## **Systematic assessment of the quality of fit of the stochastic block model for empirical networks**

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We perform a systematic analysis of the quality of fit of the stochastic block model (SBM) for 275 empirical networks spanning a wide range of domains and orders of size magnitude. We employ posterior predictive model checking as a criterion to assess the quality of fit, which involves comparing networks generated by the inferred model with the empirical network, according to a set of network descriptors. We observe that the SBM is capable of providing an accurate description for the majority of networks considered, but falls short of saturating all modeling requirements. In particular, networks possessing a large diameter and slow-mixing random walks tend to be badly described by the SBM. However, contrary to what is often assumed, networks with a high abundance of triangles can be well described by the SBM in many cases. We demonstrate that simple network descriptors can be used to evaluate whether or not the SBM can provide a sufficiently accurate representation, potentially pointing to possible model extensions that can systematically improve the expressiveness of this class of models.

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### **I. INTRODUCTION**

The stochastic block model (SBM) [\[1,2\]](#page-18-0) is an important family of generative network models used primarily for community detection [\[3\]](#page-18-0) and link prediction [\[4\]](#page-18-0). In its simplest formulation, it describes a network formation mechanism where the nodes are divided into discrete groups, and the probability of an edge existing between two nodes is given as a function of their group memberships. Many variations of this idea exist, including mixed-membership SBMs [\[5\]](#page-18-0), where nodes are allowed to belong to multiple groups, the degreecorrected SBM (DCSBM) [\[2\]](#page-18-0), where nodes are allowed to possess arbitrary degrees, as well as several extensions to other domains, such as dynamical networks [\[6–](#page-18-0)[8\]](#page-19-0) and multilayer networks  $[7,9]$ , to name a few.

SBMs also serve as generalizations of more fundamental random network models. The basic SBM has the Erdős-Rényi model [\[10\]](#page-19-0) as a special case when there is a single group, and likewise the DCSBM recovers the configuration model [\[11\]](#page-19-0) in the same situation. However, differently from these more fundamental models, the SBM possesses a set of parameters—the partition of the nodes and the affinities between groups—that is not trivially recoverable from observed networks. These parameters are *latent* information that need to be obtained via inference algorithms, which form the basis of the community detection methods that use this approach [\[3\]](#page-18-0). Furthermore, the SBM has a controllable level of complexity: by increasing the number of groups, we have the ability to express increasingly elaborate types of network structures, via arbitrary mixing patterns between the latent groups. In fact, despite its stylized nature, it can be shown that the SBM can approximate a broad class of generative models that are different from it [\[12\]](#page-19-0),

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and its inference functions similarly to fitting a histogram to numeric data in order to estimate the underlying probability density—with the node groups playing a similar role to the histogram bins. However, the expressiveness of the SBM is not absolute, especially when the networks are *sparse*, i.e., when their average degree is much smaller than the total number of nodes. In such a situation, there is no guarantee that the SBM is capable of arbitrarily approximating the true underlying model, regardless of how we infer it: By increasing the model complexity we move from a situation where we are *underfitting*, i.e., extracting patterns that do not sufficiently capture all the features of the true model, to a situation where we are *overfitting*, i.e., incorporating randomness into the model description, which is also a deviation from the true model. When we find the most adequate inference that balances statistical evidence against model complexity to prevent overfitting, we might still be missing important features of the true model, simply because it cannot be sufficiently well captured under the SBM parametrization.

Here we are not interested in evaluating the SBM as a plausible generative process of networks across all domains, since it does not represent an ultimately credible mechanism for any of them. Instead, our objective is to assess how capable it is of providing a general *effective* description of empirical networks, and in which aspects and to what extent (and not *whether*) it tends to be misspecified. Understanding the limits of the SBM representation in empirical settings is therefore a nuanced undertaking that is likely to be affected by a variety of possible sources of deviations. Since the SBM tends to yield very good comparative performance in link prediction tasks [\[13,14\]](#page-19-0), it is therefore known that it tends to outperform alternative models in capturing the structure of networks, but we still lack a more accurate assessment of its qualities and shortcomings in absolute terms.

In this work, we evaluate the quality of fit of the SBM in empirical contexts by performing *model checking* on Bayesian

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<span id="page-1-0"></span>inferences. Based on a diverse collection of 275 networks spanning various domains and several orders of size magnitude, we compare the values of many network descriptors computed on the observed network with what would be typically obtained with networks sampled from the inferred SBM. In this way, any significant discrepancy can be interpreted as a form of "residual" that points to a shortcoming of the SBM in capturing that particular network property.

Overall we find that the SBM is capable of encapsulating the network structure to a significant degree for a large fraction of the networks studied, but falls short of completely exhausting the modeling requirements in many cases. We find that for networks with very large diameter or a very slow mixing random walk [\[15\]](#page-19-0) the SBM tends to provide a poor description. This includes, for example, many transportation networks—which are typically embedded in a low-dimensional space—as well as some economic networks.<sup>1</sup> However, for other kinds of networks the quality of fit tends to be good overall.

We proceed with describing in detail the model and inference procedure (Sec.  $II$ ), our criteria to evaluate the quality of fit (Sec.  $III$ ), the network corpus used (Sec. [IV\)](#page-3-0), and the results of our analysis for it (Sec. [V\)](#page-3-0). We finalize in Sec. [VI](#page-7-0) with a conclusion.

#### **II. MODEL AND INFERENCE**

For our analysis we will use the microcanonical degreecorrected SBM (DCSBM)  $[2,17]$  $[2,17]$ , which combines arbitrary mixing patterns between groups together with arbitrary degree sequences. It has as parameters the partition of the nodes into *B* groups,  $\mathbf{b} = \{b_i\}$ , with  $b_i \in [1, B]$  being the group membership of node *i*, the degree sequence  $\mathbf{k} = \{k_i\}$ , where  $k_i$  is the degree of node  $i$ , and the edge counts between groups  $e = \{e_{rs}\}\$  (or twice that number for  $r = s$ ), given by  $e_{rs} = \sum_{ij} A_{ij} \delta_{b_i,r} \delta_{b_i,s}$ . Given these constraints, the network is generated with probability [\[17\]](#page-19-0)

$$
P(A|k, e, b) = \frac{\prod_{r < s} e_{rs}! \prod_{r} e_{rr}! \prod_{i} k_{i}!}{\prod_{i < j} A_{ij}! \prod_{i} A_{ii}! \prod_{r} e_{r}!},\tag{1}
$$

where  $A = \{A_{ij}\}\$ is the adjacency matrix of an undirected multigraph with potential self-loops, and  $e_r = \sum_s e_{rs}$ .

All the networks we will be studying are undirected simple graphs, for which the above model can give only an approx-imation. As demonstrated in Ref. [\[18\]](#page-19-0), the use of multigraph models based on the Poisson distribution (or equivalently, microcanonical models based on the pairing of half-edges, as above) cannot ascribe probabilities to simple edges (i.e.,  $A_{ij} = 1$ ) that are larger than  $1/e \approx 0.37$ . This limits the applicability of such models on networks with heterogeneous density, due to either broad degree distributions or sufficiently dense communities, which are common properties of empirical networks. To address this limitation, we use the latent multigraph model of Ref. [\[18\]](#page-19-0), where we assume that an underlying unobserved multigraph *A* is in fact responsible for

the observed simple graph *G* simply via the removal of the edge multiplicities and self-loops:

$$
P(G|A) = \prod_{i < j} \left(1 - \delta_{A_{ij},0}\right)^{G_{ij}} \delta_{A_{ij},0}^{1 - G_{ij}}.\tag{2}
$$

Note that  $P(G|A)$  can take only a value of 0 or 1, depending on whether *G* and *A* are compatible. Via this mathematical construction, the final model

$$
P(G|k, e, b) = \sum_{A} P(G|A)P(A|k, e, b)
$$
 (3)

can express both arbitrary mixing patterns between groups as well as degree correction, without the limitations of the multigraph model for networks with large local densities [\[18\]](#page-19-0). The inference of this model is performed by sampling from the posterior distribution

$$
P(A, k, e, b|G) = \frac{P(G|A)P(A|k, e, b)P(k, e, b)}{P(G)}, \quad (4)
$$

which remains tractable. Here we use the merge-split Markov chain Monte Carlo (MCMC) algorithm described in Ref. [\[19\]](#page-19-0) to efficiently sample from this distribution.

Note that for  $P(k, e, b)$  we use the nonparametric microcanonical hierarchical priors and hyperpriors described in Refs. [\[17,20\]](#page-19-0). Importantly, this kind of approach determines the appropriate model complexity (via the number of groups) according to the statistical evidence available in the data. As has been shown in these previous works, this choice guarantees that only compressive inferences are made in a manner that prevents overfitting (finding a number of groups *B* that is too large), but also with a substantial protection against underfitting (finding a number that is too small), which tends to happen when noninformative priors are used instead.

In addition to the DCSBM we will also use the configuration model as a comparison, obtained by reshuffling the edges of the obtained network while preserving its degree sequence (here we use the edge-switching MCMC algorithm [\[11\]](#page-19-0)). We note that the configuration model is an approximate special case of the DCSBM considered above when there is only a single group.<sup>2</sup> Therefore, whenever the Bayesian approach above identifies more than one group with a large probability, this automatically implies a selection of the DCSBM in lieu of the configuration model. This happens for every network that we consider in this work, meaning that the DCSBM is the favored model for all of them. Nevertheless, the configuration model serves as a good baseline to determine to what extent the quality of fit obtained with the DCSBM can be ascribed to the degree sequence alone or to the group-based mixing patterns uncovered.

#### **III. ASSESSING QUALITY OF FIT**

The approach we use to assess the quality of fit of the DCSBM is based on obtaining the *posterior predictive distribution* of certain network descriptors. More precisely, for a

<sup>&</sup>lt;sup>1</sup>See Ref.  $[16]$  for a qualitative overview of the different network classifications we consider.

<sup>2</sup>This is only approximately true since the configuration model and the latent Poisson models are not identical, but sufficiently similar for the purposes of this work [\[18\]](#page-19-0).

<span id="page-2-0"></span>scalar network descriptor  $f(G)$ , its posterior predictive distribution is given by

$$
P(y|\mathbf{G}) = \sum_{\substack{\mathbf{G}', \mathbf{A}', \mathbf{A} \\ \mathbf{k}, \mathbf{e}, \mathbf{b}}} \delta(y - f(\mathbf{G}')) P(\mathbf{G}'|\mathbf{A}')
$$

$$
\times P(\mathbf{A}'|\mathbf{k}, \mathbf{e}, \mathbf{b}) P(\mathbf{A}, \mathbf{k}, \mathbf{e}, \mathbf{b}|\mathbf{G}), \tag{5}
$$

where  $\delta(x)$  is the Dirac delta function. In other words, for each inferred parameter set  $(k, e, b)$ , weighted according to its posterior probability, we sample a new network  $G'$  from the model defined above (which can be done in time  $O(E + N)$ ) where *E* and *N* are the total number of edges and nodes, respectively, as we show in Appendix  $\overrightarrow{A}$ ), and obtain the descriptor value  $y = f(G')$ .<sup>3</sup>

We can say that a model captures well the value of a descriptor if its predictive posterior distribution ascribes high probability to values that are close to what was observed in the original network. We can obtain a compact summary of the level of agreement in two different ways. The first measures the statistical significance of the deviation, e.g., via the *z* score [\[21\]](#page-19-0)

$$
z = \frac{f(G) - \langle y \rangle}{\sigma_y},\tag{6}
$$

where  $\langle y \rangle$  and  $\sigma_y$  are the mean and standard deviation of  $P(y|G)$ . The second criterion is the relative deviation, which here we compute in two different ways,

$$
\Delta_1 = \frac{f(G) - \langle y \rangle}{f(G)}, \quad \Delta_2 = \frac{f(G) - \langle y \rangle}{f_{\text{max}} - f_{\text{min}}}, \tag{7}
$$

depending on whether the descriptor values are bounded in a well-defined interval  $[f_{min}, f_{max}]$  ( $\Delta_2$ ) or not ( $\Delta_1$ ).

The *z* score and relative deviation measure complementary aspects of the agreement between data and model, and represent different criteria which should be used together. While a high value of the *z* score can be used to reject the inferred model as a plausible explanation for the data, by itself it tells us nothing about how good an approximation it is. Conversely, the relative deviation tells us how well the descriptor is being reproduced by the model, but nothing about the statistical significance of the comparison.

In Fig. 1 we show examples that illustrate how the different criteria operate. In Figs.  $1(a)$  and  $1(b)$  we see examples that show good and bad agreements between model and data, respectively, according to both criteria simultaneously. In these cases, the conclusion is unambiguous: we either see no reason whatsoever to condemn the model, or we see a definitive reason to do so. However, in Figs.  $1(c)$  and  $1(d)$  we reach mixed conclusions. In Fig.  $1(c)$  the model typically yields different values than observed in the data, but it still ascribes a large probability to it. We cannot condemn the model as an implausible explanation for the data, but it is conceivable that the true generative model would be more concentrated on the



FIG. 1. Examples of posterior predictive distributions for some descriptors (see Table [I](#page-3-0) for definitions) using the DCSBM, together with *z* score and relative deviation. The solid black line shows the empirical value of the descriptor  $f(G)$ , and the dashed green line the mean of the predictive posterior distribution. In (a) and (b) we see examples where employing both criteria reveal unambiguously good and bad agreements, respectively, between data and model. However, in (c) we see a situation where despite a substantial disagreement with respect to the relative deviation, the *z* score indicates that the model cannot be discarded as a plausible explanation for the data. In (d) we see a situation where the *z* score points to decisive rejection of the model, but the small relative deviation allows us to accept it as an accurate approximation.

observed value. Conversely, in Fig.  $1(d)$  we see a situation where the model ascribes close to zero probability to the actual descriptor value seen in the data, but, in absolute terms, the discrepancy is quite small. Although we find evidence to condemn the plausibility of the model, we could still claim that it is a good approximation.

Overall, since we know that a model like the DCSBM cannot possibly correspond to the true generative model of empirical networks, we should expect that in situations where the network is sufficiently large, and hence there is more abundant data, the values of the *z* score will tend to be high. Here we argue that since the objective of a model like the DCSBM is to obtain a good approximation of the underlying model, not an exact representation, the ultimate criterion is a combination of the two, where we may deem the model compatible with the data when *either* the *z* score *or* the relative deviation has a sufficiently low magnitude. For the purpose of clarity and simplicity of our analysis, we will consider the thresholds  $|z| = 3$  and  $|\Delta| = 0.05$  as reasonable choices to deem the model compatible with data, although our results will not depend on these particular choices, and we will always report the full range of values.

Before continuing, some important considerations regarding model checking should be made. While an excellent model should fulfill both of the above criteria simultaneously, we need to observe that a model that maximally overfits,

<sup>&</sup>lt;sup>3</sup>The posterior predictive distribution for the configuration model is analogous, i.e.,  $P(y|\mathbf{G}) = \sum_{G'} \delta(y - f(\mathbf{G}'))P(\mathbf{G}'|\mathbf{k})$ , where  $\mathbf{k}$  are the observed degrees, and  $P(G|k)$  is the likelihood of the configuration model.



<span id="page-3-0"></span>TABLE I. Network descriptors used in this work, with their respective symbol, range of values, and how the relative deviation was computed. More details on how the descriptors are computed are given in Appendix [B.](#page-8-0)

i.e., ascribes to the observed network a probability of one, and to any other a probability of zero, will achieve the best possible performance according to both relative deviation and statistical significance. This occurs because we are using the same data to perform both the model inference and evaluate its quality, which is an invalid approach for *model selection*. Therefore, it is important to recognize the crucial difference between model checking and model selection: the latter attempts to find the model alternative that is better justified according to statistical evidence, while the former simply finds systematic discrepancies between the inferred model and data. In our analysis, protection against overfitting is obtained via Bayesian inference, and we use model checking only to evaluate the discrepancies (indeed, the fact we find discrepancies to begin with shows that we cannot be massively overfitting). Another observation is that when performing multiple comparison over many networks and descriptors, some amount of "statistically significant" deviations are always expected, even if the models inferred correspond to the true ones, unless we incorporate the fact that we are doing multiple comparisons in our criterion of statistical significance, which would be the methodologically correct approach. We will not perform such a correction in our analysis, because we do not seek to demonstrate the absolute quality of DCSBM as an ultimately plausible hypothesis for network formation. As we will see from our results, such a correction would gain us very little.

Finally, in Table I we list the network descriptors that are used in this work. Our approach requires scalar values, so we constrained ourselves to this category, and furthermore we chose quantities that can be computed quickly, so that robust statistics from the predictive posterior distributions can be obtained. Given these restrictions, we then chose descriptors that measure different aspects of the network structure, both at a local and global levels. Further details on the network descriptors are given in Appendix [B.](#page-8-0)

#### **IV. NETWORK CORPUS**

We base our analysis on a corpus containing 275 networks spanning various domains and several orders of size magnitude, as shown in Fig. 2. We have not collected every network at our disposal, but instead chosen networks that are

as diverse as possible, in both size and domain, and avoided many networks that are closely related by belonging to the same subset. In Appendix  $C$  we give more details about the data sets used.

#### **V. RESULTS**

In Fig. [3](#page-4-0) we show the summaries of the posterior predictive checks for each descriptor and network, for both models considered. We observe a wide variety of deviation magnitudes, for the same descriptors both across networks and across descriptors. As expected, the DCSBM results show systematically better agreement with the data when compared with the configuration model. Overall, the descriptors that show the



FIG. 2. (Top) Number of nodes and edges for the networks in the corpus used in this work and their domain composition (inset). (Bottom) Distribution of descriptor values for the networks in the corpus. The horizontal line marks the median values.

<span id="page-4-0"></span>

FIG. 3. Distribution of relative deviation (top), *z* score (middle), and fraction of networks reproduced (bottom) for (a) the configuration model and (b) the DCSBM, according to their respective predictive posterior distributions for each descriptor. We also show the median and mean of the absolute values for all descriptors for each network. The solid blue lines mark the negative and positive median values, and the dashed red line marks the values of  $|\Delta| = 0.05$  and  $|z| = 3$ . The fraction of networks reproduced correspond to those that have the absolute value of either  $\Delta$  or *z* below these thresholds. The points in green color correspond to the networks that are not reproduced according to this combined criterion.

worst agreement is the characteristic time of a random walk  $(\tau)$  and the diameter  $(\emptyset)$ , both of which are particularly high for networks that are embedded in two dimensions, and for which the DCSBM is an inaccurate approximation (more on this below). Nevertheless, there is no single descriptor that the DCSBM does not capture for fewer than 50% of the networks. For descriptors like *S*,  $R_r$ ,  $R_t$ , and  $\langle c \rangle$ , the difference between the DCSBM and the configuration model are relatively minor, indicating that those can be captured to a substantial degree by the degree sequence alone.

When considering all descriptors simultaneously for each network, by either the median or mean of the absolute values of the *z* score and relative deviation, we observe that a substantial majority of the networks considered show good agreement with the DCSBM, as opposed to the small minority that agree with the configuration model. The difference between the median and the mean indicates that there is a sizeable fraction of the networks where the agreement is spoiled by a few outlier descriptors—typically  $\tau$  and  $\varnothing$ .

The results obtained by the clustering coefficients are particularly interesting, since it is often the case that they are well reproduced by the DCSBM. This contrasts with what is commonly assumed, namely, that the DCSBM should not be able to capture the abundance of triangles often seen in empirical networks, because in the limit where the number of groups is much smaller than the total number of nodes, the DCSBM becomes locally tree-like [\[22\]](#page-19-0), with a vanishing probability of forming triangles. Therefore, we may imagine that the situations where there is an agreement with the DCSBM are those where the clustering values are low. However, as we see in Figs.  $4(a)$  to  $4(d)$ , this is not quite true, and we observe good agreements even when the clustering values are high. This illustrates a point made in Ref.  $[23]$ , that it is possible to obtain an abundance of triangles with the SBM simply by increasing the number of groups, in which case it can be explained as a byproduct of homophily. Indeed this is a situation we see in Figs.  $4(a)$  to  $4(d)$ , where both the relative deviation and *z*-score values can be quite small even for extremal values

<span id="page-5-0"></span>

FIG. 4. Relative deviation and *z*-score values for the global and mean local clustering coefficients,  $C_g$  and  $C_l$ , as well as diameter and characteristic time of a random walk,  $\varnothing$  and  $\tau$ , as a function of their empirical values, for every network in the corpus, when using the DCSBM. The dashed red line marks the values of  $|\Delta| = 0.05$  and  $|z| = 3$ . The size of the symbol corresponds to the logarithm of the number of edges in the network, and the darkness to the mean degree.

of clustering. However, we do notice a substantial variability between agreements, and a fair amount of instances where the DCSBM cannot capture the observed clustering values, even when they are moderate or even small. This seems to indicate that there are a variety of processes capable of resulting in high clustering values, with homophily being only one of them [\[23\]](#page-19-0). Overall, the mean local clustering values tend to be harder to reproduce than the global clustering values. In both cases, the *z* scores are systematically high, indicating that the clustering values are in general a good criterion to reject the DCSBM as a statistically plausible model, although the relative deviation values tend to be lower than what one would naively expect, meaning that the model can still serve as a reasonably accurate approximation for clustered networks in many cases.

In contrast, we observe a different behavior for the diameter and characteristic time of a random walk, which are the least well reproduced descriptors, as shown in Figs.  $4(e)$ to  $4(h)$ . For both these descriptors—which are closely related, since a network with a large diameter will also tend to result in a slow mixing random walk—it is rare to find a network with very high empirical values which the DCSBM is able to accurately describe. Therefore it seems indeed that the DCSBM offers an inadequate ansatz to describe the structure of these networks, even by optimally adjusting its complexity.

In Fig. [5](#page-6-0) we show how the model assessment depends on the size of the network. As one could expect, the *z*-score values tend to increase for larger networks, as more evidence becomes available against the plausibility of the DCSBM as the true generative model. However, the values of the relative deviation do not change appreciably for larger networks, indicating that it remains a good approximation regardless of the size of the system.<sup>4</sup>

In Fig. [6](#page-6-0) we show a summary of the fraction of all networks for which we obtain good agreement with either model, according to the network domains. Overall, we see that most domains show similar levels of agreements, except transportation and economic networks. Transportation networks are often embedded in two-dimensional spaces, resulting in large diameters and slow-mixing random walks. The economic networks considered also tend to show large values of these quantities, so the explanation for their discrepancy is the same.

#### **A. Predicting quality of fit**

Now we address the question of whether it is possible to predict the quality of fit of both models considered based solely on the empirical values of the networks descriptors. If we can isolate the descriptors which are most predictive, this would give us a general direction in which more accurate models could be constructed.

In order to evaluate the predictability, we frame it as a binary classification problem, where to each network *i* is ascribed a binary value  $y_i = 0$  if we have simultaneously  $|z_i|$ 3 and  $|\Delta_i| > 0.05$ , or otherwise  $y_i = 1$ . The feature vector for each network is composed of the empirical values of the descriptors,  $\mathbf{x}_i = (r, \langle c \rangle, C_l, C_g, \lambda_1^A, \lambda_1^H, \tau, \emptyset, R_r, R_t, S, E)$ , with the addition of the number of edges *E*. For each network *i*, we train a random forest classifier on the entire corpus with

<sup>4</sup>Sampling issues with MCMC could also contribute to the elevated *z* scores for larger networks, as we discuss in Appendix [A.](#page-8-0)

<span id="page-6-0"></span>

FIG. 5. Absolute value of the relative deviation (top), *z* score (middle) and fraction of reproduced descriptors (bottom), as a function of the number of edges, for every network in the corpus. The dashed red line marks the values of  $|\Delta| = 0.05$  and  $|z| = 3$ . The fraction of descriptors reproduced correspond to those that have the value of either  $\Delta$  or *z* below these thresholds. The points in green color correspond to the descriptors that are not reproduced according to this combined criterion.

that network removed, and evaluate the prediction score on the held-out network. We then repeat this procedure for all networks in the corpus, and evaluate how well the classifier is able to predict the binary label. We present the results of this experiment in Fig. [7](#page-7-0) (top) which shows the receiver operating characteristic (ROC) curve, where the true positive rate and the false positive rate are plotted for all threshold values used to reach a classification. The area under the ROC curve (AUC), shown in the legend, can be equivalently interpreted as the probability that a randomly chosen true positive has a prediction score higher than a randomly chosen true negative. For the DCSBM and configuration model, we obtain an AUC value of 0.91 and 0.88, respectively. This indicates a fairly high predictability, from which we can conclude that it is indeed often possible to tell whether the models will provide a good or bad agreement, based only on the descriptor values.

Further insight can be obtained by inspecting the importance of each descriptor in the overall classification. We



FIG. 6. Fraction of reproduced networks according to their domain, considering the (a) median and (b) mean values of either the *z* score, the relative deviations, or their combined values, for both models (as shown in the legend). When the combined values are used, this means that a model is deemed compatible with a network when we obtain either  $|\Delta| < 0.05$  or  $|z| < 3$ .

compute this via the so-called Gini importance [\[24\]](#page-19-0), defined as the total decrease in node "impurity" (i.e., how often a node in decision tree contributes to a decision), weighted by the proportion of samples that reach that node, averaged over all trees in the classifier.<sup>5</sup> The results can be seen in Figs.  $7(b)$ and  $7(c)$ . In both cases, we see that the number of edges is the most predictive descriptor, which is compatible with what we had already seen in Fig. 5, namely, that the larger the networks are, the easier it becomes to reject a model according to the *z* score. Otherwise, as one would expect, the importance of the remaining descriptors is largely compatible with their reproducibility shown in Fig. [3,](#page-4-0) where the descriptors that agree the least with the inferred models tend to be the most useful at predicting quality of fit beforehand.

This analysis allows us to emphasize two points: the characteristic time of a random walk  $\tau$  and the diameter  $\varnothing$ , both extremal quantities of the network structure that are closely related, are the most difficult descriptors to be captured by the DCSBM. Therefore, an extension of the model that would cater for these properties would bring the most benefit across all networks. However, beyond these two descriptors, there is no substantial difference between the ones that remain, indicating that there is no obvious direction that would bring a systematic modeling improvement over all networks. On the other hand, as we show in Appendix  $\overline{B}$ , the descriptor values and their predictive posterior deviations show nontrivial correlations, which means that if some of them are specifically

<sup>&</sup>lt;sup>5</sup>We also computed different a measure, called permutation importance, which leads to very similar results (not shown).

<span id="page-7-0"></span>

FIG. 7. Predictiveness of the quality of fit of the generative models considered, according to the empirical descriptor values, framed as a binary classification problem, as described in the text. (a) ROC curve for a leave-one-out random-forest classifier, (b) Gini feature importance for the configuration model, (c) same as (b) but for the DCSBM. Panels (d) and (e) show the best ROC AUC obtained for a set of descriptors of a given size, for the configuration model and DSCBM, respectively. Panels (f) and (g) show the same as (d) and (e), respectively, but with the number of edges excluded from the analysis.

targeted, it could potentially improve the quality of fit of other descriptors.

In order to understand what is the minimal amount of information required to predict the suitability of both models, and in this way remove the redundancy provided by the different descriptors, we computed the best ROC AUC obtained by a combination of descriptors of a given size, as shown in Figs.  $7(d)$  and  $7(e)$ . In both cases we see that the predictability is saturated by only few descriptors. $6$  In the case of the configuration model most of the predictability is already achieved by a combination of  $(C_l, \tau, E)$ . For the DCSBM we get instead

 $(r, \emptyset, E)$ . If we remove the number of edges from the set of features (since it is not informative on the actual network structure), we obtain instead  $(C_g, \lambda_1^A, \tau)$  and  $(C_l, \lambda_1^H, \emptyset)$ , for the configuration model and DCSBM, respectively. It should be emphasized that if a descriptor does not appear in the minimal set this does not mean it is not predictive of the quality of fit, only that it offers largely redundant information in that regard. Thus, for both models if we replace  $\varnothing$  with  $\tau$ or  $\lambda_1^H$  with  $\lambda_1^A$ , etc, we get similar results. This suggests that, besides spatial embeddedness (which influence  $\varnothing$  and  $\tau$  the most), the addition of explicit mechanisms for triangle formation (which affects  $C_g$ ,  $C_l$ ,  $\lambda_1^H$ ,  $\lambda_1^A$  directly) might improve the overall expressiveness of the DCSBM—which in fact has been observed in a more limited data set [\[23\]](#page-19-0).

### **VI. CONCLUSION**

We performed a systematic analysis of posterior predictive checks of the SBM on a diverse corpus of empirical networks, spanning a broad range of sizes and domains. Using a variety of network descriptors, we observed that the SBM is able to accurately capture the structure of the majority of networks in the corpus. The types of networks that show the worst agreement with DCSBM tend to possess a large diameter and a slow mixing of random walks—features that are commonly associated with a low-dimensional spatial embedding, and a violation of the "small-world" property. For the other kinds of networks the agreement tends to be fairly good, even for many networks with an abundance of triangles, in contradiction to what is commonly assumed to be possible with this class of models.

We have also identified the minimal set of network descriptors capable of predicting the quality of fit of the SBM, which is composed of the network diameter and characteristic time of a random walk as the most important, followed by clustering as a secondary feature. This points to the most productive directions in which this class of models could be improved.

It is worth emphasizing that the consistency analysis that we have performed, which compares *a posteriori* the modeling assumptions with the actual properties seen in the data, is possible only if these assumptions are made explicitly via a generative model. Community detection methods that are only descriptive in nature (such as modularity maximization  $[25]$ ) cannot be used for these purposes. Not only are these methods not guided by statistical evidence and prone to systematic overfitting, but they also provide no direct way to scrutinize the validity of their implicit assumptions [\[26\]](#page-19-0).

One of the limitations of our analysis is that it is conditioned on the set of descriptors used, and thus shortcomings or successes of the model with respect to other properties not analyzed are not uncovered. A natural extension of our work would be to consider an even broader set of descriptors that could reveal more relevant dimensions for the comparison. This kind of analysis is open ended, as there is no short supply of possible network descriptors. We hope our work will motivate further study in this direction, and with a larger variety of generative models within or beyond the SBM family.

<sup>&</sup>lt;sup>6</sup>Since we optimized exhaustively for all descriptor combinations of a given size, care should be taken to avoid overfitting, despite the leave-one-out cross-validation, because the optimization was performed on the same set of networks. Because of this, we consider always the smallest set of descriptors that reaches a ROC AUC close to the optimum, not the actual optimum, which is likely to be overfitting.

### **ACKNOWLEDGMENT**

<span id="page-8-0"></span>The computational results presented have been achieved using the Vienna Scientific Cluster (VSC).

### **APPENDIX A: POSTERIOR PREDICTIVE SAMPLING**

As described in the main text, we obtain samples from the posterior predictive distribution of Eq. [\(5\)](#page-2-0) by first sampling from the posterior distribution of Eq. [\(4\)](#page-1-0) using MCMC and then generating new networks from the inferred models. More specifically, we sample  $(A, k, e, b)$  from

$$
P(A, k, e, b|G) = \frac{P(G|A)P(A|k, e, b)P(k, e, b)}{P(G)}, \quad (A1)
$$

using the merge-split MCMC of Ref. [\[19\]](#page-19-0), together with the agglomerative initialization heuristic of Refs. [\[20,27\]](#page-19-0) and the multigraph edge moves of Ref. [\[18\]](#page-19-0). For networks of size up to  $E = 10^5$  edges we observe good equilibration of the MCMC runs, but for large networks it becomes too slow. For these large networks we settle for a point estimate of the partition *b* obtained by several runs of the initialization algorithm and keeping the best result, and then we equilibrate the chain according to  $A$  alone (which affects  $k$  and  $e$ ), which tends to happen quickly. We have verified that performing this calculation several times yields very similar results. The only noticeable outcome of this shortcut for larger networks is that it tends to reduce the variance of the posterior predictive distributions, which can potentially contribute to the elevated *z* scores we obtained in our analysis. However, since the relative deviation values we obtained did not seem to depend on the size of the network, this gives us confidence that this approach does not introduce significant biases.

Given a sample  $(A, k, e, b)$ , we are interested only in  $(k, e, b)$  (and hence samples from their marginal distribution), so we discard  $\vec{A}$  and sample a new multigraph  $\vec{A}$  from the model of Eq. [1.](#page-1-0) This can be done exactly with an efficient algorithm that works similarly to what was proposed in Refs. [\[28,29\]](#page-19-0), but is valid for the microcanonical model: Given the parameters  $(k, e, b)$  we proceed by creating for each group *r* a multiset of candidate nodes  $v_r$ , containing  $k_i$  copies of each node *i* with  $b_i = r$ . Then, for each group pair  $(r, s)$ with  $r \leq s$  and  $e_{rs} > 0$ , we repeat the following three steps for an  $e_{rs}$  number of times (or  $e_{rs}/2$  if  $r = s$ ):

(1) We sample a node *i* from the multiset  $v_r$  uniformly at random, and we remove it from the multiset.

(2) We sample a node *j* from the multiset  $v_s$  uniformly at random, and we remove it from the multiset.

(3) We add an edge  $(i, j)$  to  $A$  (i.e., increment  $A_{ij}$  by one, or two if  $i = j$ ).

The resulting multigraph *A* is sampled exactly with a probability given by Eq.  $(1)$ . Since the number of nonzero entries of *e* cannot be larger than the total number of edges *E*, the whole algorithm finishes in time  $O(N + E)$ , where *N* is the number of nodes.

Given a sample *A*, we obtain a simple graph *G* simply by removing all self-loops and truncating the edge multiplicities:

$$
G_{ij} = \begin{cases} 1, & \text{if } A_{ij} > 0 \text{ and } i \neq j, \\ 0, & \text{otherwise.} \end{cases}
$$
 (A2)

Finally, given  $G$  we compute the network descriptor  $f(G)$  of interest.

A  $C + +$  implementation of every algorithm used in this analysis is freely available as part of the graph-tool library [\[30\]](#page-19-0).

#### **APPENDIX B: NETWORK DESCRIPTORS**

Below are the definitions of the descriptors used in our analyses.

*Degree assortativity, r:* Defined as [\[31\]](#page-19-0)

$$
r=\frac{\sum_{kk'}kk'(m_{kk'}-m_km_{k'})}{\sigma_k\sigma_{k'}},
$$

where  $m_{kk'}$  is the fraction of edges with endpoints of degree  $k$ and *k*',  $m_k = \sum_{k'} m_{kk'}$ , and  $\sigma_k$  is the standard deviation of  $m_k$ .

*Mean k core,*  $\langle c \rangle$ : The *k* core is a maximal set of vertices such that its induced subgraph only contains vertices with degree larger than or equal to  $k$ . The  $k$ -core value  $c_i$  of node  $i$ is the largest value of *k* for which *i* belongs to the *k* core. The mean value is then

$$
\langle c \rangle = \frac{1}{N} \sum_i c_i.
$$

This can be computed in time  $O(N + E)$  according to the algorithm of Ref. [\[32\]](#page-19-0).

*Mean local clustering coefficient, C<sub>l</sub>: The local clustering* coefficient [\[33\]](#page-19-0) of node *i* is given by

$$
C_i = \frac{\sum_{jk} G_{ij} G_{ki} G_{jk}}{k_i (k_i - 1)}.
$$

It measures the fraction of pairs of neighbors that are also connected. The mean value is then just

$$
C_l = \frac{1}{N} \sum_i C_i.
$$

*Global clustering coefficient, Cg:* The global clustering coefficient of is given by

$$
C_g = \frac{\sum_{ijk} G_{ij} G_{ki} G_{jk}}{\sum_i k_i (k_i - 1)}.
$$

It measures the fraction of connected triads that close to form a triangle.

Leading eigenvalue of adjacency matrix,  $\lambda_1^A$ : The leading eigenvalue of the adjacency matrix is the largest value of  $\lambda$ which solves

$$
Gx=\lambda x,
$$

where  $x$  is the associated eigenvector.

*Leading eigenvalue of Hashimoto matrix,*  $\lambda_1^H$ : The leading eigenvalue of the Hashimoto (a.k.a. nonbacktracking) matrix [\[34\]](#page-19-0) is the largest value of  $\lambda$  which solves

$$
Hx=\lambda x,
$$

where  $\boldsymbol{x}$  is the associated eigenvector, and  $\boldsymbol{H}$  is an asymmetric  $E \times E$  matrix with entries defined as

$$
H_{k \to l, i \to j} = \begin{cases} 1 & \text{if } G_{kl} = G_{ij} = 1, l = i, k \neq j, \\ 0 & \text{otherwise.} \end{cases}
$$

<span id="page-9-0"></span>

FIG. 8. Absolute value of the *z* score versus absolute value of relative deviation, for every descriptor value and network in the corpus, according to (a) the configuration model and (b) the DCSBM. The dashed lines mark the values  $|z| = 3$  and  $|\Delta| = 0.05$ , and the histograms the marginal distributions. The solid blue lines mark the median values.

*Characteristic time of a random walk, τ: The characteristic* time of a random walk is obtained via the second largest eigenvalue  $\lambda_2^T \in [0, 1]$  of the transition matrix  $T$ , with entries

$$
T_{ij}=\frac{G_{ij}}{k_j},
$$

where  $k_i = \sum_j G_{ji}$ . It is defined as

$$
\tau=-\ln\lambda_2^T.
$$

If the network is disconnected, we compute  $\tau$  only on the largest component.

*Pseudodiameter,* ∅: The pseudodiameter is an approximate graph diameter. It is obtained by starting from an arbitrary source node, and finding a target node that is farthest away from the source. This process is repeated by treating the target as the new starting node, and ends when the graph distance

no longer increases. This graph distance is taken to be the pseudodiameter. The algorithm runs in time  $O(N + E)$ .

If the network is disconnected,  $\varnothing$  is taken as the maximum of pseudodiameters of the connected components.

*Node percolation profile (random removal), R<sub>r</sub>: We chose* a random node order and remove nodes sequentially from the graph according to it. If  $S_i$  is the fraction of nodes in the largest component after the *i*th removal, then the profile value is

$$
R_r = \frac{1}{N} \sum_i S_i.
$$

The value is averaged over several node orderings.

*Node percolation profile (targeted removal), R<sub>t</sub>: The com*putation is the same as  $R_r$ , but the nodes are always removed in decreasing order of the degree.



FIG. 9. (a) Kendall's correlation coefficient  $\tau$  between pairs of descriptor values across all networks in the corpus. Panels (b) and (c) show the same but for *z* score and relative deviation values, respectively, according to the DCSBM. The insets show the correlation between coefficients from each respective panel and panel (a).

<span id="page-10-0"></span>

# TABLE II. Descriptions of network data sets.



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<span id="page-18-0"></span>

*Fraction of nodes in the largest component, S:* A component is a maximal set of nodes that are connected by a path. The largest component is the component with the largest number of nodes, and *S* is the fraction of all nodes that belong to it.

In Fig. [8](#page-9-0) we show how the *z* scores and relative deviation values are related for every network descriptor, according to both models used. In Fig. [9](#page-9-0) we show Kendall's  $\tau$  correlation coefficient among the descriptor values themselves, as well as their *z* scores and relative deviations, according to the DCSBM. The insets show how the correlations among the deviations are themselves also correlated with the descriptor correlations.

#### **APPENDIX C: DATA SET DESCRIPTIONS**

Table [II](#page-10-0) gives descriptions of the network data sets used in this work. The code names in the first row correspond to the respective entries in the Netzschleuder repository [\[35\]](#page-19-0) where the networks can be downloaded. Some of the descriptions were obtained from the Colorado Index of Complex Networks [\[36\]](#page-19-0).

For all networks, the versions considered in this work were transformed into simple graphs, i.e., symmetrized versions of directed networks and/or with parallel edges and self-loops removed.

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