10^{-5}$ . It should be noted that at  $\delta_2 = 2 \times 10^{-5}$  when  $\delta_2 = 4/n_{\text{max}}$  the approximation still describes the survival probability satisfactorily. Calculated values of  $\Phi(n)$  have to be less than those obtained by the MC method, since only the first two terms of the series (4) were taken into account in Eq. (5). With an increasing of  $\delta_2$  this difference will increase due to an increasing of the contribution of the series terms. which were not taken into account. This is observed graphically when one compares the curves for  $\delta_2 = 2 \times 10^{-5}$  and  $\delta_2 = 4 \times 10^{-5}$  in Fig. 1(b).

It is interesting to compare calculations by Eq. (5) and results obtained for Q1D random walks in threedimensional topology [9]. In this model the survival probability was calculated by the Rosenstock approximation [9,10]

$$\Phi(n)^{\text{Ros}} = \exp\left[-\ln\left[\frac{1}{1-C}\right]R(n)\right], \qquad (8)$$

where R(n) is the mean number of new sites visited by a walker in <u>n</u> steps or sampling function (for  $C \ll 1$  and  $R(n) = \sqrt{8/\pi}n^{1/2}$  [19] we obtain Eq. (6)). The expression for R(n) can be obtained on the basis of a procedure developed by Montroll [19], as it was made in [9], too. In accordance with [9]

$$R(n) = \frac{n}{u_0} + \frac{2u_1}{\sqrt{\pi u_0^2}} n^{1/2} + o(n^{-1}) , \qquad (9)$$

with  $u_1 = \sqrt{2/(\pi^2 \delta_1 \delta_2^2)}$  and  $u_0$  depending on  $\delta_2$  such that it can be evaluated by numerical integration via an appropriate structure function [19]. The results obtained in accordance with Eqs. (8) and (9) are given in Fig. 1. As



FIG. 2. Decay of  $\Phi(n)$  for the trap concentration C = 0.1 at various values of the quasi-one-dimensionality degree  $\delta_2$ . The circles denote the MC simulation, and the solid lines the model results from Eq. (5). The thin line is the exact decay form of  $\Phi(n)$  for strictly 1D migration as obtained in [8]. From top to bottom:  $\delta_2 \leq 5 \times 10^{-6}$ ;  $\delta_2 = 10^{-4}$ ,  $5 \times 10^{-4}$ ,  $10^{-3}$ .

is readily seen, these data are rougher than those of Eq. (5). This is most likely due to the Rosenstock approximation being unsatisfactory for Q1D systems at long times.

For high concentrations (C = 0.1) an exact solution obtained in [2,8] for the survival probability in 1D systems has to be used in Eq. (5). Asymptotics (6) and (7) are not satisfactory for high concentrations [8]. We used results by Onipko and Galchuk [8] since it was not necessary to perform the Laplace transform, distinct from the approach suggested in [2]. As Fig. 2 shows, the calculated curves of  $\Phi(n)$  coincide with MC data if  $\delta_2 \leq 5 \times 10^{-6}$ . The approximation describes quite satisfactorily the decay of  $\Phi(n)$  for  $\delta_2 = 10^{-4}$ , although  $\delta_2$  exceeds  $1/n_{max}$ ( $\delta_2 = 2/n_{max}$ ). It is also of interest to compare the effect of a degree of quasi-one-dimensionality  $\delta_2$  on the character of decay curves of  $\Phi(n)$  at high and low concentrations.

An increase of the  $\delta_2$  parameter by a 1.5 order of magnitude from  $5 \times 10^{-6}$  to  $10^{-4}$  for C = 0.1 results in an insignificant change of the character of  $\Phi(n)$  decay curves (Fig. 2). For example, at  $\Phi(n)=0.01\Phi(0)$  [ $\Phi(n)$  is related here to  $\delta_2=5\times10^{-6}$  and  $\delta_2=10^{-4}$ ] the change is 12%, while this change is 37% for C=0.01 with an increase of  $\delta_2$  by a 1.5 order of magnitude [from  $10^{-7}$  to  $5\times10^{-6}$  Fig. 1(a)]. Thus, at high concentrations the  $\Phi(n)$  decay law is less sensitive to the changes in a degree of quasi-one-dimensionality than at low concentrations. This can be explained by taking into account the fact that a fluctuation of the trap density decreases with an increase of their concentration. The quasi-one-dimensional trapping manifests itself most effectively on long chains because a walker should make at least one off-chain hop before it will be trapped. Then due to a decrease of the trap density fluctuation the contribution of long chains into survival probability  $\Phi(n)$  will be less at high trap concentrations than at low ones, all things being equal.

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