

**Lifetime differences, direct  $CP$  violation, and partial widths in  $D^0$  meson decays to  $K^+K^-$   
and  $\pi^+\pi^-$**

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We describe several measurements using the decays  $D^0 \rightarrow K^+ K^-$  and  $\pi^+ \pi^-$ . We find the ratio of partial widths,  $\Gamma(D^0 \rightarrow K^+ K^-)/\Gamma(D^0 \rightarrow \pi^+ \pi^-)$ , to be  $2.96 \pm 0.16 \pm 0.15$ , where the first error is statistical and the second is systematic. We observe no evidence for direct  $CP$  violation, obtaining  $A_{CP}(KK) = (0.0 \pm 2.2 \pm 0.8)\%$  and  $A_{CP}(\pi\pi) = (1.9 \pm 3.2 \pm 0.8)\%$ . In the limit of no  $CP$  violation we measure the mixing parameter  $y_{CP} = -0.012 \pm 0.025 \pm 0.014$  by measuring the lifetime difference between  $D^0 \rightarrow K^+ K^-$  or  $\pi^+ \pi^-$  and the  $CP$  neutral state,  $D^0 \rightarrow K^- \pi^+$ . We see no evidence for mixing.

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The structure of the standard model has been guided by measurements of mixing and  $CP$  violation in the neutral  $K$  and  $B$  meson sectors. The standard model predictions for the rate of mixing and  $CP$  violation in the charm sector are small, with the largest predictions in both cases being  $\mathcal{O}(0.01)$ , and most predictions being  $\mathcal{O}(0.001)$  [1]. Observation of  $CP$  violation above the 1% level would be strong evidence for physics outside the standard model.

The  $SU(3)$  flavor symmetry predicts  $\Gamma(D^0 \rightarrow K^+ K^-)/\Gamma(D^0 \rightarrow \pi^+ \pi^-) = 1$  [2], while the previously measured value is  $2.80 \pm 0.20$  [3]. This deviation is most likely caused by large final state interactions. These can also give rise to a large strong phase differences between mixing and Cabibbo-suppressed  $D^0$  decays that give rise to the same final states [4]. A measure of  $CP$  violation in these decays, the direct  $CP$  violation asymmetry, is proportional to the amount of  $CP$  violation in the decays and the sine of the strong phase difference. The standard model suggests that  $CP$  violation in these decays is small since the higher-order

diagrams are suppressed, however, new physics can enhance the rate of  $CP$  violation. In this paper we present the most precise measurement to date of the ratio of partial widths,  $\Gamma(D^0 \rightarrow K^+ K^-)/\Gamma(D^0 \rightarrow \pi^+ \pi^-)$  [5]. We also present our search for direct  $CP$  violation in these decays.

In the absence of  $CP$  violation, the  $D$  meson mass eigenstates  $D_{1,2}$  are also  $CP$  eigenstates. The decay of a  $D^0$  to a  $CP$  eigenstate, such as  $K^+ K^-$  or  $\pi^+ \pi^-$ , has a purely exponential lifetime characteristic of the associated mass eigenstate. Therefore, in the limit of no  $CP$  violation, we can write the time-dependent rate of a  $D^0$  decaying to a  $CP$  eigenstate,  $f$ , as  $R(t) \propto \exp[-t\Gamma \cdot (1 - y_{CP}\eta_{CP})]$ , where  $CP|f\rangle = \eta_{CP}|f\rangle$ ,  $\Gamma$  is the average  $D^0$  width,  $y_{CP} = y = \Delta\Gamma/2\Gamma$ , and  $\Delta\Gamma$  is the width difference between the two mass eigenstates [6]. We can measure  $y_{CP}$  simply by measuring the ratio of lifetimes of the  $D^0$  decaying to a  $CP$  eigenstate ( $\tau_{CP+}$ ) and a  $CP$  neutral state such as  $K^- \pi^+$  ( $\tau$ ). Then  $y_{CP} = \tau/\tau_{CP+} - 1$ . We have used  $\tau = (\tau_{CP+} + \tau_{CP-})/2$ , and assumed that the lifetime difference is small so that the  $K^+ \pi^-$  lifetime

distribution can be fit with a single exponential.

The data were collected using the CLEO II.V upgrade [7] of the CLEO II detector [8] between February 1996 and February 1999 at the Cornell Electron Storage Ring (CESR). The data correspond to  $9.0 \text{ fb}^{-1}$  of  $e^+e^-$  collisions near  $\sqrt{s} \approx 10.6 \text{ GeV}$ . The detector consisted of cylindrical tracking chambers and an electromagnetic calorimeter immersed in a 1.5 Tesla axial magnetic field, surrounded by muon chambers. The reconstruction of displaced vertices from charm decays was made possible by the addition of a silicon vertex detector (SVX) in CLEO II.V. We utilized this improved resolution in previous searches for  $D^0-\bar{D}^0$  mixing [9] and in measurements of charmed particle lifetimes [10]. The charged particle trajectories were fit using a Kalman filter technique that takes into account energy loss as the particles pass through the material of the beam pipe and detector [11].

To measure relative efficiencies between the modes, study their backgrounds, and observe biases introduced by our methods, we use a GEANT [12]-based detector simulation of our data. We use simulations of  $e^+e^- \rightarrow c\bar{c}$  with one of the charm quarks fragmenting as a charged  $D^*$  and then decaying to a  $D^0$  and a charged  $\pi$ . The  $D^0$  then decays to  $K^+K^-$ ,  $\pi^+\pi^-$ , or  $K^+\pi^-$ . For background studies we have a simulated sample of  $e^+e^- \rightarrow q\bar{q}$  where  $q=uds$  with the quarks fragmenting and particles decaying generically guided by previous measurements. The data of this generic simulation correspond to roughly ten times the luminosity collected by the detector. For these studies the simulated samples are reconstructed and selected using the same methods as the data sample as described below.

The events are selected by searching for the decay chain  $D^{*+} \rightarrow D^0\pi_s^+$ , with subsequent decays of the  $D^0$  to  $K^+K^-$ ,  $\pi^+\pi^-$ , or  $K^-\pi^+$ . The charge of the slow pion,  $\pi_s^+$ , from the  $D^{*+}$  decay is a tag of the initial  $D^0$  flavor. Additionally, we separate signal from background using the energy release in the  $D^{*+}$  decay,  $Q \equiv M^* - M - M_\pi$ , where  $M^*$  is the candidate  $D^{*+}$  invariant mass,  $M$  is the candidate  $D^0$  invariant mass, and  $M_\pi$  is the pion mass.

All pairs of oppositely charged tracks of good quality are used to form  $D^0$  candidates assuming four particle assignments:  $K^+K^-$ ,  $K^+\pi^-$ ,  $\pi^+K^-$ , and  $\pi^+\pi^-$ . The  $D^0$  candidate is retained if any of the particle assignments has an invariant mass within 35 MeV of the  $D^0$  mass. The observed widths of  $D^0 \rightarrow K^+K^-$ ,  $\pi^+\pi^-$ , and  $K^-\pi^+$  signal candidate mass distributions are  $4.73 \pm 0.15$ ,  $4.95 \pm 0.26$ , and  $5.08 \pm 0.10 \text{ MeV}$ , respectively. The  $D^0$  daughters are constrained to come from a common vertex, and the confidence level from this constraint must be greater than 0.01%. A pion candidate with at least two SVX hits in both the  $r-\phi$  and  $r-z$  layers is combined with the  $D^0$  candidate to form a  $D^{*+}$ . The slow pion candidate is refit by constraining it to come from the intersection of the beam spot and the projection of the  $D^0$  momentum vector. This dramatically reduces the mismeasurement of the pion momentum due to multiple scattering in the beam pipe and first layer of silicon. The resulting  $Q$  distribution has a width of approximately 190 keV. We have used the same technique to measure the intrinsic width of the  $D^{*+}$  [13]. The candidate is retained if the confidence

level for the refit is greater than 0.01%,  $Q$  is less than 25 MeV, and the  $D^{*+}$  momentum is greater than  $2.2 \text{ GeV}/c$ . Finally, we require  $|\cos \theta^*| < 0.8$ , where  $\cos \theta^*$  is the angle in the  $D^0$  rest frame between a  $D^0$  daughter and the  $D^0$  direction in the laboratory frame. The signal is flat in  $\cos \theta^*$ , while the backgrounds are highly peaked at  $|\cos \theta^*| \approx 1$ . Particle identification using specific ionization is not required since the different mass hypotheses are separated by greater than 8.5 standard deviations.

The partial width measurements are obtained from binned maximum likelihood fits to the  $Q$  distribution of the  $D^{*+}$  decay. We fit in bins of momentum to eliminate potential bias due to mismodeling of the  $D^{*+}$  momentum spectrum in Monte Carlo calculations. The finite statistics of the fitting shapes are included in the statistical uncertainty of the fit. The shape of the signal is given by the shape of the  $K^-\pi^+$  candidates in the data after we have subtracted off the small background contribution. The background is determined by a fit to the data that excludes the signal region with the background shape taken from the Monte Carlo simulation. This procedure gives a one variable parameter function that describes the signal shape. All of the modes have approximately the same signal shape since the  $Q$  resolution is dominated by multiple scattering of the slow pion. A detailed discussion of this shape can be found in our paper measuring the width of the  $D^{*+}$  [13]. As a check, we also fit the  $D^0$  mass distribution using a double Gaussian for the signal shape.

We first fit the  $K\pi$  data outside of the signal region to obtain the background normalization, where we have used a threshold function of the form  $a \cdot Q^{1/2} + b \cdot Q^{3/2} + c \cdot Q^{5/2}$  to describe the background. To obtain  $R_{\pi\pi} = \Gamma(D^0 \rightarrow \pi^+\pi^-) / \Gamma(D^0 \rightarrow K^-\pi^+)$  we fit the  $Q$  distributions for the ratio of signal yields between the  $\pi\pi$  and  $K\pi$  channels, and for the normalization of the background, where we have used the signal shape and background parameters determined from the  $K\pi$  data and Monte Carlo samples, respectively. To obtain  $R_{KK} = \Gamma(D^0 \rightarrow K^+K^-) / \Gamma(D^0 \rightarrow K^-\pi^+)$  we fit the  $Q$  distributions as we did for  $R_{\pi\pi}$ , however, we add an additional component from pseudoscalar-vector decay (PV) background, where the shape is taken from Monte Carlo samples and the normalization is allowed to float. The PV background is primarily from  $D^0 \rightarrow K^-\rho^+$ ,  $\rho^+ \rightarrow \pi^+\pi^0$  where the  $\pi^0$  is nearly at rest. This background forms a broad peak in  $Q$ . The PV background is negligible in the  $\pi\pi$  and  $K\pi$  final states.

In order to maintain statistical independence, we use two different sets of Monte Carlo events. One sample is only used to determine the fitting shapes. We fit the data and the second Monte Carlo sample simultaneously to correct for small differences in acceptance between the normalization and signal modes. We perform three separate fits to the modes and report the ratios of the  $KK$  and  $\pi\pi$  signals to the  $K\pi$  signal. In these fits the signal has one free parameter, an overall normalization of the fixed shape, and the backgrounds have one free parameter for  $\pi\pi$  and  $K\pi$  and two, an extra for PV, for  $KK$ . All the background shapes are fixed in the final fits. The results of the fits are  $R_{KK} = 0.1037$

$\pm 0.0038$  and  $R_{\pi\pi} = 0.0355 \pm 0.0017$  from approximately 20000  $K^- \pi^+$ , 1900  $K^+ K^-$ , and 710  $\pi^+ \pi^-$  events. The small background rates observed in the data fits agree well with the prediction of our simulation.

We assess the systematic uncertainty due to the fitting shapes by performing a series of fits using different assumptions for the background, and several fits to the  $D^0$  mass distribution. For the latter case we obtain  $R_{KK} = 0.1035 \pm 0.0038$  and  $R_{\pi\pi} = 0.0340 \pm 0.0018$ . We estimate systematic uncertainties of 0.0017 and 0.001 due to the fitting shapes in the  $KK$  and  $\pi\pi$  modes, respectively. We vary the bin sizes,  $Q$  fit range,  $Q$  signal region used to determine the non-PV background shape, and candidate  $D^0$  mass requirement, to form a combined systematic uncertainty of 0.0005 due to these variations.

We also estimate systematic uncertainties associated with some of the event selection requirements by doing the analysis without those requirements. The variations we observe are 0.0009 in  $R_{KK}$  and 0.00095 in  $R_{\pi\pi}$  from removing the vertex confidence level requirement which roughly doubles the number of candidates in all channels, and 0.00032 in  $R_{KK}$  and 0.00016 in  $R_{\pi\pi}$  from removing the track quality requirement.

We use the  $K\pi$  data sample to study the effect of any mismodeling in the simulation of the fragmentation and the detector acceptance. These are tied together as the acceptance for the  $D^*$  daughters is not flat as a function of the angle of the tracks with respect to the beam line. This distribution depends on the momentum spectrum of the  $D^*$  due to the opening angles of the daughters. A soft  $D^*$  produces daughters with large opening angles with a higher chance for one of them to fall outside the detector acceptance. We compare the observed  $D^*$  spectrum for the  $K\pi$  channel with the tuned simulation and propagate the observed difference to the acceptance for the  $KK$  and  $\pi\pi$  modes. The effect is worse for the  $KK$  mode due its larger opening angle distribution causing more generated  $KK$  decays to be lost due to one of the  $K$  tracks going into the region close to the beam line where the acceptance is zero. For the fragmentation modeling we estimate a systematic uncertainty of 0.0014 for  $R_{KK}$  and 0.0005 for  $R_{\pi\pi}$ . We obtain relative corrections and uncertainties due to mismodeling of the detector acceptance of  $(-2.4 \pm 1.1)\%$  for  $R_{KK}$  and  $(+2.4 \pm 2.7)\%$  for  $R_{\pi\pi}$ . We apply these corrections and sum all of the systematic uncertainties in quadrature to obtain the final results  $R_{KK} = \Gamma(D^0 \rightarrow K^+ K^-) / \Gamma(D^0 \rightarrow K^- \pi^+) = 0.1040 \pm 0.0033 \pm 0.0027$  and  $R_{\pi\pi} = \Gamma(D^0 \rightarrow \pi^+ \pi^-) / \Gamma(D^0 \rightarrow K^- \pi^+) = 0.0351 \pm 0.0016 \pm 0.0017$ , where the first error is statistical and the second is systematic. These results are the most precise determinations of  $R_{KK}$  and  $R_{\pi\pi}$  to date [3].

We can combine the results, accounting for cancellations and correlations among the uncertainties to calculate  $R_{KK}/R_{\pi\pi} = 2.96 \pm 0.16(\text{stat}) \pm 0.15(\text{syst})$ . This result agrees with the world average value of  $2.80 \pm 0.20$  [3].

We can use the same procedure to search for the direct  $CP$  asymmetries

$$A_{CP} = \frac{\Gamma(D^0 \rightarrow f) - \Gamma(\overline{D^0} \rightarrow f)}{\Gamma(D^0 \rightarrow f) + \Gamma(\overline{D^0} \rightarrow f)},$$

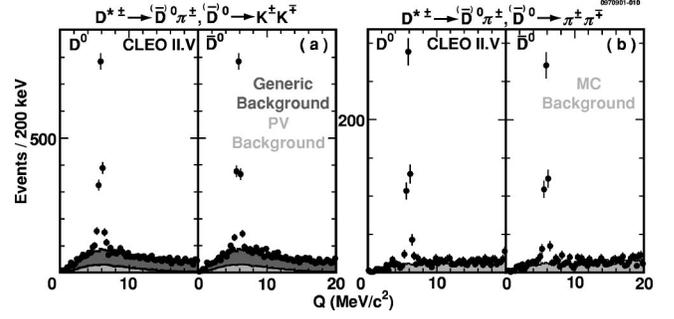


FIG. 1. The  $D^* \rightarrow D \pi_s$   $Q$  distributions for (a)  $D^0 \rightarrow K^+ K^-$  and  $\overline{D^0} \rightarrow K^+ K^-$  candidates and (b)  $D^0 \rightarrow \pi^+ \pi^-$  and  $\overline{D^0} \rightarrow \pi^+ \pi^-$  candidates. The points are the data and the histograms are the background fits.

where  $f$  can be  $K^+ K^-$  or  $\pi^+ \pi^-$ . The charge of the slow pion from the  $D^*$  decay serves as an unbiased tag of the  $D^0$  flavor since charm quarks are produced in quark-antiquark pairs at CESR and fragmentation and the  $D^*$  decay are strong processes, which conserve  $CP$ .

We measure the  $CP$  asymmetry in the same manner as the partial width analysis described above apart from the following changes. The  $K^+ K^-$  and  $\pi^+ \pi^-$  data are separated into  $D^0$  and  $\overline{D^0}$  samples based on the charge of the slow pion. However, we still normalize by the entire  $K\pi$  sample to eliminate possible bias from any asymmetry in  $D^0 \rightarrow K^- \pi^+$  decay. The  $D^*$  momentum requirement is loosened to be greater than 2.0 GeV/c since acceptance differences between modes are no longer an issue. The candidate  $D^0$  mass requirement is tightened to  $\pm 15$  MeV of the nominal  $D^0$  mass, which reduces the backgrounds by about a factor of two.

We fit the data in the same manner as in the partial width analysis, modified as described above. The  $KK$  and  $\pi\pi$   $Q$  distributions and fit results are shown in Fig. 1 and the residuals in Fig. 2. From the fits we find  $1512 \pm 47$   $D^0 \rightarrow K^+ K^-$  events,  $1511 \pm 47$   $\overline{D^0} \rightarrow K^+ K^-$  events,  $579 \pm 26$   $D^0 \rightarrow \pi^+ \pi^-$  events, and  $557 \pm 26$   $\overline{D^0} \rightarrow \pi^+ \pi^-$  events, and obtain  $A_{CP}^{KK} = 0.001 \pm 0.022$  and  $A_{CP}^{\pi\pi} = 0.020 \pm 0.032$ .

The sources of possible systematic error for the  $CP$  asymmetry measurement are the shapes used for fitting and a charge-dependent slow pion acceptance. To assess the systematic uncertainty from the fitting shapes we perform fits in which we vary the candidate  $D^0$  mass window, remove the vertex confidence level requirement, vary the width of the  $K\pi$  signal region and the  $Q$  fit region, alter the number of bins, and split the  $K\pi$  sample into two according to the charge of the associated slow pion and fit the two samples separately. We use 1/2 of the largest variation in each case, and then sum them in quadrature to obtain a systematic uncertainty due to the fitting shape of 0.0068 for  $A_{CP}^{KK}$  and 0.0069 for  $A_{CP}^{\pi\pi}$ .

A difference in slow pion acceptance for positive and negative pions can come from a number of different sources. The interaction cross section of pions with matter is different for positive and negative pions. We use the known composition of the CLEO detector and the interaction cross sections

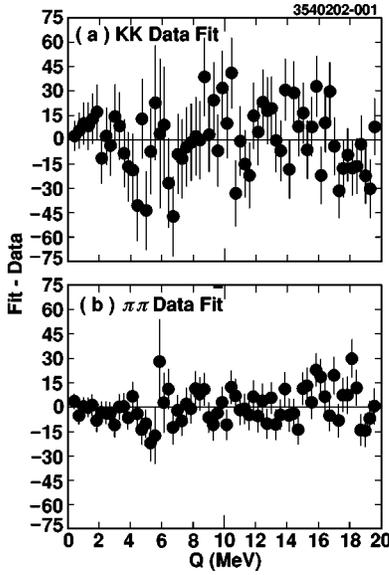


FIG. 2. The  $D^{*+} \rightarrow D \pi_s Q$  fit residual distributions for (a)  $D^0 \rightarrow K^+ K^-$  and  $\bar{D}^0 \rightarrow K^+ K^-$  candidates and (b)  $D^0 \rightarrow \pi^+ \pi^-$  and  $\bar{D}^0 \rightarrow \pi^+ \pi^-$  candidates. The fits are described in detail in the text.

to calculate the induced asymmetry as a function of momentum. We find that the bias to the asymmetry is less than 0.2%. We use the pions from  $K_s^0$  decays to search for a momentum-dependent charge bias in pion acceptance. We select the pions from  $K_s^0$  decay similarly to the method used to select the slow pions from  $D^{*+}$  decay. We compare the observed difference between the momentum spectrum for the positive and negative legs of the  $K_s^0$ , over the region of slow pion momenta from  $D^{*+} \rightarrow D^0 \pi^+$  decay, to estimate the acceptance difference for positive and negative pions to be less than 0.07%.

We have looked for a momentum-independent charge bias in track finding by generating single track Monte Carlo calculations randomly distributed in  $\theta$ ,  $\phi$ , and momentum, between 0 and 3 GeV/c. We see no significant bias, and limit the momentum-independent acceptance bias to be less than 0.48%. We translate these limits on acceptance differences and track finding biases into limits on our observed asymmetry based on the statistics of our observed data sample.

Charm quarks are expected to be produced with a small forward-backward asymmetry in  $e^+e^-$  annihilations at  $\sqrt{s} \approx 10.6$  GeV due to the interference between the photon and  $Z^0$ . The center of the luminous region was not exactly at the center of the detector, so this, coupled with the forward-backward asymmetry, induces an acceptance asymmetry. From a study of the  $K^+ \pi^-$  data and Monte Carlo samples we find an acceptance bias of  $0.014 \pm 0.014\%$ . We correct for the bias and assign the statistical error as a systematic uncertainty.

Summing all of the systematic uncertainties in quadrature and applying the correction mentioned above we arrive at the final result of  $A_{CP}^{KK} = (0.0 \pm 2.2 \pm 0.8)\%$  and  $A_{CP}^{\pi\pi} = (1.9 \pm 3.2 \pm 0.8)\%$ . We see no evidence of direct  $CP$  violation in these decays. This is the most precise measurement of these  $CP$  asymmetries to date [3,14].

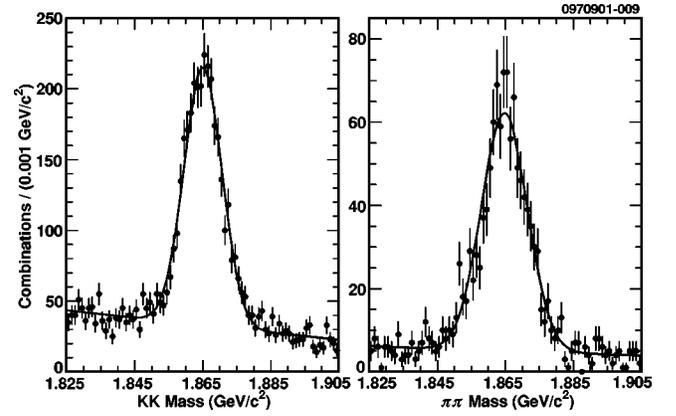


FIG. 3. The mass distribution for  $D^0 \rightarrow K^+ K^-$  (left) and  $D^0 \rightarrow \pi^+ \pi^-$  (right) candidates. The curves are the results of the fit discussed in the text.

As noted earlier we can measure the normalized mixing parameter  $y_{CP}$  by measuring the lifetime ratio between  $D^0 \rightarrow K^- \pi^+$  and  $D^0$  decay to a  $CP$  eigenstate, such as  $K^+ K^-$  or  $\pi^+ \pi^-$ :  $y_{CP} = \tau / \tau_{CP+} - 1$ . In the limit of no  $CP$  violation in the  $D$  meson sector  $y_{CP}$  is equivalent to  $y$ . We use the same data sample described above, using the decay length and momentum to determine the proper decay time. We modify the event selection criteria slightly for this analysis. We require the candidate  $D^0$  momentum to be greater than 2.3 GeV/c. We tighten the requirement on the vertex confidence level of the  $D^0$  candidate to be greater than 0.1%. Furthermore, we place an extra requirement on the data: the  $D^0$  candidate masses obtained with the three other particle assignments to the two daughters must be more than four standard deviations away from the nominal  $D^0$  mass.

We select events with a  $Q$  value within 1 MeV of the nominal value and fit their candidate  $D^0$  mass spectrum with a binned maximum likelihood fit to the sum of two Gaussians for the signal, constrained to the same central value, and a first order polynomial for the background. The data and fit results are shown in Fig. 3. The fit values are converted into a mass-dependent probability for signal and background and are used as an input to the lifetime fits. The other inputs to the lifetime fits are the measured proper decay time and its calculated uncertainty. For the  $KK$  and  $\pi\pi$  samples we fix the ratio of areas and the ratio of widths of the two Gaussians to the values determined in the  $K\pi$  fit. We perform the fits for candidate  $D^0$  mass over the range 1.825 to 1.905 GeV, and use all of these events in the lifetime fits described below.

For the signal portion of the probability distribution function for the lifetime fits we constrain the candidate  $D^0$  mass to a fixed value, which gives us a better measurement of  $y_{CP}$ . The value we constrain to is the weighted average of the  $D^0$  mass determined from the  $K\pi$ ,  $KK$ , and  $\pi\pi$  events, where each is corrected by an offset determined from Monte Carlo calculations. This offset is simply the difference between the input and measured  $D^0$  mass for each channel in the Monte Carlo sample. The offsets are  $+0.15 \pm 0.02$  MeV ( $K\pi$ ),  $+0.27 \pm 0.05$  MeV ( $KK$ ), and  $+0.10 \pm 0.09$  MeV ( $\pi\pi$ ). These offsets are caused by a distortion

in the decay vertex introduced by our fitting technique which has a bias towards smaller opening angles, lower masses, and we expect the data to have a similar bias. This mass constraint introduces a systematic bias in the lifetime measurement, but this has a small effect on  $y_{CP}$  which only depends on the ratio of lifetimes.

The candidate proper decay time,  $t$ , is given by

$$t = m \cdot \frac{(\vec{r}_{\text{dec}} - \vec{r}_{\text{prod}}) \cdot \hat{p}}{|\vec{p}|},$$

where  $\vec{r}_{\text{dec}}$  and  $\vec{p}$  are the position and momentum of the  $D^0$  candidate given by our vertex fit. We determine  $\vec{r}_{\text{prod}}$  using  $e^+e^- \rightarrow q\bar{q}$  ( $q = u d s c b$ ) events from sets of data with integrated luminosity of several  $\text{pb}^{-1}$ . The extent of the luminous region has a Gaussian width of approximately  $10 \mu\text{m}$  vertically,  $300 \mu\text{m}$  horizontally, and  $1 \text{ cm}$  along the beam direction [15]. We observe that the luminous region is stable during week long running periods and correct for changes each hour of data taking with an accuracy on the mean of a few microns. The resolution on the  $D^0$  decay point is typically  $40 \mu\text{m}$  in each dimension. The resolution in  $t$  is typically  $\sigma_t = 0.4$  in units of  $D^0$  lifetimes. We determine the proper decay time in the three dimensions separately, and combine them to arrive at the best estimate of  $t$  and  $\sigma_t$ .

We fit the lifetime distribution using an unbinned likelihood method. The signal probability distribution function (PDF) consists of an exponential convolved with a resolution function, composed of the sum of three parts, based on a simple, yet robust, physical model. For most events the calculated covariance matrix for the  $D^0$  daughters is assumed to be correct to within a global scale factor, with a Gaussian resolution function of width  $S \cdot \sigma_t$ . The scale factor,  $S$ , accounts for any common mistake in the covariance matrices, as would be present from a deficiency in the detector material description. A few percent of the events have one or more particles that have undergone a hard scatter, rendering the extrapolated vertex errors virtually meaningless. We model the contribution from these events with a single Gaussian whose normalization and width are allowed to float in the fit. For a very small fraction of events the vertex location is extremely mismeasured. These events have a nearly flat distribution in lifetime. We model this contribution with a broad Gaussian, assigning a fixed width of  $8 \text{ ps}$ . The normalization of this contribution is allowed to float in the fit. The signal PDF is multiplied, on an event-by-event basis, by the mass-dependent signal probability from the  $D^0$  candidate mass fit.

The background lifetime distribution contains two pieces: a prompt piece and a piece with nonzero lifetime. The component with nonzero lifetime comes from partially reconstructed charm decays. We model this component with a single exponential where the lifetime is another parameter of the fit. We expect the fitted value of the background lifetime to be consistent with the  $D^0$  lifetime. The relative amount of background with and without lifetime is also allowed to float in the fit. Both sorts of background are convolved with a resolution function that is modeled in the same manner as the

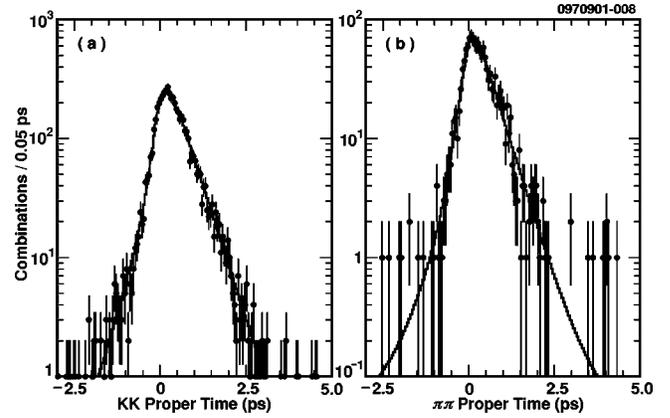


FIG. 4. The proper time distribution for all  $D^0 \rightarrow K^+ K^-$  (left) and  $D^0 \rightarrow \pi^+ \pi^-$  (right) candidates included in the fit. The curves are the fit results discussed in the text.

signal, but with an independent set of parameters. The background PDF is multiplied by one minus the signal probability from the mass fit.

The fit to the data is first done for the  $K\pi$  sample with all the parameters, except for the resolution of the very wide Gaussian, which is fixed, left to float. The fit is then repeated for the  $KK$  and  $\pi\pi$  samples with the parameters describing the signal resolution, the overall scale factor, the resolution smearing, the fraction in the “hard scattering” Gaussian, and the fraction in the “broad” Gaussian, fixed to the values found in the  $K\pi$  fit. All the background parameters are allowed to float in the  $KK$  and  $\pi\pi$  fits while for the signal the only variable parameters are the overall level and lifetime. This fit procedure is checked on a set of fully simulated Monte Carlo events representing more than ten times the amount of data for the backgrounds and signal. We observe that this procedure gives correlations between the signal lifetime and other parameters that are small and the measured signal lifetime has a unit pull. These are checked in a fast, smearing-based version of the simulation with very high statistics.

The fit results for all events included in the fit are shown in Fig. 4 and given in Table I. The fit gives resolutions and background parameters that agree with the expectations from simulated MC events. The “hard scattering” Gaussian contains about  $(4 \pm 2)\%$  of the signal, predicted to be  $2\%$  by the simulation, with a resolution of  $(0.60 \pm 0.08)$  of a  $D^0$  lifetime, predicted to be  $0.58$  by the simulation, and a negligible fraction in the “broad” Gaussian, predicted to be  $0.04\%$  in the simulation. The largest correlation for the signal lifetime, which is roughly the same in all the fits, is  $-20\%$  with the lifetime of the partially reconstructed background, which agrees very well with the simulation. All other the correlations are small; less than  $10\%$ . Small corrections to the lifetimes are computed by comparing the generated and measured values in a Monte Carlo analysis on a fully simulated sample, including backgrounds, corresponding to roughly ten times the data sample. These corrections are  $0.0006 \pm 0.0040 \text{ ps}$  in  $K^+ K^-$ ,  $-0.0011 \pm 0.0015 \text{ ps}$  in  $K^- \pi^+$ , and  $0.001 \pm 0.0058 \text{ ps}$  in  $\pi^+ \pi^-$ . Applying these corrections we obtain  $y_{CP}^{KK} = -0.019 \pm 0.030 \pm 0.010$ ,  $y_{CP}^{\pi\pi} = 0.005 \pm 0.046$

TABLE I. Summary of the lifetime fits. The parameters are those described in the text, where  $f_{\text{mis}}$  is the fraction of signal in the second and third Gaussian contributions and  $\sigma_{\text{mis}}$  is the width of the second Gaussian. Note that we have constrained the candidates to a  $D^0$  mass of 1.86514 GeV, the Monte Carlo corrected weighted average of the  $KK$ ,  $\pi\pi$ , and  $K\pi$  data. This mass constraint introduces a systematic bias in the lifetime measurement, which cancels for  $y_{CP}$  which only depends on the ratio of lifetimes. This technique yields the smallest uncertainty in  $y_{CP}$ , but is not optimal for measuring the absolute  $D^0$  lifetime.

Parameter	$K\pi$	$KK$	$\pi\pi$
Number of signal	$20272 \pm 178$	$2463 \pm 65$	$930 \pm 37$
$\tau_{\text{sig}}$ (ps)	$0.4046 \pm 0.0036$	$0.411 \pm 0.012$	$0.401 \pm 0.017$
Background frac. (%)	$8.8 \pm 0.2$	$50.7 \pm 0.7$	$29.1 \pm 1.3$
Background life frac. (%)	$81.0 \pm 4.8$	$85.7 \pm 2.9$	$32.2 \pm 7.5$
$\tau_{\text{back}}$ (ps)	$0.376 \pm 0.030$	$0.436 \pm 0.020$	$0.56 \pm 0.15$
$f_{\text{mis}}$ (%)	$3.8 \pm 0.9$	Fixed	Fixed
$\sigma_{\text{mis}}$ (ps)	$0.590 \pm 0.079$	Fixed	Fixed

$\pm 0.014$ , and combining them in a weighted average we calculate  $y_{CP} = -0.012 \pm 0.025 \pm 0.009$ , where the second error is from the Monte Carlo statistics.

We check the data for bias in several different parameters. We plot the fitted value of  $y_{CP}$  versus azimuthal angle, polar angle, date the data were collected, momentum of the candidate  $D^0$ ,  $\cos \theta^*$ , and confidence level of the vertex constraint. We find no significant biases in any of these distributions.

The kinematics of  $K\pi$ ,  $KK$ , and  $\pi\pi$   $D^0$  decays are slightly different due to the different amount of kinetic energy released. This will result in the signal resolution functions being slightly different. We have constrained all of the signal resolution functions to be the same. Studying this effect in Monte Carlo calculations and data we estimate the following systematic uncertainties: 0.007 for  $KK$ , 0.003 for  $\pi\pi$ , and 0.005 for the average.

We study the effects of background shape mismodeling by varying the amount and composition of the background. We perform these in data and Monte Carlo calculations and estimate systematic uncertainties of 0.008 for  $KK$ , 0.011 for  $\pi\pi$ , and 0.008 for the average.

We study the effect of our treatment of the proper time outlier events, which we have modeled with a wide Gaussian of fixed width. We vary the value of the width used in the wide Gaussian and also eliminate the wide Gaussian from the resolution function and impose a maximum proper time limit instead. From these studies we estimate systematic uncertainties of 0.002 for  $KK$ , 0.001 for  $\pi\pi$ , and 0.002 for the average.

We investigate the bias introduced by constraining all of the events to the same  $D^0$  mass by removing this constraint. We take the difference between the constrained and uncon-

strained fits as a systematic uncertainty: 0.005 in  $KK$ , 0.005 in  $\pi\pi$ , and 0.005 in the average. Length scale uncertainties have been studied previously by CLEO [10] and contribute negligible uncertainty to  $y_{CP}$ .

Summing all of the listed systematic uncertainties in quadrature, including the Monte Carlo statistics, we obtain the final results  $y_{CP}^{KK} = -0.019 \pm 0.029 \pm 0.016$ ,  $y_{CP}^{\pi\pi} = 0.005 \pm 0.043 \pm 0.018$ . Combining the two results we obtain  $y_{CP} = -0.012 \pm 0.025 \pm 0.014$ , which is consistent with zero and recent measurements [16].

In summary, we have used the CLEO II.V data set to obtain the world's most precise measurements of  $R_{KK} = \Gamma(D^0 \rightarrow K^+ K^-) / \Gamma(D^0 \rightarrow K^- \pi^+) = (10.40 \pm 0.33 \pm 0.27)\%$  and  $R_{\pi\pi} = \Gamma(D^0 \rightarrow \pi^+ \pi^-) / \Gamma(D^0 \rightarrow K^- \pi^+) = (3.51 \pm 0.16 \pm 0.17)\%$ , and the direct  $CP$  asymmetries  $A_{CP}^{KK} = (0.0 \pm 2.2 \pm 0.8)\%$  and  $A_{CP}^{\pi\pi} = (1.9 \pm 3.2 \pm 0.8)\%$ . We have also performed a competitive measurement of the normalized mixing parameter  $y_{CP} = -0.012 \pm 0.025 \pm 0.014$ . In all cases the first error is statistical and the second is systematic. Our partial width measurements are consistent with the previous world average, we see no evidence for direct  $CP$  violation in Cabibbo-suppressed  $D^0$  decays, and we measure a value of the mixing parameter  $y_{CP}$  consistent with zero.

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