## Search for the decay $B_{s}^{0} \rightarrow \pi^{0} \boldsymbol{\pi}^{0}$ at Belle

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#### Abstract

We report the results of the first search for the decay $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ using $121.4 \mathrm{fb}^{-1}$ of data collected at the $\Upsilon(5 \mathrm{~S})$ resonance with the Belle detector at the KEKB asymmetric-energy $e^{+} e^{-}$collider. We observe no signal and set a $90 \%$ confidence level upper limit of $7.7 \times 10^{-6}$ on the $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ decay branching fraction.


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The study of heavy-flavored hadrons decaying to hadronic final states provides an important input for understanding the interplay between strong and weak interactions. These types of decays involving weak annihilation amplitudes can be a promising place to look for disagreement between theoretical predictions and experimental observations. These decays are highly suppressed and often neglected in theoretical calculations. However,

[^0]the inclusion of rescattering effects into the theoretical framework naturally enhances their contribution [1]. Recently it was observed that the predicted branching fraction for the decay $B_{s}^{0} \rightarrow \pi^{+} \pi^{-}$, which involves topological annihilation diagrams, was substantially smaller than its measured value by the LHCb experiment [2]. This discrepancy between theoretical prediction and experimental measurement may require some models of strong interaction processes to be revisited [3]. In these aspects, searches for decays involving weak annihilation amplitudes become important and necessary.

Within the standard model (SM), the decay $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ proceeds via the $W$-exchange and "penguin" annihilation amplitudes, as shown in Fig. 1. Theoretical calculations based on the flavor diagram approach [4],


FIG. 1. $W$-exchange (top) and "penguin" annihilation (bottom) Feynman diagrams for $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$.
perturbative quantum chromodynamics [5], and QCD factorization [6] predict the branching fraction $(\mathcal{B})$ to be $(0.40 \pm 0.27) \times 10^{-6},(0.28 \pm 0.09) \times 10^{-6}$, and $(0.13 \pm$ $0.05) \times 10^{-6}$, respectively. The only measurement for this decay was made by the L3 experiment in 1995, which reported an upper limit (UL) of $\mathcal{B}<2.1 \times 10^{-4}$ at $90 \%$ confidence level (CL) [7]. The search for the decay $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ [8] described in this paper is based on a data sample of $121.4 \mathrm{fb}^{-1}$ collected at the $\Upsilon(5 \mathrm{~S})$ resonance using the Belle detector.

The Belle detector at the KEKB [9] asymmetric-energy $e^{+} e^{-}$collider is a large-solid-angle magnetic spectrometer that consists of a silicon vertex detector, a 50-layer central drift chamber, an array of aerogel threshold Cherenkov counters, a barrel-like arrangement of time-of-flight scintillation counters, and a $\mathrm{CsI}(\mathrm{Tl})$ crystal-based electromagnetic calorimeter (ECL) located inside a superconducting solenoid coil that provides a 1.5 T magnetic field. An iron flux return outside the coil is instrumented to detect $K_{L}^{0}$ mesons and identify muons. A detailed description of the Belle detector can be found elsewhere [10,11]. The analysis relies on the ECL component of the detector for the reconstruction of the photons in the $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ decay final state.

The production cross section of the $\Upsilon(5 \mathrm{~S})$ resonance at the $e^{+} e^{-}$center of mass (c.m.) energy of 10.86 GeV is $\sigma_{b \bar{b}}^{\Upsilon(5 S)}=(0.340 \pm 0.016) \mathrm{nb}$ [12], and the fraction of $b \bar{b}$ events giving rise to $B_{s}^{0}$ production modes, $B_{s}^{(*) 0} \bar{B}_{s}^{(*) 0}$, is measured to be $f_{s}=(0.201 \pm 0.031)$ [13]. There are three kinematically allowed modes of production of $B_{s}^{0}$ mesons: $B_{s}^{* 0} \bar{B}_{s}^{* 0}, B_{s}^{0} \bar{B}_{s}^{* 0}$ or $B_{s}^{* 0} \bar{B}_{s}^{0}$, and $B_{s}^{0} \bar{B}_{s}^{0}$. The production fractions from the former two are (87.0 $\pm 1.7$ )\% and $(7.3 \pm 1.4) \%$, respectively [12], while the remaining fraction is from the $B_{s}^{0} \bar{B}_{s}^{0}$ mode. The $B_{s}^{* 0}$ decays to $B_{s}^{0}$ by radiating a low-energy photon that is usually not identified due to its poor reconstruction efficiency. The number of events with $B_{s}^{0} \bar{B}_{s}^{0}$ is, therefore, estimated to be $N_{B_{s}^{0} \bar{B}_{s}^{0}}=121.4 \mathrm{fb}^{-1} \cdot \sigma_{b \bar{b}}^{\Upsilon(5 S)} \cdot f_{s}=(8.30 \pm 1.34) \times 10^{6}$.

We employ a "blind" analysis procedure to leave out the experimenter's biases and develop our analysis strategy with Monte-Carlo (MC) samples. In a "blind" analysis, the signal region is kept hidden until the selection criteria are finalized. The signal MC samples are generated with EvtGen [14] and simulated with GEANT3 [15] to model all possible detector effects. Background studies are performed with MC samples 6 times larger than the integrated luminosity of data. The analysis procedure is validated with a control sample of $B_{d}^{0} \rightarrow \pi^{0} \pi^{0}$ decays produced at the $\Upsilon(4 \mathrm{~S})$ resonance, which closely resembles the signal.

We reconstruct $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ with $\pi^{0} \rightarrow \gamma \gamma$. Photon candidates are reconstructed from ECL clusters that do not match any charged track and have energy greater than 50 (100) MeV in the ECL's barrel (end caps) region. The forward end cap, barrel, and backward end cap regions of the ECL are given by $12^{\circ}<\theta<31.4^{\circ}, 32.2^{\circ}<\theta<128.7^{\circ}$, and $131.5^{\circ}<\theta<157.2^{\circ}$, respectively, where $\theta$ is the polar angle in the laboratory frame with respect to the detector axis, in the direction opposite to the $e^{+}$beam. To remove the off-time (radiative) Bhabha and $e^{+} e^{-} \rightarrow \gamma \gamma$ events, a timing criterion based on the beam collision time is applied, which is determined at the trigger level for each candidate event. The invariant mass of the two-photon combination must lie in the range of $118 \mathrm{MeV} / c^{2}<m(\gamma \gamma)<$ $152 \mathrm{MeV} / c^{2}$, corresponding to $\pm 2.4$ standard deviations $(\sigma)$ of the invariant mass resolution around the nominal $\pi^{0}$ mass [13]. A mass-constrained fit is subsequently performed to improve the $\pi^{0}$ momentum resolution.

To further select the $B_{s}^{0}$ candidates, we apply selection criteria on their beam-energy-constrained mass $M_{\mathrm{bc}}=$ $\sqrt{\left(E_{\text {beam }}\right)^{2}-\left|\vec{p}_{\text {reco }}\right|^{2} c^{2}} / c^{2}$ and the energy difference $\Delta E^{\prime}=E_{\text {reco }}-E_{\text {beam }}+M_{\mathrm{bc}} c^{2}-m_{B_{s}^{0}} c^{2}$ in the $e^{+} e^{-}$c.m. frame, where $E_{\text {beam }}$ is the beam energy, $\vec{p}_{\text {reco }}$ and $E_{\text {reco }}$ are the momentum and energy, respectively, of the reconstructed $B_{s}^{0}$ candidate. The world average value is used for the mass of the $B_{s}^{0}$ meson, $m_{B_{s}^{0}}$ [13]. In the $\Delta E^{\prime}$ distribution, all the production channels $\left(B_{s}^{* 0} \bar{B}_{s}^{* 0}, B_{s}^{0} \bar{B}_{s}^{* 0}\right.$ or $B_{s}^{* 0} \bar{B}_{s}^{0}$, and $B_{s}^{0} \bar{B}_{s}^{0}$ ) peak at zero, while in the $\Delta E=$ $E_{\text {reco }}-E_{\text {beam }}$ variable, the peaks of $B_{s}^{* 0} \bar{B}_{s}^{* 0}$ and $B_{s}^{0} \bar{B}_{s}^{* 0}$ or $B_{s}^{* 0} \bar{B}_{s}^{0}$ are shifted to negative values by $\left(m_{B_{s}^{* 0}} c^{2}-\right.$ $\left.m_{B_{s}^{0}} c^{2}\right) / 2$ and $m_{B_{s}^{* o}} c^{2}-m_{B_{s}^{0}} c^{2}$, respectively, where, $m_{B_{s}^{* 0}}$ represents the mass of the $B_{s}^{* 0}$ meson. Therefore, to correct for these shifts in the $\Delta E$ variable, the $\Delta E^{\prime}$ variable is used for this analysis. A $B_{s}^{0}$ candidate is retained for further analysis only if it satisfies the requirement that $5.300 \mathrm{GeV} / c^{2}<M_{\mathrm{bc}}<5.434 \mathrm{GeV} / c^{2}$ and $-0.60 \mathrm{GeV}<$ $\Delta E^{\prime}<0.15 \mathrm{GeV}$.

The backgrounds near the $\Upsilon(5 S)$ resonance which can affect the analysis are: continuum $\left(e^{+} e^{-} \rightarrow q \bar{q}, q=\right.$ $u, d, s, c), B_{s}^{(*)} \bar{B}_{s}^{(*)}$ decays (referred as $b s b s$ ) and $B^{*} \bar{B}^{*}, B^{*} \bar{B}$, $B \bar{B}, B^{*} \bar{B}^{*} \pi, B^{*} \bar{B} \pi, B \bar{B} \pi$, and $B \bar{B} \pi \pi\left(B=B^{0}, B^{+}\right)$decays (referred as non-bsbs). Additional background MC studies
on the peaking background of the types $B_{s}^{0} \rightarrow \rho^{+} \rho^{-}$and $B_{s}^{0} \rightarrow K_{s}^{0} \pi^{0}$ show that their contributions are negligible. We also find no bsbs and non-bsbs background after applying all of the aforementioned selection criteria. Background MC studies, therefore, reveal the dominance of continuum background over the other types of background. Their suppression requires topological variables, which classify the signal and the continuum background based on their event shape variables in the $e^{+} e^{-}$c.m. frame.

In signal events, $B_{s}^{0}$ pairs are produced with small momenta, and the distribution of their decay products tends to be spherical. In contrast, the quark pairs of the continuum background are produced with a significant amount of momentum; therefore, their decay product distribution has a jetlike topology. A neural network algorithm (NN) [16] is employed to suppress the continuum background. The input of the NN includes sixteen modified Fox-Wolfram moments [17], and thrust axis direction, $\cos \theta_{T}$ (see Sec. 9.3 in [18]) to provide additional discrimination between the signal and the continuum background. These modified Fox-Wolfram moments are calculated based on three categories of events: the $B_{s}^{0}$ candidate daughters, the rest of events, and the missing momentum of an event. The angle $\theta_{T}$ is defined as the angle between the thrust axis of the signal $B_{s}^{0}$ candidate and the thrust axis of the remainder of the events. The NN was trained on MC samples with consistency checks to ensure no overtraining.

The choice of the selection criterion on the output of the $\mathrm{NN}, \mathcal{C}_{\mathrm{NN}}$, is determined based on a Punzi's figure-of-merit optimization [19], where the significance level is set to three standard deviations. The $\mathcal{C}_{\mathrm{NN}}$ distributions for the continuum background and the signal lies in the range of $[-1,+1]$, where the continuum backgrounds peak at -1 and the signal candidates at +1 . We require $\mathcal{C}_{\mathrm{NN}}$ to be greater than 0.90 for this analysis. This condition removes $99 \%$ of the continuum background with a signal loss of $53 \%$. To facilitate the data modeling, $\mathcal{C}_{\mathrm{NN}}$ was transformed to another variable, $\mathcal{C}_{\mathrm{NN}}^{\prime}$ using the following formula:

$$
\begin{equation*}
\mathcal{C}_{\mathrm{NN}}^{\prime}=\log \left(\frac{\mathcal{C}_{\mathrm{NN}}-\mathcal{C}_{\mathrm{NN}(\min )}}{\mathcal{C}_{\mathrm{NN}(\max )}-\mathcal{C}_{\mathrm{NN}}}\right) \tag{1}
\end{equation*}
$$

where $\mathcal{C}_{\mathrm{NN}(\text { min })}=0.90$ and $\mathcal{C}_{\mathrm{NN}(\text { max })}$ is the maximum value of $\mathcal{C}_{\mathrm{NN}}$ obtained from the NN distribution.

After applying the selection criteria described above, $10.3 \%$ of signal MC events have more than one candidate. Most of the multiple candidates are due to a signal photon reconstructed as two ECL clusters. We select the best candidate based on the sum of the $\chi^{2}$ of the massconstrained fits to the two $\pi^{0} \mathrm{~s}$. The fraction of misreconstructed events after applying all the selection criteria is found to be negligible; hence, they are not treated separately. The overall signal reconstruction efficiency in this analysis is $(12.69 \pm 0.05) \%$.

To extract the signal yield, we perform a threedimensional (3D) unbinned extended maximum likelihood (ML) fit to $M_{\mathrm{bc}}, \Delta E^{\prime}$, and $\mathcal{C}_{\mathrm{NN}}^{\prime}$. The likelihood function is defined as

$$
\begin{equation*}
\mathcal{L}_{\mathrm{fit}}=e^{-\sum_{j} n_{j}} \prod_{i}^{N}\left(\sum_{j} n_{j} P_{j}\left(\left(M_{\mathrm{bc}}\right)^{i},\left(\Delta E^{\prime}\right)^{i},\left(\mathcal{C}_{\mathrm{NN}}^{\prime}\right)^{i}\right)\right) \tag{2}
\end{equation*}
$$

where $P_{j}\left(M_{\mathrm{bc}}, \Delta E^{\prime}, \mathcal{C}_{\mathrm{NN}}^{\prime}\right)$ is the probability distribution function (PDF) of the signal or background component (specified by index $j$ ), $n_{j}$ is the yield of this component, $i$ represents the event index, and $N$ is the total number of events in the sample.

The linear correlation coefficients among $M_{\mathrm{bc}}, \Delta E^{\prime}$, and $\mathcal{C}_{\mathrm{NN}}^{\prime}$ are found to be below $3 \%$ in the signal region. Consequently, each of the 3D PDFs describing the signal and background contributions are assumed to factorize as $P_{j} \equiv P_{j}\left(M_{\mathrm{bc}}\right) P_{j}\left(\Delta E^{\prime}\right) P_{j}\left(\mathcal{C}_{\mathrm{NN}}^{\prime}\right)$. These factorized PDFs are modeled using large signal and background MC samples. The signal $M_{\mathrm{bc}}$ PDF consists of three PDFs corresponding to the three $B_{s}^{0}$ production channels. Each of them is again separately modeled from large MC samples. They are then combined according to their production fractions [12] to produce the final signal PDF for the $M_{\mathrm{bc}}$ variable. The PDF used for parametrizing $B_{s}^{0} \bar{B}_{s}^{0}$ is a sum of two Gaussian distributions with a common mean, while each of $B_{s}^{0} \bar{B}_{s}^{* 0}$ or $B^{* 0} \bar{B}_{s}^{0}$, and $B_{s}^{* 0} \bar{B}_{s}^{* 0}$, are parametrized using a sum of a Gaussian function and an empirical PDF shape known as the Crystal Ball function [20]. The signal $\Delta E^{\prime}$ variable, for all three $B_{s}^{0}$ channels, is modeled using the Crystal Ball function, which is modified for this analysis to include the asymmetric nature of the distribution about the mean position. The output from the NN is parametrized using a Gaussian and an asymmetric (bifurcated) Gaussian PDF for the signal $\mathcal{C}_{\text {NN }}^{\prime}$ variable. Unlike the signal PDF parameters for the $M_{\mathrm{bc}}$ variable, which is different for the three $B_{s}^{0}$ sources, $\Delta E^{\prime}$ and $\mathcal{C}_{\mathrm{NN}}^{\prime}$ variables take the same parameter values for the three $B_{s}^{0}$ production channels. The continuum background distribution of the $M_{\mathrm{bc}}$ variable is modeled through an empirically determined parametrized background shape referred to as the ARGUS function [21]. The continuum background is parametrized using a firstorder Chebychev polynomial and a sum of two Gaussian distributions for the $\Delta E^{\prime}$ and $\mathcal{C}_{\mathrm{NN}}^{\prime}$ variables, respectively. All of the signal parameters and the background ARGUS end point are fixed to their best fit values obtained from 1D fits to the MC simulated events. In contrast, all other background parameter values and the signal and background yields are floated. The PDFs used for modeling the signal and continuum background are listed in Table I.

To validate our analysis, we use the Belle data sample collected at the $\Upsilon(4 \mathrm{~S})$ to reconstruct the decay $B_{d}^{0} \rightarrow \pi^{0} \pi^{0}$ by applying similar event selection criteria. The results of

TABLE I. PDFs used to model the $M_{\mathrm{bc}}, \Delta E^{\prime}$, and $\mathcal{C}_{\mathrm{NN}}^{\prime}$ distributions. The notations $\mathrm{G}, \mathrm{BG}, \mathrm{CB}, \mathrm{ACB}, C P$, and A correspond to Gaussian, bifurcated Gaussian, Crystal Ball, asymmetric Crystal Ball, Chebyshev polynomial, and ARGUS functions, respectively.

| Fit component | $M_{\mathrm{bc}}$ | $\Delta E$ | $\mathcal{C}_{\mathrm{NN}}^{\prime}$ |
| :--- | :--- | :---: | :---: |
| Signal | $\mathrm{G}+\mathrm{G}\left(B_{s}^{0} \bar{B}_{s}^{0}\right)$ | ACB | $\mathrm{G}+\mathrm{BG}$ |
|  | $\mathrm{G}+\mathrm{CB}\left(B_{s}^{0} \bar{B}_{s}^{* 0}\right)$ |  |  |
|  | $\mathrm{G}+\mathrm{CB}\left(B_{s}^{* 0} \bar{B}_{s}^{* 0}\right)$ |  |  |
| Continuum | A | CP | $\mathrm{G}+\mathrm{G}$ |

the fit to $\Upsilon(4 S)$ data are shown in Fig. 2, where each fit projection is plotted after additional selection criteria are applied as described in the caption. We calculate the branching fraction, $\mathcal{B}\left(B_{d}^{0} \rightarrow \pi^{0} \pi^{0}\right)=(1.18 \pm 0.21) \times$ $10^{-6}$ (where only the statistical uncertainty is shown), which is in good agreement with our previous result [22].

The systematic uncertainties associated with the analysis are summarized in Table II. The systematic uncertainties due to the fit model are determined via ensemble investigations. To carry out an ensemble study, we generate and simulate 500,000 signal MC events. We randomly select signal events from this sample for different expected signal yields in data. In addition, background MC events are randomly extracted from the background PDFs based on the expected number of background events in the data. This MC sample that now has statistics equivalent to the expected yields in data is amplified by repeating the above procedure 1000 times. We then perform 3D unbinned extended ML fits on these 1000 pseudo-experiments to obtain pull distributions for each of the expected signal yields in data. The average deviation of a zeroth-order polynomial fit to the means of the pull values obtained for each expected signal yield from the no-bias condition is recognized as a fit bias.

TABLE II. Summary of systematic uncertainties.

| Source | Value $(\%)$ |
| :--- | :---: |
| Fit bias | -3.3 |
| Fixed PDF parametrization | ${ }_{-5.2}^{+3.5}$ |
| Fractions of $B_{s}^{* 0} \bar{B}_{s}^{* * 0)}$ | ${ }_{-3.5}^{+5.2}$ |
| Reconstruction efficiency, $\epsilon_{\text {rec }}$ | $\pm 0.4$ |
| $\mathcal{C}_{\mathrm{NN}}^{\prime}$ requirement | $\pm 3.0$ |
| $\pi^{0} \rightarrow \gamma \gamma$ selection efficiency | $\pm 4.4$ |
| $\mathcal{B}\left(\pi^{0} \rightarrow \gamma \gamma\right)$ | $\pm 0.07$ |
| $b \bar{b}$ cross section, $\sigma_{b \bar{b}}$ | $\pm 4.7$ |
| $f_{s}$ | $\pm 15.4$ |
| Total | +18.1 |

We observe a fit bias of $-3.3 \%$ and assign it as the corresponding systematic uncertainty. The uncertainty due to fixing the parameter values of the PDFs is determined by varying the best fit parameter values within $\pm 1 \sigma$ of their statistical uncertainties and measuring the deviation of the signal yields in data. We find a fractional systematic uncertainty of ${ }_{-5.2 \%}^{+3.5 \%}$ from this source. Apart from fixing the signal PDF parameters and the background PDF's ARGUS end point, we have also fixed the fractions of the $B_{s}^{0}$ production channels. We vary these fractions within $\pm 1 \sigma$ of their measured values [12] and repeat the fit. The observed relative variation ${ }_{-3.5 \%}^{+5.2 \%}$ of the signal yield is assigned as the systematic uncertainty. The systematic uncertainty of the signal reconstruction efficiency is $0.4 \%$ due to the finite number of signal MC events. The systematic uncertainty due to the efficiency of the $\mathcal{C}_{\mathrm{NN}}^{\prime}$ requirement is estimated from the control sample using a parameter, $\mathcal{R}$. It is defined as the ratio between the efficiency of the $\mathcal{C}_{\mathrm{NN}}^{\prime}$ requirement in data and MC. We assign a corresponding systematic uncertainty of $\pm 3 \%$ due to the efficiency of the $\mathcal{C}_{\mathrm{NN}}^{\prime}$ requirement.


FIG. 2. Signal enhanced projections of $M_{\mathrm{bc}}$ (left), $\Delta E^{\prime}$ (middle), and $\mathcal{C}_{\mathrm{NN}}^{\prime}$ (right) for the control sample, $B_{d}^{0} \rightarrow \pi^{0} \pi^{0}$. Each plot is generated by applying the signal region selection criteria on the two variables other than the plotted variable. The signal regions for the three variables are as follows: $5.2700 \mathrm{GeV} / c^{2}<M_{\mathrm{bc}}<5.2895 \mathrm{GeV} / c^{2},-0.23 \mathrm{GeV}<\Delta E^{\prime}<0.15 \mathrm{GeV}$, and $-3.10<\mathcal{C}_{\mathrm{NN}}^{\prime}<7.61$. The dark-filled, red (dotted), black (dash-dotted), and blue (solid) color distributions represent the signal, continuum background, rare $B_{d}^{0}$ background (backgrounds arising due to $b \rightarrow u$ transitions) and total fit function, respectively. Points with error bars represent data.


FIG. 3. Signal enhanced projections of $M_{\mathrm{bc}}$ (left), $\Delta E^{\prime}$ (middle), and $\mathcal{C}_{\mathrm{NN}}^{\prime}$ (right) for the analysis, $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$. Each plot is generated by applying the signal region selection criteria on the two variables other than the plotted variable. The signal regions for the three variables are as follows: $5.395 \mathrm{GeV} / c^{2}<M_{\mathrm{bc}}<5.434 \mathrm{GeV} / c^{2},-0.310 \mathrm{GeV}<\Delta E^{\prime}<0.140 \mathrm{GeV}$, and $-3.901<\mathcal{C}_{\mathrm{NN}}^{\prime}<7.451$. The darkfilled, red (dotted), and blue (solid) color distributions represent the signal, continuum background and total fit function, respectively. Points with error bars represent data. The peak in the $M_{\mathrm{bc}}$ distribution is due to the dominant $B_{s}^{0}$ production channel, $B_{s}^{* 0} \bar{B}_{s}^{* 0}(87.0 \%)$. The other two production channels, $B_{s}^{* 0} \bar{B}_{s}^{0}(7.3 \%)$ and $B_{s}^{0} \bar{B}_{s}^{0}(5.7 \%)$, are present but do not contain a significant number of signal events.

The systematic uncertainty for the $\pi^{0}$ selection efficiency is determined to be $2.2 \%$ per $\pi^{0}$ using the decay $\tau^{-} \rightarrow \pi^{-} \pi^{0} \nu_{\tau}$. Since this uncertainty is completely correlated for the two $\pi^{0} \mathrm{~s}$, a total systematic uncertainty of $4.4 \%$ is assigned. We assign a fractional systematic uncertainty of $0.034 \%$ on the branching fraction of $\pi^{0} \rightarrow \gamma \gamma$ [13]. Therefore, for the two $\pi^{0}$ branching fractions, a total systematic uncertainty of $0.07 \%$ is assigned.

The systematic uncertainty due to the $b \bar{b}$ production cross section at $\Upsilon(5 \mathrm{~S})$ resonance, $\sigma_{b \bar{b}}$ is estimated to be $\pm 4.7 \%$ [12]. In addition, the systematic uncertainty due to the three charmless production processes arising from $b \bar{b}$ events, $f_{\mathrm{s}}$ is assumed to be $\pm 15.4 \%$ [13]. This uncertainty on $f_{\mathrm{s}}$ is the dominant systematic uncertainty associated with any $B_{s}^{0}$ measurement at Belle.

The fit projections obtained from a 3D unbinned extended maximum likelihood fit in the signal regions are shown in Fig. 3. We obtain $5.7 \pm 5.8$ signal events and $989 \pm 32$ continuum background events in our fit to the data. The branching fraction is calculated using

$$
\begin{equation*}
\mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)=\frac{N_{\text {yield }}^{\text {sig }}}{2 \times N_{B_{s}^{0} \bar{B}_{s}^{0}} \times \epsilon^{\mathrm{rec}} \times \mathcal{B}}, \tag{3}
\end{equation*}
$$

where $N_{B_{s}^{0} \bar{B}_{s}^{0}}$ is the number of $B_{s}^{0} \bar{B}_{s}^{0}$ pairs; $\epsilon^{\text {rec }}$ and $N_{\text {yield }}^{\text {sig }}$ are the signal selection efficiency obtained from MC simulation and the signal yield obtained from the fit, respectively; and $\mathcal{B}$ is the product of the two $\pi^{0}$-decay branching fractions [13].

Incorporating the signal yield, $N_{\text {yield }}^{\text {sig }}=(5.7 \pm 5.8)$, number of $B_{s}^{0} \bar{B}_{s}^{0}$ pairs, $N_{B_{s}^{0} \bar{B}_{s}^{0}}=(8.30 \pm 1.34) \times 10^{6}$, the signal reconstruction efficiency, $\epsilon_{\text {rec }}=(12.69 \pm 0.05) \%$, and branching fraction, $\mathcal{B}\left(\pi^{0} \rightarrow \gamma \gamma\right)=(98.82 \pm 0.03) \%$ in Eq. (3), the branching fraction for $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ and its product with $f_{s}$ are calculated to be

$$
\begin{aligned}
\mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right) & =(2.8 \pm 2.8 \pm 0.5) \times 10^{-6} \\
f_{s} \times \mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right) & =(0.6 \pm 0.6 \pm 0.1) \times 10^{-6}
\end{aligned}
$$

The first uncertainty is statistical, and the second uncertainty is systematic.

Without significant signal yield, we calculate the UL on the branching fraction using a Bayesian approach. The UL on the branching fraction is estimated by integrating the likelihood function obtained from the maximum likelihood fit procedure from $0 \%$ to $90 \%$ of the area under the likelihood curve. The systematic uncertainties are incorporated by convolving the likelihood curve with a Gaussian distribution with a mean of zero and width equivalent to the total systematic uncertainty listed in Table II. The UL on the branching fraction, $\mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)$ at $90 \%$ CL [23] and the product of the branching fraction with $f_{s}$, $f_{s} \times \mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)$, is found to be

$$
\begin{aligned}
& \mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)<7.7 \times 10^{-6} \\
& f_{s} \times \mathcal{B}\left(B_{s}^{0}\right.\left.\rightarrow \pi^{0} \pi^{0}\right)<1.5 \times 10^{-6}
\end{aligned}
$$

The total systematic uncertainties associated with $\mathcal{B}\left(B_{s}^{0} \rightarrow\right.$ $\left.\pi^{0} \pi^{0}\right)$ and $f_{s} \times \mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)$ are ${ }_{-18.4 \%}^{+18.1 \%}$ and ${ }_{-10.0 \%}^{+9.5 \%}$, respectively. The results are summarized in Table III.

TABLE III. Summary of results on branching fractions and UL for $\mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)$ and $f_{s} \times \mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)$.

| Quantity | Value |
| :--- | :---: |
| $\mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)$ | $(2.8 \pm 2.8 \pm 0.5) \times 10^{-6}$ |
|  | $<7.7 \times 10^{-6}$ at $90 \% \mathrm{CL}$ |
| $f_{s} \times \mathcal{B}\left(B_{s}^{0} \rightarrow \pi^{0} \pi^{0}\right)$ | $(0.6 \pm 0.6 \pm 0.1) \times 10^{-6}$ |

To summarize, we search for the decay $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ using the final Belle data sample available at $\Upsilon(5 \mathrm{~S})$ resonance, which corresponds to an integrated luminosity of $121.4 \mathrm{fb}^{-1}$. We do not observe a significant signal yield, and thus set a $90 \%$ CL upper limit on the $B_{s}^{0} \rightarrow \pi^{0} \pi^{0}$ branching fraction of $7.7 \times 10^{-6}$. This is the most stringent UL estimated for this decay representing an order-ofmagnitude improvement over the previous result [7] by the L3 experiment in 1995.

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[23] As we use the Bayesian method, this is a "credible interval" but we use "confidence level" here following common convention.


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