## Analysis of continuous-time Markovian $\varepsilon$ -SIS epidemics on networks

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We analyze continuous-time Markovian  $\varepsilon$ -SIS epidemics with self-infections on the complete graph. The majority of the graphs are analytically intractable, but many physical features of the  $\varepsilon$ -SIS process observed in the complete graph can occur in any other graph. In this work, we illustrate that the timescales of the  $\varepsilon$ -SIS process are related to the eigenvalues of the tridiagonal matrix of the SIS Markov chain. We provide a detailed analysis of all eigenvalues and illustrate that the eigenvalues show staircases, which are caused by the nearly degenerate (but strictly distinct) pairs of eigenvalues. We also illustrate that the ratio between the second-largest and third-largest eigenvalue is a good indicator of metastability in the  $\varepsilon$ -SIS process. Additionally, we show that the epidemic threshold of the Markovian  $\varepsilon$ -SIS process can be accurately approximated by the effective infection rate for which the third-largest eigenvalue of the transition matrix is the smallest. Finally, we derive the exact mean-field solution for the  $\varepsilon$ -SIS process on the complete graph, and we show that the mean-field approximation does not correctly represent the metastable behavior of Markovian  $\varepsilon$ -SIS epidemics.

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#### I. INTRODUCTION

Over the past 100 years, many epidemic diseases have plagued humanity. Most epidemic outbreaks tend to emerge quickly, but they take much longer to disappear [1]. One of the main reasons is reinfections. A common example of a disease with reinfections is influenza, which affects large portions of the population during the winter season. The influenza virus strain keeps mutating slightly, thereby bypassing the human immune system while maintaining most of its viral properties. Another example of recurring diseases are sexually transmitted diseases, such as chlamydia and gonorrhoea. Contrary to influenza, sexually transmitted diseases do not mutate, but people can simply be reinfected after recovering from the disease.

Initially, epidemic outbreaks spread exponentially fast, because most individuals in the population are susceptible to the new disease, as happened with the COVID-19 pandemic. In a closed and well-mixed population, the number of infected individuals stabilizes after a short time and continues to oscillate around the *prevalence*, which is defined as the average number of infected individuals. Then, the epidemic process is in the *metastable* or *quasistationary state*, because the number of infected individuals remains in the vicinity of the prevalence for a long period of time, whereafter the process eventually converges to its steady state. The steady state of most epidemic processes is the overall-healthy state, which corresponds to the situation in which the disease has disappeared completely from the population.

One of the simplest stochastic epidemic models with reinfections is the continuous-time Markovian susceptibleinfected-susceptible (SIS) model, in which individuals are

The continuous-time Markovian SIS process without selfinfections on static networks has been investigated thoroughly [5]. Even on the complete graph, quantifying the average time spent in the metastable state appears challenging [6]. The major issue is that the metastable state is not stable; it collapses eventually due to a rare occurrence of successive

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either infectious (I) or healthy, but susceptible (S). The contact graph G represents the N individuals as nodes and specifies the L contacts between all pairs of individuals as links. The Markovian SIS dynamics consists of two independent Poisson processes: (i) the infected nodes can cure with rate  $\delta$ , and (ii) infected nodes can infect their connected, susceptible neighbors with rate  $\beta$ . The curing process is a nodal process, whereas the infection process evolves over the links between pairs of nodes. Occasionally, a third, independent self-infection process with self-infection rate  $\varepsilon$  is considered, which describes background or indirect infections. Infections may happen either through direct contact or indirectly, for example after touching infected surfaces or inhaling air in a closed room previously contaminated by an infected individual. The Markovian  $\varepsilon$ -SIS model consists of three independent Poisson processes: (i) the curing process with rate  $\delta$ , (ii) infection process with rate  $\beta$ , and (iii) self-infection process with rate  $\varepsilon$ . Given that the  $\varepsilon$ -SIS model consists of independent Poisson processes, we can describe the time-dependent behavior of the  $\varepsilon$ -SIS model as a continuous-time Markov chain with  $2^N$  states. However, continuous-time Markovian modeling implies that the infection and curing times are exponentially distributed. Measurements in real epidemics seem to suggest that the infection time follows a bell-shaped distribution (such as Gamma, Weibull, or log-normal distributions [2,3]). This requires non-Markovian analysis, which is, unfortunately, considerably more complex (see, e.g., [4]) and is currently insufficiently developed to compute time-dependent infection probabilities.

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curings to the absorbing or overall-healthy state [7]. Several approaches have been proposed to quantify the metastable state in the Markovian SIS model. Jacquez and Simon [8] introduced an epidemic process that prevents the original SIS process from entering the absorbing state at time t. Their reduced SIS model has a unique steady state, which can be related to the metastable state of the original SIS model. Cator and Van Mieghem [9] constructed a similar modified process: If only a single node is infected, then the modified process forbids the curing of that node. Effectively, their modified SIS process is equivalent to the SIS process, with the exception that the transitions to the absorbing state have been removed. De Oliveira and Dickman [10] proposed to store the complete time lapse of the SIS model. Once the process converges to the absorbing state, the process jumps to a randomly selected sample from the history of the SIS process. Keeling and Ross [11] and Hill et al. [12] introduced a self-infection process  $\varepsilon$  on the complete graph. In addition to the usual infection and curing processes, the nodes in the modified process can be infected by external sources, which are modeled as self-infections. Van Mieghem and Cator [13] generalized the  $\varepsilon$ -SIS model from the complete graph to general networks. Introducing a small amount of self-infections removes the absorbing state, reestablishes an irreducible Markov chain with a well-defined steady state different from the overall-healthy state and allows for a comparison between the  $\varepsilon$ -SIS model and the SIS model without self-infections. Finally, the 2<sup>N</sup>-state Markov chain, described by  $2^N$  linear differential equations, is often approximated by a mean-field approximation with N nonlinear differential equations. The simplest mean-field approximation for the SIS model on networks is called the N-Intertwined mean-field approximation (NIMFA), and it assumes that the infection state of any two nodes is uncorrelated [14–17]. A mean-field approximation is generally an adequate approximation for large, dense graphs with homogeneous transition rates and for infection rates sufficiently larger than the epidemic threshold. The accuracy of NIMFA with respect to the Markovian SIS process is studied in [18] for various graph types. A key difference is the possibility of die-outs in the stochastic model, albeit with a very small probability, whereas the mean-field model either converges to the endemic equilibrium or to the all-healthy state, and excludes the possibility of sudden die-outs.

The continuous-time Markovian  $\varepsilon$ -SIS process on the complete graph can be described by a birth and death process (BDP). BDPs can be solved by computing the probability generating function and solving the corresponding partial differential equation [19,20]. Unfortunately, solving the partial differential equation seems infeasible for  $\varepsilon$ -SIS dynamics (see Ref. [21], Appendix A). Alternatively, one may compute the Rayleigh-Ritz coefficient of the partial differential equations to derive bounds for the eigenvalues. There are also several approaches to compute the eigenvalues of BDPs exactly. One possibility is to consider the orthogonal polynomials that correspond to the BDPs, which are the Tricomi-Carlitz polynomials. The zeros of the Tricomi-Carlitz polynomials correspond to the eigenvalues of the BDP. Unfortunately, not many results are known for the zeros of the Tricomi-Carlitz polynomials [22]. Alternatively, one may

solve for the eigenvalues directly, resulting in a continued fraction expansion [23], or one can also derive bounds on the eigenvalues, e.g., by the Cauchy interlacing theorem or using a Fokker-Planck approximation [24].

The eigenvalues of Markov chains and their relation to metastability have been studied in several works. Artajelo [25] studied the second-largest eigenvalue in general Markov chains in both continuous time and discrete time. Holme and Tupikina [26] computed the exact second-largest eigenvalue in SIS epidemics for all nonisomorphic graphs with  $3 \le N \le 8$  nodes. For an arbitrary graph size *N*, exact results of the  $\varepsilon$ -SIS model can be obtained only for a few graphs, such as the complete graph and the star graph [9]. For the complete graph with homogeneous transition rates and no self-infections, several analytical results have been obtained, such as the average time before extinction [24,27–29] and the average time between the onset of the disease and the arrival at the metastable state [30,31].

In this work, we study continuous-time Markovian  $\varepsilon$ -SIS epidemics on the complete graph from an eigenvalue perspective by computing all its eigenvalues. Although we realize that the complete graph is far from a realistic setting, we derive many qualitative properties of the  $\varepsilon$ -SIS process which, we believe, may also hold for other graphs. We describe the continuous-time Markov chain for the complete graph in Sec. II. We additionally derive the general solution of the Markov chain. We introduce the concept of metastability in Sec. III, and we derive the exact mean-field solution in Sec. IV. We return to the Markovian  $\varepsilon$ -SIS process in Sec. V and provide a detailed analysis of all eigenvalues. We numerically identify the influence of the infection rate  $\beta$ , self-infection rate  $\varepsilon$ , and network size N on the eigenvalues in Sec. VI. Finally, we present conclusions in Sec. VII.

#### II. THE ε-SIS PROCESS ON THE COMPLETE GRAPH

The Markov chain of the  $\varepsilon$ -SIS process on the complete graph can be described as follows. In the Markov chain  $\mathcal{M}$ , each state M denotes the number of infected individuals in the population. Since the population consists of N individuals, the number M of infected nodes ranges from zero to N. Thus, the Markov chain  $\mathcal{M}$  has N + 1 states with the transition rates (derived in Ref. [32], p. 474)

$$M \mapsto M + 1$$
 at rate  $(\beta M + \varepsilon)(N - M)$ ,  
 $M \mapsto M - 1$  at rate  $\delta M$ ,

where  $\beta$  denotes the infection rate,  $\delta$  is the curing rate, and  $\varepsilon$  is the self-infection rate in the complete graph  $K_N$ . The Markov chain  $\mathcal{M}$  is a birth and death process with birth rate  $\lambda_k = (\beta k + \varepsilon)(N - k)$  that is quadratic in k, and death rate  $\mu_k = \delta k$  that is linear in k and is drawn in Fig. 1.

We compute the probability that *k* nodes are infected at time *t* in the Markov chain  $\mathcal{M}$ . Let M(t) be the number of infected nodes at time *t*. By introducing  $s_k(t) = \Pr[M(t) = k]$  as the probability that the number M(t) of infected nodes at time *t* equals *k*, the following differential equations describe



FIG. 1. The Markov chain  $\mathcal{M}$  of the  $\varepsilon$ -SIS process on the complete graph is a finite birth and death process with birth rate  $\lambda_k = (\beta k + \varepsilon)(N - k)$  and death rate  $\mu_k = \delta k$ . The state M = 0, ..., N of the Markov chain denotes the number of infected nodes in the graph. Furthermore,  $\beta$  is the infection rate,  $\delta$  is the curing rate, and  $\varepsilon$  is the self-infection rate.

the exact dynamics of the Markov chain  $\mathcal{M}$ :

$$\begin{aligned} \frac{d \, s_0}{dt} &= \mu_1 s_1(t) - \lambda_0 s_0(t), \\ \frac{d \, s_k}{dt} &= -(\lambda_k + \mu_k) s_k(t) + \lambda_{k-1} s_{k-1}(t) + \mu_{k+1} s_{k+1}(t), \\ k &= 1, \dots, N-1. \\ \frac{d \, s_N}{dt} &= \lambda_{N-1} s_{N-1}(t) - \mu_N s_N(t), \end{aligned}$$

define the scaled birth rate  $\tilde{\lambda}_k = \lambda_k/\delta = (\tau k + \varepsilon^*)(N - k)$ and the scaled death rate  $\tilde{\mu}_k = \mu_k/\delta = k$ , where the effective infection rate  $\tau = \beta/\delta$  and the effective self-infection rate  $\varepsilon^* = \varepsilon/\delta$ . We emphasize that the scaled time  $\tilde{t}$ , the effective infection rate  $\tau$ , and the effective self-infection rate  $\varepsilon^*$  are dimensionless variables. Introducing the  $(N + 1) \times 1$ vector  $\mathbf{s} = (s_0, \dots, s_N)^T$ , the linear differential equations can be written in matrix notation:

$$\frac{d\,\mathbf{s}^T}{d\tilde{t}} = \mathbf{s}^T P,\tag{1}$$

If the curing rate  $\delta > 0$ , the  $\varepsilon$ -SIS process can be simplified by introducing the rescaled time  $\tilde{t} = \delta t$ . We additionally

$$P = \begin{pmatrix} -\tilde{\lambda}_0 & \tilde{\lambda}_0 \\ \tilde{\mu}_1 & -(\tilde{\lambda}_1 + \tilde{\mu}_1) & \tilde{\lambda}_1 \\ \tilde{\mu}_2 & -(\tilde{\lambda}_2 + \tilde{\mu}_2) \end{pmatrix}$$

Together with the initial condition  $\mathbf{s}_0 = \mathbf{s}(0)$ , Eq. (1) describes the exact dynamics of the continuous-time Markovian  $\varepsilon$ -SIS process on the complete graph  $K_N$ , which can be solved to find

$$\mathbf{s}^{T}(\tilde{t}) = \mathbf{s}_{0}^{T} e^{P\tilde{t}}.$$
(3)

Using the probability vector of the number of infected nodes  $\mathbf{s}(\tilde{t})$  at time  $\tilde{t}$ , one may compute the average fraction of infected cases  $y(\tilde{t})$ , commonly known as the *prevalence*, as

$$y(\tilde{t}) = \frac{1}{N} \sum_{k=0}^{N} k s_k(\tilde{t}).$$
 (4)

In the remainder of this work, we omit the tilde for the scaled time  $\tilde{t}$  for readability.

We intend to show that metastability in the  $\varepsilon$ -SIS process is directly linked to the eigenvalues of the transition matrix *P*. We start by denoting the eigenvalues  $\xi_1, \ldots, \xi_{N+1}$ , the right-eigenvectors  $\mathbf{v}_1, \ldots, \mathbf{v}_{N+1}$ , and the left-eigenvectors  $\mathbf{w}_1, \ldots, \mathbf{w}_{N+1}$  of the  $(N + 1) \times (N + 1)$  transition matrix *P*. The eigenvalues of the transition matrix *P* are real, because *P* 

where *P* is the  $(N + 1) \times (N + 1)$  tridiagonal *transition matrix*.

$$\begin{array}{cccc} \ddots & & & \\ \ddots & & & \\ \ddots & & & \tilde{\lambda}_{N-2} \\ & \tilde{\mu}_{N-1} & -(\tilde{\lambda}_{N-1} + \tilde{\mu}_{N-1}) & \tilde{\lambda}_{N-1} \\ & & \tilde{\mu}_N & -\tilde{\mu}_N \end{array} \right).$$

$$(2)$$

is similar to a symmetric matrix  $\tilde{P}$  and similarity preserves the eigenvalues. The transformed matrix  $\tilde{P}$  is computed in Appendix A. The eigenvalues of *P* and  $\tilde{P}$  cannot be computed analytically for N > 4 because it involves finding the roots of a characteristic polynomial with degree *N*. Thus one resorts to numerical methods to obtain the eigenvalues. Once the eigenvalues are known, the corresponding eigenvectors can be computed analytically (Ref. [32], Appendix A.6.3).

Since all eigenvalues are real-valued, we may rank them in decreasing order  $\xi_1 \ge \xi_2 \ge \cdots \ge \xi_{N+1}$ . Given that the tridiagonal matrix  $\tilde{P}$  is symmetric and all off-diagonal terms are nonzero,<sup>1</sup> all eigenvalues of *P* are distinct (Ref. [33], Lemma 7.7.1). The same conclusion follows by computing the probability generating function of the  $\varepsilon$ -SIS process and concluding that the resulting differential equation is of Sturm-Liouville type [21], which is known to have simple eigenvalues. The

<sup>&</sup>lt;sup>1</sup>Here, we assume that the curing rate  $\delta$ , the effective infection rate  $\tau$  and the effective self-infection rate  $\varepsilon^*$  are nonzero.

transition matrix *P* of the  $\varepsilon$ -SIS Markov chain has a unique, largest eigenvalue  $\xi_1 = 0$ , which corresponds to the steady state  $\pi$ . The remaining eigenvalues  $\xi_2, \ldots, \xi_{N+1}$  are negative and distinct. The solution (3) can be written as

$$\mathbf{s}(t) = \pi + \sum_{k=2}^{N+1} c_k e^{\xi_k t} \mathbf{w}_k,$$
(5)

where  $\xi_k$  is the eigenvalue corresponding to the righteigenvector  $\mathbf{w}_k$  of the  $\varepsilon$ -SIS process, and the constant  $c_k =$  $\mathbf{v}_k^T \mathbf{s}(0)$  projects the initial vector  $\mathbf{s}(0)$  on the kth lefteigenvector  $\mathbf{v}_k$  [Ref. [11], Eq. (2.5)]. The vector  $\mathbf{s}(t)$ , whose components  $s_k(t)$  specify the probability that k nodes at time t are infected, is decomposed in Eq. (5) into N + 1 eigenstates of which the corresponding eigenvectors  $\mathbf{w}_1, \ldots, \mathbf{w}_{N+1}$ span the N + 1 vector space. Each eigenvector  $\mathbf{w}_k$  in Eq. (5) is weighted by the coefficient  $c_k e^{\xi_k t}$ . The eigenvalue  $\xi_k$  resembles a rate and has units 1/time. The contribution of eigenvector  $\mathbf{w}_k$  to the solution  $\mathbf{s}(t)$  decays exponentially over time with a decay rate equal to the eigenvalue  $\xi_k$  (the contribution decays because  $\xi_k < 0$ ). The eigenvalue  $\xi_k$  is thus inversely proportional to the average time that the corresponding eigenvector  $\mathbf{w}_k$  significantly contributes to the solution  $\mathbf{s}(t)$ .

In particular, the second-largest eigenvalue  $\xi_2$  (sometimes called the *convergence rate*, spectral gap, mixing rate, or decay parameter) is inversely proportional to the average time required to converge toward the steady state [25,34]. The convergence rate for continuous-time Markov chains and BDPs is thoroughly analyzed in probability theory. For an overview of bounds of the convergence rate in Markov chains and BDPs, we refer the reader to Van Doorn *et al.* [35] and Artalejo [25] and references therein. If the effective self-infection rate  $\varepsilon^* = 0$ , then the average time of convergence E[ $T_{\text{extinction}}$ ] =  $-1/\xi_2$  to the steady state (or equivalently, the *extinction time*, survival time, or absorption time) on the complete graph has the following exact relationship [27]:

$$E[T_{\text{extinction}}] = \sum_{i=1}^{N} \sum_{j=0}^{i-1} \frac{(N-i+j)!}{i(N-i)!} \tau^{j}.$$

Nearly all works consider the SIS process without selfinfections. Most proofs for the convergence rate (e.g., the proof in [27]) are based on the hitting time distribution of the absorbing state. By introducing the self-infection process, the absorbing state no longer exists and the proofs therefore do not hold for the  $\varepsilon$ -SIS process. Fortunately, the introduction of the self-infection process makes the Markov chain of the  $\varepsilon$ -SIS process irreducible, implying that the steady state exists and is also unique. The existence of the steady state greatly simplifies the analysis of the  $\varepsilon$ -SIS process for large times, because we can analyze the behavior of the conceptually simpler steady state instead of the more complicated metastable state.

## III. METASTABILITY IN THE ε-SIS PROCESS

Our primary motivation for researching the eigenvalues of the transition matrix P in Eq. (2) is the observation of plateau-behavior in the  $\varepsilon$ -SIS process [21], which is illustrated



FIG. 2. Plateau-behavior in the Markovian  $\varepsilon$ -SIS process on the complete graph. Each curve is computed using the exact solution (5), and the parameters are N = 30 nodes, effective infection rate  $\tau = 2.5\tau_c^{(1)} = 2.5/(N-1)$ , the process starts with a single infected node, and the effective self-infection rate  $\varepsilon^* =$  $\{10^{-2}, 10^{-3}, 10^{-4}, 10^{-5}, 10^{-6}, 10^{-7}, 0\}$ . For  $\varepsilon^* \leq 10^{-5}$ , the background colors indicate the current phase of the  $\varepsilon$ -SIS process: (I) initial phase, (II) metastable behavior, and (III) convergence to the steady state.

in Fig. 2. Each curve in Fig. 2 is computed based on the exact solution (5), where initially a single node is infected. For small effective self-infection rates  $\varepsilon^* \leq 10^{-5}$ , Fig. 2 depicts roughly three regimes for the time-varying prevalence: (I) initial phase, (II) metastable behavior, and (III) convergence to the steady state. Phase (I) is characterized by the fast convergence to the metastable state. In the metastable phase (II), the prevalence *y* stays nearly constant for an extended period of time. Finally, phase (III) shows the exponential convergence to the steady state  $\pi$ .

Plateau-behavior is generally caused by metastability of the dynamical process, where the infection and curing processes are temporarily in equilibrium (the physical explanation is presented in [7]). We consider the following definition of metastability for general dynamical processes that possess a steady state  $\pi$ .

Definition III.1. A dynamical process is metastable if the process stays in a certain state for an extended period of time before converging to the steady state  $\pi$ .

We quantify metastability for the  $\varepsilon$ -SIS process using the eigenvalue ratio, which was first introduced by Jacquez and Simon (Ref. [8], p. 85).

Definition III.2. The eigenvalue ratio  $\rho$  is defined as

$$\rho = \frac{\xi_3}{\xi_2}.\tag{6}$$

The eigenvalue ratio  $\rho$  is an indicator for the existence of plateaus [8]. If the eigenvalue ratio  $\rho$  is large, then the second-largest eigenvalue  $\xi_2$  is much larger than the other eigenvalues  $\xi_3, \ldots, \xi_{N+1}$ . The influence of the right-eigenvectors  $\mathbf{w}_2, \ldots, \mathbf{w}_{N+1}$  on the solution  $\mathbf{s}(t)$  in (5), weighted by the exponentials  $e^{\xi_k t}$ , will therefore converge much faster to zero for the eigenvalues  $\xi_3, \ldots, \xi_{N+1}$  than for the second-largest eigenvalue  $\xi_2$ .

The height of the plateaus in Fig. 2 equals the prevalence y, which is defined in Eq. (4). Plateau-behavior, as shown in Fig. 2, is only clearly observed if the steady-state prevalence  $y_{\infty}$  is sufficiently different from the prevalence in the metastable state and if the effective self-infection rate  $\varepsilon^*$  is sufficiently small. Our definition of metastability in the  $\varepsilon$ -SIS process is as follows:

Definition III.3. The  $\varepsilon$ -SIS process is metastable if the eigenvalue ratio  $\rho \gg 1$  and the prevalence  $y_{\infty}$  in the steady state is sufficiently different from the prevalence in the metastable state.

We analyze the case when the metastable prevalence y and the steady-state prevalence  $y_{\infty}$  are equal in Appendix B. We find the power-law relation  $\varepsilon^* \sim \tau^{-N}$  if  $\tau$  is sufficiently larger than the epidemic threshold  $\tau_c$ . For further details, see Appendix B.

If the  $\varepsilon$ -SIS process is metastable, the average time spent in the metastable state is roughly equivalent to the average time required to converge to the steady state, because the average time between the onset of the disease and the arrival at the metastable state is relatively short (see Fig. 2 for an example).

## IV. MEAN-FIELD APPROXIMATION OF THE ε-SIS PROCESS

The majority of the research on epidemiology is based on mean-field approximations of stochastic processes. However, the behavior of the Markovian  $\varepsilon$ -SIS process is intrinsically different from its mean-field approximation, especially regarding the metastable state. Recently, Prasse *et al.* [36] solved the continuous-time mean-field SIS process on the complete graph  $K_N$  with arbitrary initial conditions. We derive a similar result in Theorem IV.1 for the  $\varepsilon$ -SIS process with self-loops, whereby we added the infection rates  $\tau_{ii}$  from each node *i* to itself. We refer to Appendix C for the derivation of the mean-field equations, which approximate the exact, Markovian solution s(t) by the mean-field state vector  $s_{MF}(t)$ .

Theorem IV.1. Consider the mean-field approximation of the  $\varepsilon$ -SIS process on the complete graph with homoge-

neous transition rates and self-loops, given by Eq. (C2) in Appendix C, with arbitrary initial conditions. Then the viral state  $\mathbf{s}_{MF}(t)$  is equal to

$$\mathbf{s}_{\mathrm{MF}}(t) = c_1(t)\mathbf{z}_1 + c_2(t)\mathbf{z}_2 \tag{7}$$

at every time *t*. Here,  $\mathbf{z}_1$  and  $\mathbf{z}_2$  are orthonormal agitation modes, which are given by

$$\mathbf{z}_1 = \frac{1}{\sqrt{N}} \mathbf{u},\tag{8}$$

where  $\mathbf{u} = (1, ..., 1)^T$  denotes the  $N \times 1$  all-one vector, and

$$\mathbf{z}_2 = \left\| \left( I - \mathbf{z}_1 \mathbf{z}_1^T \right) \mathbf{s}_{\mathrm{MF}}(0) \right\|_2^{-1} \left( I - \mathbf{z}_1 \mathbf{z}_1^T \right) \mathbf{s}_{\mathrm{MF}}(0).$$
(9)

The functions  $c_l(t) = \mathbf{z}_l^T \mathbf{s}_{MF}(t) \in \mathbb{R}$ , where l = 1, 2, are the projection of the viral state vector  $\mathbf{s}_{MF}(t)$  on the agitation mode  $\mathbf{z}_l$ . The scalar function  $c_1(t)$  equals

$$c_{1}(t) = \frac{1}{2\tau\sqrt{N}} \Big[ \tau N - 1 - \varepsilon^{*} + w_{\varepsilon^{*}} \tanh\left(\frac{w_{\varepsilon^{*}}}{2}t + \Upsilon_{1,\varepsilon^{*}}(c_{1}(0))\right) \Big]$$

with the viral slope

$$w_{\varepsilon^*} = \sqrt{(1 + \varepsilon^* - \tau N)^2 + 4\varepsilon^* \tau N}$$

and the constant

$$\Upsilon_{1,\varepsilon^*}(c_1(0)) = \operatorname{arctanh}\left(\frac{1}{w_{\varepsilon^*}} \left(2\tau\sqrt{N}\mathbf{z}_1^T\mathbf{s}_{\mathrm{MF}}(0) - \tau N + 1 + \varepsilon^*\right)\right).$$

The scalar function  $c_2(t)$  equals

$$c_{2}(t) = \Upsilon_{2,\varepsilon^{*}}(c_{2}(0)) \exp\left(-\frac{1+\varepsilon^{*}+\tau N}{2}t\right)$$
  
 
$$\times \operatorname{sech}\left(\frac{w_{\varepsilon^{*}}}{2}t + \Upsilon_{1,\varepsilon^{*}}(c_{1}(0))\right),$$



FIG. 3. The exact Markovian solution (solid lines), the mean-field approximation without self-loops (dashed lines), and exact mean-field solution with self-loops from Theorem IV.1 (dotted lines) of the  $\varepsilon$ -SIS process on the complete graph with N = 30 nodes for various effective self-infection rates  $\varepsilon^*$  for (a) effective infection rate  $\tau = 2.5/(N-1)$  and (b) effective infection rate  $\tau = 1/(N-1)$ . The mean-field approximation is generally a loose upper bound for the Markovian  $\varepsilon$ -SIS dynamics.



FIG. 4. All eigenvalues  $\xi_k$  of the transition matrix *P* on the complete graph with effective infection rate  $\tau = 2/(N-1)$ , and effective self-infection rate  $\varepsilon^* = 0.01/N$  on (a) linear-linear scale and (b) log-log scale for k = 1, ..., N + 1. The eigenvalues have been normalized with respect to the number of nodes *N*. Normalized index 0 corresponds to the largest eigenvalue  $\xi_1 = 0$ , and normalized index 1 corresponds to the smallest eigenvalue  $\xi_{N+1}$ . The horizontal dash-dotted line indicates the lower bound for  $\xi_{N+1}$  from Theorem V.1. The largest eigenvalue  $\xi_1 = 0$  is not shown in (b) because of the logarithmic axis. The inset in (a) zooms in for  $0 \leq (k-1)/N \leq 0.05$  and shows the critical index  $k_2$  as a vertical dash-dotted line. In this case, the critical index  $k_1 = 0$ .

with the constant

$$\Upsilon_{2,\varepsilon^*}(c_2(0)) = \mathbf{z}_2^T \mathbf{s}_{MF}(0) \cosh\left(\Upsilon_{1,\varepsilon^*}(c_1(0))\right).$$
(10)

*Proof.* See Appendix C.

As derived in Appendix C, the mean-field prevalence upperbounds the Markovian prevalence y(t). Theorem IV.1 states that the mean-field dynamics on the complete graph can be reduced from N equations to two agitation modes, where one is related to the initial condition and the other to the steady state. Hence, the metastable state [Phase (II) in Fig. 2] does not exist under the mean-field approximation, because the existence of only two agitation modes does not allow for an intermediate, transient regime.

We plot the Markovian prevalence, the mean-field prevalence and the mean-field prevalence with self-loops, based on Theorem IV.1, in Fig. 3. Figure 3 illustrates that the mean-

field approximation vastly overestimates the time-dependent fraction of infected nodes of the Markovian  $\varepsilon$ -SIS process, both with and without self-loops. If the effective infection rate  $\tau$  is larger than the epidemic threshold  $\tau_c$ , Fig. 3(a) shows that, for N = 30, the discrepancy between the mean-field prevalence and the Markovian prevalence is large everywhere, except at very small timescales or if  $\varepsilon^*$  is very large. Figure 3(b) is situated around the epidemic threshold  $\tau_c$ , where the mean-field approximation is known to have the worst accuracy [18]. In the limit  $N \to \infty$ , the mean-field approximation error converges to zero, which we further illustrate in Sec. VIC. Given that the metastable state is a key epidemiological quantity and that the metastable state does not exist under the mean-field approximation, we focus in the remainder of this work on the Markovian  $\varepsilon$ -SIS process.



FIG. 5. The critical indices  $k_1$  (a) and  $k_2$  (b) for different values of the effective infection rate  $\tau$  and the effective self-infection rate  $\varepsilon^*$  on the complete graph  $K_N$  with N = 20 nodes.

In addition to approximating the Markovian  $\varepsilon$ -SIS process by a mean-field approximation, various other approximation methods exist. We investigate an eigenmode approximation of the Markovian  $\varepsilon$ -SIS process on  $K_N$  in Appendix D. Unfortunately, for an accurate approximation, the number of eigenmodes scales proportional to the number of nodes N, rendering the approximation method infeasible for large networks.

### V. SPECTRUM ANALYSIS OF THE E-SIS PROCESS

Since the eigenvalues  $\xi_k$  are key for the characteristic timescales of the  $\varepsilon$ -SIS process, we focus on the determination of the eigenvalues  $\xi_k$  of the transition matrix *P* from Eq. (2). The eigenvalues are computed in several parameter limits in Theorem V.1.

*Theorem V.1.* The eigenvalues  $\xi_k$  of the transition matrix *P* satisfy<sup>2</sup>

$$\begin{aligned} &\text{for } \varepsilon^* \to \infty, \quad \xi_k = -(k-1)\varepsilon^* - (k-1)[\tau(N+1-k)+1] + O\left(\frac{1}{\varepsilon^*}\right), \\ &\text{for } \tau \to \infty, \quad \xi_k = \begin{cases} -\frac{1}{4}\tau(N-1)(N+1) \pm \sqrt{\tau}(N+1)\sqrt{\frac{N-1}{2}} + O(1) & \text{if } N \text{ odd}, k = \frac{N+2\pm 1}{2}, \\ -\tau\left(\frac{N}{2}-1\right)\left(\frac{N}{2}+1\right) + \frac{N^3}{12} + \frac{N^2}{12} - \frac{5N}{6} - \frac{1}{3} \\ &\pm \sqrt{1+\left(\frac{N}{2}\right)^3\left(\frac{N}{2}+1\right)^2\left(\frac{N}{2}-1\right)} + O\left(\frac{1}{\sqrt{\tau}}\right) & \text{if } N \text{ even}, k = \frac{N+2\pm 2}{2}, \varepsilon^* = 1, \\ -\tau(k-1)(N-k+1) \\ &+ \left(-\varepsilon^*(N-k+1) + \frac{(k-1)(N-k+1)(N+1)}{(2k-N-3)(2k-N-1)}\right) + O\left(\frac{1}{\sqrt{\tau}}\right) & \text{otherwise}, \end{cases} \end{aligned}$$

$$\text{for } \varepsilon^* < \frac{1}{N}, \quad \xi_{N+1} \gtrsim \begin{cases} -\left(\frac{1}{2x}+1+\frac{x}{2}\right)N & \text{for } x \ge 1, \\ -2N & \text{for } x < 1, \end{cases}$$

for k = 1, 2, ..., N + 1 and where  $x = \tau / \tau_c^{(1)}$  is the normalized effective infection rate.

Proof. See Appendix E.

Theorem V.1 states that the eigenvalue ratio  $\rho$  for large effective self-infection rates  $\varepsilon^*$  equals  $\rho = \xi_3/\xi_2 \approx 2$ . Thus metastability is not expected if the self-infection process dominates the infection and curing processes. For large but finite effective infection rates  $\tau$  and effective self-infection rates  $\varepsilon^*$ , the asymptotic expansions in Theorem V.1 are only valid if the second term is strictly smaller than the first term, the third is smaller than the second, etc. For example, the second term in the expansion  $\tau \to \infty$  must be strictly smaller than the first term, which only holds if  $\tau > N + 1$  (see Appendix E for the derivation). Hence, the expansions in Theorem V.1 provide valuable insights into the eigenvalues  $\xi_k$ , even for finite values of the effective infection rate  $\tau$  and effective self-infection rate  $\varepsilon^*$ .

We present here a full numerical analysis of the eigenvalues for finite parameter values. We compute the eigenvalues of the transition matrix *P* in MATLAB using the command eig. Figure 4 illustrates the normalized eigenvalues  $\xi_k/N$  versus the normalized index (k - 1)/N of the  $\varepsilon$ -SIS process for k = 1, ..., N + 1. An interesting observation from Fig. 4(a) is the negligible dependence of the network size *N* on the normalized eigenvalues  $\xi_k/N$ . We further investigate the influence of the network size *N* by simulations in Sec. VI. Figure 4(b) shows that the second-largest eigenvalue  $\xi_2$  deviates significantly from the other eigenvalues  $\xi_3, ..., \xi_{N+1}$ . The difference between the second-largest eigenvalue  $\xi_2$  and the third-largest eigenvalue  $\xi_3$  is the precise reason why we used the eigenvalue ratio  $\rho$  in our Definition III.3 of the metastable state in the  $\varepsilon$ -SIS process.

The inset of Fig. 4(a) shows a "staircase" for certain eigenvalues of the  $\varepsilon$ -SIS process. We propose to subdivide the eigenvalues into three regimes. We *define* the critical index  $k_1$  as the stair of the staircase, and the critical index  $k_2$  as the end of the staircase. If the staircase does not exist, we take  $k_1 = k_2 = 0$ . The main plot of Fig. 4(a) illustrates that the eigenvalues  $\xi_k$  with  $k > k_2$  roughly follow a quadratic relation between the normalized eigenvalues  $\xi_k/N$  and the normalized



FIG. 6. The normalized eigenvalues  $\xi_k/N$  on the complete graph  $K_N$  with N = 20 nodes and  $\tau = 10$  for different effective self-infection rates between  $\varepsilon^* = 10$  and  $\varepsilon^* = 20$ . By varying the effective self-infection rate  $\varepsilon^*$ , one eigenvalue of the nearly degenerate pair changes significantly, whereas the other remains largely unchanged.

<sup>&</sup>lt;sup>2</sup>Contrary to our general consensus that the eigenvalues  $0 = \xi_1 > \xi_2 > \cdots > \xi_{N+1}$  are ordered, the eigenvalues  $\xi_k$  in Theorem V.1 are not necessarily ordered.

index (k-1)/N. Theorem V.1 states that the eigenvalues  $\xi_k$  converge to  $\tau(k-1)(N-k+1)$  if the effective infection rate  $\tau$  tends to infinity, which explains the nearly quadratic form for  $k > k_2$ . On the other hand, the inset of Fig. 4(a) shows staircase behavior for  $k_1 < k \leq k_2$ . The staircase is the result of two nearly degenerate (but strictly distinct) eigenvalues of the  $\varepsilon$ -SIS process. If the effective infection rate  $\tau$  tends to infinity, the eigenvalues converge to  $\xi_k \rightarrow \tau(k-1)(N-k+1)$ , which are degenerate because  $\xi_k = \xi_{N-k+2}$  for all  $k = 1, \ldots, N + 1$ . Even for large, finite effective infection rates  $\tau$ , there is a small region  $k_1 < k \leq k_2$  where staircase behavior is observed, which is illustrated in Fig. 4.

The staircase behavior from Fig. 4(a) is not always observed. Figure 5 illustrates the dependence of the critical indices  $k_1$  and  $k_2$  on the effective infection rate  $\tau$  and the effective self-infection rate  $\varepsilon^*$ . Staircases are observed if the critical index  $k_2 > 0$  in Fig. 5, whereas the critical index  $k_1$  indicates the start of the staircase and is only nonzero for a small region in the ( $\tau$ ,  $\varepsilon^*$ )-space. If the effective infection rate  $\tau$  is below the epidemic threshold  $\tau_c$ , both indices  $k_1 = k_2 = 0$  and no staircases can be observed. In the limit  $\varepsilon^* \to \infty$ , Theorem V.1 shows that staircases do not exist, because  $\xi_k \to -(k-1)\varepsilon^*$ for k = 1, ..., N + 1. In the intermediate regime, where  $0 < \varepsilon^* < \infty$  and  $\tau > \tau_c$ , the critical index  $k_2$  is often nonzero, indicating staircase behavior. However, Fig. 5 additionally shows small blue regions, where staircase behavior is not observed. The blue regions are centered around  $\varepsilon^* = \tau (n + \frac{1}{2})$ , where n = 0, 1, 2, ... Upon further inspection, the staircases seem best visible for  $\varepsilon^* = \tau n$ , where n = 1, 2, ...

For a given number of k infected nodes, the total effective infection rate equals  $\tau k(N - k)$  and the total effective self-infection rate  $\varepsilon^*(N - k)$ . The effect of both infection processes is equally strong if its rates are equal, which implies that  $\varepsilon^* = \tau k$ . Thus, the staircases are best visible if the total rate of the infection process and self-infection process are equally large. We argue that one eigenvalue of the nearly degenerate pair is due to the infection process. Then,



FIG. 7. The eigenvalues of the transition matrix *P* for the complete graph with N = 30 nodes and effective infection rate  $\tau = 2.5\tau_c^{(1)}$  for varying effective self-infection rates  $\varepsilon^*$  on a log-log scale. Subfigure (a) shows the second-to-fifth-largest eigenvalues  $\xi_2, \xi_3, \xi_4, \xi_5$  and smallest eigenvalue  $\xi_{N+1}$  as a function of the effective self-infection rate  $\varepsilon^*$ . The vertical line indicates the relation  $\varepsilon^* = \tau$ . Subfigure (b) is equivalent to (a), but here the eigenvalue ratio  $\rho$  is colored green (light) and the area with the other eigenvalues is colored blue (dark). Finally, (c) shows a plot of the eigenvalue ratio  $\rho$  vs the effective self-infection rate  $\varepsilon^*$ .

by varying either  $\tau$  or  $\varepsilon^*$ , one of the eigenvalues of the pair must remain unchanged. Figure 6 shows the normalized eigenvalues for varying effective self-infection rates  $\varepsilon^*$  between  $\tau$  and  $2\tau$ . Indeed, while the effective self-infection rate  $\varepsilon^*$  is varied, one eigenvalue of the nearly degenerate pair changes whereas the other stays approximately constant. Thus, a plausible explanation for the staircase behavior is a balance between self-infections and infections between nodes. While Fig. 6 supports our explanation for the staircase behavior, an analytic proof remains an open research question.

To summarize, we believe that the eigenvalues  $\xi_k$  in the regime  $2 < k \leq k_1$  are related to self-infection-dominated behavior, the regime  $k_1 < k \leq k_2$  describes when the influence of the infection and self-infection process is equally strong, and the regime  $k > k_2$  pertains to the infection-dominated behavior. Equation (11) summarizes our findings for the eigenvalues  $\xi_k$  of the  $\varepsilon$ -SIS process:

 $\xi_k \approx \begin{cases} 0 & \text{for } k = 1, \\ \text{convergence rate} & \text{for } k = 2, \\ \text{roughly linear} & \text{for } 2 < k \leq k_1, \\ \text{staircases} & \text{for } k_1 < k \leq k_2, \\ \text{roughly quadratic} & \text{for } k_2 < k \leq N + 1. \end{cases}$ (11)

### VI. NUMERICAL SIMULATIONS

We perform numerical simulations to examine how the entire set of eigenvalues changes with the effective self-infection rate  $\varepsilon^*$  the effective infection rate  $\tau$  and the network size N.

#### A. The influence of the effective self-infection rate $\varepsilon^*$

First, we examine the influence of the effective selfinfection rate  $\varepsilon^*$  on the eigenvalues by fixing the network size *N* and the effective infection rate  $\tau$ . Figure 7 shows the absolute value of the eigenvalues of the transition matrix *P* for varying effective self-infection rates  $\varepsilon^*$ .

The  $\varepsilon$ -SIS process is governed by two infection processes: the infection process and the self-infection process. All events in the continuous-time Markovian *ɛ*-SIS process are independent, such that the rate to transition from a completely susceptible population to one infected node is solely governed by the self-infection process and may happen for every node independently, leading to a total rate  $\varepsilon^* N$ . Simultaneously, if a single node is infected, another node is infected with rate  $\tau(N-1)$ . If  $\varepsilon^*N < \tau(N-1)$ , the process is dominated by the infection process, whereas the process is dominated by self-infections if  $\varepsilon^* N > \tau (N-1)$ . For large network sizes N, the influence of the self-infection process and the infection process is equally large if  $\tau \approx \varepsilon^*$ . The vertical line in Fig. 7 indicates the relation  $\tau = \varepsilon^*$ . The minimum of the eigenvalue ratio  $\rho$  and the relation  $\tau = \varepsilon^*$  coincide in Fig. 7(c). At the intersection point  $\tau = \varepsilon^*$ , the minimal eigenvalue ratio  $\rho$  is approximately 1. Then both eigenvalues are approximately equal and exhibit staircases in the eigenvalue spectrum, which was discussed in detail in Sec. V.

The second-largest eigenvalue  $\xi_2$  changes significantly from  $\varepsilon^* = 0$  to  $\varepsilon^* = \tau$ , but the remaining eigenvalues  $\xi_3, \ldots, \xi_{N+1}$  stay nearly constant. To the right of the vertical line in Fig. 7, the remaining eigenvalues increase as well. In the limit of large effective self-infection rates  $\varepsilon^*$ , it holds



FIG. 8. Illustration of the explosive phase transition for effective self-infection rates  $\varepsilon^* > 0$  and no phase transition for  $\varepsilon^* = 0$  for the second-largest eigenvalue  $\xi_2$ . The remaining eigenvalues  $\xi_3, \ldots, \xi_{N+1}$  are nearly indistinguishable for different  $\varepsilon^*$ . The network size equals N = 30.

that  $\xi_k \approx -(k-1)\varepsilon^*$  according to Theorem V.1. The smallest eigenvalue  $\xi_{N+1} \approx -\varepsilon^* N$  and the second-largest eigenvalue  $\xi_2 \approx -\varepsilon^*$  are shown by dashed lines in Fig. 7. Thus, the eigenvalues  $\xi_k$  increase linearly with the effective self-infection rate  $\varepsilon^*$ , because the self-infection process dominates the other processes and drives the  $\varepsilon$ -SIS process to the all-infected state.

If the effective self-infection rate  $\varepsilon^* < \tau$ , Fig. 7 shows that the self-infection process barely influences the characteristic timescales of the  $\varepsilon$ -SIS process, which was also observed in Fig. 2. A seemingly contradictory result was recently obtained by Van Mieghem [7], who showed that the steady-state prevalence  $y_{\infty}$  exhibits an explosive phase transition at certain small effective self-infection rates  $\varepsilon^*$ . The difference is that we consider a fixed effective infection rate  $\tau$  and vary the effective self-infection rate  $\varepsilon^*$ , which contrasts [7], where the effective infection rate  $\tau$  is varied for fixed self-infection rates  $\varepsilon^* = 0$  and  $\varepsilon^* > 0$ . Performing a similar analysis as [7] on the eigenvalues, Fig. 8 illustrates that the second-largest eigenvalue  $\xi_2$  is heavily influenced by the effective self-infection rate  $\varepsilon^*$ . For any finite effective self-infection rate  $\varepsilon^* > 0$ , there exists a phase transition for some effective infection rate  $\tau$ , where the second-largest eigenvalue  $\xi_2$  converges to a constant for large effective infection rates  $\tau$ . In the limit  $\varepsilon^* \to 0$ , no such transition is observed, which is in agreement with [7]. The other eigenvalues  $\xi_3, \ldots, \xi_{N+1}$  remain largely unaffected by considering the limit  $\varepsilon^* \to 0$ .

### B. The influence of the effective infection rate $\tau$

Analogously to the effective self-infection rate  $\varepsilon^*$ , we analyze the influence of the effective infection rate  $\tau$  on the eigenvalues of the transition matrix *P* in Fig. 9. The vertical line in Fig. 9 illustrates the mean-field epidemic threshold  $\tau_c^{(1)} = \frac{1}{N-1}$ , which is slightly smaller than the true epidemic threshold [9,37]. Below the epidemic threshold  $\tau_c$ , the eigenvalue ratio  $\rho$  in Fig. 9(c) is small. Around the epidemic



FIG. 9. The eigenvalues of the transition matrix *P* for the complete graph with N = 30 nodes and effective self-infection rate  $\varepsilon^* = 10^{-5}$  for varying effective infection rates  $\tau$ . Subfigure (a) shows the second-to-fifth-largest eigenvalues  $\xi_2$ ,  $\xi_3$ ,  $\xi_4$ ,  $\xi_5$  and smallest eigenvalue  $\xi_{N+1}$  as a function of the effective infection rate  $\tau$ . The vertical line indicates the mean-field epidemic threshold  $\tau_c^{(1)}$ , which is slightly smaller than the true epidemic threshold  $\tau_c$ . Subfigure (b) is equivalent to (a), but here the eigenvalue ratio  $\rho$  is colored green (light) and the area with the other eigenvalues is colored blue (dark). Finally, (c) shows a plot of the eigenvalue ratio  $\rho$  vs the effective infection rate  $\tau$ . Finally, (a) illustrates the lower bound for the smallest eigenvalue  $\xi_{N+1}$  from Theorem V.1 by a dashed line.

threshold  $\tau_c$ , the eigenvalue ratio  $\rho$  increases rapidly, as illustrated in Fig. 9(c).

For large effective infection rates  $\tau$ , the second-largest eigenvalue  $\xi_2$  converges to  $-\varepsilon^* N$  whereas the remaining eigenvalues  $\xi_3, \ldots, \xi_{N+1}$  increase linearly with the effective infection rate  $\tau$ , which is in line with Theorem V.1. Hence, the eigenvalue ratio  $\rho$  tends to infinity if  $\tau \to \infty$  and the system is considered metastable according to our Definition III.3. For large effective infection rates  $\tau$  and starting with a nonzero number of infected nodes, the remaining nodes will be infected extremely quickly. The spreading is only slowed down by the transition from 0 to 1 infected node. The metastable state here is the all-healthy state, which takes a considerable amount of time to leave, whereafter the process converges extremely fast to the all-infected state. The convergence rate from the metastable state to the steady state, which equals minus the second-largest eigenvalue  $-\xi_2$ , exactly equals the rate to jump from 0 to 1 infected node, which is given by the scaled birth rate  $\tilde{\lambda}_0 = \varepsilon^* N$ .

The epidemic threshold  $\tau_c$  in the Markovian  $\varepsilon$ -SIS process exhibits a second-order phase transition.<sup>3</sup> Van Mieghem and Cator (Ref. [13], p. 9) derived the following relation for the epidemic threshold  $\tau_c$  for  $\varepsilon$ -SIS dynamics on the complete graph for large network sizes  $N \gg 1$  and for small effective

<sup>&</sup>lt;sup>3</sup>A first-order, abrupt phase transition at the epidemic threshold  $\tau_c$  exhibits a discontinuity in the first derivative of the steady-state prevalence  $y_{\infty}$ . A second-order, continuous phase transition is an interval [a, b] in which the behavior of the  $\varepsilon$ -SIS process gradually changes from nearly exponential die-outs to long-lived epidemic outbreaks.



FIG. 10. Several formulas for the suggested epidemic threshold  $\tau_c$  for various network sizes *N*.

self-infection rates  $\varepsilon^* < \frac{1}{N}$ :

$$\tau_{c,\infty} = \frac{1}{N+2 - 2\sqrt{N+1}}.$$
 (12)

We expand both approximations for the epidemic threshold  $\tau_c$  for large network sizes  $N \gg 1$ :

$$\begin{aligned} \tau_c^{(1)} &= \frac{1}{N-1} = \frac{1}{N} \bigg[ 1 + \frac{1}{N} + O\bigg(\frac{1}{N^2}\bigg) \bigg], \\ \tau_{c,\infty} &= \frac{1}{N+2 - 2\sqrt{N+1}} \\ &= \frac{1}{N} \bigg[ 1 + \frac{2}{\sqrt{N}} + \frac{2}{N} - \frac{1}{N\sqrt{N}} + O\bigg(\frac{1}{N^2}\bigg) \bigg]. \end{aligned}$$

Figure 9 illustrates that the eigenvalues  $\xi_2$ ,  $\xi_3$ , and  $\xi_4$  reach a minimum at a certain infection rate  $\tau$ . We verify our hypothesis that the minimum of  $\xi_2$ ,  $\xi_3$ , or  $\xi_4$  coincides with the true epidemic threshold by plotting the mean-field threshold  $\tau_c^{(1)}$ , Van Mieghem and Cator's threshold  $\tau_{c,\infty}$ , and the numerically obtained effective infection rates  $\tau$  for which the eigenvalues  $\xi_2$ ,  $\xi_3$  and  $\xi_4$  attain a minimum in Fig. 10. Additionally, we compute the steady-state prevalence  $y_{\infty}$  for the  $\varepsilon$ -SIS process, and we take the derivative of the steady-state prevalence  $y_{\infty}$  with respect to the effective infection rate  $\tau$ . Then the epidemic threshold follows as the effective infection rate  $\tau$ for which  $dy_{\infty}/d\tau$  is maximal.<sup>4</sup> The maximum of  $dy_{\infty}/d\tau$ indicates for which effective infection rate  $\tau$  the steady-state prevalence grows the fastest, which is presumably a good indicator of the epidemic threshold. Figure 10 depicts that the effective infection rate  $\tau$  where the minimum  $\xi_3$  is attained is very close to Van Mieghem and Cator's threshold based on the maximum of  $dy_{\infty}/d\tau$ . Interestingly, Van Mieghem and Cator's threshold  $\tau_{c,\infty}$  is always larger than the mean-field threshold  $\tau_c^{(1)}$ , but always lower than the other estimates.

Figure 11 shows the steady-state prevalence  $y_{\infty}$  and the eigenvalue ratio  $\rho$  for varying infection rates  $\tau$ . Above the epidemic threshold  $\tau_c$ , the eigenvalue ratio  $\rho$  increases



FIG. 11. The steady-state prevalence  $y_{\infty}$ , the mean-field steadystate prevalence  $y_{\infty}^{(1)}$ , and the eigenvalue ratio  $\rho$  with N = 30 nodes and effective self-infection rate  $\varepsilon^* = 10^{-5}$  for varying effective infection rate  $\tau$ . The vertical line indicates the mean-field epidemic threshold  $\tau_c^{(1)} = \frac{1}{N-1}$ , which is slightly smaller than the true epidemic threshold  $\tau_c$ .

significantly and the time-dependent  $\varepsilon$ -SIS process will show metastable behavior. Below the epidemic threshold  $\tau_c$ , metastability is never observed because the eigenvalue ratio  $\rho$  is small. We conclude here that the epidemic threshold  $\tau_c$ not only describes for which infection rates  $\tau$  the epidemic persists or dies out, but it is additionally a good descriptor of whether the  $\varepsilon$ -SIS process exhibits metastable behavior.

For noncomplete graphs, we expect that a similar conclusion will hold, but the  $2^N$ -sized Markov chain for general graphs prevents us from rigorously demonstrating this claim. For connected Erdös-Rényi graphs with  $N \leq 12$  nodes and a randomly chosen link-connectivity p, simulations indicate that the eigenvalue ratio  $\rho$  and the steady-state prevalence  $y_{\infty}$  show a similar plot as for the complete graph in Fig. 11.

#### C. The influence of the network size N

Finally, we investigate the effect of the network size N on the  $\varepsilon$ -SIS process. In the limit  $N \to \infty$ , the Markovian  $\varepsilon$ -SIS process exactly converges to the mean-field approximation [38]. Figure 12 shows the steady-state prevalence  $y_{\infty}$  for various network sizes and the mean-field prevalence  $y_{\infty}^{(1)} = 1 - 1/x$ . By increasing the network size N, the steady-state prevalence  $y_{\infty}$  converges to the mean-field approximation. One of the methods in Sec. VI B to estimate the epidemic threshold  $\tau_c$ is based on the computation of the derivative  $dy_{\infty}/d\tau$ , which is plotted in Fig. 12(b). We estimate the epidemic threshold  $\tau_c$  based on the peak of  $dy_{\infty}/d\tau$ , which converges to the mean-field threshold x = 1 if the network size N increases to infinity. All curves have a  $1/x^2$  tail, which agrees with the mean-field steady-state prevalence  $dy_{\infty}^{(1)}/d\tau = 1/x^2$ .

We further investigate the eigenvalues of the  $\varepsilon$ -SIS process in Fig. 13, where we scale the effective infection rate  $\tau/N$ and the effective self-infection rate  $\varepsilon^*/N$ . If the effective infection rate  $\tau = 1.5\tau_c^{(1)}$  is larger than the epidemic threshold  $\tau_c$ , as shown in Fig. 13(a), then by fixing k, all eigenvalues  $\xi_k$  converge to a constant value in the limit  $N \to \infty$ . For

<sup>&</sup>lt;sup>4</sup>We further investigate the maximum of  $dy_{\infty}/d\tau$  as a function of the network size N in Sec. VIC.



FIG. 12. The normalized effective infection rate  $x = \tau/\tau_c^{(1)}$  vs the steady-state prevalence  $y_{\infty}$  for varying network sizes N and the mean-field steady-state prevalence  $y_{\infty}^{(1)} = 1 - \frac{1}{r}$  and  $\varepsilon^* = 0.01/N$ .

example, the second-largest eigenvalue  $\xi_2$  stays nearly constant for  $N \gtrsim 100$ . Closer to the epidemic threshold, namely for  $\tau = 1.2\tau_c^{(1)}$  in Fig. 13(b), convergence occurs near  $N \gtrsim 600$ .

The mean-field approximation of the  $\varepsilon$ -SIS process is often used in network theory to reduce the computational complexity of the 2<sup>N</sup>-sized Markov chain. In the limit of the network size N approaching infinity, the meanfield approximation becomes the exact solution, at least for the complete graph [38]. In practice, however, finitesized networks are also approximated by mean-field methods, introducing an approximation error. It is known that meanfield methods perform poorly around the epidemic threshold, because the assumed independence of the stochastic variables does not hold. We illustrate here that around the epidemic threshold, the number of nodes N in the complete graph required to obtain a reasonable accuracy with the mean-field method tends to increase closer to the epidemic threshold. We illustrate the aforementioned statement by focusing on the second-largest eigenvalue  $\xi_2$  of the transition matrix *P*. We define the *critical network size*  $N_c$  as the smallest network size *N* for which the second-largest eigenvalue  $\xi_2$  has a relative error of less than  $10^{-6}$  compared to  $\xi_2$  in the thermodynamic limit. Figure 14 exemplifies that the critical network size  $N_c$ increases if the epidemic threshold is approached from above.

### VII. CONCLUSION

In this paper, we analyzed the continuous-time Markovian  $\varepsilon$ -SIS process on the complete graph with N nodes. The transition matrix corresponding to the underlying Markov chain has N + 1 distinct eigenvalues, of which the largest eigenvalue is zero and corresponds to the steady state and the remaining eigenvalues are all negative. Metastable behavior can be observed in the  $\varepsilon$ -SIS process if the ratio between the second-largest and third-largest eigenvalue of the transition matrix is sufficiently large. The remaining eigenvalues are nearly



FIG. 13. The eigenvalues of the transition matrix *P* for the complete graph with effective self-infection rate  $\varepsilon^* = 0.01/N$ . Both subfigures show the second-to-fifth-largest eigenvalues  $\xi_2, \xi_3, \xi_4, \xi_5$  and smallest eigenvalue  $\xi_{N+1}$  as a function of the network size *N*. Subfigure (a) corresponds to effective infection rate  $\tau = 1.5\tau_c^{(1)} = 1.5/(N-1)$  and (b) to  $\tau = 1.2\tau_c^{(1)} = 1.2/(N-1)$ .



FIG. 14. The critical network size  $N_c$  as a function of the normalized effective infection rate  $x = \tau/\tau_c^{(1)}$ . The analysis is conducted on the complete graph with effective self-infection rate  $\varepsilon^* = 0.01/N$ . The critical network size  $N_c$ , which indicates when the  $\varepsilon$ -SIS process can be safely approximated using a mean-field method, increases if the normalized effective infection rate x approaches the epidemic threshold  $x_c = 1$ .

degenerate for large effective infection rates, which results in staircases in the eigenvalue plot. The staircases are particularly related to the complete graph, and they are best visible if  $\varepsilon^* = n\tau$ , where n = 1, 2, ... The epidemic threshold can be accurately estimated using the effective infection rate for which the third-largest eigenvalue of the transition matrix is the smallest. We additionally showed that the epidemic threshold not only distinguishes between small and large epidemic outbreaks, but additionally describes when the  $\varepsilon$ -SIS process may exhibit metastable behavior.

Even though we confined ourselves in this paper to the complete graph, we believe that some of our results are physical characteristics of  $\varepsilon$ -SIS dynamics and may hold for any graph. Monte Carlo simulations show metastable behavior for the  $\varepsilon$ -SIS process on general graphs [21], in non-Markovian SIS epidemics [39], and also in SIS processes on time-varying or adaptive networks [40]. Therefore, our results on the complete graph may describe the general physical behavior of metastability in  $\varepsilon$ -SIS dynamics on networks.

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## APPENDIX A: SIMPLIFYING THE TRANSITION MATRIX P

To prove that the (asymmetric) transition matrix P has real eigenvalues, we show that P is similar to a symmetric matrix. Here, we follow Ref. [32] (Appendix A.6.3). We omit the tildes on the scaled birth rate  $\lambda_k$  and the scaled death rate  $\mu_k$  in this Appendix for readability. Let H be a diagonal matrix with values  $(h_1, h_2, \ldots, h_{N+1})$  on the diagonal. Then the similarity transform  $\tilde{P} = HPH^{-1}$  equals

$$\tilde{P} = \begin{pmatrix} -\lambda_0 & \lambda_0 \frac{h_1}{h_2} \\ \mu_1 \frac{h_2}{h_1} & -(\lambda_1 + \mu_1) & \lambda_1 \frac{h_2}{h_3} \\ \mu_2 \frac{h_3}{h_2} & -(\lambda_2 + \mu_2) & \ddots \\ & & \ddots \\ & & \ddots \\ & & & \mu_{N-2} \frac{h_N}{h_{N-1}} & -(\lambda_{N-1} + \mu_{N-1}) & \lambda_{N-1} \frac{h_N}{h_{N+1}} \\ & & & \mu_{N-1} \frac{h_{N+1}}{h_N} & -\mu_N \end{pmatrix}.$$
(A1)

To assure the symmetry  $\tilde{P} = \tilde{P}^T$ , we require that  $\tilde{P}_{ij} = \tilde{P}_{ji}$  for all  $1 \le i, j \le N + 1$ . Then we find the following condition:

$$\frac{h_{k+1}}{h_k}\lambda_k = \frac{h_k}{h_{k+1}}\mu_{k-1}$$
(A2)

for all k = 2, ..., N. Solving recurrence relation (A2) and taking  $h_1 = 1$ , we find that

$$h_k = \sqrt{\prod_{i=1}^{k-1} \frac{\mu_{i-1}}{\lambda_i}}$$

which implies that  $\tilde{P} = HPH^{-1}$  is a symmetric matrix. The transformed matrix  $\tilde{P}$  equals

$$\tilde{P} = \begin{pmatrix} -\lambda_0 & \sqrt{\lambda_0 \mu_1} \\ \sqrt{\lambda_0 \mu_1} & -(\lambda_1 + \mu_1) & \sqrt{\lambda_1 \mu_2} \\ & \sqrt{\lambda_1 \mu_2} & -(\lambda_2 + \mu_2) & \ddots \\ & & \ddots \\ & & \ddots \\ & & & \sqrt{\lambda_{N-2} \mu_{N-1}} & -(\lambda_{N-1} + \mu_{N-1}) & \sqrt{\lambda_{N-1} \mu_N} \\ & & & \sqrt{\lambda_{N-1} \mu_N} & -\mu_N \end{pmatrix}$$

The transformed matrix  $\tilde{P}$  is symmetric and has therefore real eigenvalues. Since the similarity transform preserves the eigenvalues, we conclude that *P* has real eigenvalues.

For the numerical computation of the eigenvalues of the transition matrix P, the best practice is to consider the transformed matrix  $\tilde{P}$  instead of the original matrix P, because symmetric matrices have favorable properties for most eigenvalue algorithms [41]. We further improve the numerical procedure by removing the steady state  $\pi$  corresponding to the eigenvalue  $\xi = 0$ . Here, we follow Ref. [35] (Secs. 3 and 4). Given the transition matrix P, we define  $\bar{P}$  as the reduced  $N \times N$ -matrix with entries  $\bar{p}_{ij} = p_{ji} - p_{0i}$  for i, j = 1, 2, ..., N. We now prove that the eigenvalues of  $\bar{P}$  are equal to the eigenvalues of P, except for the removed eigenvalue zero. Let  $\xi$  be an eigenvalue of P, and let  $\mathbf{v} = (v_0, v_1, ..., v_N)$  be the corresponding eigenvector. Furthermore, we denote  $\mathbf{u}$  as the  $(N + 1) \times 1$  all-ones vector. Since  $P\mathbf{u} = \mathbf{0}$ , it follows that

$$\bar{P}^T(\bar{\mathbf{v}} - v_0\bar{\mathbf{u}}) = \xi(\bar{\mathbf{v}} - v_0\bar{\mathbf{u}}),$$

where  $\bar{\mathbf{v}} = (v_1, \dots, v_N)$  and  $\bar{\mathbf{u}}$  is the  $N \times 1$  all-ones vector. Thus  $\xi$  is also an eigenvalue of  $\bar{P}$ , unless  $\bar{\mathbf{v}} - v_0 \bar{\mathbf{u}} = \mathbf{0}$ , which is only true if  $\bar{\mathbf{v}}$  is constant, thus corresponding to  $\xi = 0$ . Hence, the eigenvalues of  $\bar{P}$  equal the eigenvalues of P, and the matrix  $\bar{P}$  equals

$$\bar{P} = \begin{pmatrix} -(\lambda_0 + \lambda_1 + \mu_1) & \mu_2 - \lambda_0 & -\lambda_0 & \dots & -\lambda_0 & -\lambda_0 & -\lambda_0 \\ \lambda_1 & -(\lambda_2 + \mu_2) & \mu_3 & & & \\ & \ddots & & & \\ & & \ddots & & \\ & & & \lambda_{N-2} & -(\lambda_{N-1} + \mu_{N-1}) & \mu_N \\ & & & & \lambda_{N-1} & -\mu_N \end{pmatrix}.$$

The reduced matrix  $\overline{P}$  can be transformed into a tridiagonal matrix. Let T be the upper triangular matrix with ones at and above the diagonal, and the remaining terms are zero. Then

$$T = \begin{pmatrix} 1 & 1 & 1 & \cdots & 1 & 1 \\ 1 & 1 & \cdots & 1 & 1 \\ & \ddots & & \vdots & \vdots \\ & & & \ddots & \vdots & \vdots \\ & & & & 1 & 1 \\ & & & & & 1 \end{pmatrix}, \quad T^{-1} = \begin{pmatrix} 1 & -1 & & & \\ & 1 & -1 & & \\ & & \ddots & \ddots & \\ & & & 1 & -1 \\ & & & & & 1 \end{pmatrix},$$

such that the reduced, transformed matrix  $\overline{P} = T\overline{P}T^{-1}$  equals

The reduced, transformed matrix  $\overline{P}$  is asymmetric and can be converted to a symmetric matrix  $\overline{P}$  using the same transformation matrix H as for the original transition matrix P. The main advantages of the reduced, transformed, symmetric matrix  $\overline{P}$  compared to the original transition matrix P are that numerical methods to obtain eigenvalues (i) are more efficient because the matrix is symmetric, (ii) prevent complex eigenvalues due to symmetry, and (iii) are less prone to rounding errors as the zero eigenvalue is removed.



FIG. 15. Illustration of the effective self-infection rate  $\varepsilon^*$  for which the metastable prevalence *y* and the steady-state prevalence  $y_{\infty}$  are equal. The power-law tail is observed for  $\tau > 0.3 \approx 9\tau_c^{(1)}$  and the results are obtained on a graph with N = 30 nodes.

## APPENDIX B: EQUAL METASTABLE AND STEADY-STATE PREVALENCE

The time-dependent behavior of the  $\varepsilon$ -SIS process, which is illustrated in Fig. 2, shows that the final steady-state prevalence  $y_{\infty}$  can be lower or higher than the prevalence y in the metastable state. For a given network size N and effective infection rate  $\tau$ , we believe that there always exists some  $\varepsilon^* > 0$  for which the prevalence in the metastable state and the steady-state prevalence  $y_{\infty}$  are equal. We numerically determine this self-infection rate  $\varepsilon^*$  and depict the result in Fig. 15. We only show effective infection rates  $\tau$  that are larger than the epidemic threshold  $\tau_c$ , because the metastable state does not exist below the threshold. If the effective infection rate  $\tau$  is sufficiently large, Fig. 15 depicts a power-law decay with exponent  $\alpha = -27.6$ .

We find an explanation for this result in Theorem 4 by Van Mieghem [7], which states that the epidemic threshold  $\tau_c$  for sufficiently small  $\varepsilon^*$  is bounded by

$$\frac{1}{e} \left( \frac{10^{-s}}{\varepsilon^* (N-1)!} \right)^{\frac{1}{N-1}} < \tau_c < \left( \frac{10^{-s}}{\varepsilon^* (N-1)!} \right)^{\frac{1}{N-1}},$$

where  $10^{-s}$  specifies an agreed fraction of infected nodes that determines whether an outbreak has taken place. Considering that the final steady-state prevalence  $y_{\infty}$  is of order 1, we choose s = 0. Naturally, the equation can be rewritten in terms of the effective self-infection rate  $\varepsilon^*$ :

$$\frac{1}{e^{1-N}(N-1)!}\tau^{1-N} \leqslant \varepsilon^* \leqslant \frac{1}{(N-1)!}\tau^{1-N}.$$
 (B1)

Considering the example in Fig. 15 with N = 30 nodes, our estimation (B1) states that

$$10^{-19}\tau^{-29} \leq \varepsilon^* \leq 10^{-31}\tau^{-29}.$$

The estimation is very similar to the fitted function from Fig. 15, but for larger effective self-infection rates  $\varepsilon^*$ , the approximated solution (B1) starts to deviate.

### APPENDIX C: MEAN-FIELD ε-SIS

The well-known N-intertwined mean-field approximation (NIMFA) [16] assumes that any two stochastic variables X and Y are uncorrelated: E[XY] = E[X]E[Y]. Applying the mean-field approximation to the  $\varepsilon$ -SIS process on the complete graph, the governing equations become (Ref. [32], p. 462)

$$\frac{d\mathbf{s}_{\mathrm{MF}}(t)}{dt} = \varepsilon \mathbf{u} - (\delta + \varepsilon)\mathbf{s}_{\mathrm{MF}}(t) + \mathrm{diag}\left[\mathbf{u} - \mathbf{s}_{\mathrm{MF}}(t)\right]\tilde{B}\mathbf{s}_{\mathrm{MF}}(t),$$
(C1)

where  $\mathbf{s}_{MF}(t) = (s_1(t), \dots, s_N(t))^T$  is the  $N \times 1$  viral state vector,  $\mathbf{u} = (1, \dots, 1)^T$  is the  $N \times 1$  all-ones vector, and  $\tilde{B} = \beta(\mathbf{u}\mathbf{u}^T - I)$ , where I is the  $N \times N$  identity matrix. As an upper bound to Eq. (C1), we consider the  $\varepsilon$ -NIMFA process with self-loops, thus  $\beta_{ii} = \beta > 0$ . Furthermore, we rescale time  $\tilde{t} = \delta t$ , and introducing the effective infection rate  $\tau = \beta/\delta$ and the effective self-infection rate  $\varepsilon^* = \varepsilon/\delta$ , we find, after dropping the tildes,

$$\frac{d\mathbf{s}_{\mathrm{MF}}(t)}{dt} = \varepsilon^* \mathbf{u} - (1 + \varepsilon^*) \mathbf{s}_{\mathrm{MF}}(t) + \mathrm{diag} \left[\mathbf{u} - \mathbf{s}_{\mathrm{MF}}(t)\right] B\mathbf{s}_{\mathrm{MF}}(t),$$
(C2)

where  $B = \tau \mathbf{u} \mathbf{u}^T$ . We adopt the approach in [36] to obtain the solution of the  $\varepsilon$ -SIS process on the complete graph. Since the Markovian  $\varepsilon$ -SIS process is non-negatively correlated [42] and we introduced self-loops  $\beta_{ii} > 0$ , the solution of (C2) is a strict upper bound for Markovian  $\varepsilon$ -SIS dynamics.

*Proof of Theorem IV.1.* We prove Theorem IV.1 in three steps. First, in Appendix C1, we show that the viral state  $\mathbf{s}_{MF}(t)$  is in a two-dimensional subspace at every time *t*. More specifically, we show that  $\mathbf{s}_{MF}(t) = c_1(t)\mathbf{z}_1 + c_2(t)\mathbf{z}_2$  for two  $N \times 1$  agitation modes  $\mathbf{z}_1, \mathbf{z}_2$  and two scalar functions  $c_1(t), c_2(t) \in \mathbb{R}$ . Second, in Appendix C2, we obtain the closed-form expression for the function  $c_1(t)$ . Third, given the function  $c_1(t)$ , we obtain the function  $c_2(t)$  in Appendix C3.

### 1. The viral state is in a two-dimensional subspace

With the definition of the agitation mode  $z_1$  in (8), we can write the infection rate matrix *B* as

$$B = \tau N \mathbf{z}_1 \mathbf{z}_1^T$$
.

Thus,  $\varepsilon$ -NIMFA on the complete graph (C2) becomes

$$\frac{d\mathbf{s}_{MF}(t)}{dt} = \varepsilon^* \sqrt{N} \mathbf{z}_1 - (1 + \varepsilon^*) \mathbf{s}_{MF}(t) + \text{diag}(\sqrt{N} \mathbf{z}_1) - \mathbf{s}_{MF}(t) \tau N \mathbf{z}_1 \mathbf{z}_1^T \mathbf{s}_{MF}(t).$$

Suppose that (7) holds at time t. Then, we obtain that

$$\frac{d\mathbf{s}_{\text{MF}}(t)}{dt} = \varepsilon^* \sqrt{N} \mathbf{z}_1 - (1 + \varepsilon^*) c_1(t) \mathbf{z}_1 - (1 + \varepsilon^*) c_2(t) \mathbf{z}_2$$
$$+ \operatorname{diag}[(\sqrt{N} - c_1(t)) \mathbf{z}_1 - c_2(t) \mathbf{z}_2] \tau N \mathbf{z}_1 \mathbf{z}_1^T [c_1(t) \mathbf{z}_1 + c_2(t) \mathbf{z}_2].$$

Since  $\mathbf{z}_1^T \mathbf{z}_1 = 1$  and  $\mathbf{z}_1^T \mathbf{z}_2 = 0$ , it follows that

$$\frac{d\mathbf{s}_{\mathrm{MF}}(t)}{dt} = (\varepsilon^* \sqrt{N} - (1 + \varepsilon^*)c_1(t))\mathbf{z}_1 - (1 + \varepsilon^*)c_2(t)\mathbf{z}_2 + \tau Nc_1(t)\operatorname{diag}[(\sqrt{N} - c_1(t))\mathbf{z}_1 - c_2(t)\mathbf{z}_2]\mathbf{z}_1$$

which is equivalent to

$$\begin{aligned} \frac{d\mathbf{s}_{\mathrm{MF}}(t)}{dt} &= (\varepsilon^* \sqrt{N} - (1 + \varepsilon^*)c_1(t))\mathbf{z}_1 - (1 + \varepsilon^*)c_2(t)\mathbf{z}_2 \\ &+ \tau N \left(\sqrt{N}c_1(t) - c_1^2(t)\right) \mathrm{diag}(\mathbf{z}_1)\mathbf{z}_1 \\ &- \tau N c_1(t)c_2(t) \mathrm{diag}(\mathbf{z}_2)\mathbf{z}_1. \end{aligned}$$

From the definition of the agitation mode  $\mathbf{z}_1$  in (8), we obtain that diag $(\mathbf{z}_1)\mathbf{z}_1 = \mathbf{z}_1/\sqrt{N}$  and diag $(\mathbf{z}_2)\mathbf{z}_1 = \mathbf{z}_2/\sqrt{N}$ . Thus, we arrive at

$$\frac{d\mathbf{s}_{\mathrm{MF}}(t)}{dt} = (\varepsilon^* \sqrt{N} - (1 + \varepsilon^*)c_1(t))\mathbf{z}_1 - (1 + \varepsilon^*)c_2(t)\mathbf{z}_2 + \tau \sqrt{N} (\sqrt{N}c_1(t) - c_1^2(t))\mathbf{z}_1 - \tau \sqrt{N}c_1(t)c_2(t)\mathbf{z}_2,$$

which simplifies to

$$\frac{d\mathbf{s}_{\mathrm{MF}}(t)}{dt} = \left(\varepsilon^* \sqrt{N} + (\tau N - 1 - \varepsilon^*)c_1(t) - \tau \sqrt{N}c_1^2(t)\right)\mathbf{z}_1 - (1 + \varepsilon^* + \tau \sqrt{N}c_1(t))c_2(t)\mathbf{z}_2.$$
(C3)

Hence, the  $N \times 1$  viral state vector  $\mathbf{s}_{MF}(t)$  is equal to the linear combination (7) of only two agitation modes  $\mathbf{z}_1$ ,  $\mathbf{z}_2$  at every time *t*. Thus, solving  $\varepsilon$ -NIMFA on the complete graph simplifies to obtaining a closed-form expression for the functions  $c_1(t)$  and  $c_2(t)$ .

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## 2. First agitation mode

Since

$$\frac{dc_l(t)}{dt} = \mathbf{z}_l^T \frac{d\mathbf{s}_{\rm MF}(t)}{dt}$$
(C4)

for both l = 1, 2, we obtain for the scalar function  $c_1(t)$  from (C3) that

$$\frac{dc_1(t)}{dt} = \varepsilon^* \sqrt{N} + (\tau N - 1 - \varepsilon^*)c_1(t) - \tau \sqrt{N}c_1^2(t).$$
 (C5)

The differential equation (C5) is separable,

$$\frac{dc_1(t)}{-\tau\sqrt{N}c_1^2(t) + (\tau N - 1 - \varepsilon^*)c_1(t) + \varepsilon^*\sqrt{N}} = dt.$$

Hence, it follows that

$$\frac{dc_1(t)}{c_1^2(t) + \mu_1 c_1(t) - \mu_2} = -\tau \sqrt{N} dt$$
(C6)

with the constants

$$\mu_1 = \frac{1 + \varepsilon^* - \tau N}{\tau \sqrt{N}} \tag{C7}$$

and

$$\mu_2 = \frac{\varepsilon^*}{\tau}.\tag{C8}$$

We obtain from (C6) that

$$\int \frac{dc_1(t)}{c_1^2(t) + \mu_1 c_1(t) - \mu_2} = -\tau \sqrt{Nt} + K(c_1(0))$$

for some constant  $K(c_1(0)) \in \mathbb{R}$ . By integration, it follows that

$$\frac{2}{\sqrt{-4\mu_2-\mu_1^2}}\tan^{-1}\left(\frac{\mu_1+2c_1(t)}{\sqrt{-4\mu_2-\mu_1^2}}\right) = -\tau\sqrt{N}t + K(c_1(0)),$$

which yields that

$$\frac{\mu_1 + 2c_1(t)}{\sqrt{-4\mu_2 - \mu_1^2}} = \tan\left(\frac{1}{2}\sqrt{-4\mu_2 - \mu_1^2}(-\tau\sqrt{N}t + K(c_1(0)))\right)$$

We isolate for  $c_1(t)$  and arrive at

$$c_1(t) = -\frac{1}{2}\mu_1 + \frac{1}{2}\sqrt{-4\mu_2 - \mu_1^2} \tan\left(\frac{1}{2}\sqrt{-4\mu_2 - \mu_1^2}\left(-\tau\sqrt{N}t + K(c_1(0))\right)\right).$$

With the imaginary unit  $i = \sqrt{-1}$ , it follows that

$$c_1(t) = -\frac{1}{2}\mu_1 + \frac{1}{2}i\sqrt{\mu_1^2 + 4\mu_2}\tan\left(\frac{1}{2}i\sqrt{\mu_1^2 + 4\mu_2}(-\tau\sqrt{N}t + K(c_1(0)))\right)$$

Finally, with the relation  $i \tan(ix) = -\tanh(x)$  of the tangent and hyperbolic tangent and with  $-\tanh(x) = \tanh(-x)$  for all  $x \in \mathbb{R}$ , we arrive at

$$c_1(t) = -\frac{1}{2}\mu_1 + \frac{1}{2}\sqrt{\mu_1^2 + 4\mu_2} \tanh\left(\frac{1}{2}\sqrt{\mu_1^2 + 4\mu_2}(\tau\sqrt{N}t - K(c_1(0)))\right).$$
(C9)

To further simplify (C9), we obtain with the definition of the constants  $\mu_1$  and  $\mu_2$  in (C7) and (C8) that

$$\sqrt{\mu_1^2 + 4\mu_2} = \sqrt{\frac{(1 + \varepsilon^* - \tau N)^2}{\tau^2 N}} + 4\frac{\varepsilon^*}{\tau},$$

which simplifies to

$$\sqrt{\mu_1^2 + 4\mu_2} = \frac{1}{\tau\sqrt{N}} w_{\varepsilon^*},$$
 (C10)

where we define the *viral slope*  $w_{\varepsilon^*}$  for  $\varepsilon$ -NIMFA as

$$w_{\varepsilon^*} = \sqrt{(1+\varepsilon^*-\tau N)^2+4\varepsilon^*\tau N}.$$

If  $\varepsilon^* = 0$ , then the viral slope  $w_{\varepsilon^*}$  equals  $w_{\varepsilon^*} = |w|$ , where w is the viral slope of the NIMFA model, defined in [43]. With (C10) and the definition of  $\mu_1$  in (C7), the function  $c_1(t)$  in (C9) becomes

$$c_1(t) = \frac{1}{2} \frac{\tau N - 1 - \varepsilon^*}{\tau \sqrt{N}} + \frac{1}{2} \frac{1}{\tau \sqrt{N}} w_{\varepsilon^*} \tanh\left(\frac{w_{\varepsilon^*}}{2} t + \Upsilon_{1,\varepsilon^*}(c_1(0))\right)$$

with the constant  $\Upsilon_{1,\varepsilon^*}(c_1(0)) = -\frac{1}{2} \frac{1}{\tau\sqrt{N}} w_{\varepsilon^*} K(c_1(0))$ . Thus, it holds that

$$c_1(t) = \frac{1}{2\tau\sqrt{N}} \bigg[ \tau N - 1 - \varepsilon^* + w_{\varepsilon^*} \tanh\bigg(\frac{w_{\varepsilon^*}}{2}t + \Upsilon_{1,\varepsilon^*}(c_1(0))\bigg)\bigg].$$
(C11)

The initial condition of the projection  $c_1(t)$  is given by  $c_1(0) = \mathbf{z}_1^T \mathbf{s}_{\text{MF}}(0)$ , which yields for the constant  $\Upsilon_{\varepsilon^*}(c_1(0))$  that

$$\Upsilon_{1,\varepsilon^*}(c_1(0)) = \operatorname{arctanh}\left(\frac{1}{w_{\varepsilon^*}} \left(2\tau \sqrt{N} \mathbf{z}_1^T \mathbf{s}_{\mathrm{MF}}(0) - \tau N + 1 + \varepsilon^*\right)\right).$$

## 3. Second agitation mode

From (C4) for l = 2, it follows with (C3) that the scalar function  $c_2(t)$  obeys

$$\frac{dc_2(t)}{dt} = -(1 + \varepsilon^* + \tau \sqrt{N}c_1(t))c_2(t).$$
(C12)

The closed-form expression for the function  $c_1(t)$  is given in (C11), and the differential equation (C12) becomes

$$\frac{dc_2(t)}{dt} = -(1+\varepsilon^*)c_2(t) - \frac{1}{2} \bigg[ \tau N - 1 - \varepsilon^* + w_{\varepsilon^*} \tanh\bigg(\frac{w_{\varepsilon^*}}{2}t + \Upsilon_{1,\varepsilon^*}(c_1(0))\bigg) \bigg] c_2(t)$$
(C13)

$$= -\frac{1+\varepsilon^*+\tau N}{2}c_2(t) - \frac{w_{\varepsilon^*}}{2} \tanh\left(\frac{w_{\varepsilon^*}}{2}t + \Upsilon_{1,\varepsilon^*}(c_1(0))\right)c_2(t).$$
(C14)

The remaining steps are similar to Ref. [36] (Proof of Theorem 4). Since

$$\frac{d\ln\left(c_2(t)\right)}{dt} = \frac{1}{c_2(t)}\frac{dc_2(t)}{dt}$$

we obtain that

$$\frac{d\ln(c_2(t))}{dt} = -\frac{1+\varepsilon^* + \tau N}{2} - \frac{w_{\varepsilon^*}}{2} \tanh\left(\frac{w_{\varepsilon^*}}{2}t + \Upsilon_{1,\varepsilon^*}(c_1(0))\right).$$
(C15)

The integral of the hyperbolic tangent equals [44]

$$\int \tanh(\xi)d\xi = \ln(\cosh(\xi)) + C$$

for an arbitrary constant  $C \in \mathbb{R}$ , where  $\cosh(\xi)$  denotes the hyperbolic cosine. Hence, we obtain with the chain rule from (C15) that

$$\ln\left(c_{2}(t)\right) = -\frac{1+\varepsilon^{*}+\tau N}{2}t - \frac{w_{\varepsilon^{*}}}{2}\frac{2}{w_{\varepsilon^{*}}}\ln\left(\cosh\left(\frac{w_{\varepsilon^{*}}}{2}t+\Upsilon_{1,\varepsilon^{*}}(c_{1}(0))\right)\right) + C$$

for some constant  $C \in \mathbb{R}$ . Thus, exponentiation yields that the function  $c_2(t)$  equals

$$c_{2}(t) = \exp(C) \exp\left(-\frac{1+\varepsilon^{*}+\tau N}{2}t\right) \left[\cosh\left(\frac{w_{\varepsilon^{*}}}{2}t+\Upsilon_{1,\varepsilon^{*}}(c_{1}(0))\right)\right]^{-1}$$
$$= \Upsilon_{2,\varepsilon^{*}}(c_{2}(0)) \exp\left(-\frac{1+\varepsilon^{*}+\tau N}{2}t\right) \operatorname{sech}\left(\frac{w_{\varepsilon^{*}}}{2}t+\Upsilon_{1,\varepsilon^{*}}(c_{1}(0))\right),$$

where we denote the hyperbolic secant by  $\operatorname{sech}(\xi) = \cosh(\xi)^{-1}$ . The constant  $\Upsilon_{2,\varepsilon^*}(c_2(0))$  follows from the initial condition  $c_2(0) = \mathbf{z}_2^T \mathbf{s}_{MF}(0)$  as

$$\Upsilon_{2,\varepsilon^*}(c_2(0)) = \mathbf{z}_2^T \mathbf{s}_{\mathrm{MF}}(0) \cosh\left(\Upsilon_{1,\varepsilon^*}(c_1(0))\right).$$
(C16)



FIG. 16. The time-varying prevalence y(t) for the complete graph with N = 30 nodes, effective infection rate  $\tau = 2.5\tau_c^{(1)}$ , and effective self-infection rate  $\varepsilon^* = 10^{-6}$ . The solid line indicates the solution (5), and the dashed lines show the approximation (D1) for various values of m.

## APPENDIX D: EIGENVALUE TRUNCATION OF THE *e*-SIS PROCESS

If the metastable state exists, Fig. 2 depicts roughly three regimes for the time-varying prevalence: (I) initial phase, (II) metastable behavior, and (III) convergence to the steady state. Since the behavior is rather limited, we expect that the time-dependent dynamics of the  $\varepsilon$ -SIS process can be accurately approximated. Apart from the largest eigenvalue  $\xi_1 = 0$  and the second-largest eigenvalue  $\xi_2$ , the majority of the eigenvalues are largely clustered. The (N + 1)-sized linear process (1) can perhaps be approximated accurately with only  $m \ll (N + 1)$  eigenvalues and eigenvectors. This methodology has been applied successfully for the mean-field SIS model around the epidemic threshold [43], but to the best of our knowledge, no results have been obtained for the Markovian SIS process.

We approximate the solution (5) by considering only the *m* largest eigenvalues and corresponding eigenvectors:

$$\tilde{\mathbf{s}}(t) = \sum_{k=1}^{m} c_k e^{\xi_k t} \mathbf{w}_k,$$
(D1)

where  $c_k = \mathbf{v}_k^T \mathbf{s}(0)$ . Approximating the exact solution (5) by the approximation (D1) introduces an error,

$$e(t) = \|\mathbf{s}(t) - \mathbf{\tilde{s}}(t)\| = \left\|\sum_{k=m+1}^{N+1} c_k e^{\xi_k t} \mathbf{w}_k\right\| \leqslant \sum_{k=m+1}^{N+1} e^{\xi_k t} \|c_k \mathbf{w}_k\|,$$

where  $\|\cdot\|$  denotes a vector norm and we used the vector inequality  $\|\mathbf{a} + \mathbf{b}\| \leq \|\mathbf{a}\| + \|\mathbf{b}\|$ . Further, since eigenvectors are normalized with norm 1,

$$e(t) \leqslant \sum_{k=m+1}^{N+1} e^{\xi_k t} \|c_k \mathbf{w}_k\| < e^{\xi_{m+1} t} \sum_{k=m+1}^{N+1} |c_k|.$$

Thus the error scales as  $e(t) = O(e^{\xi_{m+1}t})$ . Figure 16 shows the solution (5) and the approximation (D1) for various choices



FIG. 17. The relation between the critical number of eigenvalues  $m^*$  and the number of nodes N in the graph. The critical number  $m^*$  is defined as the smallest integer m for which the initial prevalence  $\tilde{y}(0)$  of the approximated solution (D1) differs at most  $10^{-3}$  from the initial prevalence y(0) of the exact solution (1).

of *m*. The original solution (5) is recovered accurately using only m = 8 from the total of 31 eigenvalues.

We define the critical number of eigenvalues  $m^*$  as the smallest m for which  $|y(0) - \tilde{y}(0)| \leq 10^{-3}$ , where y(0) and  $\tilde{y}(0)$  describe the initial prevalence of the exact solution (5) and approximated solution (D1), respectively. The critical number  $m^*$  is an integer between 1 and N + 1. Figure 17 shows an apparent linear relationship between the critical number  $m^*$  and the network size N. Even though an accurate approximation is possible, the number of required eigenvalues grows linearly with the network size N, effectively rendering the approximation method infeasible for large networks. This contrasts results for the mean-field  $\varepsilon$ -SIS process with arbitrary initial conditions but homogeneous parameters, where the number of required equations reduces to only two (see Appendix C).

### APPENDIX E: EIGENVALUE APPROXIMATIONS AND BOUNDS

This Appendix contains all proofs for the eigenvalues  $\xi_k$  of the transition matrix P in Eq. (2) in several parameters limits. Most proofs use the symmetric transition matrix  $\tilde{P}$ , which is derived in Appendix A. We start by presenting Theorem E.1.

Theorem E.1 (based on Ref. [45], pp. 366–372, and Ref. [46], pp. 303–321). Given a Hermitian matrix  $P = A + \alpha B$ , where A and B are Hermitian matrices and  $\alpha$  is a small parameter, such that the element  $a_{ij}$  is strictly larger than  $b_{ij}$ for all  $\alpha$  and all  $1 \leq i, j \leq N$ . We assume that the eigenvalues  $\xi_k^{(0)}$  and eigenvectors  $\mathbf{x}_k^{(0)}$  of the unperturbed matrix A can be computed easily. (a) If an eigenvalue  $\xi_k^{(0)}$  of *A* is *simple*, then the eigenvalue  $\xi_k$  of *P* is up to fourth order

$$\begin{split} \xi_{k} &= \xi_{k}^{(0)} + \alpha W_{kk} + \alpha^{2} \sum_{\substack{l=1\\l \neq k}}^{N} \frac{|W_{lk}|^{2}}{\xi_{k}^{(0)} - \xi_{l}^{(0)}} \\ &+ \alpha^{3} \sum_{\substack{l=1\\l \neq k}}^{N} (W_{ll} - W_{kk}) \left( \frac{W_{kl}}{\xi_{k}^{(0)} - \xi_{l}^{(0)}} \right)^{2} \\ &+ \alpha^{3} \sum_{\substack{l=1\\l \neq k}}^{N} \sum_{\substack{m=1\\m \neq k}}^{N} \frac{W_{kl} W_{lm} W_{mk}}{(\xi_{k}^{(0)} - \xi_{l}^{(0)}) (\xi_{k}^{(0)} - \xi_{m}^{(0)})} + O(\alpha^{4}), \end{split}$$

where  $W_{ij} = (\mathbf{x}_i^{(0)})^T B \mathbf{x}_j^{(0)}$ .

(b) If an eigenvalue  $\xi_k^{(0)}$  of *A* is *twofold single degenerate*,<sup>5</sup> then the eigenvalue  $\xi_k$  of *P* is up to third order

$$\xi_{k} = \xi_{k}^{(0)} + \alpha \xi_{k}^{(1)} + \alpha^{2} \sum_{\substack{l=1\\l\neq k_{1}\\l\neq k_{2}}}^{N} \frac{|W_{lk}'|^{2}}{\xi_{k}^{(0)} - \xi_{l}^{(0)}} + O(\alpha^{3}),$$

where the eigenvalues  $\xi_k^{(1)}$  are determined from the eigenvalue equation

$$\begin{pmatrix} W_{k_1,k_1} & W_{k_1,k_2} \\ W_{k_2,k_1} & W_{k_2,k_2} \end{pmatrix} \begin{pmatrix} \alpha_k \\ \beta_k \end{pmatrix} = \xi_k^{(1)} \begin{pmatrix} \alpha_k \\ \beta_k \end{pmatrix},$$
(E1)

where  $W_{ij} = (\mathbf{x}_i^{(0)})^T B \mathbf{x}_j^{(0)}$ . The indices  $k_1$  and  $k_2$  correspond to the degenerate eigenvalues  $\xi_k^{(0)} = \xi_{k_1}^{(0)} = \xi_{k_2}^{(0)}$ . The corrected zero-order eigenvectors are  $\mathbf{x}_k^{\prime(0)} = \alpha_{k_1} \mathbf{x}_{k_1}^{(0)} + \beta_{k_2} \mathbf{x}_{k_2}^{(0)}$ . The second-order correction term requires the W' matrix, which has elements  $W'_{ij} = (\mathbf{x}_i^{\prime(0)})^T B \mathbf{x}_j^{\prime(0)}$ .

(c) If an eigenvalue  $\xi_k^{(0)}$  of *A* is *twofold double degenerate*,<sup>6</sup> then the eigenvalue  $\xi_k$  of *P* is up to third order

$$\begin{split} \xi_{k} &= \xi_{k}^{(0)} + \frac{\alpha}{2} \big( W_{k_{1},k_{1}} + W_{k_{2},k_{2}} \pm \sqrt{\big( W_{k_{1},k_{1}} - W_{k_{2},k_{2}} \big)^{2} + 4W_{k_{1},k_{2}}W_{k_{2},k_{1}} \big)} \\ &+ \frac{\alpha^{2}}{2} \big( M_{k_{1},k_{1}} + M_{k_{2},k_{2}} \pm \sqrt{\big( M_{k_{1},k_{1}} - M_{k_{2},k_{2}} \big)^{2} + 4M_{k_{1},k_{2}}M_{k_{2},k_{1}} \big)} + O(\alpha^{3}), \end{split}$$

where  $W_{ij} = (\mathbf{x}_i^{(0)})^T B \mathbf{x}_i^{(0)}$  and

$$M_{ij} = \sum_{\substack{m=1 \ m \neq i \\ m \neq j}}^{N} \frac{W_{i,m}W_{m,j}}{\xi_i^{(0)} - \xi_m^{(0)}}$$

*Sketch of the proof.* Instead of providing a proof, we sketch the idea of the proof here. An actual proof can be based on [45] and/or [46].

The primary goal of this theorem is to approximate the eigenvalues of a matrix P, for which exact computations are generally infeasible. Defining a small variable  $\alpha$ , we split up the Hermitian (or here, symmetric) matrix P into  $P = A + \alpha B$ , where A is a diagonal matrix and B contains the remaining, symmetric terms. We emphasize that both A and B may contain functions of  $\alpha$ , as long as the element  $a_{ij}$  is larger than  $b_{ij}$  for all i, j in the limit  $\alpha \rightarrow 0$ . The division of P into A and B is generally not unique, but the number of choices is heavily restricted by the small parameter  $\alpha$ .

If the eigenvalue  $\xi_k$  of the matrix A is nondegenerate, one can perform an ordinary eigenvalue expansion that can be found in any textbook covering perturbation expansions of linear operators [part (a)]. Some eigenvalues may appear multiple times in the spectrum of A, which complicates the analysis. In our case, each eigenvalue exists at most twice, thus we confine ourselves to twofold degeneracy. The procedure can be easily generalized to n-fold degeneracy.

If the eigenvalues  $\xi_k^{(0)}$  are twofold-degenerate, the corresponding eigenvectors are not determined up to a scalar value, but are only known to be in the span of two vectors. We can choose the eigenvectors freely, as long as they belong to the span and are orthonormal to each other. If the eigenvalues are distinct at the first order, that is, the degeneracy is lifted at the first order [part (b)], we can determine the zeroth-order eigenvectors and use that basis as if we would perform a regular expansion as in part (a), with the exception that the summation over all terms excludes both the current index *k* as well as the index k' corresponding to the same zero-order eigenvalue  $\xi_k^{(0)} = \xi_{k'}^{(0)}$  and that we use the zeroth-order eigenvectors  $\mathbf{x}'_k^{(0)}$  instead of the original  $\mathbf{x}_k^{(0)}$ .

If the first-order correction of the eigenvalue still maintains the degeneracy, the second-order correction of the eigenvalues must be computed [part (c)]. As before, if the eigenvalues are no longer degenerate after adding the second-order correction, the zeroth-order eigenvectors can be determined. The procedure can be repeated up to higher orders, which is outside of the scope of this theorem.

## 1. The limit $e^* \to \infty$

*Proof.* The symmetric transition matrix  $\tilde{P}$  from Eq. (A1) can be rewritten as

$$\tilde{P} = \tilde{A} + \tilde{B},$$

<sup>&</sup>lt;sup>5</sup>An eigenvalue  $\xi$  of some matrix *A* is *twofold single degenerate* if the eigenvalue  $\xi$  appears twice in the spectrum of *A* and is no longer degenerate after adding the first correction term.

<sup>&</sup>lt;sup>6</sup>An eigenvalue  $\xi$  of some matrix *A* is *twofold double degenerate* if the eigenvalue  $\xi$  appears twice in the spectrum of *A* and is no longer degenerate after adding the second-order correction term.

where  $\tilde{A}$  is a diagonal matrix with elements

$$\tilde{A}_{kk} = -\varepsilon^*(N-k)$$
 for  $k = 0, 1, \dots, N$ 

and  $\tilde{B}$  is a symmetric, tridiagonal matrix with elements

$$\tilde{B}_{kk} = -\tau k(N-k) - k,$$
  
$$\tilde{B}_{k-1,k} = \sqrt{k(N-k+1)[(k-1)\tau + \varepsilon^*]}.$$

We define the small parameter  $\alpha = \frac{1}{\sqrt{\varepsilon^*}}$  (which is small, because  $\varepsilon^* \to \infty$ ), such that  $\alpha^2 \tilde{P} = A + \alpha B$ , where we defined  $A = \alpha^2 \tilde{A}$  and  $B = \alpha \tilde{B}$ . The matrix *A* is a diagonal matrix with

elements

$$A_{kk} = -(N - k)$$

and the matrix *B* is then tridiagonal with elements

$$B_{kk} = -\alpha k [\tau (N - k) + 1],$$
  
$$B_{k-1,k} = \sqrt{k(N - k + 1)[(k - 1)\tau \alpha^2 + 1]}$$

The eigenvalues of *A* are simply  $a_k = -(N - k)$  for k = 0, 1, ..., N and the corresponding eigenvector  $\mathbf{x}_k = \mathbf{e}_k$ , where  $\mathbf{e}_k$  is the  $(N + 1) \times 1$  all-zeros vector, except at entry *k*, where it is 1. Since  $\mathbf{x}_k = \mathbf{e}_k$ , we immediately find that  $W_{lk} = \mathbf{x}_l^T B \mathbf{x}_k = B_{kl}$ . In the same manner, we find  $a_k - a_l = k - l$ . Given the uniqueness of the eigenvalues  $a_k$ , we follow part (a) of Theorem E.1 to find

$$\alpha^{2}\xi_{k} = -(N-k) + \alpha B_{kk} + \alpha^{2} \sum_{\substack{l=1\\l\neq k}}^{N} \frac{(B_{kl})^{2}}{k-l} + \alpha^{3} \sum_{\substack{l=1\\l\neq k}}^{N} (B_{ll} - B_{kk}) \left(\frac{B_{lk}}{k-l}\right)^{2} + \alpha^{3} \sum_{\substack{l=1\\l\neq k}}^{N} \sum_{\substack{m=1\\m\neq k\\m\neq l}}^{N} \frac{B_{lk}B_{ml}B_{km}}{(k-l)(k-m)} + O(\alpha^{4}).$$

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The value  $B_{kl}$  is only nonzero if  $l \in \{k - 1, k, k + 1\}$ . We conclude that  $B_{lk}B_{ml}B_{km} = 0$  if  $k \neq l \neq m$ . Thus

$$\alpha^{2}\xi_{k} = -(N-k) + \alpha B_{kk} + \alpha^{2} \sum_{\substack{l=1\\l\neq k}}^{N} \frac{(B_{kl})^{2}}{k-l} + \alpha^{3} \sum_{\substack{l=1\\l\neq k}}^{N} (B_{ll} - B_{kk}) \left(\frac{B_{lk}}{k-l}\right)^{2} + O(\alpha^{4}).$$

The  $\alpha^2$ -term can be worked out as follows:

$$\begin{split} \sum_{\substack{l=1\\l\neq k}}^{N} \frac{(B_{kl})^2}{k-l} &= \left(\frac{B_{k,k-1}^2}{k-(k-1)} + \frac{B_{k,k+1}^2}{k-(k+1)}\right) \\ &= [\tau\alpha^2(k-1)+1](N-k+1)k - (\tau\alpha^2k+1)(N-k)(k+1) \\ &= \tau\alpha^2[(k-1)(N-k+1)k - k(N-k)(k+1)] + [(N-k+1)k - (N-k)(k+1)] \\ &= \tau\alpha^2k(3k-2N-1) + (2k-N). \end{split}$$

The  $\alpha^3$ -term can be worked out as follows:

$$\begin{split} \sum_{\substack{l=1\\l\neq k}}^{N} (B_{ll} - B_{kk}) \left(\frac{B_{lk}}{k-l}\right)^2 &= \alpha \left( (2\tau k - \tau N - \tau - 1) \frac{B_{k,k-1}^2}{[k - (k-1)]^2} + (2\tau k - \tau N + \tau - 1) \frac{B_{k,k+1}^2}{[k - (k+1)]^2} \right) \\ &= \alpha (2\tau k - \tau N - \tau - 1) [\tau \alpha^2 (k-1) + 1] (N - k + 1) k \\ &+ \alpha (2\tau k - \tau N + \tau - 1) (\tau \alpha^2 k + 1) (N - k) (k+1) \\ &= -8\alpha^3 k^3 \tau^2 + 3\alpha^3 k^2 \tau^2 + 3\alpha^3 k^2 \tau - 6\alpha k^2 \tau + 9\alpha^3 k^2 N \tau^2 - \alpha^3 k \tau^2 - \alpha^3 k \tau \\ &+ 2\alpha k - 2\alpha^3 k N^2 \tau^2 - \alpha^3 k N \tau^2 - 2\alpha^3 k N \tau + 6\alpha k N \tau - \alpha N^2 \tau + \alpha N \tau - \alpha N \\ &= -6\alpha k^2 \tau + 2\alpha k + 6\alpha k N \tau - \alpha N^2 \tau + \alpha N \tau - \alpha N + O(\alpha^3). \end{split}$$

Assembling the results, we find

$$\alpha^{2}\xi_{k} = -(N-k) + \alpha^{2}[-\tau k(N-k) - k] + \alpha^{2}(\tau \alpha^{2}k(3k-2N-1) + (2k-N)) + \alpha^{4}(-6k^{2}\tau + 2k + 6kN\tau - N^{2}\tau + N\tau - N) + O(\alpha^{4}).$$

Using  $\alpha^2 = 1/\varepsilon^*$ , we find

$$\begin{split} \xi_k &= -\varepsilon^* (N-k) - [\tau k (N-k) + k] + \left( \frac{1}{\varepsilon^*} \tau k (3k - 2N - 1) + (2k - N) \right) \\ &+ \frac{1}{\varepsilon^*} (-6k^2 \tau + 2k + 6kN \tau - N^2 \tau + N \tau - N) + O\left(\frac{1}{\varepsilon^*}\right). \end{split}$$

The index k = 0, 1, ..., N can be transformed so as to make sure that the eigenvalues  $\xi_k$  are descending:  $0 = \xi_1 > \xi_2 > \cdots > \xi_{N+1}$ . We define the index  $\hat{k} = N - k + 1$ , which takes values in  $\hat{k} = 1, 2, ..., N + 1$  such that we recover

$$\begin{split} \xi_{\hat{k}} &= -(\hat{k}-1)\varepsilon^* - (\hat{k}-1)[\tau(N+1-\hat{k})+1] \\ &+ (\tau(2N+2\hat{k}N-2-5\hat{k}-3\hat{k}^2)+N-2-2\hat{k})\frac{1}{\varepsilon^*} + O\bigg(\frac{1}{\varepsilon^*}\bigg), \end{split}$$

which concludes our proof. Presumably the  $O(\alpha^4)$ -terms will bring additional terms for the  $O(\frac{1}{\varepsilon^*})$ -term in the final solution, thus our estimate is only correct until  $O(\frac{1}{\varepsilon^*})$ .

## 2. The limit $\tau \to \infty$

If the effective infection rate  $\tau$  tends to infinity, then the dynamics of the  $\varepsilon$ -SIS process simplifies to an SI process. Metastability cannot be observed in the SI process, because the number of infected nodes only increases, until all nodes are infected. The exact time-dependent solution of the continuous-time Markovian SI process on any fixed graph with heterogeneous infection rates is given in [47].

*Proof.* The symmetric transition matrix  $\tilde{P}$  from Eq. (A1) can be rewritten as

$$\tilde{P} = \tilde{A} + \tilde{B},$$

where  $\tilde{A}$  is a diagonal matrix with elements

$$\tilde{A}_{kk} = -\tau k(N-k)$$
 for  $k = 0, 1, \dots, N$ 

and  $\tilde{B}$  is a symmetric, tridiagonal matrix with elements

$$\tilde{B}_{kk} = -\varepsilon^* (N-k) - k,$$
  
$$\tilde{B}_{k-1,k} = \sqrt{k(N-k+1)[(k-1)\tau + \varepsilon^*]}.$$

We define the small parameter  $\alpha = \frac{1}{\sqrt{\tau}}$  (which is small, because  $\tau \to \infty$ ), such that  $\alpha^2 \tilde{P} = A + \alpha B$ , where we defined  $A = \alpha^2 \tilde{A}$  and  $B = \alpha \tilde{B}$ . The matrix A is then diagonal with elements

$$A_{kk} = -k(N-k)$$

and the matrix B is then tridiagonal with elements

$$B_{kk} = -\alpha(\varepsilon^*(N-k)+k),$$
  
$$B_{k-1,k} = \sqrt{k(N-k+1)(k-1+\varepsilon^*\alpha^2)}.$$

The eigenvalues of *A* are simply  $a_k = -k(N - k)$  for k = 0, 1, ..., N and the corresponding eigenvector  $\mathbf{x}_k = \mathbf{e}_k$ , where  $\mathbf{e}_k$  is the  $(N + 1) \times 1$  all-zeros vector, except at entry *k* where it is 1. We distinguish between networks with even and odd sizes and treat the case  $\varepsilon^* = 1$  with special care.

### a. Case 1: Even size N

Consider a graph with an even number of nodes *N*. Then the matrix *A* has one simple eigenvalue  $a_k$  with index k = N/2. The remaining eigenvalues are twofold-degenerate. Let  $a_k$  and  $a_{N-k}$  be a pair of degenerate eigenvalues. Then the eigenvalue equation is according to part (b) from Theorem E.1,

$$\begin{pmatrix} B_{k,k} & B_{k,N-k} \\ B_{N-k,k} & B_{N-k,N-k} \end{pmatrix} \begin{pmatrix} \alpha_k \\ \beta_k \end{pmatrix} = \xi_k^{(1)} \begin{pmatrix} \alpha_k \\ \beta_k \end{pmatrix}.$$
 (E2)

Since the number of nodes *N* is even, we know that  $B_{k,N-k} = B_{N-k,k} = 0$  for all  $k \neq N/2$ . Thus the eigenvalue correction equals

$$\xi_{k}^{(1)} = B_{k,k} = -\alpha(\varepsilon^{*}(N-k) + k),$$
  

$$\xi_{N-k}^{(1)} = B_{N-k,N-k} = -\alpha(\varepsilon^{*}k + (N-k)),$$

which are different for all  $k \neq \frac{N}{2}$  and all even network sizes N, except when  $\varepsilon^* = 1$ . If  $\varepsilon^* \neq 1$ , the eigenvectors  $(\alpha_k \ \beta_k)^T$  equal the elementary vectors  $\mathbf{e}_k$ , which implies<sup>7</sup> that part (a) from Theorem E.1 can be used instead of part (b).

Applying part (a) from Theorem E.1, we find

$$\alpha^{2}\xi_{k} = a_{k} + \alpha B_{kk} + \alpha^{2} \sum_{\substack{l=1\\l\neq k}}^{N} \frac{(B_{lk})^{2}}{a_{k} - a_{l}} + O(\alpha^{3})$$
$$= -k(N-k) - \alpha^{2}(\varepsilon^{*}(N-k) + k) + \alpha^{2} \left(\frac{(B_{k-1,k})^{2}}{a_{k} - a_{k-1}} + \frac{(B_{k+1,k})^{2}}{a_{k} - a_{k+1}}\right) + O(\alpha^{3})$$

<sup>&</sup>lt;sup>7</sup>Please consult the sketch of the proof of Theorem E.1 for the complete reasoning.

$$= -k(N-k) - \alpha^{2} (\varepsilon^{*}(N-k)+k) + \alpha^{2} \left( \frac{k(N-k+1)(k-1+\varepsilon^{*}\alpha^{2})}{2k-N-1} + \frac{(k+1)(N-k)(k+\varepsilon^{*}\alpha^{2})}{N-1-2k} \right) + O(\alpha^{3})$$
  
=  $-k(N-k) + \alpha^{2} \left( -\varepsilon^{*}(N-k) + \frac{k(N+1)(N-k)}{(2k-N-1)(2k-N+1)} \right) + O(\alpha^{3}).$ 

Using  $\alpha^2 = \frac{1}{\tau}$ , we find the following relationship for the eigenvalues:

$$\xi_k = -\tau k(N-k) + \left(-\varepsilon^*(N-k) + \frac{k(N+1)(N-k)}{(2k-N-1)(2k-N+1)}\right) + O\left(\frac{1}{\sqrt{\tau}}\right).$$
(E3)

For  $k = \frac{N}{2}$ , we simply find

$$\xi_{N/2} = -\tau \frac{N^2}{4} - \frac{N}{2} \left( \varepsilon^* + \frac{N}{2} (N+1) \right) + O\left(\frac{1}{\sqrt{\tau}}\right).$$
(E4)

From Eq. (E4), we may conclude that (E3) is only valid if  $\tau > \frac{2\varepsilon^*}{N}$  and  $\tau > N + 1$ .

## b. Case 2: Odd size N and $\varepsilon^* \neq 1$

For odd network sizes *N*, all eigenvalues  $a_k$  are twofold-degenerate. All eigenvalues can be computed using (E3) provided that  $\varepsilon^* \neq 1$ . However, special attention is required for eigenvalues  $a_k$  with indices  $k = \frac{N+1}{2}$  and  $k = \frac{N-1}{2}$ . In that case, the eigenvalue equation (E1) becomes

$$\begin{pmatrix} -\frac{\alpha}{2}[\varepsilon^*(N+1)+N-1] & \frac{N+1}{2}\sqrt{\frac{N-1}{2}}+\varepsilon^*\alpha^2 \\ \frac{N+1}{2}\sqrt{\frac{N-1}{2}}+\varepsilon^*\alpha^2 & -\frac{\alpha}{2}[\varepsilon^*(N-1)+N+1] \end{pmatrix} \begin{pmatrix} \alpha_{(N-1)/2} \\ \beta_{(N-1)/2} \end{pmatrix} = \xi^{(1)}_{(N-1)/2} \begin{pmatrix} \alpha_{(N-1)/2} \\ \beta_{(N-1)/2} \end{pmatrix}$$

whose eigenvalues are distinct:

$$\xi_{(N-1)/2}^{(1)} = \frac{\alpha(N-\varepsilon^*)}{2} \pm \sqrt{\left(\frac{\alpha(N-\varepsilon^*)}{2}\right)^2 + \frac{1}{8}(N+1)^2(N-1)}$$

The corresponding eigenvectors are

$$\binom{\alpha_{(N-1)/2}}{\beta_{(N-1)/2}} = \binom{\alpha(N-\varepsilon^*) \pm \sqrt{\alpha^2(N-\varepsilon^*)^2 + (N+1)^2\left(\frac{N-1}{2} + \varepsilon^*\alpha^2\right)}}{(N+1)\sqrt{\frac{N-1}{2} + \varepsilon^*\alpha^2}}.$$

We will not continue our analysis for the  $O(\alpha^2)$  terms, because the computations are tedious. Our final result is

$$\alpha^{2}\xi_{(N-1)/2} = -\frac{N-1}{2}\frac{N+1}{2} + \alpha\frac{\alpha(N-\varepsilon^{*})}{2} \pm \sqrt{\left(\frac{\alpha(N-\varepsilon^{*})}{2}\right)^{2} + \frac{1}{8}(N+1)^{2}(N-1) + O(\alpha^{2})}.$$

Using  $\alpha^2 = \frac{1}{\tau}$ , we find the following relationship for the eigenvalues:

$$\xi_{(N-1)/2} = -\frac{1}{4}\tau(N-1)(N+1) \pm \sqrt{\tau}(N+1)\sqrt{\frac{N-1}{2}} + O(1).$$
(E5)

The relation (E5) is a valid perturbation expansion if  $\tau > \frac{8}{N-1}$ . Since the computations of the second-order terms are tedious, we have omitted them here. The key observation is that the eigenvalue  $\xi_{(N-1)/2}$  scales with an  $O(\sqrt{\tau})$  term, which is not the case for even-sized networks. We have not found any physical or intuitive reasoning why this is the case.

## c. Case 3: Even size N and $\varepsilon^* = 1$

In this case, we construct the *M*-matrix for  $k \neq N/2$ ,

$$M = \begin{pmatrix} M_{kk} & M_{k,N-k} \\ M_{N-k,k} & M_{N-k,N-k} \end{pmatrix},$$

where

$$M_{k,N-k} = \sum_{\substack{m=1\\m\neq k\\m\neq N-k}}^{N} \frac{W_{k,m}W_{m,N-k}}{\xi_{k}^{(0)} - \xi_{m}^{(0)}}.$$

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We find

$$M_{kk} = \frac{k(N-k)(N+1)}{(2k-N-1)(2k-N+1)} + k + \alpha^2 \frac{k^2 + (N-k)^2 + N}{(2k-N-1)(2k-N+1)}$$

In most cases,  $M_{k,N-k} = 0$  because the product  $W_{k,m}W_{m,N-k}$  is only nonzero for m = k - 1 or m = k + 1, and m = N - k + 1 or m = N - k - 1. Hence, the product is only non-zero if  $k = \frac{N}{2} - 1$ . Thus, for  $k \neq \frac{N}{2} - 1$ , the second-order correction of the eigenvalues is unique and the eigenvalues follow as

$$\alpha^{2}\xi_{k} = -k(N-k) - \alpha^{2}N + \alpha^{2} \left( \frac{k(N-k)(N+1)}{(2k-N-1)(2k-N+1)} + k + \alpha^{2} \frac{k^{2} + (N-k)^{2} + N}{(2k-N-1)(2k-N+1)} \right) + O(\alpha^{3}).$$

Using  $\alpha^2 = \frac{1}{\tau}$ , we find the following relationship for the eigenvalues:

$$\xi_k = -\tau k(N-k) + k - N + \frac{k(N-k)(N+1)}{(2k-N-1)(2k-N+1)} + O\left(\frac{1}{\sqrt{\tau}}\right).$$
(E6)

For  $k = \frac{N}{2} - 1$ , additional work is required because the off-diagonal terms are nonzero. The *M*-matrix is in this case

$$M = \begin{pmatrix} \frac{1}{12}(N-2)(N^2+3N+8) & \frac{N}{2}\left(\frac{N}{2}+1\right)\sqrt{\left(\frac{N}{2}-1+\alpha^2\right)\left(\frac{N}{2}+\alpha^2\right)} \\ \frac{N}{2}\left(\frac{N}{2}+1\right)\sqrt{\left(\frac{N}{2}-1+\alpha^2\right)\left(\frac{N}{2}+\alpha^2\right)} & \frac{1}{12}(N+2)(N^2-N+4) \end{pmatrix}$$

whose eigenvalues are

$$\xi_k^{(2)} = \frac{N^3}{12} + \frac{N^2}{12} + \frac{N}{6} - \frac{1}{3} \pm \sqrt{1 + \left(\frac{N}{2}\right)^3 \left(\frac{N}{2} + 1\right)^2 \left(\frac{N}{2} - 1\right)}$$

thus the total eigenvalue expansion equals

$$\xi_k = -\tau \left(\frac{N}{2} - 1\right) \left(\frac{N}{2} + 1\right) + \frac{N^3}{12} + \frac{N^2}{12} - \frac{5N}{6} - \frac{1}{3} \pm \sqrt{1 + \left(\frac{N}{2}\right)^3 \left(\frac{N}{2} + 1\right)^2 \left(\frac{N}{2} - 1\right)}$$

# d. Case 4: Odd size N and $\varepsilon^* = 1$

For the case  $k \neq \frac{N-1}{2}$ , the eigenvalues follow from (E6) and the result for  $k = \frac{N-1}{2}$  is described in (E5).

Finally introducing  $\tilde{k} = k + 1$ , such that the index  $\tilde{k}$  runs from 1 to N + 1, finalizes the proof.

3. Limit 
$$\varepsilon^* \to 0$$

For self-infection rates  $\varepsilon^* < \frac{1}{N}$ , Gershgorin's circle theorem leads to a tight bound for the smallest eigenvalue  $\xi_{N+1}$ . The case  $\tau > \tau_c$  appeared earlier in Ref. [48] (Corollary 3).

*Proof.* We denote the (not necessarily ordered) eigenvalues  $\xi_1, \ldots, \xi_{N+1}$  of the transition matrix *P* from Eq. (2). Given the scaled birth rate  $\tilde{\lambda}_k$  and death rate  $\tilde{\mu}_k$ , Gershgorin's circle theorem provides the following bounds for the eigenvalues:

$$|\xi_k + \lambda_k + \tilde{\mu}_k| \leqslant \lambda_{k-1} + \tilde{\mu}_{k+1}.$$

Substituting the scaled birth rate  $\tilde{\lambda}_k = (\tau k + 1)(N - k)$  and the scaled death rate  $\tilde{\mu}_k = k$ , we find

$$f(k) \leqslant \xi_k \leqslant 1 + \varepsilon^* + \tau (2k - N - 1), \tag{E7}$$

where the lower bound f(k) equals

$$f(k) = -\tau[(k-1)(N-k+1) + k(N-k)] -\varepsilon^*[2N-2k+1] - [2k+1].$$
(E8)

The lower bound f(k) is negative for all k = 1, ..., N + 1. 1. The upper bound is most negative for k = 1, hence  $\min_k \xi_k \leq 1 + \varepsilon^* - \tau(N-1)$ . Using  $\tau = x\tau_c^{(1)} = \frac{x}{N-1}$ , we find  $\min_k \xi_k \leq 1 - x + \varepsilon^* < 0$  above the epidemic threshold (x > 1). For the remaining eigenvalues  $\xi_k$ , the upper bound is larger than zero, which is not confining. Simulations also indicate that the upper bound is very loose. Instead, we focus on the lower bound f(k).

### a. Above the epidemic threshold

The lower bound f(k) is the smallest when  $\frac{df}{dk} = 0$ , and we find<sup>8</sup>

$$\frac{df}{dk}(\hat{k}) = -\tau(2N - 4\hat{k} + 2) + 2\varepsilon^* - 2 = 0.$$
 (E9)

Solving for  $\hat{k}$  gives

$$\hat{k} = \frac{1 - \varepsilon^*}{2\tau} + \frac{N+1}{2}.$$

Since we consider the case  $\varepsilon^* < \frac{1}{N}$ , the effective self-infection rate  $\varepsilon^*$  can be neglected,

$$\hat{k} = \frac{1}{2\tau} + \frac{N+1}{2}.$$
 (E10)

Hence,  $\hat{k}$  is at least larger than N/2. The lower bound becomes

$$f(\hat{k}) = -\frac{1}{2\tau} - \frac{3}{2} - N - \frac{\tau N(N-1)}{2} + \varepsilon^* \left[\frac{1}{\tau} - N\right].$$

<sup>8</sup>We assume here that the function f(k) is varying slowly, which means that the difference between f(k + 1) and f(k) is small for all k. Since the function f(k) is a quadratic function in k, the function f(k) indeed varies sufficiently slowly.

Using the normalized effective infection rate  $x = \tau (N - 1)$ , we approximately find

$$f(\hat{k}) \approx -\left(\frac{1}{2x} + 1 + \frac{x}{2}\right)N + \left(\frac{1}{2x} - \frac{3}{2}\right).$$
 (E11)

The lower bound (E11) holds for all eigenvalues  $\xi_k$  where  $1 \leq k \leq N + 1$ . Thus we may conclude that

$$\xi_{N+1}\gtrsim -\left(\frac{1}{2x}+1+\frac{x}{2}\right)N.$$

### b. Below the epidemic threshold

If  $\tau < \tau_c^{(1)} = \frac{1}{N-1}$ , the lower bound f(k) is the smallest for  $\hat{k} = N - 1$ . We find

$$f(\hat{k}) = -\tau(3N - 5) - 2N + 1.$$

Using  $\tau = \frac{x}{N-1}$  and x < 1, we find approximately

$$f(\hat{k}) \approx -2N + 1 - 3x.$$
 (E12)

Equation (E12) holds for all eigenvalues  $\xi_k$ , with  $1 \le k \le N + 1$ , thus we conclude that

$$\xi_{N+1} \gtrsim -2N.$$

### 4. The limit $\tau \rightarrow 0$

If the effective infection rate  $\tau = 0$ , then the  $\varepsilon$ -SIS model reduces to a birth and death process with linear rates that can be solved exactly for any time *t* and any number of initially infected nodes [19]. The eigenvalue ratio  $\rho$  in the limit of small effective infection rates  $\tau$  is approximately  $\rho = \xi_3/\xi_2 \approx 3/2$ , because the second-largest eigenvalue  $\xi_2 \approx -2$  and  $\xi_3 \approx -3$ . Thus, metastability cannot be observed for small effective infection rates  $\tau$ .

#### 5. Final considerations

Apart from limit cases and bounds, the second-largest eigenvalue  $\xi_2$  can be computed using the following approach.

*Theorem E.2* (Van Doorn *et al.* [35]). The convergence rate  $-\xi_2$  equals

$$\max_{\mathbf{d}>\mathbf{0}} \min_{1 \le k \le N} \alpha_k = -\xi_2 = \min_{\mathbf{d}>\mathbf{0}} \max_{1 \le k \le N} \alpha_k, \quad (E13)$$

where **d** =  $(d_1, ..., d_N)$ , and  $d_i > 0$  for i = 1, ..., N and

$$\alpha_k = \tau \left[ \left( \frac{d_{k+1}}{d_k} - 1 \right) k^2 + \left( 1 - \frac{d_{k+1}}{d_k} \right) kN + 2k - N - 1 \right]$$
$$+ \varepsilon^* \left[ \left( \frac{d_{k+1}}{d_k} - 1 \right) k + \left( 1 - \frac{d_{k+1}}{d_k} \right) N + 1 \right]$$
$$+ \left[ \left( 1 - \frac{d_{k-1}}{d_k} \right) k + \frac{d_{k-1}}{d_k} \right]$$

for k = 1, ..., N and  $d_0 = d_{N+1} = 0$ .

Theorem E.2 associates the computation of the secondlargest eigenvalue  $\xi_2$  of the transition matrix *P* to the finding of a suitable, positive vector **d**. Equation (E13) illustrates that choosing any positive vector **d** directly provides lower and upper bounds for the convergence rate  $\xi_2$ . Unfortunately, simply generating random values for the vector **d** does not provide sharp bounds for the second-largest eigenvalue  $\xi_2$ .

As an example, we consider  $\mathbf{d} = \mathbf{u}$ , where  $\mathbf{u}$  is the  $N \times 1$  all-ones vector. According to Theorem E.2, the convergence rate  $\xi_2$  is then bounded by

$$-\tau(N-1) + 1 + \varepsilon^* \leq -\xi_2 \leq \tau(N-1) + 1 + \varepsilon^*$$
. (E14)

If the effective infection rate  $\tau$  is larger than the mean-field epidemic threshold  $\tau_c^{(1)} = \frac{1}{N-1}$ , the lower bound in (E14) is negative and therefore not confining. For  $\tau < \tau_c$ , the lower bound in (E14) appears to be a loose bound. In the limit  $\tau \rightarrow 0$ , the convergence rate equals  $-\xi_2 = 1 + \varepsilon^*$ , which agrees with Theorem V.1. The upper bound is positive in both cases, and is at least 1, but it appears to be a loose bound as well.

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