Self-consistent optimization of the *z* expansion for *B*-meson decays

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We discuss the self-consistency imposed by the analyticity of regular parts of form factors, appearing in the z expansion for semileptonic B-meson decays, when fitted in different kinematic regions. Relying on the uniqueness of functions defined by analytic continuation, we propose four metrics which measure the departure from the ideal analytic self-consistency. We illustrate the process using Belle data for $B \rightarrow D\ell\nu_{\ell}$ with the two kinematic regions chosen as the five low-z and the five high-z bins. For this specific example, the metrics provide consistent indications that some choices (order of truncation, Boyd-Grinstein-Lebed or Bourrely-Caprini-Lellouch) made in the form of the z expansion can be optimized. However, other choices (z origin, location of isolated poles and threshold constraints) appear to have very little effect on these metrics. On the other hand, changing the kinetic regions affects the results and should also be considered in the optimization process. We briefly discuss the implication for optimization of the z expansion for nucleon form factors relevant for neutrino oscillation experiments.

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I. INTRODUCTION

Experimental differential decay rates for exclusive semileptonic decays of B mesons [1–9], combined with ab initio lattice QCD calculations of the hadronic form factors [10-24] provide reliable numerical estimations of the Cabibbo-Kobayashi-Maskawa matrix elements $|V_{ub}|$ and $|V_{cb}|$. Accurate lattice calculations are only possible for a large enough invariant square of the 4-momentum of the leptons, denoted q^2 , or more specifically when the recoil energy of the final state meson is significantly smaller than the inverse lattice spacing. In order to predict the shape of the differential decay rate over the entire kinematic range from reliable lattice results in the high- q^2 region, an analytic continuation method developed in the context of kaon decays [25,26] has been adapted for B-meson decays by Boyd, Grinstein, and Lebed (BGL) [27] and Bourrely, Caprini, and Lellouch (BCL) [28]. The method is often called the z expansion. The basic idea is to map the branch cut in the complex q^2 plane onto the boundary of the unit disk in z with the rest of the cut complex q^2 plane being mapped into the interior of the disk. The goal is to find parametrizations of the form factors for specific processes where the effects of thresholds and isolated poles can to some extent be separated from a smooth behavior in the

kinematic range. Ideally, after the mapping, the kinematic range becomes a small interval near the origin and a few terms in the Taylor expansion provide reliable results. General strategies for combining the lattice and experimental data are discussed in Ref. [11].

The extrapolation of lattice results with computationally accessible q^2 to the full kinematic range relevant for experimental analysis has been performed for various decay modes and by various collaborations [10-24]. Specific choices will be reviewed below. In general, the agreement with the overall shape of the experimental differential decay rate provides a strong guidance to select reasonable procedures. If one assumes the standard model is correct then an *ab initio* calculation in the full kinematic range should reproduce the shape of the experimental data. Under this assumption, the only unknown quantity is V_{ch} , a Cabibbo-Kobayashi-Maskawa matrix element. The z expansion being a compact and model-independent method is very important to summarize the experimental results, especially as it does not depend on the binning procedure. Recent experiments provide fits of their data using the zexpansion. This amounts to continuous functions that allow comparisons among experiments with different binnings. For semileptonic decays of B mesons involving tree-level virtual W^{\pm} bosons, the form factor can be expressed in term of analytic functions in the entire kinematic interval. An important implication is that a (perfect) knowledge of the analytic function in any open set in the complex q^2 plane uniquely determines the function in the whole interval provided that no singularities or cuts prevent the analytical continuation [29].

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In this article, we discuss the self-consistency imposed by analyticity on regular parts of form factors when fitted in different kinematic regions. In Sec. II, we review the BGL and BCL parametrizations. In Sec. III, we consider existing goodness of fit measures (χ^2 and the Akaike information criterion (AIC)] and define four dimensionless metrics which measure the departure from ideal analytic selfconsistency. These metrics are "cost functions" for which a large value indicates an inconsistent parametrization conflicting with the assumed analyticity as defined mathematically in [29]. We illustrate the idea by calculating these four metrics for $B \to D\ell \nu_\ell$ using partial decay widths provided by the Belle Collaboration [6] with the two kinematic regions chosen as the five low-z and the five high-z bins. The numerical results are analyzed in Sec. IV where we discuss the possibility of discriminating among a certain number of choices (order of truncation, BGL or BCL, z origin, and threshold constraints) made in the z expansion. In Appendix D, we repeat the calculations with the two kinematic regions chosen as the six low-zand the four high-z bins. The results are summarized in the conclusions where we also comment on new methods of determining the order of truncation of the z expansion [24,30]. We also briefly mention possible applications for optimization of lattice nucleon form factors reviewed in [31] and relevant for neutrino experiments such as DUNE.

II. BGL AND BCL PARAMETRIZATIONS

In the following, we focus on different parametrizations of the form factor that describe the decays $B^0 \to D^- \ell^+ \nu_{\ell}$ and $B^+ \to D^0 \ell^+ \nu_{\ell}$, with $\ell = \{e, \mu\}$. In the isospin limit, these processes can be described by a differential decay rate that depends on the hadronic recoil variable $w \equiv (m_B^2 + m_D^2 - q^2)/(2m_B m_D)$ [6],

$$\frac{d\Gamma}{dw} = K(w^2 - 1)^{3/2} f_+(w)^2, \tag{1}$$

with

$$K = \frac{G_F^2 m_D^3}{48\pi^3} |V_{cb}|^2 (m_B + m_D)^2 \frac{4r}{(1+r)^2} \eta_{\rm EW}^2, \quad (2)$$

where G_F is the Fermi coupling constant, m_B and m_D are the masses of the *B* and *D* mesons, respectively, $r = m_D/m_B$, and η_{EW} represents the electroweak corrections.

The two parametrizations of the vector form factor $f_+(w)$ that we investigate are the BGL and BCL parametrizations. Both parametrizations use the *z* expansion which takes the real kinematic range and embeds it into a complex domain, where the process of analytical continuation defines a unique regular (analytic and single-valued) function. The mapping variable is $z(q^2, t_0)$, where

$$z(q^2, t_0) = \frac{\sqrt{t_+ - q^2} - \sqrt{t_+ - t_0}}{\sqrt{t_+ - q^2} + \sqrt{t_+ - t_0}},$$
(3)

 q^2 is the momentum transfer, $q^2 = m_B^2 + m_D^2 - 2wm_Bm_D$, and $t_+ = (m_B + m_D)^2$. This change of coordinates maps the cut complex q^2 plane onto the unit disk. At threshold, $q^2 = t_+$ and z = -1. The cut is mapped into the boundary of the disk. The variable t_0 determines where the z expansion is centered about. We consider two t_0 values $t_0 = t_{opt} = (m_B + m_D)(\sqrt{m_B} - \sqrt{m_D})^2$ and $t_0 = t_- = (m_B - m_D)^2$. The choice of t_0 should not appreciably affect the z expansion fit results, but it can be used to adjust the systematic uncertainties. Following the original authors, $t_0 = t_{opt}$ is used with BCL [28] and $t_0 = t_-$ is used with BGL [27]. The choice $t_0 = t_{opt}$ puts z in the range $z \in [-0.0323, 0.0323]$, and the choice $t_0 = t_-$ puts z in the range $z \in [0.0, 0.0646]$.

We define the BGL parametrization as $f_{+,BGL}$, with the explicit form used for a lattice calculation [16] and the analysis of the Belle data [6], both for $B \rightarrow D\ell \nu_{\ell}$,

$$f_{+,\text{BGL}}(z) \equiv \frac{1}{\phi_{+}(z)} \sum_{n=0}^{N} a_{+,n} z^{n}, \qquad (4)$$

with

$$\phi_{+}(z) = 1.1213(1+z)^{2}(1-z)^{1/2} \\ \times \left[(1+r)(1-z) + 2\sqrt{r}(1+z) \right]^{-5}.$$
 (5)

The outer function, $\phi_+(z)$, is to some extent arbitrary but must be analytic and nonzero for |z| < 1 in order to enforce the unitarity condition on $a_{+,n}$ [6,27]. We then define the BCL parametrization as $f_{+,BCL}$, with the explicit form used for a lattice calculation [12] $B \to \pi \ell \nu_{\ell}$ but with $m_{B^*}^2$ replaced by $m_{B^*}^2$,

$$f_{+,\text{BCL}}(z) \equiv \frac{1}{1 - q^2(z)/m_{B_c^*}^2} \sum_{k=0}^{K-1} b_{+,k} \left[z^k - (-1)^{k-K} \frac{k}{K} z^K \right].$$
(6)

One difference between the BGL and BCL parametrizations is the use of the threshold condition discussed in Appendix C. Here, BGL does not use the threshold condition, while BCL does use it. The free parameters $a_{+,n}$ and $b_{+,n}$ are fitted by using least square fitting methods [32] and must satisfy the following unitarity conditions [27]:

$$\sum_{n=0}^{N} |a_{+,n}|^2 \le 1,$$
(7)

and [12,28]

$$\sum_{j,k=0}^{K} B_{jk} b_{+,j} b_{+,k} \le 1.$$
(8)

We have checked that the inequalities are satisfied [33]. For $f_{+,BGL}$ we consider N = 0, 1, 2 and for $f_{+,BCL}$ we consider K = 1, 2, 3. It is important to note that N is not the number of parameters while K is, so to avoid confusion, the number of parameters used in the fit will be denoted n_p for both parametrizations, where $n_p = N + 1$ and $n_p = K$. The z^K term that is attached to every $b_{+,k}$ comes from the threshold condition which will be discussed in more detail in Appendix C.

The BGL parametrization sometimes includes a Blaschke factor $P_+(z)$ as well, which contains the information about the pole at $q_*^2 \equiv m_{B_*}^2 = 40.02 \text{ GeV}^2$. However, it has been shown that the Blaschke factor does not appreciably affect the z fit for the BGL analysis of $B \to D\ell \nu_\ell$ due to the pole being very far away from the kinematical region [16]. For this reason the Blaschke factor has been set to 1 in [6,16] and our definition in Eq. (4) follows this choice, while BCL replaces the commonly used outer function and Blaschke factor with a prefactor that has a pole at the same location as the Blaschke factor. The pole q_*^2 corresponds to $z(q_*^2, t_-) = -0.308$ and $z(q_*^2, t_{opt}) = -0.337$. The construction of the B_{jk} matrix can be found in [28], and we calculate the values B_{00} , B_{01} , B_{02} , and B_{03} for $B \to D\ell \nu_{\ell}$ and display them in Table I. The remaining B_{mn} values can be calculated using the following relations [28]:

$$B_{i(i+k)} = B_{0k}, (9)$$

and

$$B_{jk} = B_{kj}.\tag{10}$$

As discussed above, the outer functions are to some extent arbitrary so long as they are analytic and nonzero in the z range that we are interested in. In addition, the Blaschke factor is sometimes set equal to 1 [6,16] because the poles appear to be sufficiently far away from the semileptonic kinematic region. It is then interesting to investigate the parametrizations of Eqs. (4) and (6) with their prefactors set equal to 1. To differentiate the form factors when there are no prefactors, we denote Eq. (4) with

TABLE I. The matrix elements B_{jk} which are used in the BCL unitarity condition for $n_p = 1, 2, 3$.

<i>B</i> ₀₀	B_{01}	B_{02}	<i>B</i> ₀₃
0.0118	-0.0028	-0.0069	0.0038

no prefactors as $f_{+,NN}$ and we denote Eq. (6) with no prefactors as $f_{+,NT}$, where NN stands for no prefactor and no threshold, while NT stands for no prefactor with threshold. The explicit forms of $f_{+,NN}$ and $f_{+,NT}$ can be found below,

$$f_{+,\rm NN}(z) = \sum_{n=0}^{N} a_{+,n} z^n,$$
 (11)

and

$$f_{+,\rm NT}(z) = \sum_{n=0}^{N-1} b_{+,n} \bigg[z^n - (-1)^{n-N} \frac{n}{N} z^N \bigg], \quad (12)$$

where NN uses $t_0 = t_-$ and NT uses $t_0 = t_{opt}$.

III. NEW METRICS AND RESULTS

In this section, we consider several tests that compare the goodness of fit for the Belle data with various parametrizations and introduce new metrics to compare the selfconsistency of these parametrizations.

A. χ^2 test

We use the LsqFit Python library [32] to perform the fits of the different models with the Belle data, and the resulting fit parameters can be found in Appendix A. The LsqFit library also provides the χ^2 and reduced- χ^2 , χ^2_{ν} , and these values are provided for BGL, BCL, NN, and NT with one, two, and three parameters, denoted 1p, 2p, and 3p, respectively, in Table II. We also provided the *p*-values which take into account the number of fitting parameters.

It is clear that the 2p and 3p cases have very significantly lower χ^2 and χ^2_{ν} then the 1p case. However it is less clear that the χ^2 , χ^2_{ν} , and *p*-values provide significant discrimination between 2p and 3p for a given model or discrimination among the different models. For a given model, the χ^2_{ν} of the 2p are typically 5%–10% smaller than for the 3p and the *p*-values typically 10% larger for 2p. However, the difference of theses quantities among the 2p are smaller. Note also that all the 3p fits are very close to each other and have the same χ^2 with three significant digits.

To further investigate the differences between the form factors using BGL and BCL with 2p and 3p and give an idea of the role played by the quadratic terms (curvature), we set the form factor with BCL 3p as a reference and calculate the relative deviation with other parametrizations options. The results can be found in Fig. 1. We see that the differences between BCL 3p and BGL 3p are an order of magnitude smaller than the differences for either of the 2p cases, which also helps explain why the global metrics had identical values for BGL and BCL with 3p.

It is also possible to estimate the role of the curvature considering the value of the quadratic term for the largest

TABLE II. The χ^2 , χ^2_{ν} , AIC, C_0 , C_1 , D_1 , and D_2 values calculated from the fits performed on the differential decay width data with the BGL, BCL, NN, and NT parametrizations with 1p, 2p, and 3p. The *p*-values are also calculated using BGL and BCL with 1p, 2p, and 3p for the 5L5H partition.

	χ^2	$\chi^2_{ u}$	<i>p</i> -value	AIC	C_0	C_1	D_1	D_2
BGL 1p	99	11	2.5×10^{-17}	101	15.76	39.44	41.85	10.95
BGL 2p	4.56	0.57	0.803	8.56	1.18	5.16	5.78	4.76
BGL 3p	4.55	0.65	0.715	10.55	17.56	91.76	104.64	117.68
BCL 1p	33.3	3.7	0.0001	35.3	3.37	7.63	8.02	1.90
BCL 2p	4.64	0.58	0.795	8.64	0.63	2.72	3.00	1.74
BCL 3p	4.55	0.65	0.715	10.55	15.23	78.96	89.85	100.37
NN 1p	135	15	1.1×10^{-24}	137	24.50	64.26	68.54	19.02
NN 2p	4.88	0.61	0.770	8.88	1.85	7.82	8.88	8.21
NN 3p	4.55	0.65	0.715	10.55	19.10	100.16	115.43	126.79
NT 1p	135	15	1.1×10^{-24}	137	24.50	64.31	68.54	19.02
NT 2p	5.04	0.63	0.753	9.04	2.11	8.83	10.03	9.55
NT 3p	4.55	0.65	0.715	10.55	19.81	104.04	119.70	131.76

value of z in units of the constant term. It amounts to 2% in the BGL case and 1% in the BCL case. The difference between the 2p and 3p series can be obtained from data presented in Appendix A: $0.0001 - 0.0056z + 0.07z^2$ for BGL and $0.002 + 0.13z - 8z^2$ for BCL. In order to allow meaningful comparison, we can divide these difference of series by the zeroth-order coefficient for 3p and obtain $0.00801282 - 0.448718z + 5.60897z^2$ for BGL and $0.00258065 + 0.167742z - 10.3226z^2$ for BCL. For the largest values of z, the three orders contribute, respectively, in absolute value as (0.00801282, 0.0289872, 0.0234071) for BGL and (0.00258065, 0.00541806, 0.0107694) for BCL. As these numbers represent relative changes due to the curvature, it is clear that the effects are small for BGL and even smaller for BCL.

B. Akaike information criterion

A test that enables quantitative comparisons between models with differing numbers of parameters that are not rigorously possibly without Bayesian techniques is the



FIG. 1. The relative deviation of our reference form factor parametrization BCL 3p with the other parametrization options, plotted as a function of q^2 .

AIC. The AIC value is defined like an augmented χ^2 value, where the augment is adding a $2n_p$ term [34,35],

$$AIC = 2n_p + \chi^2. \tag{13}$$

The inclusion of a penalty which is linear in n_p is used to discourage overfitting, and the factor 2 in front of n_p is discussed in [36]. Therefore, the preferred model will be the model that has the lowest AIC value. Changes in the AIC values Δ AIC when the number of degrees of freedom ν is changed by $\Delta\nu$ can be considered significant if $|\Delta$ AIC $|/AIC| > |\Delta\nu|/\nu$ [37].

These AIC values are displayed in Table II. The $2n_p$ term for 1p will not compensate for the drastically larger χ^2 value compared to the 2p and 3p fits, clearly showing that 1p is not descriptive enough. The AIC values in Table II show that all the 2p cases satisfy the inequality regarding ν , which indicates that the 2p case is preferred over the 3p case. However, the AIC values between BGL and BCL for the same number of parameters is still too close to determine anything significant about which of the models does a better job of fitting the Belle data. We proceed to define our own metrics to find one that is able to distinguish between the different parametrization options that we consider.

C. Self-consistency metrics

In the form factor expressions, the polynomials in z are approximations of analytic functions in the kinematic range. The absence of singularities or cuts in that range implies that the exact knowledge of the function in an open region can uniquely determine the function in another region. This can be achieved by analytic continuation [29]. In the current context, if an analytic function is defined on an open segment of the real z axis corresponding to the kinematic range and if we partition this segment in a region

 \mathcal{H} corresponding to a high-*z* (or equivalently high-*w* or low- q^2) part and the complementary region \mathcal{L} in the low-*z* region, it is then clear that ideally the perfect knowledge of the function in \mathcal{H} uniquely determines the function in \mathcal{L} and vice versa.

In practice, if we use experimental data, we know the function at a finite number of points with a limited accuracy. It is expected that, if we obtain a polynomial approximation in \mathcal{H} using the data in \mathcal{H} , we call $f_+^{\text{high}}(z)$ and extend this polynomial to \mathcal{L} , and if we obtain $f_+^{\text{low}}(z)$ by swapping the roles of \mathcal{H} and \mathcal{L} , then the discrepancy

$$\Delta f(z) \equiv f_+^{\text{high}}(z) - f_+^{\text{low}}(z) \tag{14}$$

is nonzero and provides a measure of the inconsistency of the continuations due to imperfect knowledge of the function in addition to the uncertainty in the data.

A rough global measure of the inconsistency of a specific method used to obtain the polynomial approximation could be the L^2 norm of $\Delta f(z)$. This quantity depends on the units of the form factor and the range of z in the integral. For a decent approximation, one would expect that $(\Delta f(z))^2$ would be of the order of the average experimental variance $\bar{\sigma}_{exp}^2$ and we could expect to get a quantity of order one by dividing by the length of the z interval and the average experimental variance $\bar{\sigma}_{exp}^2 = 0.00199$. For these reasons we start with the dimensionless quantity

$$C_0 \equiv \frac{1}{\bar{\sigma}_{\exp}^2 |z_{\max} - z_{\min}|} \int_{z_{\min}}^{z_{\max}} (\Delta f(z))^2 dz.$$
(15)

A more refined metric denoted C_1 can be obtained by weighting *locally* with the inverse local variance $\sigma_{\exp}^2(z)$ obtained from the experimental data by interpolating with LsqFit,

$$C_{1} \equiv \frac{1}{|z_{\max} - z_{\min}|} \int_{z_{\min}}^{z_{\max}} \frac{(\Delta f(z))^{2}}{\sigma_{\exp}^{2}(z)} dz.$$
 (16)

When the experimental form factors are provided as binned data with n_b bins, we can define a discrete version of C_1 as

$$D_1 = \frac{1}{n_{\rm bin}} \sum_{i=1}^{n_{\rm bin}} \left(\frac{\Delta f_i}{\sigma_i} \right)^2, \tag{17}$$

with $\Delta f_i = \Delta f(z_i)$, z_i being in the middle of the *i*th bin. This can be calculated in a straightforward way without the need of interpolations. If the bins are narrow enough, we expect that $D_1 \simeq C_1$. The general form of D_1 is reminiscent of a χ square, however, σ_i^2 is not the variance of Δf_i .

Given that the experimental binned data may involve significant correlations among the bins, we can pursue the analogy and generalize D_1 to

$$D_2 = \frac{1}{n_{\text{bin}}} \sum_{i,j=1}^{n_{\text{bin}}} \Delta f_i \mathcal{C}_{ij}^{-1} \Delta f_j, \qquad (18)$$

with C_{ij} the covariance matrix of the binned data for the form factor.

The covariance matrix C_{ij} is calculated using sampled bootstrap form factor data points. Using the GVAR Python library [38], we generated $M = 10^4$ bootstrap differential decay width datasets generated from the Belle data using the underlying covariance matrix. Then using Eqs. (1) and (2), we converted the generated differential decay width data into data describing the form factor, with f^i being the set of random form factor data in the *i*th bin. Finally, we calculated C_{ij} using

$$C_{ij} = \frac{1}{M} \sum_{m=1}^{M} \left(f_m^i - \bar{f}^i \right) \left(f_m^j - \bar{f}^j \right), \tag{19}$$

where f_m^i is the *m*th data point in the *i*th bin and \bar{f}^i is the mean of the data in the *i*th bin. And the σ_i^2 from Eq. (19) is the diagonal entries of the covariance matrix, $\sigma_i^2 = C_{ii}$.

The Belle data have $n_{\text{bin}} = 10$ bins, and we split the data in half between the \mathcal{L} and \mathcal{H} regions. We denote this choice 5L5H. In Appendix D, we consider a six low-z four high-z division denoted 6L4H. We then fit the free parameters $a_{+,n}$ and $b_{+,n}$ to the \mathcal{L} region data and the \mathcal{H} region data separately. If we had perfect knowledge of f_+ in the \mathcal{L} region we could reconstruct it in the \mathcal{H} region and vice versa. An example of this is shown for BGL with 2p in Fig. 2 and for BCL with 2p in Fig. 3.

We carry out this method for BGL and BCL with 2p and 3p, use the resulting parameters to plot the form factor $f_+(z)$, and convert the Belle data from differential decay width data to form factor data and include it in Figs. 4 and 5. Both BGL and BCL have more overlap between the fits for 2p than for 3p. In the 3p fits case, the error band is very small



FIG. 2. The results of our fits using the BGL parametrization. The orange line is the mean value of our fit and the lighter orange region is the 1- σ error band, the black triangles are the Belle data with error bars, and the dashed arrow indicates whether region \mathcal{L} (left) or region \mathcal{H} (right) was used in the fit.



FIG. 3. Same as Fig. 2 but for the BCL parametrization.



FIG. 4. Plots of the form factor $f_+(z)$ vs z using the results of our fits from regions \mathcal{L} and \mathcal{H} for the 5L5H partition. The results of our BGL fit with 2p (left), the results of our BCL fit with 2p (right). The blue region indicates the fit from region \mathcal{H} and the green region indicates the fit from region \mathcal{L} . In both cases, the solid lines are for the z values used in the fit region and the dashed lines are for the z values used in the extrapolated region.



FIG. 5. Same as Fig. 4 but for 3p.

in the \mathcal{H} region and the error band increases in size as it moves into the \mathcal{L} region, but the \mathcal{L} fit has large error bands in both the \mathcal{L} and \mathcal{H} regions.

Now using the results of the fits for every parametrization with 1p, 2p, and 3p, we are able to calculate the discrete and continuous metrics. The values of C_0 and C_1 that we calculate are listed in Table II for all parametrizations with 1p, 2p, and 3p, and D_1 and D_2 are similarly shown in Table II.

Since $\Delta f(z)$ reflects the discrepancy between fits and extrapolations, the most self-consistent model is the one with the lowest values of C_0 , C_1 , D_1 , and D_2 . We see that the 2p values are all considerably lower than the corresponding 1p or 3p values and are clearly preferred, which agrees with what we obtained with the AIC metrics. We define the mean values of the C_0 , C_1 , D_1 , and D_2 for only the 2p models as $\bar{C}_0^{2p} = 1.443$, $\bar{C}_1^{2p} = 6.133$, $\bar{D}_1^{2p} = 6.923$, and $\bar{D}_{2}^{2p} = 6.065$. Comparing the C_0 , C_1 , and D_1 values from the 2p models with these mean values, the values from the BCL parametrization lie 55%-56% below the mean values, the values from the BGL parametrization lie 15%–18% below, the values from the NN parametrization lie 27%–28% above the mean values, and the values from the NT parametrization lie 43%-46% above the mean values. For the D_2 values from the 2p models, BCL lies 71% below the mean value, BGL lies 21% below the mean value, NN lies 35% above the mean value, and NT lies 57% above the mean value. For each metric with 2p compared to the mean values with 2p, the BCL parametrization is significantly lower than the mean and appears to be the most self-consistent from the point of view of analyticity.

We have also considered the effects of relaxing the threshold condition in BCL in C and changing the values of t_0 , namely, using t_{opt} for BGL or t_- for BCL in Appendix B. The tables make clear that these choices affect the metrics by at most a few percent and are essentially irrelevant.

IV. ANALYSIS OF SELF-CONSISTENCY METRICS

In this section, we discuss the relations among the new metrics. First we will compare different metrics for a given parametrization and then we will compare the same metric but for different parametrizations.

A. Comparing the metrics for a given parametrization

The first comparison that we make is between C_1 and C_0 . The ratios C_1/C_0 can be found in Table III. The values in the 2p row are consistent within ~3% of each other, and the values in the 3p row are within $\lesssim 1\%$ of each other. In other words, the two metrics are proportional, with a proportionality constant which depends mostly on the number of parameters used in the fit. Note that the absolute normalization of C_0 or C_1 is not important. From the standard deviations in the binned data [6], we have

$$\frac{1}{n_{\rm bin}} \sum_{i=1}^{n_{\rm bin}} \frac{1}{\sigma_i^2} \simeq \frac{2.81}{\left(\frac{1}{n_{\rm bin}} \sum_{i=1}^{n_{\rm bin}} \sigma_i\right)^2},\tag{20}$$

which partially explains that C_1 is larger than C_0 . We next look at D_1/C_1 . Since D_1 is a discrete version of C_1 , we expect relative differences of the order of $1/n_{\text{bin}} = 0.1$. We see that this is the case in Table IV.

TABLE III. The ratio of C_1 to C_0 calculated with BGL, BCL, NN, and NT for 2p and 3p for the 5L5H partition.

	$C_{1,\mathrm{BGL}}/C_{0,\mathrm{BGL}}$	$C_{1,\mathrm{BCL}}/C_{0,\mathrm{BCL}}$	$C_{1,\rm NN}/C_{0,\rm NN}$	$C_{1,\mathrm{NT}}/C_{0,\mathrm{NT}}$
2p	4.373	4.318	4.227	4.185
3p	5.226	5.185	5.244	5.252

TABLE IV. The ratio of D_1 to C_1 calculated with BGL, BCL, NN, and NT for 2p and 3p for the 5L5H partition.

	$D_{1,\mathrm{BGL}}/C_{1,\mathrm{BGL}}$	$D_{1,\mathrm{BCL}}/C_{1,\mathrm{BCL}}$	$D_{1,\rm NN}/C_{1,\rm NN}$	$D_{1,\mathrm{NT}}/C_{1,\mathrm{NT}}$
2p	1.120	1.103	1.136	1.136
3p	1.140	1.138	1.154	1.151

TABLE V. The ratio of D_2 to D_1 calculated with BGL, BCL, NN, and NT for 2p and 3p for the 5L5H partition.

	$D_{2,\mathrm{BGL}}/D_{1,\mathrm{BGL}}$	$D_{2,\mathrm{BCL}}/D_{1,\mathrm{BCL}}$	$D_{2,\rm NN}/D_{1,\rm NN}$	$D_{2,\mathrm{NT}}/D_{1,\mathrm{NT}}$
2p	0.824	0.580	0.925	0.952
3p	1.125	1.117	1.098	1.101

In summary, we found that the three metrics C_0 , C_1 , and D_1 provide consistent estimates of the departure from analyticity. For instance, we could just consider D_1 , which is easier to calculate from experimental binned data. So far, we have ignored correlations among the bins. Table II of Ref. [6] shows that these correlations are significant, which motivated the introduction of D_2 . The ratios D_2/D_1 are provided in Table V. We see that for 2p the ratios have a stronger dependence on the parametrization which amplifies the discrimination (a lower D_1 means an even lower D_2). On the other hand, for the suboptimal choice 3p, the ratio is about 1.1 in the four cases and D_2 does not provide new information.

The values in the 3p row are consistent within $\lesssim 1\%$ of the other values in the 3p row, however, we see that the D_2/D_1 ratio for 2p is smaller for BCL than it is for BGL, NN, or NT by roughly 43%. It is great that these comparisons show consistency, and they show that the D_2 metric provides the most information to discriminate between the parametrizations.

B. Comparing the same metric for different parametrizations

Since the NN and NT parametrizations are not ever mentioned in the literature, at this point we ignore them and focus again on BGL and BCL as they were shown to be preferred over NN or NT by every metric we considered. Now comparing the same metric between the BGL and BCL parametrizations, the ratios $C_{i,\text{BCL}}/C_{i,\text{BGL}}$ for i = 0, 1are given in Table VI and $D_{j,\text{BCL}}/D_{j,\text{BGL}}$ for j = 1, 2 in

TABLE VI. The ratio of $C_{i,BCL}$ to $C_{i,BGL}$ calculated for both 2p and 3p with both choices of t_0 for the 5L5H partition.

	$C_{0,\mathrm{BCL}}/C_{0,\mathrm{BGL}}$		$C_{1,\mathrm{BCL}}/C_{1,\mathrm{BGL}}$	
	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$
2p	0.536	0.545	0.523	0.532
3p	0.862	0.866	0.855	0.860

TABLE VII. The ratio of $D_{j,BCL}$ to $D_{j,BGL}$ calculated for both 2p and 3p with both choices of t_0 for the 5L5H partition.

	$D_{1,\mathrm{BCL}}/D_{1,\mathrm{BGL}}$		$D_{2,\mathrm{BCL}}/D_{2,\mathrm{BGL}}$	
	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$
2p	0.515	0.524	0.360	0.370
3p	0.854	0.859	0.849	0.854

Table VII. For this, we recalculated the BGL parametrization with $t_0 = t_{opt}$ and the BCL parametrization with $t_0 = t_-$, in order to have the *z* expansion consistent when comparing different parametrizations.

It is important is observe that the values in Tables VI and VII for the 3p rows are all consistent within $\lesssim 1\%$ of the other values in the 3p rows. It is also important to see that the values in the 2p rows are very similar for $C_{0,BCL}/C_{0,BGL}$, $C_{1,BCL}/C_{1,BGL}$, and $D_{1,BCL}/D_{1,BGL}$, however, there is a decrease of roughly 36% in the 2p values for $D_{2,BCL}/D_{2,BGL}$. This shows the consistency of our defined metrics and shows that C_0 , C_1 , and D_1 offer a similar amount of information compared to D_2 , which possibly contains more information about the fits because it is the only metric to differ in these categories when compared.

V. CONCLUSION

In conclusion, we investigated the BGL and BCL parametrizations of the form factor used in the differential decay rate of $B \rightarrow D\ell \nu_{\ell}$. With the experimental binned data collected by the Belle Collaboration [6], we found that the standard χ^2 and χ^2_{ν} do not provide us with enough information to distinguish between BGL vs BCL or 2p vs 3p. The AIC clearly favors 2p over 1p or 3p but the differences between BGL and BCL are too small to be meaningful.

We introduced four metrics or cost functions (C_0 , C_1 , D_1 , and D_2) that measure the discrepancy between fits and extrapolations of the regular parts of form factors in the high and low parts of the kinematic range. Given the analyticity of these regular parts, a perfect fit in one region would provide a unique and perfect analytical continuation in the other region and vice versa. The first metric (C_0) is a dimensionless L^2 norm of the discrepancy. C_1 is a locally

weighted version of C_0 that favors the kinematic regions with smaller experimental uncertainties. D_1 is a discretized version of C_1 which can be implemented directly from the experimental binned data. D_2 is an extension of D_1 that incorporates the correlations among the bins. In view of the significant bin correlations [6], D_2 should be a better measure than D_1 . C_0 , C_1 , and D_1 provide very similar and consistent discriminations, while D_2 somehow amplifies the discriminations for 2p.

All the new metrics strongly favor 2p over 3p. A possible interpretation is that the experimental uncertainties prevent an accurate determination of the quadratic corrections and that one partially extrapolates the experimental noise which is not an analytical function of z. This is consistent with the large uncertainties on the quadratic coefficients already found for global fits as displayed in Appendix A.

All the metrics favor 2p over 1p. Except for D_2 in BCL, the new metrics are significantly larger for 1p. This is clearly consistent with the very low *p*-values for the 1p global fits provided in Table II, which implies that corrections to the constant approximation are significant and result in significantly different constant approximations in the high- and low-*z* regions. Note that for BCL it appears that 1p is a better approximation than 1p for the other parametrizations.

Focusing on the 2p results, we find a finer resolution among parametrizations. For all the new metrics, we observe smaller values for BCL than for the other parametrizations. In addition, BGL does better than no prefactor. It is possible that, in the case considered here, the BCL prefactor captures the features of the actual form factor in a slightly better way. This is hinted at by the fact that a constant approximation has a significantly smaller χ^2 for BCL. This observation may be anecdotal and study of other cases should shed more light on the question. We also found that other choices, such as the value of t_0 or the imposition of a threshold condition, have a marginal impact on the values of the new metrics.

Given that the experimental data came in ten *w* bins, we found it natural to split the interval into 5L5H values to investigate the self-consistency. However, the size of the error bars on the form factors are typically larger for low *w* which motivated a comparison with the case of splitting the Belle data into the 6L4H *w* values. The metrics are given in Appendix D. Some of the features are very similar. The new metrics strongly favor 2p over 3p and 1p. The D_2 metric differs from the original analysis of the 5L5H dataset split, which showed the BCL 1p and 2p as being too close to favor one over the other. We also see the finest resolution in the 2p results among all the parametrizations. In the 6L4H case, we find the smallest values with the BGL parametrization which is unlike the 5L5H case, which saw the

smallest values with the BCL parametrization. In the 6L4H case, we find that the four new metrics are in most cases significantly lower than in the 5L5H case, implying lower discrepancies between the low-*w* data and the high-*w* data with this other subdivision.

It should be emphasized that all the metrics measure discrepancies among fits and not closeness to data. It might be possible to include them in augmented χ^2 [34,35], however, determining the coefficient in front of the metric is a nontrivial task. It should also be noted that very recently, Bayesian inference methods have been used to deal with the truncation question [24] and applied to $B_s \rightarrow K \ell \nu_{\ell}$ [23]. These methods consider higher-order expansions and provide results in agreement with other calculations based on unitarity [30]. It would be very interesting to repeat our analysis using this Bayesian inference procedure for the two sets of bins considered here separately and compare alternative higher-order expansions with our metrics.

So far our calculations of the metrics have been limited to one set of experimental data [6] for $B \to D\ell \nu_{\ell}$ and it is premature to draw general conclusions. Applying the method to other processes involving the z expansion should help identify more general properties. The z expansion has also been used extensively in the study of nucleon form factors. Various neutrino-deuteron scattering experiments have been combined to extract the z expansion of the isovector axial nucleon form factor from experiment [39]. The z expansion has also been used to parametrize lattice calculations of the same quantity, see for instance [40–48] and more references in a recent review article [31]. These parametrizations have been used to incorporate nucleon effects in the calculations of neutrino-nucleus cross section [49]. The method that we proposed can be applied to nucleon form factors as long as one can perform new fits in distinct kinematic regions. This is feasible for binned data, but if extrapolation procedures are involved. such as the continuum limit in lattice calculations, all the details of the existing procedure need to be repeated in kinematic subregions.

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APPENDIX A: OUR CALCULATED FIT PARAMETERS

For completeness, we list the fit parameters that were the result of our fits to the Belle data. For 1p, 2p, and 3p, we show the BGL and BCL fit parameters in Table VIII and the NN and NT parameters in Table X. We also provide the ratios of the fit parameters for BGL and BCL in Table IX and for NN and NT in Table XI. Finally, we provide the parameters for BGL, BCL, NN, and NT using the \mathcal{L} and \mathcal{H} fits in Tables XII–XV. For BGL and BCL, one can see that, for 1p, the fits in the \mathcal{H} region are closer to the global fits; for 2p, the fits in the \mathcal{H} and \mathcal{L} regions are more consistent with each other and the global fits; for 3p, slightly larger variations are observed but the global fits are somehow averages of the \mathcal{H} and \mathcal{L} fits.

TABLE VIII.	The $a_{+,n}$ and b_{+}	$_n$ values that came t	from the global fit of BGL and BCL	•
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	BGL			BCL		
	$a_{+,0}$	$a_{+,1}$	<i>a</i> _{+,2}	$b_{+,0}$	$b_{+,1}$	$b_{+,2}$
1p	0.00804(19)			0.703(16)		
2p	0.01238(41)	-0.0654(58)		0.773(19)	-2.41(42)	
3p	0.01248(67)	-0.071(31)	0.07(37)	0.775(20)	-2.28(64)	-8(28)

TABLE IX. The ratios $a_{+,n+1}/a_{+,n}$ and $b_{+,n+1}, b_{+,n}$ using the central values of the parameters from Table VIII.

	BGL		BCL	
	$a_{+,1}/a_{+,0}$	$a_{+,2}/a_{+,1}$	$\overline{b_{+,1}/b_{+,0}}$	$b_{+,2}/b_{+,1}$
2p	-5.281		-3.112	
<u>3p</u>	-5.710	-0.99998	-2.942	3.716

TABLE X. The $a_{\pm,n}$ and $b_{\pm,n}$ values that came from the global fit of NN and NT.

	NN			NT		
	$a_{+,0}$	$a_{+,1}$	<i>a</i> _{+,2}	$b_{+,0}$	$b_{+,1}$	$b_{+,2}$
1p	0.666(17)			0.666(17)		
2p	1.154(37)	-7.22(53)		0.921(23)	-7.14(53)	
3p	1.181(61)	-8.7(2.8)	18(33)	0.917(24)	-7.56(82)	18(32)

TABLE XI. The ratios $a_{+,n+1}/a_{+,n}$ and $b_{+,n+1}, b_{+,n}$ using the central values of the parameters from Table X.

	NN		NT	
	$\overline{a_{+,1}/a_{+,0}}$	$a_{+,2}/a_{+,1}$	$\overline{b_{+,1}/b_{+,0}}$	$b_{+,2}/b_{+,1}$
2p	-6.257		-7.750	
3p	-7.404	-2.088	-8.236	-2.376

TABLE XII. The $a_{\pm,n}$ values that came from the \mathcal{L} and \mathcal{H} fits of BGL.

		BGL \mathcal{L} fit			BGL \mathcal{H} fit		
	$a_{+,0}$	$a_{+,1}$	<i>a</i> _{+,2}	$a_{+,0}$	$a_{+,1}$	$a_{+,2}$	
1p	0.00825(22)			0.00804(20)			
2p	0.01238(41)	-0.0663(62)		0.01240(40)	-0.0658(57)		
3p	0.01232(75)	-0.063(36)	-0.04(44)	0.01264(69)	-0.079(32)	0.16(37)	

TABLE XIII. The $b_{+,n}$ values that came from the \mathcal{L} and \mathcal{H} fits of BCL.

	BCL \mathcal{L} fit			BCL \mathcal{H} fit		
	$b_{+,0}$	$b_{+,1}$	$b_{+,2}$	$b_{+,0}$	$b_{+,1}$	$b_{+,2}$
1p	0.715(18)			0.703(17)		
2p	0.772(19)	-2.45(45)		0.774(19)	-2.45(42)	
3p	0.774(20)	-2.22(66)	-17(33)	0.774(20)	-2.45(67)	-1(28)

TABLE XIV. The $a_{\pm,n}$ values that came from the \mathcal{L} and \mathcal{H} fits of NN.

		NN \mathcal{L} fit		NN \mathcal{H} fit		
	$a_{+,0}$	$a_{+,1}$	<i>a</i> _{+,2}	$a_{+,0}$	$a_{+,1}$	<i>a</i> _{+,2}
1p	0.929(26)			0.708(17)		
2p	1.227(64)	-10.2(2.1)		1.159(53)	-7.29(85)	
3p	1.12(16)	0.7(14.0)	-237(299)	1.15(26)	-7(11)	-3(108)

TABLE XV. The $b_{+,n}$ values that came from the \mathcal{L} and \mathcal{H} fits of NT.

		NT \mathcal{L} fit			NT \mathcal{H} fit		
	$b_{+,0}$	$b_{+,1}$	$b_{+,2}$	$b_{+,0}$	$b_{+,1}$	$b_{+,2}$	
1p	0.929(26)			0.708(17)			
2p	0.896(28)	-10.3(2.1)		0.923(28)	-7.15(83)		
3p	0.889(30)	-14.7(6.0)	-241(306)	0.923(36)	-7.2(3.8)	-3(104)	

APPENDIX B: INVESTIGATING THE t₀ PARAMETER

The BGL and BCL parametrizations use different choices for t_0 , although the value of t_0 does not affect the size of the *z* range, but it does affect the center of the *z* range. We investigated the effect of t_0 on our metrics by calculating all the metrics using both choices of t_0 . The χ^2 , χ^2_{ν} , and AIC values can be found in Table XVI, and the C_0 , C_1 , D_1 , and D_2 values can be found in Table XVII.

We find that the choice of t_0 has negligible effects on the χ^2 , χ^2_{ν} , and AIC metrics at all with the precision that we consider. However, the C_0 and C_1 as well as the D_1 and D_2 metrics have some minor differences based on the choice of t_0 , but the differences are on the order of 1%. This confirms that the main role of t_0 is to set the central value of the *z* range.

		χ ²	$\chi^2_{ u}$		AIC	
	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$
BGL 1p	99	99	11	11	101	101
BGL 2p	4.56	4.56	0.57	0.57	8.56	8.56
BGL 3p	4.55	4.55	0.65	0.65	10.55	10.55
BCL 1p	33.3	33.3	3.7	3.7	35.3	35.3
BCL 2p	4.64	4.64	0.58	0.58	8.64	8.64
BCL 3p	4.55	4.55	0.65	0.65	10.55	10.55
NN 1p	135	135	15	15	137	137
NN 2p	4.88	4.88	0.61	0.61	8.88	8.88
NN 3p	4.55	4.55	0.65	0.65	10.55	10.55
NT 1p	135	135	15	15	137	137
NT 2p	4.96	5.04	0.62	0.63	8.96	9.04
NT 3p	4.55	4.55	0.65	0.65	10.55	10.55

TABLE XVI. Same as Table II, but every value is calculated with both choices of t_0 for the 5L5H partition.

	C_0		C_1		D_1		D_2	
	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$
BGL 1p	15.76	15.38	39.44	38.44	41.85	41.76	10.95	10.64
BGL 2p	1.18	1.16	5.16	5.11	5.78	5.72	4.76	4.69
BGL 3p	17.56	17.58	91.76	91.83	104.64	104.54	117.68	117.57
BCL 1p	3.37	3.37	7.62	7.63	8.02	8.02	1.90	1.90
BCL 2p	0.63	0.63	2.70	2.72	2.98	3.00	1.71	1.74
BCL 3p	15.13	15.23	78.44	78.96	89.39	89.85	99.85	100.37
NN 1p	24.50	24.50	64.26	64.26	68.54	68.54	19.02	19.02
NN 2p	1.85	1.87	7.82	7.88	8.88	8.95	8.21	8.29
NN 3p	19.10	19.19	100.16	100.66	115.43	115.81	126.79	127.23
NT 1p	24.50	24.50	64.31	64.31	68.54	68.54	19.02	19.02
NT 2p	2.09	2.11	8.72	8.83	9.92	10.03	9.42	9.55
NT 3p	19.69	19.81	103.42	104.04	119.70	119.70	131.16	131.76

TABLE XVII. Same as Table II but every value is calculated with both choices of t_0 for the 5L5H partition.

APPENDIX C: BCL WITH NO THRESHOLD CONDITION

The threshold condition from [28] comes from $z(t_+, t_0) = -1$ which can be seen in Eq. (3) and from the fact that $(z+1) \sim \text{const} \times (q^2 - t_+)^{1/2}$ near z = -1. Then the threshold condition is

$$\left[\frac{df_+}{dz}\right]_{z=-1} = 0. \tag{C1}$$

We investigate the effect of the threshold condition by reproducing our results using the BCL parametrization with no threshold condition, which we call BCL^* and has the form

$$f_{+,\text{BCL}^*}(z) = \frac{1}{1 - q^2(z)/m_{B_c}^2} \sum_{n=0}^N b_{+,n} z^n.$$
(C2)

Using this BCL*, we recalculate all the values in Tables XVI and XVII and the results are displayed in Tables XVIII and XIX.

These values are mostly identical to the BCL values shown in Tables XVI and XVII, with the only differences appearing in C_0 , C_1 , D_1 , and D_2 .

TABLE XVIII. Same as Table XVI but using BCL* for the 5L5H partition.

	χ^2		$\chi^2_{ u}$		AIC	
	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$
BCL* 1p	33.3	33.3	3.7	3.7	35.3	35.3
BCL* 2p	4.64	4.64	0.58	0.58	8.64	8.64
BCL* 3p	4.55	4.55	0.65	0.65	10.55	10.55

TABLE XIX. Same as Table XVII but using BCL* for the 5L5H partition.

	C_0		C_1		D_1		<i>D</i> ₂	
	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$	$t_0 = t_{-}$	$t_0 = t_{\rm opt}$
BCL* 1p	3.37	3.37	7.62	7.63	8.04	8.04	1.93	1.93
BCL* 2p	0.63	0.61	2.65	2.55	2.96	2.84	1.61	1.54
BCL* 3p	17.03	17.88	88.14	77.00	102.68	89.58	109.22	96.28



FIG. 6. Plots of the form factor $f_+(z)$ vs z using the results of our fits from regions \mathcal{L} and \mathcal{H} for the 6L4H partition. The results of our BGL fit with 2p (left), the results of our BCL fit with 2p (right). The blue region indicates the fit from region \mathcal{H} and the green region indicates the fit from region \mathcal{L} . In both cases, the solid lines are for the z values used in the fit region and the dashed lines are for the z values used in the extrapolated region.



FIG. 7. Same as Fig. 6 but for 3p.

APPENDIX D: ADJUSTING THE $\mathcal L$ AND $\mathcal H$ REGIONS

As explained in the conclusions, we also investigated splitting the data into the six lowest w values and the four highest w values, which we abbreviate as 6L4H. The plots in Figs. 6 and 7 show the results of the 6L4H fits for both BGL and BCL with both 2p and 3p.

In Fig. 6, we see much smaller error bands in comparison with Fig. 4 with both the low-*w* fits and the high-*w* fits. We see the same effect in Fig. 7 in comparison to Fig. 5 but to a much larger extent. It is also interesting how the low-*w* fits in Figs. 4

C_0	C_1	D_1	D_2
9.29	23.26	24.47	6.35
0.24	0.65	0.68	0.19
4.09	3.62	3.63	6.12
2.50	5.66	5.90	1.39
0.62	0.66	0.63	0.26
4.72	3.40	3.32	5.20
13.34	34.98	37.26	10.39
0.22	0.98	1.10	0.72
3.68	3.89	3.99	7.13
13.34	35.01	37.26	10.39
0.25	1.14	1.29	0.98
3.53	4.04	4.18	7.56
	9.29 0.24 4.09 2.50 0.62 4.72 13.34 0.22 3.68 13.34 0.25	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$

TABLE XX. The C_0 , C_1 , D_1 , and D_2 values for BGL, BCL, NN, and NT with 1p, 2p, and 3p. for the 6L4H partition.

and 5 underestimate the data in the \mathcal{H} region, but when we adjust the \mathcal{L} region then the 2p fits underestimate the data less in \mathcal{H} , but the 3p \mathcal{L} fits now overestimate the data in \mathcal{H} .

We also recalculated the metrics found in Table II but with 6L4H and display the results in Table XX. Every metric in Table XX is lower for the corresponding metric in Table II. The metrics from the 3p parametrizations had the most significant decrease with the change to 6L4H.

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