

Dephasing of Majorana qubits due to quasistatic disorder

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Quantum bits based on Majorana zero modes are expected to be robust against certain noise types, and hence provide a quantum computing platform that is superior to conventional qubits. This robustness is not complete, though: imperfections can still lead to qubit decoherence and hence to information loss. In this work, we theoretically study Majorana-qubit dephasing in a minimal model: in a Kitaev chain with quasistatic disorder. Our approach, based on numerics as well as first-order nondegenerate perturbation theory, provides a conceptually simple physical picture and predicts Gaussian dephasing. We show that as system parameters are varied, the dephasing rate due to disorder oscillates out-of-phase with respect to the oscillating Majorana splitting of the clean system. In our model, first-order dephasing sweet spots are absent if disorder is uncorrelated. We describe the crossover between uncorrelated and highly correlated disorder, and show that dephasing measurements can be used to characterize the disorder correlation length. We expect that our results will be utilized for the design and interpretation of future Majorana-qubit experiments.

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I. INTRODUCTION

Theoretical proposals [1–3] suggest that Majorana zero modes (MZMs) can be engineered in quasi-one-dimensional semiconducting-superconducting hybrid systems [4,5]. The past decade has witnessed intense experimental activities to establish MZMs [4,6–20]. It is expected that MZMs could serve as building blocks in experiments demonstrating topologically protected quantum memories, quantum dynamics, or even quantum computing [21–28]. In that context, understanding the decoherence of Majorana qubits [26,29–40] is an important task.

The minimal model hosting MZMs (see Fig. 1) is the Kitaev chain [41]. It can be used to describe the dephasing process of a Majorana qubit. The ground state of a finite-length topological Kitaev chain hosts two MZMs at the two ends of the chain, implying that the ground state is approximately twofold degenerate, with one ground state being of even fermion parity and the other being of odd fermion parity. In a chain with a finite length, a small energy splitting ε_0 separates the two ground states. If random components, such as disorder [29,42], are incorporated in the model, then the splitting becomes a random variable. To encode a single qubit with MZMs, two wires and hence four MZMs are needed [43]. In such a two-wire Majorana qubit, the random splittings in the two wires add up to a random Larmor frequency of the qubit, leading to qubit dephasing.

In this work, we theoretically study dephasing of Majorana qubits in the presence of slow charge noise. A key target in topological quantum computing is the experimental demonstration of a topologically protected quantum memory based on MZMs; hence it is imperative to understand the potential

sources of qubit decoherence, to assess future device functionality and provide optimization guidelines. Furthermore, qubit dephasing measurement is an established tool to reveal the noise structure of the qubit's environment [44–47]; understanding dephasing is important for that application, too. In our work, we focus on the model of quasistatic disorder [48–53], a minimal model of slow (low-frequency) charge noise or $1/f$ noise [40,44,46,47,54–57], which has been a very important source of qubit dephasing both in semi- and superconductor environments.

Naturally, Majorana qubit dephasing due to weak and slow (quasistatic) charge noise is determined by the probability distribution of the splitting; see, e.g., our Sec. IV. First we study that splitting distribution for uncorrelated disorder by both numerical and analytical methods, and show that it is Gaussian for weak disorder. We argue that this result is consistent with the log-normal splitting-envelope distribution found by Ref. [29] in our Appendix C.

Having the splitting distribution at hand, we use it to characterize the dephasing of a Majorana qubit subject to weak quasistatic disorder. In simple models, the time dependence of qubit dephasing often follows a Gaussian function [26,58]. Here we show that this is also the case for the qubit studied here. Our key results for the spatially uncorrelated disorder model are that (i) we provide an analytical formula for the dephasing susceptibility [see Eq. (29)] and the dephasing time [see Eq. (35)], (ii) we reveal an out-of-phase oscillation between the splitting of the clean system and the dephasing susceptibility to disorder [see Fig. 2(b) and Figs. 3(a) and 3(b)], and (iii) we highlight the absence of dephasing sweet spots in our model (see Sec. IV). Finally, we show that the spatial correlation length of the disorder has a strong impact

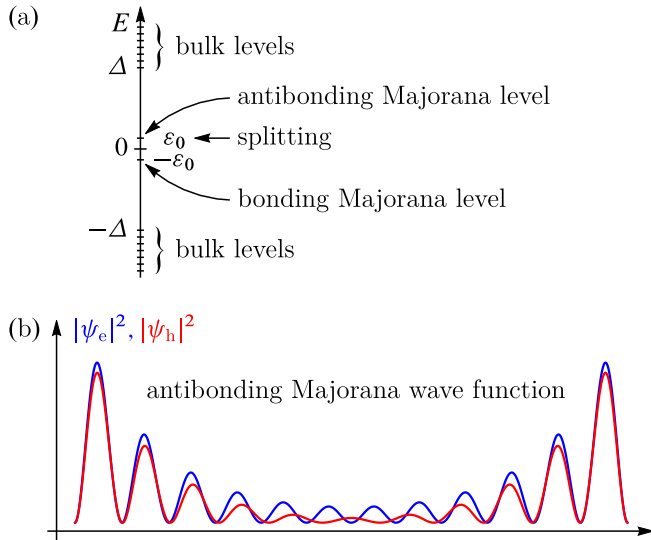


FIG. 1. Spectrum and Majorana wave functions in a topological superconductor wire. (a) Schematic spectrum of the Bogoliubov–de Gennes matrix. (b) Electron and hole components of the antibonding Majorana wave function.

on a dephasing experiment. As a consequence, we expect that in future Majorana-qubit experiments, measuring the dephasing time as function of control parameters (e.g., chemical potential) will provide information about the spatial structure of noise.

The rest of the paper is organized as follows. In Sec. II, we show numerical results for the splitting distribution of the disordered chain, highlighting the Gaussian distribution of the splitting, and the out-of-phase oscillation between the clean splitting and the splitting susceptibility to disorder. In Sec. III, we use the continuum version of the Kitaev chain, together with mode matching and first-order perturbation theory, to establish the semianalytical description of the splitting distribution, and to derive approximate analytical results for that. Furthermore, we compare the results of the two models. In Sec. IV, we relate the splitting distribution and the dephasing dynamics of a Majorana qubit based on two Kitaev chains. In Sec. V, we show that the parameter dependence of the dephasing time is sensitive to the correlation length of the disorder. We discuss implications and follow-up ideas in Sec. VI, and conclude in Sec. VII.

II. DISORDER-INDUCED SPLITTING DISTRIBUTION IN THE KITAEV CHAIN

We use the Kitaev-chain tight-binding model [41] to numerically investigate Majorana qubit dephasing. In this section, we numerically determine the disorder-induced distribution of the signful splitting (see definition below), using the Kitaev chain. We anticipate that this distribution is Gaussian for weak quasistatic disorder (see below within this section for details), and that the Majorana qubit dephasing time T_2^* in a two-chain setup is determined by the standard deviation σ_{ϵ_0} of the signful splitting ϵ_0 [see Eq. (35) in Sec. IV].

TABLE I. Parameter values used in the numerical and analytical calculations.

Parameter/scale	Notation	Value
Continuum model:		
Effective mass (of InAs)	m	$0.023m_e$
Chemical potential	μ_C	1 meV
Superconducting gap	Δ_C	$200 \mu\text{eV}$
Kitaev chain:		
Normal hopping amplitude	t	6.62 eV
Lattice constant	a	0.5 nm
Chemical potential	μ_K	-13.3 eV
Superconducting pairing potential	Δ_K	8.14 meV
Length scales:		
Fermi wavelength	λ_F	$511a$
Fermi wave number	k_F	$0.0123/a$
Superconductor coherence length	ξ	$814a$
Inverse coherence length	κ	$0.00123/a$

The Hamiltonian of a finite-length Kitaev chain in real space reads [41]

$$H_K = - \sum_{n=1}^N (\mu_K + \delta\mu_n^{(K)}) c_n^\dagger c_n - t \sum_{n=1}^{N-1} (c_n^\dagger c_{n+1} + \text{H.c.}) - \Delta_K \sum_{n=1}^{N-1} (c_n c_{n+1} + \text{H.c.}), \quad (1)$$

where c_n^\dagger and c_n are the electron creation and annihilation operator on site n , respectively, t is the hopping amplitude, μ_K is the chemical potential, Δ_K is the superconducting pair potential, and N is the number of sites. We model disorder as a random on-site potential, independent on each site, drawn from Gaussian distribution with zero mean and standard deviation σ_μ , that is, $\delta\mu_n^{(K)} \sim \mathcal{N}(0, \sigma_\mu)$. For a discussion of the relation between this model and disorder in real samples, see Sec. VI.

We obtain the splitting ϵ_0 using the Bogoliubov–de Gennes (BdG) transformation [59], i.e., by numerically finding the smallest positive eigenvalue of the corresponding real-space $2N \times 2N$ BdG Hamiltonian [60]. We will calculate the splitting ϵ_0 for a clean system, i.e., in the absence of any disorder, as well as for random on-site disorder realizations. In the latter case, ϵ_0 becomes a random variable—with a Gaussian distribution for weak disorder, as shown in Fig. 2(a) and discussed below.

A clean Kitaev chain has a splitting that decreases in an oscillatory fashion as the chain length is increased [61–64]. This is shown in Fig. 2(b), where we plot the numerically calculated length dependence of ϵ_0 (red solid line) for a parameter set shown in the “Kitaev chain” section of Table I.

Now we introduce disorder and study the splitting distribution. Figure 2(a) shows three examples of that distribution, for a fixed disorder strength $\sigma_\mu = 100 \mu\text{eV}$, for three different chain lengths $N = 2000, 3000, 4000$. The figure clearly shows a Gaussian character for all three distributions. Furthermore, the figure also shows the trend that both the mean and the standard deviation of these distributions decrease as the chain length increases.

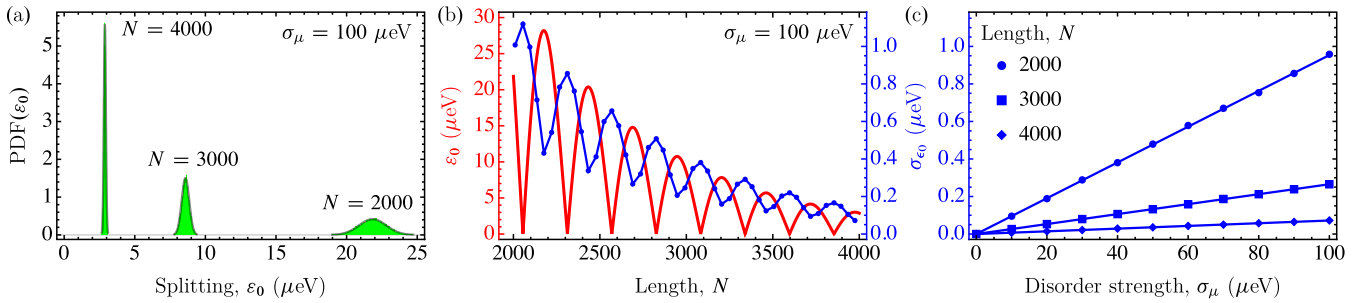


FIG. 2. Splitting, its probability distribution, and its standard deviation from the Kitaev chain model. (a) Numerically obtained probability density functions (pdf's) of the splitting for three different lengths in disordered system. Gray lines are fitted Gaussian pdf's. (b) Splitting of the clean system (red line) and standard deviation of the signful splitting (blue points) are shown as a function of chain length, for the disorder strength $\sigma_\mu = 100$ μeV . Out-of-phase oscillation can be observed between the splitting and its standard deviation. (c) Standard deviation of the signful splitting is shown as a function of the strength of the on-site disorder for three different lengths. The dependence on the disorder strength is linear for the shown range. Results for disordered systems are calculated using 10 000 realizations.

Figure 2(c) shows a more systematic analysis of the length and disorder-strength dependence of the standard deviation σ_{ϵ_0} of the signful splitting ϵ_0 . Note the difference between the *signful splitting* ϵ_0 and the *splitting* ϵ_0 . The signful splitting is defined by $\epsilon_0 \equiv \epsilon_o - \epsilon_e$, where ϵ_o (ϵ_e) is the energy of the odd (even) ground state. We have defined the splitting [see Fig. 1(a)] as the absolute value of the signful splitting, i.e., $\epsilon_0 \equiv |\epsilon_0|$. The distinction between ϵ_0 and ϵ_0 is motivated by the observation that the dephasing dynamics is related to the signful splitting ϵ_0 ; see Eq. (35).

For all lengths displayed in Fig. 2(c), the standard deviation σ_{ϵ_0} of the signful splitting shows a clear linear dependence on the disorder strength σ_μ . This linear dependence motivates the definition of the dimensionless *dephasing susceptibility to disorder*, $\chi = \sigma_{\epsilon_0}/\sigma_\mu$. In Eq. (28), we will provide an approximate analytical formula for this susceptibility.

The Gaussian character of the splitting distribution, and the linear dependence of the splitting standard deviation on the disorder strength, can be qualitatively understood in three steps. We briefly summarize these here, and will use these

considerations in the next section in our quantitative derivations.

(i) The bonding and antibonding Majorana levels, see Fig. 1(a), are particle-hole symmetric partners of each other. This implies that disorder (or any other perturbation) can not couple them directly.

(ii) Therefore, there is no need to use degenerate or quasidegenerate perturbation theory to describe the leading-order effect of disorder on the energy levels. It is sufficient to do first-order nondegenerate perturbation theory for, say, the antibonding level. This explains the linear dependence of the splitting standard deviation on the disorder strength.

(iii) The first-order perturbative description implies that the first-order energy correction $\delta\epsilon_0^{(1)}$ due to disorder in our model (independent random on-site energies) is a sum of many independent random variables for long chains, $N \gg 1$, and hence the central limit theorem ensures the Gaussian character of that energy correction.

Finally, we point out an out-of-phase oscillation effect between the standard deviation σ_{ϵ_0} of the signful splitting and the clean splitting ϵ_0 . In Fig. 2(b), the blue points show the

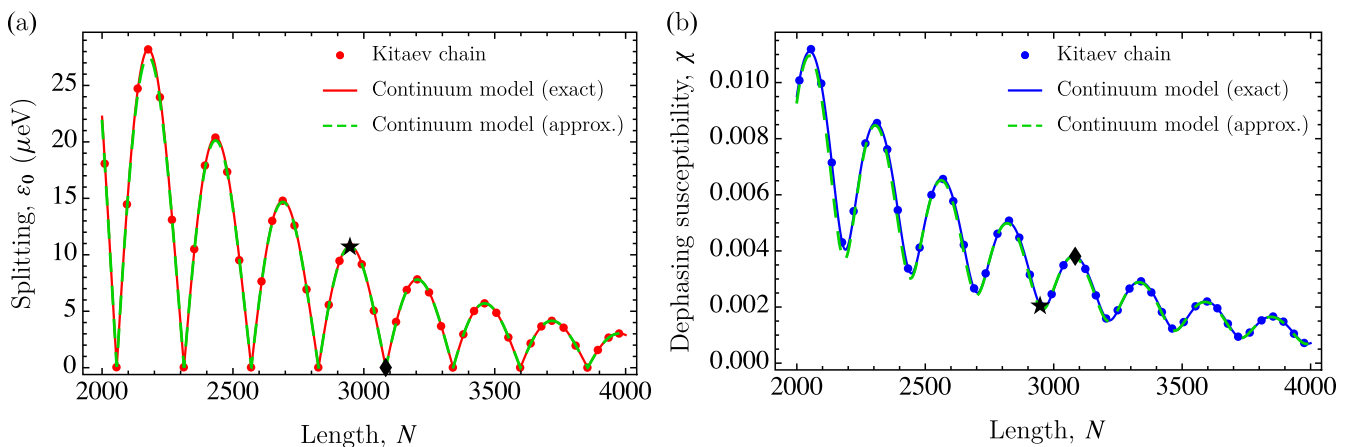


FIG. 3. Comparison of continuum-model (lines) and Kitaev-model (dots) results for the splitting and its standard deviation. (a) Splitting as a function of length for the clean system. (b) Dephasing susceptibility to disorder—that is, the ratio $\chi = \sigma_{\epsilon_0}/\sigma_\mu$ of the standard deviation of the splitting and the strength of the on-site disorder—is shown as a function of the length. See Table I for parameter values. In both panels, star and diamond denote two specific chain lengths, for which the dephasing curves are shown in Fig. 4.

length dependence of the standard deviation σ_{ε_0} of the signful splitting, for the disorder strength $\sigma_\mu = 100 \mu\text{eV}$. Note that the y axis for these blue points is the right y axis which is also colored blue. Figure 2(b) shows that the splitting standard deviation σ_{ε_0} oscillates and decays as the length increases, similarly to the splitting of the clean system. However, there is an out-of-phase oscillation between the splitting and its disorder-induced standard deviation: e.g., the standard deviation has a maximum wherever the splitting reaches zero.

Note that with our numerical approach, it is straightforward to estimate the standard deviation σ_{ε_0} of the splitting; however, Majorana-qubit dephasing is determined by the standard deviation σ_{ε_0} of the signful splitting [Eq. (35)]. To estimate the latter, we do the following. If the expectation value of the splitting ε_0 is much larger than its standard deviation, then $\sigma_{\varepsilon_0} \approx \sigma_{\varepsilon_0}$; hence we use the splitting statistics to estimate σ_{ε_0} . If the above condition does not hold, then we convert the statistics of the splitting to the statistics of the signful splitting, and from the latter we estimate σ_{ε_0} , as described in Appendix B.

III. SPLITTING IN THE CONTINUUM VERSION OF THE KITAEV CHAIN

Numerical computation of the splitting from the Kitaev chain model or other tight-binding models can be computationally expensive for larger system size. To establish a more efficient calculational tool, and to enable analytical results for the splitting absolute value and its standard deviation, here we study the continuum version of the Kitaev chain. These analytical results serve also as a benchmark against which the numerical results can be checked.

First, we use mode matching to obtain the BdG wave function of the quasi-zero-energy mode in a clean (disorder-free) wire [63]. Second, we use this wave function and first-order nondegenerate perturbation theory to determine the standard deviation σ_{ε_0} of the splitting.

A. Splitting and antibonding Majorana wave function in a clean wire

The continuum model has the following momentum-space Hamiltonian [29]:

$$\mathcal{H}_C(k) = \left(\frac{\hbar^2 k^2}{2m} - \mu_C \right) \sigma_z - \Delta'_C \hbar k \sigma_x, \quad (2)$$

where m is the effective mass, μ_C is the chemical potential, and σ_x and σ_z are Pauli matrices acting in Nambu space. The index C stands for ‘‘continuum.’’ For future use, we define

$$k_F = \sqrt{2m\mu_C}/\hbar, \quad (3a)$$

$$v_F = \hbar k_F/m, \quad (3b)$$

$$\Delta_C = \Delta'_C \hbar k_F, \quad (3c)$$

$$\xi = \hbar v_F/\Delta_C, \quad (3d)$$

where k_F is the Fermi wave number, v_F is the Fermi velocity, Δ_C is the superconducting gap, and ξ is the superconductor

coherence length. We will describe a finite-length wire with length L and hard-wall boundary conditions. The relation of this Hamiltonian and the Kitaev chain Hamiltonian is detailed in Appendix A.

We use mode-matching to determine the Majorana antibonding state and its energy (the splitting). The first step is to establish the evanescent modes close to zero energy in a homogeneous system. With that aim, we insert the standard plane-wave ansatz to the BdG equation $\mathcal{H}_C(-i\partial_x)\psi(x) = \varepsilon\psi(x)$ defined by Eq. (2).

It is straightforward to show that this approach yields four evanescent solutions for energies $0 \leq \varepsilon < \sqrt{\Delta_C^2 - (\Delta_C^2/2\mu_C)^2}$, with complex wave numbers $k_1 = K + i\kappa$, $k_2 = -K + i\kappa$, $k_3 = K - i\kappa$, and $k_4 = -K - i\kappa$. Here

$$K = \frac{1}{\hbar} \sqrt{m \left(\sqrt{\mu_C^2 - \varepsilon^2} + \mu_C - \frac{\Delta_C^2}{2\mu_C} \right)}, \quad (4a)$$

$$\kappa = \frac{1}{\hbar} \sqrt{m \left(\sqrt{\mu_C^2 - \varepsilon^2} - \mu_C + \frac{\Delta_C^2}{2\mu_C} \right)}. \quad (4b)$$

Furthermore, K and κ are positive numbers for $0 \leq \varepsilon < \sqrt{\Delta_C^2 - (\Delta_C^2/2\mu_C)^2}$. The corresponding non-normalized wave functions have the form

$$\psi_{k_i}(x) = \begin{pmatrix} u_{k_i} \\ v_{k_i} \end{pmatrix} e^{ik_i x} = \begin{pmatrix} \Delta'_C \hbar k_i \\ \frac{\hbar^2 k_i^2}{2m} - \mu_C - \varepsilon \end{pmatrix} e^{ik_i x}, \quad (5)$$

where u_{k_i} and v_{k_i} represent the electron and hole components of the wave function in the momentum space.

The antibonding Majorana wave function must be a linear superposition of the four evanescent modes at a given energy:

$$\psi(x) = \begin{pmatrix} \psi_e(x) \\ \psi_h(x) \end{pmatrix} = \sum_{i=1}^4 \alpha_i \begin{pmatrix} u_{k_i} \\ v_{k_i} \end{pmatrix} e^{ik_i x}, \quad (6)$$

where $\psi_e(x)$ and $\psi_h(x)$ are the electron and hole components of the Majorana bound state; furthermore, the α_i 's are complex coefficients. This wave function $\psi(x)$ must satisfy the hard-wall boundary conditions:

$$\psi(0) = \psi(L) = \begin{pmatrix} 0 \\ 0 \end{pmatrix}. \quad (7)$$

This condition is fulfilled by coefficient vectors satisfying the following homogeneous linear set of equations:

$$\mathcal{M} \begin{pmatrix} \alpha_1 \\ \alpha_2 \\ \alpha_3 \\ \alpha_4 \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}, \quad (8)$$

where the ε -dependent matrix \mathcal{M} is defined as

$$\mathcal{M} = \begin{pmatrix} u_{k_1} & u_{k_2} & u_{k_3} & u_{k_4} \\ v_{k_1} & v_{k_2} & v_{k_3} & v_{k_4} \\ u_{k_1} e^{ik_1 L} & u_{k_2} e^{ik_2 L} & u_{k_3} e^{ik_3 L} & u_{k_4} e^{ik_4 L} \\ v_{k_1} e^{ik_1 L} & v_{k_2} e^{ik_2 L} & v_{k_3} e^{ik_3 L} & v_{k_4} e^{ik_4 L} \end{pmatrix}. \quad (9)$$

As follows from Eq. (8), for a given length L , the condition

$$\det[\mathcal{M}(\varepsilon)] = 0 \quad (10)$$

gives the energy ε_0 of the antibonding Majorana state. In general, Eq. (10) leads a transcendental equation, which can be solved numerically: $\varepsilon_{0,\text{num}}$. Power-series expansion of $\det(\mathcal{M})$ in ε up to second order provides an analytical

$$\varepsilon_0(L) \approx \begin{cases} \Delta_C \frac{2k_F}{\sqrt{k_F^2 - 1/\xi^2}} e^{-L/\xi} \left| \sin\left(\sqrt{k_F^2 - 1/\xi^2} L\right) \right|, & \text{if } k_F > 1/\xi, \\ \Delta_C \frac{k_F}{\sqrt{1/\xi^2 - k_F^2}} e^{-(1/\xi - \sqrt{1/\xi^2 - k_F^2})L}, & \text{if } k_F < 1/\xi. \end{cases} \quad (12a)$$

$$(12b)$$

If $k_F > 1/\xi$ (i.e., when $\mu_C > \Delta_C/2$), the splitting has an oscillatory part, but if $k_F < 1/\xi$, the splitting decreases purely exponentially as the length increases. For a physically feasible parameter set shown in Table I, including a chemical potential (e.g., set by a gate voltage) $\mu_C = 1$ meV, we obtain $k_F > 1/\xi$. This is the case we focus on from now on. To reach $k_F < 1/\xi$, the chemical potential needs to be suppressed as $0 < \mu_C < 0.1$ meV; we do not treat this case here.

Our result (12a) is in fact a generalization of an earlier result; see below Eq. (5) in Ref. [62] and Eq. (18) in Ref. [65]. The only difference is the appearance of $\sqrt{k_F^2 - 1/\xi^2}$ in our result. The earlier result can be obtained by taking the limit $k_F \gg 1/\xi$ of our formula (12), i.e., by applying the approximation $\sqrt{k_F^2 - 1/\xi^2} \approx k_F$.

Next, we describe the antibonding Majorana wave function. To simplify the description, we utilize the symmetries of the setup. The clean system has inversion symmetry. The corresponding operator has the form $\Pi = \pi \otimes \sigma_z$, where π is the inversion with respect to the point $x = L/2$, acting in real space, and σ_z acts in Nambu space. Inversion symmetry, together with the assumption that the antibonding Majorana energy level is nondegenerate, implies that

$$\begin{pmatrix} \psi_e(x) \\ \psi_h(x) \end{pmatrix} = \begin{pmatrix} \pm \psi_e(L-x) \\ \mp \psi_h(L-x) \end{pmatrix}. \quad (13)$$

where A_e and A_h are normalization factors,

$$\phi_e = \arctan \left[\frac{pe^{-\kappa L} \sin(KL)}{1 + pe^{-\kappa L} \cos(KL)} \right], \quad (17a)$$

$$\phi_h = \arctan \left[\frac{-pe^{-\kappa L} \sin(KL)}{1 - pe^{-\kappa L} \cos(KL)} \right] \quad (17b)$$

are phases, and $p = +1$ ($p = -1$) corresponds to the behavior under inversion, that is, to the upper (lower) sign in Eq. (13).

solution that in the limit of $L \gg 1/k_F$, ξ reads

$$\varepsilon_0(L) \approx 2\Delta_C k_F e^{-L/\xi} \left| \frac{\sin\left(\sqrt{k_F^2 - 1/\xi^2} L\right)}{\sqrt{k_F^2 - 1/\xi^2}} \right|, \quad (11)$$

where we use $K|_{\varepsilon=0} = \sqrt{k_F^2 - 1/\xi^2}$ and $\kappa|_{\varepsilon=0} = 1/\xi$.

Depending on the relative magnitude of k_F and $1/\xi$, from Eq. (11) we obtain

The Hamiltonian Eq. (2) also has bosonic time-reversal symmetry with the operator $T = (\mathbb{1} \otimes \sigma_z)\mathcal{K}$, where \mathcal{K} is the complex conjugation, and it fulfills the relation $T^2 = 1$. Time-reversal symmetry restricts the form of the nondegenerate energy eigenstate as

$$T \begin{pmatrix} \psi_e(x) \\ \psi_h(x) \end{pmatrix} = \begin{pmatrix} \psi_e^*(x) \\ -\psi_h^*(x) \end{pmatrix} = e^{i\varphi} \begin{pmatrix} \psi_e(x) \\ \psi_h(x) \end{pmatrix}, \quad (14)$$

where φ depends on the global phase of the wave function. For concreteness, we fix this global phase such that $\varphi = 0$. Given an eigenstate ψ with an arbitrary global phase, the eigenstate with $\varphi = 0$ is obtained as $\psi(x) + T\psi(x)$. This choice $\varphi = 0$ leads to

$$\text{Im}[\psi_e(x)] = 0, \quad (15a)$$

$$\text{Re}[\psi_h(x)] = 0. \quad (15b)$$

Equations (7), (13), and (15) constrain the form of the wave function:

$$\psi(x) = \begin{pmatrix} A_e \{e^{-\kappa x} \sin(Kx - \phi_e) + pe^{-\kappa(L-x)} \sin[K(L-x) - \phi_e]\} \\ iA_h \{e^{-\kappa x} \sin(Kx - \phi_h) - pe^{-\kappa(L-x)} \sin[K(L-x) - \phi_h]\} \end{pmatrix}, \quad (16)$$

In the limit $L \gg \xi > 1/k_F$, the following approximations can be made:

$$p = \text{sgn}[\sin(\sqrt{k_F^2 - 1/\xi^2} L)], \quad (18a)$$

$$\phi_e = -\phi_h = \phi \approx e^{-L/\xi} \left| \sin\left(\sqrt{k_F^2 - 1/\xi^2} L\right) \right|, \quad (18b)$$

$$A_e = A_h = A \approx \frac{1}{\sqrt{\frac{\xi}{2} \left(1 - \frac{\cos(\phi/2)}{k_F^2 \xi^2}\right)}}. \quad (18c)$$

We note that p changes sign where the splitting vanishes.

To obtain Eq. (18), we compare the wave function in Eqs. (6) and in (16), yielding

$$\phi_e = \arctan \left[\frac{\alpha_1 u_{k_1} + \alpha_2 u_{k_2}}{i(\alpha_2 u_{k_2} - \alpha_1 u_{k_1})} \right]. \quad (19)$$

The coefficient vector $(\alpha_1, \alpha_2, \alpha_3, \alpha_4)^T$ is the null space of the matrix \mathcal{M} , which we find analytically by Gauss elimination. Comparing Eqs. (17) and (19) up to leading order in $e^{-L/\xi}$, we find Eq. (18). Furthermore, we find the approximate formula for the phases in Eq. (18 b) using Eqs. (17), by taking the leading-order approximation in $e^{-L/\xi}$, and utilizing Eq. (18). To obtain Eq. (18), we assumed that the electron and hole character of the wave function has exactly equal probability in the limit of $L \gg \xi > 1/k_F$, which results in $A_e = A_h$. We derived Eq. (18) from the norm of the wave function in Eq. (16) by taking the limit for $L \rightarrow \infty$. We will use Eqs. (18) to derive an approximate analytical formula for the dephasing susceptibility to disorder shown in Eq. (29).

B. Standard deviation of the splitting

Now we describe the broadening of the splitting distribution due to on-site disorder. The full Hamiltonian of the disordered system can be written as

$$H_C = \mathcal{H}_C(-i\hbar\partial_x) + H_{\text{dis}}, \quad (20)$$

where $H_{\text{dis}} = \delta\mu_C(x)\sigma_z$ is the disorder Hamiltonian, representing disorder in the chemical potential. We model disorder as a collection of potential steps, where the lengths of the steps are equal and denoted by a_{dis} :

$$\delta\mu_C(x) = \sum_{i=1}^{N_{\text{dis}}} \delta\mu_i^{(C)} \Xi_i(x), \quad (21)$$

where

$$\Xi_i(x) = \begin{cases} 1, & i-1 \leq x/a_{\text{dis}} < i, \\ 0, & \text{otherwise.} \end{cases} \quad (22)$$

This model is a natural analog of the disorder model we used in the Kitaev chain, with the identification $a = a_{\text{dis}}$, where a is the lattice constant of the Kitaev chain.

We regard disorder as a perturbation, and calculate the first-order energy shift. Naively, one should do degenerate perturbation theory, since the antibonding and bonding Majorana energies are close to each other. However, disorder does not couple them; hence nondegenerate perturbation theory is sufficient. The proof of this is as follows.

Due to the particle-hole symmetry of the BdG Hamiltonian, $\langle \psi | H_{\text{dis}} | P\psi \rangle = \langle P\psi | H_{\text{dis}} | \psi \rangle = 0$, where $|\psi\rangle$ and $|P\psi\rangle$ are the positive and negative energy solution of the BdG Hamiltonian, and $P = (\mathbb{1} \otimes \sigma_x)\mathcal{K}$ is the operator of the

particle-hole symmetry. This can be seen by

$$\begin{aligned} \langle \psi | H | P\psi \rangle &= -\langle \psi | PH | \psi \rangle \\ &= -\langle P\psi | H | \psi \rangle^* = -\langle H\psi | P\psi \rangle = -\langle \psi | HP\psi \rangle, \end{aligned} \quad (23)$$

where the first equation is implied by the fact that H anti-commutes with P , the second equation is a consequence of the antiunitary property of P , the third equation is obtained by flipping the scalar product, and the fourth equation is implied by H being Hermitian.

Applying the relation $\langle \psi | H_{\text{dis}} | P\psi \rangle = 0$ to the antibonding $|\psi\rangle$ and bonding $|P\psi\rangle$ Majorana wave functions, we conclude that they are uncoupled and therefore the first-order disorder-induced shift of the signful splitting is $\delta\varepsilon_0^{(1)} = \langle \psi | H_{\text{dis}} | \psi \rangle$. Using Eqs. (6) and (21), this shift can be written as

$$\langle \psi | H_{\text{dis}} | \psi \rangle = \sum_{i=1}^{N_{\text{dis}}} \delta\mu_i^{(C)} \Theta_i, \quad (24)$$

where

$$\Theta_i = \int_{(i-1)a_{\text{dis}}}^{ia_{\text{dis}}} [|\psi_e(x)|^2 - |\psi_h(x)|^2] dx. \quad (25)$$

In analogy with our disorder model in the Kitaev chain, discussed in Sec. II, we assume independence and a normal distribution for the chemical potential disorder, which we denote as $\delta\mu_i^{(C)} \sim \mathcal{N}(0, \sigma_\mu)$, where σ_μ is the disorder strength, and the $\delta\mu_i^{(C)}$'s are independent of each other. From Eq. (24), we conclude that the disorder matrix element also follows a Gaussian distribution:

$$\langle \psi | H_{\text{dis}} | \psi \rangle \sim \mathcal{N}(0, \sigma_{\varepsilon_0}), \quad (26)$$

where $\sigma_{\varepsilon_0} = \sigma_\mu \sqrt{\sum_{j=1}^{N_{\text{dis}}} \Theta_j^2}$, the standard deviation of the distribution of the signful splitting.

Let us suppose that $|\psi_e(x)|^2 - |\psi_h(x)|^2$ varies slowly on the scale of a_{dis} . This implies

$$\begin{aligned} \sum_{i=1}^{N_{\text{dis}}} \Theta_i^2 &\approx \sum_{i=1}^{N_{\text{dis}}} a_{\text{dis}}^2 [|\psi_e(ia_{\text{dis}})|^2 - |\psi_h(ia_{\text{dis}})|^2]^2 \\ &\approx a_{\text{dis}} \int_0^L [|\psi_e(x)|^2 - |\psi_h(x)|^2]^2 dx. \end{aligned} \quad (27)$$

Therefore, the dephasing susceptibility to disorder is obtained as

$$\chi \equiv \frac{\sigma_{\varepsilon_0}}{\sigma_\mu} = \sqrt{a_{\text{dis}} \int_0^L [|\psi_e(x)|^2 - |\psi_h(x)|^2]^2 dx}. \quad (28)$$

By solving Eq. (8) numerically, we obtain the values $\alpha_{i,\text{num}}$ of α_i . Substituting these numerical values $\alpha_{i,\text{num}}$ and $\varepsilon_{0,\text{num}}$ into Eqs. (4) and (16), we obtain the semianalytical wave function $\psi_e(x)$ and $\psi_h(x)$. After normalization, Eq. (28) can be performed. The results, shown in Fig. 3(b) as ‘‘exact,’’ are discussed below.

As an alternative to the above semianalytical approach, an approximate analytical formula can be obtained by substituting the form of the wave function in Eqs. (16) into Eq. (28) using Eqs. (18). After the integration over x , and taking series

expansion in $\kappa e^{-\kappa L}$, the dephasing susceptibility to disorder in the limit of $L \gg \xi > 1/k_F$ can be written as

$$\chi = \sqrt{\frac{a_{\text{dis}}}{2\xi}} e^{-L/\xi} \sqrt{\frac{8L}{\xi} - 3 + \left(\frac{4L}{\xi} + 3\right) \cos(2\sqrt{k_F^2 - 1/\xi^2}L)}. \quad (29)$$

Equation (29) is the key result of our work. It reveals that the dephasing susceptibility (and hence the dephasing rate) as a function of system parameters exhibits oscillations that are out-of-phase with the oscillations of the clean splitting given in Eq. (12). This is apparent as Eq. (12) contains a sine whereas Eq. (29) contains a cosine.

Furthermore, Eq. (29) also suggests the absence of dephasing sweet spots in this setting: the long expression below the square root in Eq. (29) is always positive due to the condition $L \gg \xi$.

In conclusion, we have described a semianalytical procedure, and an approximate analytical procedure, to estimate the disorder-induced broadening of the distribution of the signful splitting in a continuum model of a 1D topological superconductor.

C. Comparing the results of the two models

Here, we show the correspondence of the Kitaev chain and continuum model results for the splitting ε_0 , the disorder-induced standard deviation σ_{ε_0} , and the dephasing susceptibility χ . In Appendix A, we show how to connect the parameters of the continuum and discrete (Kitaev) models.

In Fig. 3(a), we plot the splitting of the clean system as a function of the chain length. Red points show the numerical result from the Kitaev chain model, whereas the red solid line shows the semianalytical exact result from the continuum model, obtained by solving Eq. (10) numerically. The dashed green line shows the result of Eq. (12). Parameter values are those listed in Table I. In Fig. 3(a), the Kitaev chain result (red points) and the exact result from the continuum model (red solid line) are indistinguishable. The analytical approximate result (green dashed line) shows a slight deviation from the other two data sets for a short chain, but becomes indistinguishable from those for long chains.

In Fig. 3(b), we plot the dephasing susceptibility, that is, the ratio of splitting standard deviation σ_{ε_0} and the disorder strength σ_{μ} , as a function of the chain length. The Kitaev model result (points) is obtained numerically, using 10 000 random disorder realizations for each length. Here again, the two models show satisfactory agreement.

IV. SIGNFUL SPLITTING DISTRIBUTION AND MAJORANA QUBIT DEPHASING

In this section, we complete our primary task, and describe the dephasing dynamics of a Majorana qubit subject to quasistatic disorder.

A. Noise model: Quasistatic disorder

Let us start this description by defining our noise model of quasistatic disorder, and relating it to device physics. Electrical potential fluctuations are generically present in qubit

devices, and often dominate qubit decoherence. In many experiments, this noise has been found to follow a frequency-dependent power spectrum $S(f) \propto 1/f$. Due to dominance of the low-frequency component, one can refer to this type of noise as slow charge noise.

In this work, we account for the most prominent feature of this noise, i.e., that it detunes the electrostatic potential felt by the electrons in the Majorana wire. Regarding the spatial structure of the noise, we first focus on short-range correlations (Sec. IV), but later we also describe Majorana qubit dephasing as the spatial correlation length is varied (Sec. V).

Regarding the temporal structure of noise, we follow numerous earlier works by applying the quasistatic approximation. To define the quasistatic approximation, we first recall how a dephasing-time experiment (*Ramsey experiment*) is performed. First, a balanced superposition of the two computational basis states, with a Bloch vector aligned with, say, the x axis, is prepared. Then, this state is allowed to evolve freely for a waiting time τ_w much shorter than the dephasing time. After time τ_w , the qubit is measured in the x basis. This is often called one “shot” of the experiment. This shot is repeated many ($N_{\text{rep}} \gg 1$) times to gain statistics and eliminate shot noise, and the whole sequence of N_{rep} shots is repeated for $N_{\tau} \gg 1$ different, stepwise increasing values of the waiting time. Typically, the largest τ_w value is a few times greater than the dephasing time.

As applied to this scheme, the quasistatic approximation of noise is composed of two assumptions: (i) for each run, the noise is considered time-independent, i.e., it is static disorder, and (ii) for the N_{rep} shots with a single waiting time, the different static disorder configurations acting during the different runs provide a good statistical coverage of all disorder configurations.

B. Majorana qubit dephasing

Consider a Majorana qubit encoded in two identical topological superconducting wires. All parameters are assumed to be equal, including the disorder strength. The two wires are assumed to be decoupled from each other (no tunneling between the two wires). We restrict our attention to the globally even ground state of this setup, which is spanned by the basis states $|0\rangle \equiv |e_1, e_2\rangle$ and $|1\rangle \equiv |o_1, o_2\rangle$, where the names e and o refer to the even and odd fermion parities of the corresponding states, and the indices 1 and 2 refer to the first and second wire, respectively.

To perform a qubit dephasing experiment, one usually creates an initial state $|\psi_i\rangle$ that is a balanced superposition of the two basis states, e.g., with a qubit polarization vector along the x direction

$$|\psi_i\rangle = \frac{1}{\sqrt{2}}(|0\rangle + |1\rangle). \quad (30)$$

The qubit polarization vector (Bloch vector) for this state is

$$\vec{p} \equiv \langle \psi_i | \vec{\sigma} | \psi_i \rangle = (1, 0, 0), \quad (31)$$

where $\vec{\sigma} = (\sigma_x, \sigma_y, \sigma_z)$ is the vector of Pauli matrices. Note that the preparation of this initial state itself can be corrupted by disorder, a complication that we disregard here.

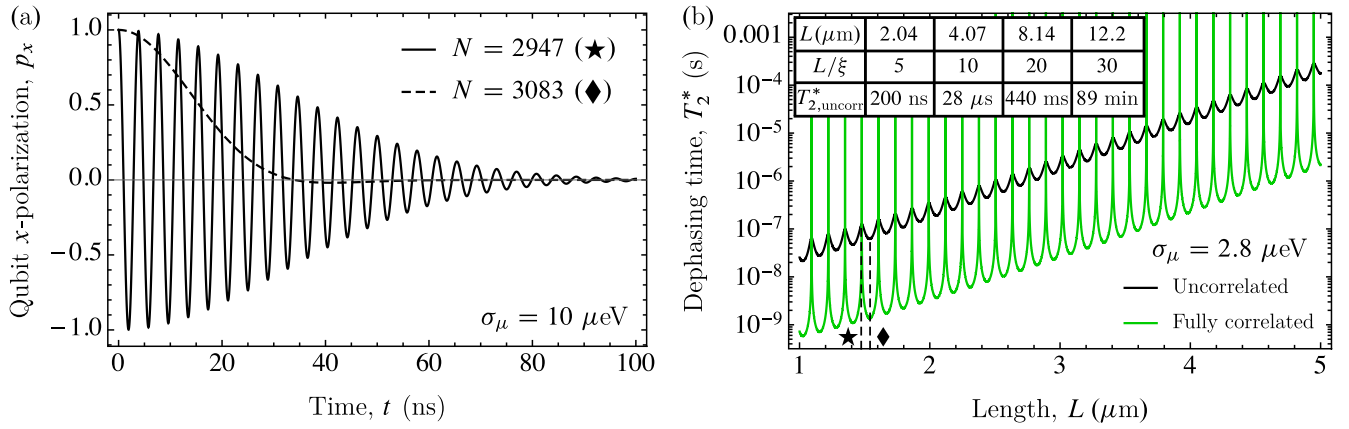


FIG. 4. Dephasing of a Majorana qubit due to quasistatic disorder. (a) Dephasing curves. The x component of the disorder-averaged polarization vector as a function of time for two different lengths (see star and diamond in Fig. 3). The envelopes of the curves follow Gaussian dephasing. The finiteness of the mean of the signful splitting is responsible for the oscillations of the solid line. The out-of-phase oscillation between the clean splitting and dephasing susceptibility is illustrated: the longer chain (dashed line) shows faster dephasing but no Larmor precession. (b) Inhomogeneous dephasing time as a function of the length for disorder strength $\sigma_\mu = 2.8 \mu\text{eV}$ when disorder is uncorrelated (black) and when fully correlated (green). Table inset shows the results corresponding to uncorrelated disorder for specific lengths.

After preparation, the relative phase between the two basis states evolves in time due to the excess energy $\epsilon_0^{(1)}$ of o_1 with respect to e_1 in wire 1, and the excess energy $\epsilon_0^{(2)}$ of o_2 with respect to e_2 in wire 2. In particular, the time-dependent wave function, up to an irrelevant global phase, reads

$$|\psi(t)\rangle = \frac{1}{\sqrt{2}} (|0\rangle + e^{-i(\epsilon_0^{(1)} + \epsilon_0^{(2)})t/\hbar} |1\rangle). \quad (32)$$

Then, the quasistatic assumption implies that on average, for a large number of measurements, the qubit polarization vector evolves in time as

$$\langle \vec{p}(t) \rangle \equiv \langle \psi(t) | \vec{\sigma} | \psi(t) \rangle = \int d\epsilon \rho(\epsilon) \begin{pmatrix} \cos(\epsilon t/\hbar) \\ -\sin(\epsilon t/\hbar) \\ 0 \end{pmatrix}, \quad (33)$$

where $\epsilon = \epsilon_0^{(1)} + \epsilon_0^{(2)}$ is the random qubit energy splitting, and $\rho(\epsilon)$ is its pdf.

We illustrate the dephasing dynamics by calculating $\langle p_x(t) \rangle$, the x component of the disorder-averaged polarization vector as the function of time. We will refer to this function as the *dephasing curve*. We evaluate the dephasing curve based on the observation that the signful splitting has a Gaussian pdf in the parameter range we consider. Based on Eq. (33), this implies the following well-known result [58,66] for the dephasing curve:

$$\langle p_x(t) \rangle = e^{-\left(\frac{\epsilon_{0,c}}{\hbar} t\right)^2} \cos\left(\frac{2\epsilon_{0,c}}{\hbar} t\right), \quad (34)$$

where $\epsilon_{0,c}$ is the clean splitting. This result implies that dephasing follows Gaussian decay, and this decay is characterized by the timescale

$$T_2^* = \frac{\hbar}{\sigma_{\epsilon_0}} = \frac{\hbar}{\sigma_\mu \chi}, \quad (35)$$

which is often called the *inhomogeneous dephasing time*.

Figure 4(a) shows two dephasing curves for the parameter set shown in Table I, the solid line showing fast oscillations

(i.e., Larmor precession), and the dashed line showing no oscillations. The dashed line corresponds to the diamond ($N = 3083$) in Fig. 3, with chain length fine-tuned such that the clean splitting vanishes. The solid line corresponds to the star ($N = 2947$) in Fig. 3, with chain length fine-tuned such that the clean splitting has a local maximum. For both chain lengths, the pdf of the signful splitting is Gaussian. However, the mean of the signful splitting (which is the same as the clean splitting $\epsilon_{0,c}$) is zero for the $N = 3083$ case, and finite for the $N = 2947$ case, the latter being responsible for the oscillations in Fig. 4(a). This figure also illustrates the out-of-phase relation between the clean splitting and dephasing susceptibility [see, e.g., Fig. 2(b)]: the smaller the clean splitting, the faster the dephasing.

It is also interesting to note that the oscillation (Larmor precession) induced by the finite clean splitting, as shown by the solid line in Fig. 4(a), has a much smaller timescale than the dephasing time. It would be interesting to study in detail how this fast Larmor precession influences the fidelity of quantum gates, e.g., based on braiding of MZMs [21].

The black solid line Fig. 4(b) shows the inhomogeneous dephasing time as a function of the length. (The green solid line will be discussed in the next section.) The dephasing time is calculated analytically by substituting the approximate formula of χ given by Eq. (29) into Eq. (35). Aside from the oscillations seen in Fig. 4(b), the dependence of the dephasing time on the chain length is dominated by the exponential factor $T_2^* \propto e^{L/\xi}$. The figure corresponds to a disorder strength $\sigma_\mu = 2.8 \mu\text{eV}$, which implies a dephasing time $T_2^* = 200$ ns for $L/\xi = 5$. The oscillatory nature of the black solid result in Fig. 4(b) is responsible for the feature of Fig. 4(a) that the shorter chain (star) has a longer T_2^* than the longer chain (diamond).

The inset of Fig. 4(b) shows the calculated inhomogeneous dephasing time values for specific chain lengths. We use this table, in particular the inhomogeneous dephasing time value $T_2^* = 200$ ns at $L/\xi = 5$, to relate our results the earlier dephasing-time estimates of Ref. [35] (see Table I therein).

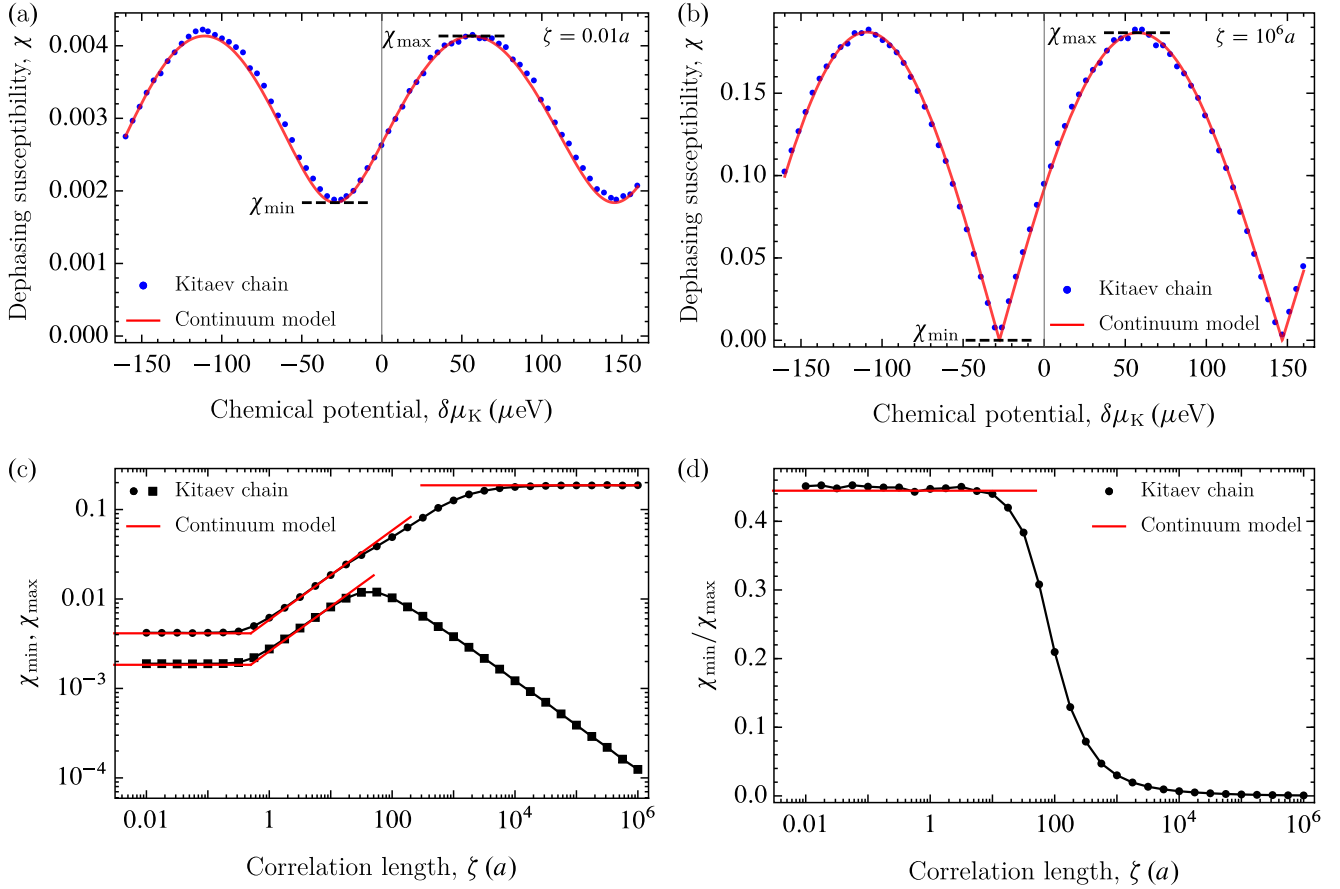


FIG. 5. Effect of the spatial correlation of disorder on dephasing. (a), (b) Dephasing susceptibility as a function of the chemical potential for uncorrelated disorder (a) and for fully correlated disorder (b). Points show results from the Kitaev chain model with spatially correlated disorder. Solid lines show analytical results of the continuum model in the two limits. (c) Adjacent minimum and maximum values of the dephasing susceptibility as a function of the correlation length. (d) Ratio of the minimum and the maximum values as a function of the correlation length. This ratio is constant in the uncorrelated and weakly correlated regimes ($\zeta \lesssim 10a$), and goes to zero as the correlation length increases. The latter feature corresponds to the dephasing sweet spots of the fully correlated regime seen in (b).

Reference [35] predicts this T_2^* value from intrinsic sources, without any disorder in the sample. Therefore, our parameter value $\sigma_\mu = 2.8 \mu\text{eV}$ provides an estimate for the crossover disorder strength, that is, the disorder strength above which dephasing due to quasistatic disorder dominates the intrinsic dephasing mechanisms of a clean system (homogeneous $1/f$ charge noise, phonons, equilibrium quasiparticles).

Experimental data indicate that the typical energy scale of local electrostatic fluctuations in state-of-the-art semiconductor quantum devices is of the order of a few μeV ; see, e.g., Table II of Ref. [35]. This suggests that the mechanism we describe here will be relevant for early-stage Majorana-qubit experiments.

V. DEPHASING DYNAMICS AS A PROBE OF SPATIAL DISORDER CORRELATIONS

Up to this point, we have focused on the case where the on-site disorder is uncorrelated between different sites. This model represents short-range-correlated disorder that leads to the absence of a dephasing sweet spot. On the other hand, if dephasing is caused by the fluctuation of a global control parameter, e.g., the chemical potential, then a dephasing sweet

spot is expected when the clean splitting has a maximum as the function of that parameter. This is exemplified, e.g., by Eq. (5) of Ref. [35].

In this section, we go beyond the uncorrelated disorder model to highlight the relation of the spatial correlations of the disorder and the dephasing curve. To this end, we generalize our disordered Kitaev chain model by regarding the on-site energies as correlated normal random variables, described by a multivariate normal distribution with zero means and the covariance matrix

$$\Sigma_{ij} = \sigma_\mu^2 e^{-|i-j|a/\zeta}, \quad (36)$$

where i, j are site indices and ζ is the correlation length. For further details of the model see Appendix D; for a discussion between our model and disorder in real devices see Sec. VI. Parameter ζ controls the spatial correlation in the disorder realizations: $\zeta \lesssim a$ indicates uncorrelated disorder; furthermore, $\zeta \gtrsim L$ corresponds to the homogeneous, fully correlated case.

Figures 5(a) and 5(b) show the dephasing susceptibility as a function of the chemical potential for $\zeta = 0.01a$ (uncorrelated disorder) and $\zeta = 10^6a$ (fully correlated disorder), respectively. The system size is fixed to 3000 sites. The chemical potential $\delta\mu_K$ is measured from μ_K given in Table I. The

blue points correspond to a numerical calculation based on the Kitaev chain model with correlated on-site energy disorder. The red line of Fig. 5(a) shows the analytically obtained dephasing susceptibility of the continuum model, see Eq. (29), with $a_{\text{dis}} = a$. The red line in Fig. 5(b) shows the dephasing susceptibility of the continuum model against homogeneous chemical potential disorder, which can be obtained by taking the derivative of the clean splitting formula in Eq. (12) with respect to μ_K . We find furthermore that the corresponding lengthy formula can be approximated (not shown) as

$$\chi_{\text{fcorr}} = \frac{2L}{\xi} \frac{k_F^2}{k_F^2 - 1/\xi^2} e^{-L/\xi} \left| \cos \left(\sqrt{k_F^2 - 1/\xi^2} L \right) \right|. \quad (37)$$

The main observations in Figs. 5(a) and 5(b) are as follows: (i) the dephasing susceptibility oscillates as a function of the chemical potential in both panels, (ii) the magnitude of the oscillations is greater in the case of fully correlated disorder [Fig. 5(b)] than in the case of uncorrelated disorder [Fig. 5(a)], and (iii) the case of fully correlated disorder [Fig. 5(b)] exhibits dephasing sweet spots, where the dephasing susceptibility χ vanishes. See, e.g., at $\delta\mu_K \approx -30 \mu\text{eV}$.

In a future Majorana-qubit dephasing experiment, oscillations such as those shown in Figs. 5(a) and 5(b) are directly observable, e.g., by tuning the chemical potential via a back-gate voltage. Here we argue that such an oscillatory data set can be used to infer the disorder correlation length.

To illustrate this opportunity, we take two adjacent extrema in Figs. 5(a) and 5(b): a minimum (χ_{min}) and a maximum (χ_{max}), located closest to $\delta\mu_K = 0$ (see labels in figure). In Fig. 5(c), we show how χ_{min} and χ_{max} evolve as functions of the correlation length ζ . Results from the Kitaev chain model are depicted by black markers, whereas red lines show the analytical results in the uncorrelated regime ($\zeta \lesssim a$), the weakly correlated regime ($a \lesssim \zeta \lesssim 10a$; derivation discussed below), and the fully correlated regime ($\zeta \gtrsim L$). For uncorrelated disorder ($\zeta \lesssim a$), the extremal dephasing susceptibilities are approximately constants. However, for weakly correlated disorder ($a \lesssim \zeta \lesssim 10a$), χ_{min} and χ_{max} increase as the correlation length is increased. This means that qubit dephasing is more sensitive to correlated disorder than to the uncorrelated one. In the fully correlated disorder limit ($\zeta \gtrsim L$), χ_{max} saturates, whereas χ_{min} decreases, in accordance with the dephasing sweet spot seen for this limit in Fig. 5(b).

To introduce a procedure which estimates the correlation length from the dephasing curves, we plot the ratio $\chi_{\text{min}}/\chi_{\text{max}}$ as a function of the correlation length in Fig. 5(d). In the uncorrelated and weakly correlated regimes ($\zeta \lesssim 10a$), the ratio is a constant; for longer correlation length, it tends to zero due to the existence of dephasing sweet spots.

The ratio $\chi_{\text{min}}/\chi_{\text{max}}$ as a function of the correlation length is monotonic, which provides an opportunity to characterize the correlation length experimentally, in the following way: one can measure the dephasing curves by varying the chemical potential and determine the corresponding dephasing times. By fine-tuning the chemical potential, two adjacent minima and maxima of the dephasing times can be determined: $T_{2,\text{min}}^*$ and $T_{2,\text{max}}^*$. The ratio of the extremal dephasing times equals the inverse ratio of the extremal dephasing susceptibilities, i.e., $T_{2,\text{max}}^*/T_{2,\text{min}}^* = \chi_{\text{min}}/\chi_{\text{max}}$, which can be

seen from Eq. (35). Using Fig. 5(d), one can infer the correlation length, or at least can distinguish between short-range and long-range disorder correlations.

In order to support our numerical results in the uncorrelated, weakly, and fully correlated regimes in Fig. 5(c), we provide analytical results from the continuum model, shown as the three solid red line segments in Fig. 5(c). In the uncorrelated limit ($\zeta \lesssim a$), we make use of the extrema of Eq. (29) with $a_{\text{dis}} = a$ in order to determine χ_{min} and χ_{max} . In the fully correlated regime ($\zeta \gtrsim L$), we take the maximum of Eq. (37).

To obtain an analytical result for the weakly correlated regime ($a \lesssim \zeta \lesssim 10a$) from the continuum model, we make use of Eq. (29), which expresses the dephasing susceptibility χ as function of the parameter a_{dis} . [Recall that in our continuum model, the disorder is modeled by series of potential steps of length a_{dis} ; see Eq. (21).] We substitute $a_{\text{dis}} = 2\zeta$ in Eq. (29) to express χ as the function of the disorder correlation length ζ . In what follows, we argue why we identify a_{dis} with 2ζ .

The covariance function of the disorder in the continuum model can be written as

$$C_C(x, y) = \begin{cases} \sigma_\mu^2, & \text{if } \lfloor x/a_{\text{dis}} \rfloor = \lfloor y/a_{\text{dis}} \rfloor, \\ 0, & \text{otherwise.} \end{cases} \quad (38)$$

Matching is based on the following relation

$$\frac{\int_0^\infty x C_C(x, 0) dx}{\int_0^\infty C_C(x, 0) dx} = \frac{\int_0^\infty x C_K(x, 0) dx}{\int_0^\infty C_K(x, 0) dx}, \quad (39)$$

where we use the continuum form of the covariance matrix Σ_{ij} [cf. Eq. (36)]:

$$C_K(x, y) = \sigma_\mu^2 e^{-|x-y|/\zeta}. \quad (40)$$

Equation (39) leads to $a_{\text{dis}} = 2\zeta$. By substituting it into Eq. (29), we find good agreement between the numerical (black points) and analytical (red solid lines) results; see the weakly correlated disorder regime ($a \lesssim \zeta \lesssim 10a$) in Fig. 5(c).

Finally, we discuss the dephasing time for fully correlated disorder, which is shown in Fig. 4(b) by the green solid line, as a function of the length. The dephasing time is calculated analytically by substituting Eq. (37) into Eq. (35). Results from fully correlated disorder oscillate in phase with the results arising from uncorrelated disorder (black solid line). Dephasing sweet spots appear as singularities, showing diverging dephasing time. This is the consequence of our limited dephasing model which is based on the linear approximation $1/T_2^* \propto \sigma_\mu$; see Eq. (35). A higher-order approach would resolve the singular behavior.

In conclusion, our results in Fig. 4 provide important practical insights on how to optimize a Majorana qubit setup for a dephasing experiment. The effect of homogeneous charge noise, that is, a uniform random shift of the chemical potential, can be mitigated by fine-tuning the chemical potential: the dephasing time can be strongly enhanced by such a fine-tuning. If the noise is not spatially homogeneous, then the magnitude of the improvement depends on the correlation length of the disorder: for uncorrelated noise, fine-tuning could yield at most a factor of two improvement [cf. Fig. 5(a)], but this improvement factor gradually increases for increasing disorder correlation length.

VI. DISCUSSION

In the main part of this work, we used a model of short-range-correlated disorder. This is admittedly a minimal model of disorder in real samples; nevertheless we find it important and relevant to provide the corresponding results, because (i) this is a conceptually simple, canonical model, often used in the literature, applied to effects ranging from Anderson localization to Majorana physics [29,42], and (ii) these results also serve as a benchmark for more realistic models. Also, the level of disorder in state-of-the-art hybrid nanowires seems to be too strong to allow for the clear observation of Majorana zero modes, which suggests that disorder will likely play a dominant role also in the initial Majorana-qubit experiments, e.g., qubit dephasing time measurements. In real nanowire samples, disorder might arise due to various physical mechanisms [67], e.g., fluctuating charge traps in the substrate, atoms, ions, and molecules contaminating the wire surface, impurity atoms built in to the crystal upon growth, electron scattering on rough or oxidized wire surfaces and the core-shell interface, inhomogeneous strain patterns due to thermal expansion coefficient mismatch and metal deposition (shell, gates, contacts), gate-voltage fluctuations, etc. It is an important ongoing effort to mitigate these mechanisms; alternatively, it is useful to characterize and control their effects on Majorana qubit decoherence.

In our dephasing calculation, we have chosen the quasistatic approximation also for its conceptual simplicity and widespread use in the literature [49,51]. In real devices, classical or quantum noise often follows a characteristic noise spectrum, e.g., $1/f$ noise [35,40,46,47,55–57,68–70], Johnson-Nyquist noise, quantum noise of phonons [35,36], gate-voltage fluctuations [31,35,39,40], etc. Going beyond the quasistatic approximation by incorporating these frequency-dependent noise features would be an important addition to this work. An especially appealing task is to describe the combined effect of static spatial disorder and fluctuating electric fields; this direction might actually reveal connections between actual device physics and the minimal model used in our present work. A conceptually different but equally important information loss mechanism for Majorana qubits is quasiparticle poisoning [27,32,33,71–73].

In this work, we focused on the case of low disorder, in the hope that material growth and device fabrication advances will convey qubit experiments in that parameter range. Current devices might have much stronger disorder [74–76], and it is an interesting extension of our work to study how Majorana qubit dephasing occurs in the presence of strong disorder. A further natural extension of our work is to step-by-step move from the Kitaev chain minimal model to more realistic real-space models, e.g., from 1D Rashba wire [1,2,77] to 3D Schrödinger-Poisson models [76,78,79], and beyond.

One of the key result of the paper is that the dephasing susceptibility oscillates as a function of system parameters out-of-phase with respect to the oscillations of the clean splitting. This is shown in Fig. 2(b). How robust is this result upon relaxing the simple hard-wall boundary condition leading to the result in Fig. 2(b)? We have performed numerical simulations exploring this question, by extending our model in two ways: (1) We have relaxed the hard-wall boundary condition

to a confinement potential that has a steplike dependence at the two edges of the 1D topological superconductor, and (2) we have added a homogeneous electric field, that is, a chemical potential that varies linearly with position. In the parameter range we studied, the two quantities were following the same type of out-of-phase oscillations as shown in Fig. 2(b). We see it as an interesting follow-up question to understand this robustness.

VII. CONCLUSIONS

We have studied the Majorana splitting of the disordered topological Kitaev chain, serving as a minimal model of dephasing of Majorana qubits. Focusing on the case of spatially uncorrelated disorder, we characterized the Gaussian probability distributions of the signful splitting, using numerics as well as simple semianalytical and approximate analytical techniques. We established a Gaussian decay envelope for the dephasing curve, as a consequence of the Gaussian distribution of the signful splitting. We have found that the standard deviation of the signful splitting, and hence the dephasing rate, oscillates as a function of system parameters out-of-phase with respect to the oscillations of the clean splitting. We have also pointed out the absence of dephasing sweet spots in the case of spatially uncorrelated quasistatic disorder. Furthermore, we have described how Majorana qubit dephasing changes as a function of disorder correlation length, and argued that dephasing measurements can be used to characterize the disorder correlation length. We expect that our results will be used in the design and interpretation of future experiments, aiming to demonstrate topologically protected quantum memory, quantum dynamics, or quantum computing, based on Majorana zero modes.

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APPENDIX A: CONNECTING THE CONTINUUM MODEL WITH THE KITAEV CHAIN

The Kitaev chain [Eq. (1)] is a discretized version of the continuum model [Eq. (21)], and vice versa, the continuum model can be obtained from the Kitaev chain via the envelope function approximation. The relation of the two models is outlined in Ref. [62], but for the sake of self-containedness, we describe it here in detail. We match the parameters of the two models via matching their dispersion relations as shown in Fig. 6.

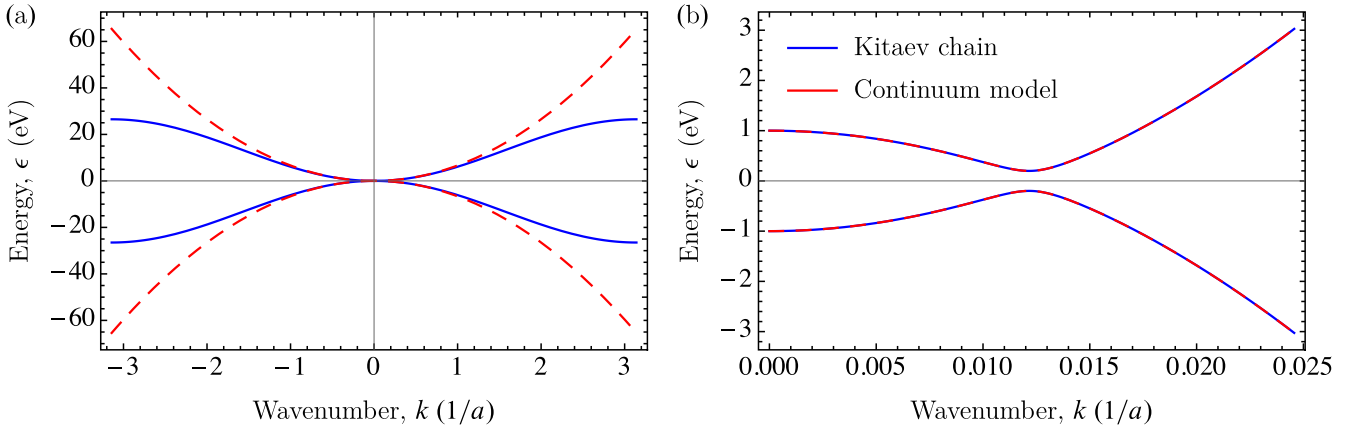


FIG. 6. The Kitaev chain band structure (blue) and the continuum model band structure (red), (a) over the 1D Brillouin zone, (b) in the vicinity of zero energy, with parameters chosen such that the band structures match each other in the vicinity of zero energy. See Table I for parameter values.

To match the two models, we recall the BdG Hamiltonian of the Kitaev chain in momentum space, which reads

$$\mathcal{H}_K(k) = [-2t \cos(ka) - \mu_K] \sigma_z + 2\Delta_K \sin(ka) \sigma_y. \quad (\text{A1})$$

Here we note that the momentum-space superconducting term $2\Delta_K \sin(ka) \sigma_y$ of the Kitaev chain is proportional to σ_y , whereas the corresponding term $\Delta'_C \hbar k \sigma_x$ of the continuum model [Eq. (2)] is proportional to σ_x . This difference is irrelevant and can be transformed away with a unitary transformation in Nambu space, since the rest of both Hamiltonians is proportional to σ_z .

Matching the continuum model and the Kitaev chain model is based on the following criteria:

(1) The lengths in the two models are naturally matched as $L = Na$, where L is the length of the wire and N is the number of sites in the lattice model.

(2) In the absence of superconductivity, the effective mass in the vicinity of $k = 0$ has to be the same in the two models, which yields the condition

$$m = \frac{\hbar^2}{2ta^2}. \quad (\text{A2a})$$

(3) In the absence of the superconducting terms, the minima of the bulk spectra have to be at the same energy. This is achieved by adjusting the chemical potentials in the following way:

$$\mu_C = 2t + \mu_K. \quad (\text{A2b})$$

(4) The low-energy (close to zero energy) spectra of the two models will be similar if the minimum of the bulk normal band is just slightly below zero energy; formally this can be written as

$$0 < 1 + \frac{\mu_K}{2t} \ll 1. \quad (\text{A2c})$$

(5) In the presence of superconductivity, the superconducting gaps have to be equal, a condition approximately satisfied by the identification

$$\Delta'_C = \frac{\Delta_K a}{\hbar} \sqrt{2 - \frac{\mu_K}{t}}. \quad (\text{A2d})$$

We note that here we have already assumed that Eqs. (A2a) and (A2b) are fulfilled. Equation (A2d) is an approximation in the sense that we match energy gaps of the two models that are opened at k_F , i.e., at the wave number where the band touches zero in the absence of the superconductivity. The actual gap (i.e., the energy difference minimized over the wave number) is in general located at a slightly different wave number k_0 , but in the limit of Eq. (A2c), $k_0 \approx k_F$.

Based on the above criteria, we choose the parameter values listed in Table I to compare the results of the Kitaev chain and the continuum model.

The energy dispersion of the Kitaev chain (blue solid) and that of the continuum model (red dashed) are compared over the 1D Brillouin zone in Fig. 6(a), and in the vicinity of the Brillouin zone center and the Fermi wave number in Fig. 6(b).

Below, we will need the following relations between the parameters of the two models:

$$k_F = \frac{\sqrt{2 + \frac{\mu_K}{t}}}{a}, \quad (\text{A3a})$$

$$\Delta_C = \Delta_K \sqrt{4 - \left(\frac{\mu_K}{t}\right)^2}, \quad (\text{A3b})$$

$$\xi = \frac{2ta}{\Delta_K \sqrt{2 - \frac{\mu_K}{t}}}. \quad (\text{A3c})$$

We obtain Eq. (A3a) from Eq. (3a) by substituting Eqs. (A2a) and (A2b). We get Eq. (A3b) from Eq. (3c) by substituting Eqs. (A2d) and (A3a). We obtain Eq. (A3c) from Eq. (3d) by combining Eqs. (3b), (A2a), (3a), and (3c).

APPENDIX B: INFERRING THE STANDARD DEVIATION OF THE SIGNFUL SPLITTING FROM SAMPLES OF THE SPLITTING

Figure 2(b) shows the standard deviation of σ_{ϵ_0} of the signful splitting ϵ_0 of a Kitaev chain due to disorder. How did we compute σ_{ϵ_0} ? The smallest non-negative eigenvalue of the BdG matrix is the *absolute value* of the signful splitting; hence its standard deviation taken over many disorder realizations

does not provide σ_{ϵ_0} . Here, we provide an indirect way to compute σ_{ϵ_0} by assuming that the signful splitting is normally distributed, an assumption in accordance with our result (26). Under that assumption, the absolute value of the signful splitting has a *folded normal distribution*. We have an easy access to samples of the splitting by using BdG Hamiltonian, and by following the procedure outlined below, we are able to compute σ_{ϵ_0} from samples of the splitting.

Let us use a general notation for easier readability. The probability density function of the normal random variable X , representing the signful splitting, reads

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}}, \quad (\text{B1})$$

where m is the mean and σ is the standard deviation of X . The probability density function of the random variable $|X|$, representing the splitting, is

$$f_{|X|}(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}} \left(1 + e^{\frac{2mx}{\sigma^2}}\right), \quad (\text{B2})$$

which is often called a folded normal distribution.

We estimate the parameters m and σ (mean and standard deviation of signful splitting) from a sample $\{x_i | i = 1, \dots, n\}$ of $|X|$ (the splitting). Here n is the size of the sample. Our estimation is based on the maximum likelihood estimation procedure. The log-likelihood of the distribution estimated from the sample $\{x_i\}$ can be written as

$$\begin{aligned} l(\{x_i\}; m, \sigma) &= \ln \left[\prod_{i=1}^n f_{|X|}(x_i) \right] = \\ &= -\frac{n}{2} \ln(2\pi\sigma^2) - \sum_{i=1}^n \frac{(x_i - m)^2}{2\sigma^2} \\ &+ \sum_{i=1}^n \ln \left(1 + e^{\frac{2mx_i}{\sigma^2}}\right). \end{aligned} \quad (\text{B3})$$

To estimate the value of m and σ , we need to find the maximum point of the likelihood function; hence we take $\partial_m l(\{x_i\}; m, \sigma) = 0$ and $\partial_\sigma l(\{x_i\}; m, \sigma) = 0$, which lead to

$$m = \frac{1}{n} \sum_{i=1}^n x_i \tanh\left(\frac{mx_i}{\sigma}\right), \quad (\text{B4a})$$

$$\sigma^2 = m^2 + \frac{1}{n} \sum_{i=1}^n x_i^2 - \frac{2m}{n} \sum_{i=1}^n x_i \tanh\left(\frac{mx_i}{\sigma}\right). \quad (\text{B4b})$$

From Eqs. (B4a) and (B4b), we get

$$\sigma^2 = \left(\frac{1}{n} \sum_{i=1}^n x_i^2\right) - m^2. \quad (\text{B5})$$

In general, the coupled Eqs. (B4a) and (B4b) have to be solved. However, in our case, to determine the standard deviation of the signful splitting, Eq. (B5), it is sufficient as know the square of the signful splitting mean m : it is equal to the square of the splitting of the clean system $\epsilon_{0,c}$. This implies

the formula

$$\sigma_{\epsilon_0} = \sqrt{\left(\frac{1}{n} \sum_{i=1}^n \epsilon_{0,i}^2\right) - \epsilon_{0,c}^2}, \quad (\text{B6})$$

where the $\epsilon_{0,i}$'s are splittings in disordered realizations and $\epsilon_{0,c}$ is the splitting for the clean system. We used this result to compute the data in Fig. 2(b).

APPENDIX C: COMPARISON WITH THE RESULTS OF BROUWER *ET AL.* [29]

In the main text, we predict a normal distribution for signful splitting ϵ_0 . On the other hand, the key result of Ref. [29] is that the *splitting envelope* $\epsilon_{0,\max}$ [for clarification, see their Fig. 1(c)] has a log-normal distribution. Although the two quantities (signful splitting and splitting envelope) are not the same, they are in fact interrelated. In this Appendix, we identify a parameter range where both our results and the results of Ref. [29] are valid, and establish the relation of these results. Our comparison suggests that the two unknown constants appearing in the analytical results of Ref. [29] (C_m and C_v ; see below) are actually zero.

The main result of Ref. [29] is as follows. The quantity $\ln(\epsilon_{0,\max}/2\Delta_C)$ has a normal distribution with mean and variance given by their Eq. (16), that is,

$$\langle \ln(\epsilon_{0,\max}/2\Delta_C) \rangle = -L[1/\xi - 1/2l] + C_m, \quad (\text{C1a})$$

$$\text{var} \ln(\epsilon_{0,\max}/2\Delta_C) = L/2l + C_v. \quad (\text{C1b})$$

Here, C_m and C_v are the unknown constants, that is, unknown order-of-unity corrections independent of L , l , and ξ . (Even though these constants are not displayed in Eq. (16) of Ref. [29], they are introduced in the text following that equation.) Furthermore, $l = \hbar^2 v_F^2 / \gamma$ is the mean free path, where γ corresponds to the disorder strength in their model, which is identified with our model as $\gamma = a_{\text{dis}} \sigma_\mu^2$.

Their results stand if the following conditions are satisfied:

$$1/k_F \ll \xi, \quad (\text{C2a})$$

$$\xi < 2l, \quad (\text{C2b})$$

$$\epsilon_{0,\max} \ll \min(\Delta_C, \hbar/\tau), \quad (\text{C2c})$$

where $\tau = \hbar v_F / l$. On the other hand, our result for the clean splitting (12) is valid if $L \gg \xi$, and our result for the dephasing susceptibility to disorder (29) is valid if $L \gg \xi$ and if disorder is weak.

First, we assume that the parameter ranges of validity of the two results have some overlap, and show that in such a common parameter range, the two results are consistent. Second, we provide an example for the common parameter range where both results should be valid and hence should be consistent with each other.

To show the consistency of the two results, we suppose that

$$l \gg \xi, \quad (\text{C3a})$$

$$L/2l \ll 1, \quad (\text{C3b})$$

$$C_m = 0, \quad (\text{C3c})$$

$$C_v = 0. \quad (\text{C3d})$$

Equation (C3) stands for weak disorder, whereas Eq. (C3 b) together with Eq. (C3 d) provides that $\ln(\varepsilon_{0,\max}/2\Delta_C)$ has a standard deviation much smaller than 1. Furthermore, the choice of C_m and C_v in Eqs. (C3) and (C3 d) is required to match the result of Ref. [29] with our results.

Our results, together with Eqs. (C3), imply that the splitting envelope $\varepsilon_{0,\max}$ approximately follows a normal distribution with mean and standard deviation as follows:

$$\langle \varepsilon_{0,\max} \rangle = 2\Delta_C e^{-L/\xi}, \quad (\text{C4a})$$

$$\sigma_{\varepsilon_{0,\max}} = \Delta_C \sqrt{\frac{2L}{l}} e^{-L/\xi} = \sigma_\mu \frac{\sqrt{2La_{\text{dis}}}}{\xi} e^{-L/\xi}. \quad (\text{C4b})$$

We obtained Eq. (C4) from Eq. (12) by taking the limit $k_F \ll 1/\xi$ and by omitting the sinusoidal oscillatory part in the latter. Furthermore, we obtained Eq. (C4 b) from Eq. (29) by taking the limit $L \gg \xi$, and by substituting the cosine term with -1 . The latter substitution is needed because the disorder-induced standard deviation of the splitting has a local minimum whenever the clean splitting has a local maximum [see Fig. 3(a)].

The key mathematical statement we use to show the consistency of Eq. (C1) and Eq. (C4) is the following: If X is a log-normal random variable such that $\ln X$ is a normal random variable with mean μ and standard deviation σ [that is, $\ln X \sim \mathcal{N}(\mu, \sigma)$], and the standard deviation fulfills $\sigma \ll 1$, then X is approximately a normal random variable with mean e^μ and standard deviation $e^\mu \sigma$ [that is, $X \sim \mathcal{N}(e^\mu, e^\mu \sigma)$]. This follows from the fact that the exponential function can be well approximated around any point by its linear series expansion in a sufficiently small environment of the point. We apply this approximation to Eq. (C1) using the assumptions of Eq. (C3). This procedure yields Eq. (C4), implying that our result is consistent with the earlier result.

Finally, we provide an example for the common parameter range where both results are valid. Equation (C2) is satisfied for the parameter set in Table I. In the weak-disorder limit, Eq. (C2 b) is fulfilled. For weak disorder, $\hbar/\tau \ll \Delta_C$; furthermore, using Eq. (C4), the condition $\varepsilon_{0,\max} \ll \hbar/\tau$ is

equivalent to the condition

$$\ln(2l/\xi) \ll L/\xi. \quad (\text{C5})$$

In addition, Eq. (C5) and Eq. (C3 d) can be combined as

$$\ln(2l/\xi) \ll L/\xi \ll 2l/\xi. \quad (\text{C6})$$

For weak disorder, there is a finite interval for the system length L where Eq. (C6) is fulfilled. For example, for the parameter values given in Table I, and for disorder strength $\sigma_\mu = 10 \mu\text{eV}$, Eq. (C6) is evaluated as

$$14\,500 \ll L/a \ll 4.32 \times 10^{10}. \quad (\text{C7})$$

Note that our numerical results shown in the main text correspond to system lengths that are one order of magnitude smaller than the lower end of this interval.

To conclude, we have established the consistency between the earlier analytical results of Ref. [29] for the statistics of the splitting envelope, and our analytical results for the statistics of the signful splitting described in the main text. To ensure this consistency, we had to assume that the order-of-unity constant offset parameters C_m and C_v , which were not calculated in Ref. [29], are actually zero. This indirect determination of the offset parameters is a useful by-product of the comparison.

APPENDIX D: CORRELATED DISORDER

In Sec. V, we study the effect of the spatial correlations of the disorder on dephasing. To determine the dephasing susceptibility of the disordered Kitaev chain, we have to generate numerous spatially correlated disorder realizations. In this Appendix, we show a method allowing us to do that in an efficient way.

The vector of the on-site energies $\delta\mu_K$ is an N -dimensional random variable vector described by a multivariate normal distribution, i.e., $\delta\mu_K \sim \mathcal{N}(\mathbf{0}, \Sigma)$, where

$$[\Sigma]_{ij} = \sigma_\mu^2 e^{-|i-j|a/\zeta} \quad (\text{D1})$$

is the covariance matrix. Each component of $\delta\mu_K$ has zero mean and standard deviation σ_μ ; furthermore the length scale of the correlations is ζ .

The Cholesky decomposition of Σ has the form

$$\Sigma = \mathbf{L}\mathbf{L}^\top, \quad (\text{D2})$$

where \mathbf{L} is a lower triangular matrix. We note that Σ is a real-valued symmetric positive-definite matrix. Let \mathbf{Z} be an N -dimensional standard normal random vector. All components of \mathbf{Z} are independent and each is a zero-mean unit-variance normally distributed random variable. It is straightforward to see that $\delta\mu_K = \mathbf{L}\mathbf{Z}$ follows the desired distribution with the covariance matrix described in Eq. (D1). Thus to generate correlated random samples of on-site disorder, one can first generate uncorrelated samples (according to \mathbf{Z}), and then multiply them by the matrix \mathbf{L} .

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