

Quantum speedup of classical mixing processes

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Most approximation algorithms for #P-complete problems (e.g., evaluating the partition function of a monomer-dimer or ferromagnetic Ising system) work by reduction to the problem of approximate sampling from a distribution π over a large set \mathcal{S} . This problem is solved using the *Markov chain Monte Carlo* method: a sparse, reversible Markov chain P on \mathcal{S} with stationary distribution π is run to near equilibrium. The running time of this random walk algorithm, the so-called *mixing time* of P , is $O(\delta^{-1} \log 1/\pi_*)$ as shown by Aldous, where δ is the spectral gap of P and π_* is the minimum value of π . A natural question is whether a speedup of this classical method to $O(\sqrt{\delta^{-1}} \log 1/\pi_*)$ is possible using *quantum walks*. We provide evidence for this possibility using quantum walks that *decohere* under repeated randomized measurements. We show that (i) decoherent quantum walks always mix, just like their classical counterparts, (ii) the mixing time is a robust quantity, essentially invariant under any smooth form of decoherence, and (iii) the mixing time of the decoherent quantum walk on a periodic lattice \mathbb{Z}_n^d is $O(nd \log d)$, which is indeed $O(\sqrt{\delta^{-1}} \log 1/\pi_*)$ and is asymptotically no worse than the diameter of \mathbb{Z}_n^d (the obvious lower bound) up to at most a logarithmic factor.

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I. INTRODUCTION

A. Markov chain Monte Carlo and quantum walks

A rich theory has been developed for computing approximate solutions to problems in combinatorial enumeration and statistical physics which are #P-complete and therefore unlikely to have efficiently computable exact solutions. Among the highlights are randomized polynomial-time algorithms for approximating the permanent of a non-negative matrix [1], the volume of a convex polytope [2], and the partition functions of monomer-dimer and ferromagnetic Ising systems [3,4]. At the heart of these algorithms, and consuming much of the running time, is a subroutine for approximate sampling from a distribution π over a large set \mathcal{S} of states. This problem is solved using the *Markov chain Monte Carlo* (MCMC) method: a *random walk* $(x_t)_{t=0,1,2,\dots}$ on \mathcal{S} (i.e., a random sequence of state transitions) is simulated from an arbitrary initial state x_0 to a random state x_T distributed close to π . The random walk is generated by a *Markov chain* P on \mathcal{S} (i.e., an $|\mathcal{S}| \times |\mathcal{S}|$ column-stochastic matrix of state transition probabilities) whose *stationary* (fixed-point) distribution is π . The time required to guarantee closeness to π , or *mixing*, is the so-called *mixing time* τ_{mix} . Bounding τ_{mix} reduces to estimating the *spectral gap* $\delta := 1 - |\lambda|$, where λ is the second-largest eigenvalue of P in magnitude, by Aldous' inequality [5],

$$\delta^{-1} \leq \tau_{\text{mix}} \leq \delta^{-1} (\ln 1/\pi_*), \quad (1)$$

where $\pi_* := \min_x \pi(x) > 0$.

Example 1: Monomer-dimer systems. Suppose we are given as input a graph G , consisting of vertices V and edges E , and a positive value λ . Let \mathcal{S}_k be the set of all k -matchings in G , where a *k-matching* is a k -subset of edges which are vertex-disjoint. A pair of vertices connected by an edge in the

matching is a *dimer*; an isolated vertex is a *monomer*. Define the *Gibbs distribution*

$$\pi(x) := \frac{1}{Z(\lambda)} \lambda^{|x|} \quad (2)$$

on the set $x \in \mathcal{S} := \cup_k \mathcal{S}_k$, where $Z(\lambda) = \sum_{x \in \mathcal{S}} \lambda^{|x|} = \sum_k |\mathcal{S}_k| \lambda^k$ is the *partition function*. Our computational task is to evaluate $Z(\lambda)$ as accurately as possible.

We cannot hope to compute $Z(\lambda)$ exactly, because this problem is #P-complete for any $\lambda > 0$. However, consider that we can write $Z(\lambda)$ as a telescoping product,

$$Z(\lambda) = \frac{Z(\lambda_r)}{Z(\lambda_{r-1})} \times \frac{Z(\lambda_{r-1})}{Z(\lambda_{r-2})} \times \dots \times \frac{Z(\lambda_1)}{Z(\lambda_0)} \times Z(\lambda_0) \quad (3)$$

for some sequence $\lambda = \lambda_r > \lambda_{r-1} > \dots > \lambda_1 > \lambda_0 = 0$, where $Z(\lambda_0) = 1$ is trivial to compute and the ratios $Z(\lambda_{i+1})/Z(\lambda_i)$ are easily seen to equal the expectation values $E_{x \leftarrow \pi_i}[(\lambda_{i+1}/\lambda_i)^{|x|}]$. (Here π_i is the Gibbs distribution for λ_i .) It can be shown that r need not be very large to guarantee that the expectation values $E_{x \leftarrow \pi_i}[(\lambda_{i+1}/\lambda_i)^{|x|}]$ are bounded and well-estimated using relatively few samples from the distributions π_i .

Thus we have reduced the problem of approximating $Z(\lambda)$ efficiently to the problem of sampling efficiently from the Gibbs distribution π . The sampling problem can be solved by running a random walk $(x_t)_{t=0,1,2,\dots}$ on the Markov chain P with the following transition rule: at state $x_t \in \mathcal{S}_k$, choose an edge $e = (u, v) \in E$ uniformly at random. If e is not in x_t and $x' := x_t \cup \{e\}$ is a valid matching, set $x_{t+1} := x' \in \mathcal{S}_{k+1}$; if e is in x_t , set $x_{t+1} := x \setminus \{e\} \in \mathcal{S}_{k-1}$ with probability $1/\lambda$. Otherwise, either move to a "nearby" state $x_{t+1} \in \mathcal{S}_k$ (i.e., set x_{t+1} so that $|x_{t+1} \cap x_t| = k - 1$) or simply stay put (i.e., set $x_{t+1} = x_t$). It is not too hard to see that P has π as its stationary distribution; moreover, its spectral gap is $\Omega(|V||E|)$ [3,6]. An upper bound on the mixing time follows from Aldous' inequality (1).

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Notice that Aldous' inequality (1) is tight with respect to the spectral gap. However, it is also somewhat unsatisfactory in that although the *diameter* (i.e., the maximum distance between any two vertices) of the graph underlying P —the obvious lower bound for sampling from its stationary distribution—scales like $\sqrt{\delta^{-1}}$, the dependence is δ^{-1} in Aldous' inequality due to the Gaussian-like spreading behavior of random walks. In MCMC algorithms like the monomer-dimer example above, the additional factor of $\sqrt{\delta^{-1}}$ can increase the running time significantly. Removing this factor would imply considerable improvement of such algorithms in both the *known* upper bounds (since existing estimates of δ could be used in conjunction with a sharper inequality than Aldous') and the *true* upper bounds (since Aldous' inequality is tight, and thus a sharper inequality could only come from a faster approximate sampling method). Thus a natural question is whether there is a way to modify the standard MCMC method to obtain a speedup to $O(\sqrt{\delta^{-1}} \log 1/\pi_*)$. This seems unlikely using classical randomized methods: Chen, Lovasz, and Pak [7] have shown that we can sometimes speed up mixing by *lifting* (essentially, by locally “re-routing”) a Markov chain, but this requires both knowledge of the chain's global structure and its use in solving an *NP*-hard flow problem to find good routes. However, although lifting the chain seems unlikely to be practical, an idea that might work is *quantizing* the chain.

Why might a quantized Markov chain, or *quantum walk*, help us reach π more quickly? Two reasons are (i) a quantum walk is as simple to realize as its classical counterpart (i.e., it is computable locally and online at each step of the walk, unlike a classical lifting of the chain), and (ii) there is empirical (and some theoretical) evidence that quantum walks tend to propagate and “spread” probability mass across \mathcal{S} in time on the order of the diameter on precisely the same low-dimensional graphs that trip up their classical counterparts. Based on these observations, the possibility of obtaining a quantum speedup for the mixing problem has been pursued by Nayak *et al.* [8,9], Aharonov *et al.* [10], Moore and Russell [11], Gerhardt and Watrous [12], and Richter [13].

We remark that a quantum speedup theorem of the sort we seek has already been proven for the *hitting* problem, in which we search (rather than sample from) the states of a Markov chain: Szegedy [14] proved a quadratic quantum speedup for the *hitting time* (i.e., the time to detect the presence of a “marked” state) of any symmetric Markov chain, generalizing considerably the celebrated search algorithm of Grover [15] and implying quantum speedups for structured search problems such as element distinctness [16], triangle finding [17], matrix product verification [18], and group commutativity testing [19]. It is this success which inspires us to investigate the possibility of a quantum speedup for the mixing problem with the goal of transferring the speedup to MCMC algorithms.

B. Our contributions

We present evidence of a possible quantum MCMC speedup to $O(\sqrt{\delta^{-1}} \log 1/\pi_*)$ using quantum walks that *decohere* under repeated randomized measurements. Decoher-

ence (in small amounts) was first identified as a way to improve spreading and mixing properties in numerical experiments performed by Kendon and Tregenna [20] and in analytical estimates by Fedichkin *et al.* [21–23]. On the other hand, high rates of decoherence in quantum walks have been shown to degrade mixing properties substantially by the quantum Zeno effect [24]. For an excellent survey of these and other aspects of decoherent quantum walks, see Kendon [25].

Our technical contributions are as follows: First, we show that for any symmetric Markov chain P , we can generate an arbitrarily good approximation to the uniform stationary distribution π of P by subjecting the continuous-time quantum walk $U_{ct}(P) = \exp(iP)$ to reasonably “smooth” decoherence. Thus decoherent quantum walks (which are nonunitary) offer a way of circumventing an obstacle first identified by Aharonov *et al.* [10], who observed that *unitary* quantum walks often converge (in the time-averaged sense) to highly non-uniform distributions.

Second, we show that the optimal mixing time of a decoherent quantum walk is a robust quantity, in that it remains essentially invariant under any sufficiently smooth form of decoherence. In particular, decoherent quantum walks undergoing repeated Cesaro-averaged [10–12] and Bernoulli-Poisson-averaged [20,24] measurements are nearly equivalent in mixing efficiency. The proof applies more generally to a game involving time-dependent Markov chains (not necessarily describing quantum phenomena) and may be of independent interest.

Third, we prove a theorem on threshold mixing of quantum walks on (Cartesian) graph powers in order to show that the decoherent continuous-time quantum walk $U_{ct}(P(G))$ on a periodic lattice $G = \mathbb{Z}_n^d$ [where $P(G)$ denotes the standard “simple random walk” Markov chain on G] can be used to generate a good approximation to the uniform stationary distribution π of P in time $O(nd \log d)$. This upper bound is asymptotically no worse than the diameter of \mathbb{Z}_n^d (the obvious lower bound) up to at most a logarithmic factor and is $O(\sqrt{\delta^{-1}} \log 1/\pi_*)$ for both high-dimensional *and* low-dimensional lattices (unlike its classical counterpart). For $d = 1$, this proves a conjecture of Kendon and Tregenna [20] based on numerical experiments [26]. For $d \geq 1$, it extends the results of Fedichkin *et al.* [21–23] by confirming $O(n)$ and $O(d \log d)$ scaling (suggested by their analytical estimates and numerical experiments in regimes of both high and low decoherence) of the fastest-mixing walk, which they conjectured to be decoherent rather than unitary. We briefly discuss the prospects for extending this result to the discrete-time Grover walk $U_{dt}(P(\mathbb{Z}_n^d))$ [27–29].

Previously, mixing speedups had been proven only for the unitary quantum walks of Nayak *et al.* [8,9] and Aharonov *et al.* [10] (on the cycle) and of Moore and Russell [11] (on the hypercube). Thus our work shows that introducing a small amount of decoherence to a quantum walk can simultaneously force convergence to the uniform distribution while preserving a quantum mixing speedup, an advantageous combination for algorithmic applications [30].

II. PRELIMINARIES

A. Markov chains

Let P be a *Markov chain* (column-stochastic matrix) on the set \mathcal{S} ($|\mathcal{S}|=N$) which is *irreducible* (strongly connected); then it has a unique distribution π which is *stationary* (i.e., satisfies $P\pi=\pi$). Moreover, π is strictly positive: $\pi_* := \min_x \pi(x)$ satisfies $\pi_* > 0$. In particular, if P is symmetric, then π is the uniform distribution u (the N -dimensional column vector with each component equal to $1/N$). If G is an undirected graph, we denote the standard (“simple random walk”) Markov chain on G by $P(G)$.

A Markov chain P which is both irreducible and *aperiodic* (nonbipartite) is by definition *ergodic* and satisfies

$$P^t \rightarrow \pi 1^\dagger = [\pi \cdots \pi] \quad \text{as } t \rightarrow \infty, \quad (4)$$

where 1^\dagger is the N -dimensional row vector with each component equal to 1. We can thus define the (*threshold*) *mixing time*

$$\tau_{\text{mix}} := \min \left\{ T: \frac{1}{2} \|P^t - \pi 1^\dagger\|_1 \leq \frac{1}{2e} \quad \forall t \geq T \right\}, \quad (5)$$

where $\|\cdot\|_1$ is the matrix 1-norm. The mixing is *perfect* if $P^t = \pi 1^\dagger$. Let $\delta := 1 - |\lambda|$ be the *spectral gap* of P , where λ is the second-largest eigenvalue of P in magnitude. We say that P is *reversible* if the matrix DPD^{-1} is symmetric, where D is the diagonal matrix $D(x,x) = \sqrt{\pi(x)}$. A precise statement of Aldous’ inequality (1) is as follows.

Theorem 2 (Aldous [5]). Let P be a reversible, ergodic Markov chain with stationary distribution π and spectral gap δ . Then its mixing time satisfies

$$\delta^{-1} \leq \tau_{\text{mix}} \leq \delta^{-1} (\ln 1/\pi_*).$$

We will also use the *maximum pairwise column distance* $d(P) := \max_{x,x'} \frac{1}{2} \|P(\cdot,x) - P(\cdot,x')\|_1$ to estimate the mixing time. It is related to the matrix 1-norm distance by the inequality:

$$\frac{1}{2} \|P - \pi 1^\dagger\|_1 \leq d(P) \leq \|P - \pi 1^\dagger\|_1. \quad (6)$$

The following propositions (see [13] for proofs) can be used to estimate the mixing time of P given a common lower bound on most of the entries in each column.

Proposition 3. If $d(P) \leq \alpha$, then $\tau_{\text{mix}} \leq \lceil \log_{1/\alpha} 2e \rceil$.

Proposition 4. If at least βN entries in each column of P are bounded below by γ/N , where $\beta > \frac{1}{2}$ and $\gamma > 0$, then $d(P) \leq 1 - \gamma[1 - 2(1 - \beta)]$.

B. Quantum walks

Henceforth, a *quantum walk* is a pair $\langle U, \omega_T \rangle$ where the *transition rule* U is a unitary operator acting on a finite-dimensional Hilbert space and the *measurement rule* ω_T is a T -parametrized family of probability density functions on $[0, \infty)$ characterizing the (random) time at which a total measurement is performed on the Hilbert space [31].

The unitary transition rule determines the orbit of a pure quantum state, or *wave function* (l_2 -normalized complex vec-

tor), just as a Markov chain (or stochastic transition rule) determines the orbit of a classical distribution (l_1 -normalized non-negative vector). Let P be the Markov chain used in an MCMC algorithm; in particular, P is reversible. Two natural quantizations of P are (i) the $S \times S$ unitary *continuous-time walk* given by $U_{ct}(P) = \exp(iDPD^{-1})$ [32,33], where $H = DPD^{-1}$ is the time-independent *Hamiltonian*, and (ii) the $S^2 \times S^2$ unitary discrete-time *Grover-Szegedy walk* given by $U_{dt}(P) = (RS)^2$ [14], where S is the involution $\sum_{x,y \in \mathcal{S}} |x,y\rangle \mapsto |y,x\rangle$, R is the reflection $\sum_{x \in \mathcal{S}} |x\rangle \langle x| \otimes (2|p_x\rangle \langle p_x| - I)$, and $|p_x\rangle$ is the vector $\sum_{y \in \mathcal{S}} \sqrt{P(y,x)} |y\rangle$. The quantization $U_{ct}(P)$ was used by Childs *et al.* [34] to solve in polynomial time a natural oracle problem for which no classical polynomial-time algorithm exists. The quantization $U_{dt}(P)$ was used by Szegedy [14] to prove a quadratic quantum speedup for the hitting time of any symmetric Markov chain.

Measurement collapses the wave function to a classical distribution according to the map $|\phi\rangle \mapsto \sum_x |x\rangle \langle x| \phi|^2$. For a quantum walk $\langle U, \omega_T \rangle$, the Markov chain *generated* by the quantum walk is given by $\hat{P}_T(y,x) := E_{t \sim \omega_T} [\langle y|U^t|x\rangle|^2]$, where E denotes expected value. We say that the quantum walk *threshold mixes* if the Markov chain it generates mixes in time $O(1)$. Examples of measurement rules from the literature include the point distribution $\delta_T(t) := \delta(t-T)$ where δ is the delta function [8,9], the uniform distribution $\bar{\mu}_T := \frac{1}{T} \chi_{[0,T]}$ and its discrete-time counterpart $\bar{\nu}_T := \frac{1}{T} \chi_{[0,T-1]}$ where χ is the characteristic function [10], the exponential distribution $\tilde{\mu}_T(t) := \frac{1}{T} \exp(-t/T)$, and the geometric distribution $\tilde{\nu}_T(t) := \frac{1}{T} (1 - \frac{1}{T})^t$. The exponential and geometric distributions are *memoryless* and describe the interarrival time between measurements in a Poisson process with measurements occurring at rate $\lambda = 1/T$ and a Bernoulli process with measurements occurring with probability $p = 1/T$ at each time step, respectively. These processes coincide with the decoherence models of Alagic and Russell [24] and Kendon and Tregenna [20], respectively.

Nayak *et al.* [8,9] and Aharonov *et al.* [10] showed that the so-called *Hadamard walks* $\langle U_{Had}, \delta_T \rangle$ and $\langle U_{Had}, \bar{\nu}_T \rangle$ on the cycle Z_n threshold-mix quadratically faster than the classical random walk, although their definitions of threshold mixing are slightly different than ours. Moore and Russell [11] showed that the continuous-time walk $\langle U_{ct}(P(G)), \delta_T \rangle$ and the Grover walk $\langle U_{dt}(P(G)), \delta_T \rangle$ on the hypercube $G = Z_2^d$ mix perfectly and almost perfectly, respectively, in time $T = O(d)$. They also showed that the continuous-time quantum walk on the hypercube with measurement $\bar{\mu}_T$ does not mix to the uniform stationary distribution π of $P(Z_2^d)$ in the limit $T \rightarrow \infty$. Gerhardt and Watrous [12] showed the same for a continuous-time quantum walk on the symmetric group with measurement $\bar{\mu}_T$. See the survey by Kempe [35] for further results on quantum walks.

We remark that there is a nice way to use the quantization $U_{dt}(P)$ to solve the mixing problem in time $O(1/\sqrt{\delta\pi_*})$ which, although prohibitively costly in π_* , exhibits the desired dependence on δ . Consider the stationary eigenstate $|\tilde{\pi}\rangle := \sum_{x \in \mathcal{S}} \sqrt{\pi(x)} |x\rangle |p_x\rangle$ of $U_{dt}(P)$. It is clear that we can retrieve a good approximation to the classical distribution π by

generating and then measuring a good approximation to $|\tilde{\pi}\rangle$ [36]. Magniez *et al.* [37] observe that $O(1/\sqrt{\delta})$ steps of phase estimation on $U_{dt}(P)$ enable us to reflect about $|\tilde{\pi}\rangle$. By alternating this reflection with a reflection about $|\tilde{z}\rangle := |z\rangle|p_z\rangle$ where z is the initial walk state (in particular, $\langle \tilde{z} | \tilde{\pi} \rangle \geq \sqrt{\pi_*}$), we can generate $|\tilde{\pi}\rangle$ from $|\tilde{z}\rangle$ in time $O(1/\sqrt{\delta\pi_*})$. In fact, this algorithm is described by Magniez *et al.* [37] as a *hitting* algorithm (i.e., generate the *unknown* state $|\tilde{z}\rangle$ from the fixed initial state $|\tilde{\pi}\rangle$); the idea of running a quantum hitting algorithm *in reverse* as a mixing algorithm was suggested by Childs [38].

III. MIXING PROPERTIES OF DECOHERENT QUANTUM WALKS

A. Two types of convergence

Let \hat{P}_T be the Markov chain generated by a quantum walk $\langle U, \omega_T \rangle$. Then repeating the quantum walk T' times in succession generates the Markov chain $(\hat{P}_T)^{T'}$. The following lemma (a variant of Theorem 3.4 in Aharonov *et al.* [10]) and theorem describe the asymptotic behavior of \hat{P}_T and $(\hat{P}_T)^{T'}$ in the limits $T \rightarrow \infty$ and $T' \rightarrow \infty$, respectively. For concreteness we will take $U = U_{ct}$ in this subsection and the next; it is a simple exercise to extend the results to discrete-time walk variants. Although stated explicitly for quantum walks, the results apply to any time-independent quantum dynamics on a finite-dimensional Hilbert space subjected to random destructive measurements.

Lemma 5 (The limit $T \rightarrow \infty$). Let P be a symmetric Markov chain and ω_T be a family of distributions satisfying $E_{t \leftarrow \omega_T}[e^{i\theta t}] \rightarrow 0$ as $T \rightarrow \infty$ for any $\theta \neq 0$. In the limit $T \rightarrow \infty$, the Markov chain \hat{P}_T generated by the quantum walk $\langle U_{ct}(P), \omega_T \rangle$ approaches the Markov chain Π with entries

$$\Pi(y, x) := \sum_j \left| \sum_{k \in C_j} \langle y | \phi_k \rangle \langle \phi_k | x \rangle \right|^2, \quad (7)$$

where $\{\lambda_k, |\phi_k\rangle\}$ is the spectrum of P and $\{C_j\}$ is the partition of these indices k obtained by grouping together the k with identical λ_k .

Proof. Decomposing the quantum walk along spectral components gives us

$$\hat{P}_T(y, x) = E_{t \leftarrow \omega_T} \left[\left| \sum_k \langle y | \phi_k \rangle \langle \phi_k | x \rangle e^{i\lambda_k t} \right|^2 \right]. \quad (8)$$

Writing $|\cdot|^2$ as a product of complex conjugates, the right hand side becomes

$$E_{\omega_T} \left[\left(\sum_k \langle y | \phi_k \rangle \langle \phi_k | x \rangle \right) \left(\sum_l \langle \phi_l | y \rangle \langle x | \phi_l \rangle \right) e^{i(\lambda_k - \lambda_l)t} \right]. \quad (9)$$

By linearity of expectation, this is equivalent to

$$\left(\sum_k \langle y | \phi_k \rangle \langle \phi_k | x \rangle \right) \left(\sum_l \langle \phi_l | y \rangle \langle x | \phi_l \rangle \right) E_{\omega_T} [e^{i(\lambda_k - \lambda_l)t}]. \quad (10)$$

Now by assumption, $E_{\omega_T} [e^{i(\lambda_k - \lambda_l)t}]$ vanishes as $T \rightarrow \infty$ for all $\lambda_k \neq \lambda_l$, so we have

$$\begin{aligned} \hat{P}_T(y, x) &\rightarrow \sum_k \langle y | \phi_k \rangle \langle \phi_k | x \rangle \left(\sum_{l: \theta_l = \theta_k} \langle \phi_l | y \rangle \langle x | \phi_l \rangle \right) \\ &= \sum_j \left| \sum_{k \in C_j} \langle y | \phi_k \rangle \langle \phi_k | x \rangle \right|^2 \\ &= \Pi(y, x) \end{aligned} \quad (11)$$

in the limit $T \rightarrow \infty$.

It can be inferred from this lemma that most quantum walks converge to a distribution ρ other than the uniform stationary distribution $\pi = u$, and that ρ is not even independent of the initial walk state [39]. There are exceptions to this rule, for example quantum walks with distinct eigenvalues on Cayley graphs of Abelian groups (as observed by Aharonov *et al.* [10]), but they are not likely to arise in MCMC applications where the Markov chains have little structure. How then are we to sample from u using quantizations of these Markov chains? Here is where decoherence helps.

Theorem 6 (The limit $T' \rightarrow \infty$). Let P be a symmetric, irreducible Markov chain and ω_T be a family of distributions satisfying $E_{t \leftarrow \omega_T}[e^{i\theta t}] \rightarrow 0$ as $T \rightarrow \infty$ for any $\theta \neq 0$. For T sufficiently large (but fixed), the T' -repeated quantum walk $\langle U_{ct}(P), \omega_T \rangle$ generates a Markov chain $(\hat{P}_T)^{T'}$ approaching $u^{1\dagger}$ in the limit $T' \rightarrow \infty$.

Proof. We need to show that for T sufficiently large, the Markov chain \hat{P}_T is ergodic with uniform stationary distribution.

That the uniform distribution is stationary is clear: each of the $P_t(y, x) := |\langle y | e^{iPt} | x \rangle|^2$ has uniform stationary distribution since the uniform classical state is invariant under unitary quantum operations and under total measurement of the system; thus any probabilistic combination \hat{P}_T of them has uniform stationary distribution.

To show that \hat{P}_T is ergodic for all sufficiently large T , it is sufficient (by Lemma 5) to prove that Π is ergodic. (The latter implies the former because the ergodic matrices form an open subset of the set of stochastic matrices.) Why is Π ergodic? Because the 1-eigenspace of P is precisely the space spanned by u , so it follows from Lemma 5 [by consideration of only this nondegenerate eigenspace in the expression (7)] that $\Pi(y, x) \geq 1/N^2$ for every x, y .

In fact, each of the P_t (and so \hat{P}_T and Π as well) is symmetric [40]. To see this, write out the Taylor series for $\exp(iPt)$ and note that every positive integer power P^k is symmetric [since $P^2(x, y) = \sum_z P(x, z) \cdot P(z, y) = \sum_z P(y, z) \cdot P(z, x) = P^2(y, x)$]. This property will be quite useful in the next subsection: it will allow us to use Theorem 1 to relate the spectral gap and the mixing time of \hat{P}_T .

B. Invariance of the mixing time

Consider the quantum walks $\langle U_{ct}(P), \tilde{\mu}_T \rangle$ and $\langle U_{ct}(P), \tilde{\mu}'_T \rangle$ where P is a symmetric Markov chain. We show that these two quantum walks mix with essentially the same efficiency. The result extends beyond $\tilde{\mu}, \tilde{\mu}'$ to any pair of measurements ω, ω' which are sufficiently smooth or have nontrivial overlap as distributions. Let \tilde{P}_T and \tilde{P}'_T be the Markov chains

generated by the walks with measurement $\bar{\mu}$ and $\tilde{\mu}$, respectively, and let $\bar{\delta}_T$ and $\tilde{\delta}_T$ be their respective spectral gaps.

Lemma 7 (Spectral gap inequality). Let $\bar{\delta}_T$ and $\tilde{\delta}_T$ be defined as above. Then for any $k \geq 1$ we have the inequality:

$$e^{-1}\bar{\delta}_T \leq \tilde{\delta}_T \leq k(1 - e^{-k})\bar{\delta}_{kT} + 2e^{-k}. \quad (12)$$

Proof. Suppose we want to simulate \bar{P}_T by \tilde{P}_T . Scaling the distribution $\bar{\mu}_T$ by $\alpha := 1/e$ allows us to “fit it inside” the distribution $\tilde{\mu}_T$ (i.e., $e^{-1}\bar{\mu}_T \leq \tilde{\mu}_T$ pointwise), so we can express $\tilde{\mu}_T$ as the probabilistic combination $\alpha\bar{\mu}_T + (1-\alpha)\nu$ for some distribution ν , so that

$$\tilde{P}_T = E_{t \leftarrow \tilde{\mu}_T}[P_t] = \alpha E_{\bar{\mu}_T}[P_t] + (1-\alpha)E_\nu[P_t] = \alpha\bar{P}_T + (1-\alpha)Q, \quad (13)$$

where Q is stochastic with uniform stationary distribution. It follows that

$$\|\tilde{P}_T|_{u^\perp}\|_2 \leq 1/e\|\bar{P}_T|_{u^\perp}\|_2 + (1-1/e)\|Q_{u^\perp}\|_2 \quad (14)$$

which implies that $\tilde{\delta}_T \geq (1/e)\bar{\delta}_T$ since $\|Q|_{u^\perp}\|_2 \leq 1$.

Suppose we want to simulate \tilde{P}_T by \bar{P}_{kT} . Then the basic approach is the same, but since the support of $\tilde{\mu}_T$ is not compact we have to be careful. Scaling the distribution $\tilde{\mu}_T$ by $\beta := 1/k$ allows us to fit it inside the distribution $\bar{\mu}_{kT}$ up to the point $t=kT$, and the probability mass in $\tilde{\mu}_T$ past $t=kT$ is only $\Pr_{t \leftarrow \tilde{\mu}_T}[t > kT] = e^{-k}$. So we can write

$$\tilde{\mu}_T = (1 - e^{-k}) \cdot \tilde{\mu}_T^{\text{head}} + e^{-k} \cdot \tilde{\mu}_T^{\text{tail}}, \quad (15)$$

where $\tilde{\mu}_T^{\text{head}}$ and $\tilde{\mu}_T^{\text{tail}}$ are the conditional distributions of $\tilde{\mu}_T$ such that $t \leq kT$ and $t > kT$, respectively; thus

$$\tilde{P}_T = (1 - e^{-k}) \cdot \tilde{P}_T^{\text{head}} + e^{-k} \cdot \tilde{P}_T^{\text{tail}}, \quad (16)$$

where $\tilde{P}_T^{\text{head}}$ and $\tilde{P}_T^{\text{tail}}$ are the expectations of P_t with respect to $\tilde{\mu}_T^{\text{head}}$ and $\tilde{\mu}_T^{\text{tail}}$, respectively. Since we can fit $\tilde{\mu}_T^{\text{head}}$ inside $\bar{\mu}_{kT}$ if we scale it by $1/k$, we can write

$$\bar{P}_{kT} = \frac{1}{k}\tilde{P}_T^{\text{head}} + \left(1 - \frac{1}{k}\right)Q, \quad (17)$$

where Q is stochastic with uniform stationary distribution. The above equations yield

$$\bar{P}_{kT} = \frac{1}{k(1 - e^{-k})}(\tilde{P}_T - e^{-k}\tilde{P}_T^{\text{tail}}) + \left(1 - \frac{1}{k}\right)Q. \quad (18)$$

From the triangle inequality, $\|\bar{P}_{kT}|_{u^\perp}\|_2$ is at most

$$\frac{\|\tilde{P}_T|_{u^\perp}\|_2 + e^{-k}\|\tilde{P}_T^{\text{tail}}|_{u^\perp}\|_2}{k(1 - e^{-k})} + \left(1 - \frac{1}{k}\right)\|Q|_{u^\perp}\|_2. \quad (19)$$

Rearranging terms and simplifying, we have

$$\frac{(1 - \|\tilde{P}_T|_{u^\perp}\|_2) - 2e^{-k}}{k(1 - e^{-k})} \leq 1 - \|\bar{P}_{kT}|_{u^\perp}\|_2. \quad (20)$$

Theorem 8 (Equivalence of measurements). Let P be a symmetric Markov chain. Then (i) if the T' -repeated quan-

tum walk $\langle U_{ct}(P), \bar{\mu}_T \rangle$ threshold-mixes, then the $T' \cdot O(\log N)$ -repeated quantum walk $\langle U_{ct}(P), \tilde{\mu}_T \rangle$ threshold-mixes; (ii) if the T' -repeated quantum walk $\langle U_{ct}(P), \tilde{\mu}_T \rangle$ threshold-mixes, then the $T' \cdot O(\log T' \log N)$ -repeated quantum walk $\langle U_{ct}(P), \bar{\mu}_{T \cdot O(\log T')} \rangle$ threshold-mixes.

Proof. To see (i), note that our assumption implies that \bar{P}_T mixes in time T' . Therefore $\bar{\delta}_T = \Omega(1/T')$ by Theorem 2, and from Lemma 7 it follows that $\tilde{\delta}_T = \Omega(1/T')$. Applying Theorem 2 again, we obtain for \tilde{P}_T a mixing time of $O(T' \log N)$.

The proof of (ii) is almost as straightforward. Our assumption implies that \tilde{P}_T mixes in time T' , so $\tilde{\delta}_T = \Omega(1/T')$ by Theorem 2. Set k to be the smallest integer for which $\tilde{\delta}_T \geq 3e^{-k}$; in particular, $k = \Theta(\log \tilde{\delta}_T^{-1}) = O(\log T')$. By Lemma 7,

$$\bar{\delta}_{kT} \geq \frac{1}{k(1 - e^{-k})}(\tilde{\delta}_T - 2e^{-k}) \geq \frac{1}{k(1 - e^{-k})}(e^{-k}). \quad (21)$$

Asymptotically, the right hand side is

$$\Theta\left(\frac{\tilde{\delta}_T}{\log \tilde{\delta}_T^{-1}}\right) = \Theta\left(\frac{1}{T' \log T'}\right). \quad (22)$$

Applying Theorem 2 again, we obtain for \bar{P}_{kT} a mixing time of $O(T' \log T' \log N)$.

It should be readily apparent that this equivalence holds for any two measurement rules with finite expectation and significant overlap for most T . We also remark that although the above lemma and theorem are stated in terms of quantum walks, the proofs indicate that they are merely statements about an abstract game involving a collection of symmetric Markov chains $\{P_t\}_{t \geq 0}$ and a T -parametrized family of probability measures $\{\omega_T\}$, where we seek to minimize the “cost function” $T \cdot T'$.

IV. QUANTUM SPEEDUP FOR PERIODIC LATTICES

A. A near-diameter upper bound

The classical random walk on the periodic lattice \mathbb{Z}_n^d (with $N = n^d$ vertices and diameter $\lfloor n/2 \rfloor d$) has uniform stationary distribution $\pi = u$ and spectral gap $\delta = \Theta(\frac{1}{dn^2})$. It threshold-mixes in time $\Theta(n^2 d \log d)$, which is $O(\sqrt{\delta}^{-1} \log 1/\pi_*)$ only when \mathbb{Z}_n^d is quite high-dimensional: in particular, when d is roughly of order n^2 or larger. We show that a few repetitions of the continuous-time quantum walk can bring this down to $O(nd \log d)$, which is $O(\sqrt{\delta}^{-1} \log 1/\pi_*)$ for any $d \geq 1$ and $n \geq 2$ and asymptotically no worse than the diameter of \mathbb{Z}_n^d up to at most a logarithmic factor. First we prove a lemma governing mixing of various decoherent quantum walks on the cycle \mathbb{Z}_n .

Lemma 9 (Mixing on cycles). Let \mathbb{Z}_n be the cycle on $n \geq 2$ vertices. The continuous-time walks $\langle U_{ct}(P(\mathbb{Z}_n)), \omega_T \rangle$ with measurement $\omega \in \{\delta, \bar{\mu}, \tilde{\mu}\}$ threshold-mix for any $T \in \mathcal{I} := [\frac{2}{3}\frac{n}{2}, \frac{n}{2}]$, and the Hadamard walks $\langle U_{Had}(\mathbb{Z}_n), \omega_T \rangle$ with measurement $\omega \in \{\bar{\nu}, \tilde{\nu}\}$ threshold-mix for any $T \in \mathcal{J} := [\frac{2}{3}\frac{n}{\sqrt{2}}, \frac{n}{\sqrt{2}}]$.

Proof. Consider first the continuous-time walk. To prove that it threshold-mixes with any of the measurements $\omega \in \{\delta, \bar{\mu}, \tilde{\mu}\}$ for any $T \in \mathcal{I}$, it suffices by Proposition 3 to show that for every $t \in \mathcal{T}' := [\frac{3n}{5}, \frac{4n}{5}]$, $d(P_t)$ is bounded below 1 by a positive constant, where P_t is the Markov chain generated with measurement $\omega = \delta$. (Indeed, this easily implies that $d(\bar{P}_T)$ and $d(\tilde{P}_T)$ are bounded below 1 by a smaller positive constant, where \bar{P}_T and \tilde{P}_T are the Markov chains generated with measurement $\omega = \bar{\mu}$ and $\omega = \tilde{\mu}$, respectively.)

Let $|\phi_t\rangle$ and $|\psi_t\rangle$ be the wave functions at time t for the continuous-time walks on \mathbb{Z} and \mathbb{Z}_n , respectively, starting from the origin (without loss of generality, since \mathbb{Z} and \mathbb{Z}_n are vertex-transitive). Then for each $\bar{y} \in \mathbb{Z}_n$ we have

$$\langle \bar{y} | \psi_t \rangle = \sum_{y \equiv \bar{y} \pmod n} \langle y | \phi_t \rangle. \quad (23)$$

Childs [38] shows that $\langle y | \phi_t \rangle = (-i)^y J_y(t)$, where J_y is a Bessel function of the first kind. In particular, for $|y| \gg 1$ the quantity $|J_y(t)|$ is (i) exponentially small in $|y|$ for $t < (1 - \epsilon) \cdot |y|$ and (ii) of order $|y|^{-1/2}$ for $t > (1 + \epsilon) \cdot |y|$. For every $t < \frac{4n}{5}$, property (i) implies that the only term in the above summand that is non-negligible is the $\langle y | \phi_t \rangle$ with $|y| < \frac{n}{2}$ (call it \hat{y} and note that $\bar{y} \leftrightarrow \hat{y}$ is a 1-1 correspondence), so we can use property (ii) to conclude that up to a negligible correction

$$|\langle \bar{y} | \psi_t \rangle| \approx |\langle \hat{y} | \phi_t \rangle| = \Theta(1/\sqrt{n}) \quad (24)$$

for every $|\hat{y}| \gg 1$ and $t > (1 + \epsilon) \cdot |\hat{y}|$. In particular, the nearly $\frac{3}{5}n$ different \bar{y} with $1 \ll |\hat{y}| \leq \frac{3}{5} \frac{n}{2}$ satisfy $|\langle \bar{y} | \psi_t \rangle| = \Omega(1/\sqrt{n})$, and therefore $P_t(\bar{y}, \bar{0}) = \Omega(1/n)$, for every $t \in \mathcal{T}'$. So by Proposition 4, $d(P_t)$ is bounded below 1 by a positive constant.

For the Hadamard walk, the wave function is no longer characterized by Bessel functions, but it retains the same essential asymptotic spreading behavior as its continuous-time counterpart (see Nayak *et al.* [8,9]), and the argument above works with little modification. A caveat is the emergence of a parity problem: if n is even, then \mathbb{Z}_n is bipartite and the wave function is supported only on vertices of the same parity at each integer time step. Hence the Hadamard walk with $\omega = \delta$ threshold-mixes only on vertices of the same parity, but with time-averaged measurement $\omega = \bar{v}$ or $\omega = \tilde{v}$ parity is broken and threshold-mixing occurs on the entire vertex set.

Although the argument above relies on the asymptotic behavior of the wave function as $n \rightarrow \infty$, this is clearly the difficult case: if n is bounded, then it suffices to show only that there exists a time (or a pair of consecutive time steps, in the case of the Hadamard walk) in which the wave function is supported on at least 2/3 of the vertices.

For the Hadamard walk with measurement $\omega = \tilde{v}$, Lemma 9 resolves a conjecture of Kendon and Tregenna [20] based on numerical experiments. Let G^d denote the d th (Cartesian) power of a graph G . Examples are the d -dimensional standard lattice (the d th power of a line) and the d -dimensional

periodic lattice (the d th power of a cycle). The following theorem shows how to extend a threshold-mixing result from G to G^d .

Theorem 10 (Mixing on graph powers). Suppose the continuous-time quantum walk $\langle U_{ct}(P(G)), \delta_T \rangle$ threshold-mixes. Then the $O(\log d)$ -repeated walk $\langle U_{ct}(P(G^d)), \delta_{Td} \rangle$ threshold-mixes.

Proof. The Hamiltonian $H' = P(G^d)$ is related to the Hamiltonian $H = P(G)$ by the identity

$$H' = \frac{1}{d} \sum_{j=1}^d I^{\otimes(j-1)} \otimes H \otimes I^{\otimes(d-j)}. \quad (25)$$

Since H' commutes with the identity I , which can introduce at most a global phase factor to the system, the Markov chain P'_t generated by the walk $\langle U_{ct}(P(G^d)), \delta_t \rangle$ is the d th tensor power of the Markov chain P_{td} generated by the walk $\langle U_{ct}(P(G)), \delta_{td} \rangle$. By assumption, $d(P_T) \leq \alpha$ for some constant $\alpha < 1$. By Proposition 3, we can choose $T' = O(\log d)$ to ensure that

$$\frac{1}{2} \|(P_T)^{T'} - u1^\dagger\|_1 \leq d((P_T)^{T'}) \leq \frac{1}{6d^2}. \quad (26)$$

Then at least $n^{\sqrt[3]{2/3}}$ entries in each column of $(P_T)^{T'}$ are bounded below by $\frac{1-1/2d}{n}$, otherwise we would have the contradiction

$$\begin{aligned} \frac{1}{2} \|(P_T)^{T'} - u1^\dagger\|_1 &= 1 - \sum_y \min \left\{ (P_T)^{T'}(y, x), \frac{1}{n} \right\} \\ &> (1 - \sqrt[3]{2/3}) \frac{1}{2d} \geq \frac{1}{6d^2}, \end{aligned} \quad (27)$$

where the first equation uses the identity $\frac{1}{2} \|p - q\|_1 = 1 - \sum_k \min\{p_k, q_k\}$ for distributions p, q and the last inequality uses simple algebra along with the fact that for any $d \geq 1$

$$\left(1 - \frac{1/3}{d}\right)^d \geq \frac{2}{3}. \quad (28)$$

Since $(P'_{Td})^{T'} = ((P_T)^{T'})^{\otimes d}$, at least $(n^{\sqrt[3]{2/3}})^d = \frac{2}{3} n^d$ of the entries in each column of $(P'_{Td})^{T'}$ are bounded below by $(\frac{1-1/2d}{n})^d \geq \frac{1}{2n^d}$. It follows from Proposition 4 that $(P'_{Td})^{T'}$ threshold-mixes in time $O(1)$. We have the following corollary for the d th power \mathbb{Z}_n^d of the cycle \mathbb{Z}_n .

Corollary 11 (Mixing on periodic lattices). Let \mathbb{Z}_n^d be the d -dimensional periodic lattice with $n \geq 2$ vertices per side. The $O(\log d)$ -repeated continuous-time quantum walk $\langle U_{ct}(P(\mathbb{Z}_n^d)), \omega_{nd/2} \rangle$ with measurement $\omega \in \{\delta, \bar{\mu}, \tilde{\mu}\}$ threshold-mixes.

Proof. Combining Lemma 9 with Theorem 10, we conclude that $\langle U_{ct}(P(\mathbb{Z}_n^d)), \delta_T \rangle$ threshold-mixes for any $T \in [\frac{2nd}{3}, \frac{nd}{2}]$. It is easy to see (cf. Lemma 9) that this is sufficient to imply the stated corollary not only for the measurement $\omega = \delta$, but also for the time-averaged measurements $\bar{\mu}$ and $\tilde{\mu}$.

For $d \geq 1$, this extends the results of Fedichkin *et al.* [21–23] by confirming $O(n)$ and $O(d \log d)$ scaling (sug-

gested by their analytical estimates and numerical experiments in regimes of both high and low decoherence) of the fastest-mixing walk, which they conjectured to be decoherent rather than unitary.

B. The Grover walk

An important question is whether there is a T' -repeated Grover walk $\langle U_{dt}(P(\mathbb{Z}_n^d)), \omega_T \rangle$ that threshold-mixes for $T = O(nd)$ and $T' = O(\log d)$. Szegedy [14] showed that the *phase gap* (minimum nonzero eigenvalue phase from $[-\pi, \pi]$ in absolute value) of $U_{dt}(P)$ is $\Omega(\sqrt{\delta})$ and exploited this property to prove a quadratic quantum speedup for the *hitting time* of any symmetric Markov chain P . A natural adaptation of his argument to the *mixing time* setting would be something like the following: since the phase gap of $U_{dt}(P)$ is $\theta = \Omega(\sqrt{\delta})$, we expect to see by decomposing the action of $U_{dt}(P)$ along spectral components that roughly $O(1/\theta) = O(\sqrt{\delta^{-1}})$ time steps suffice for the orbit of any initial basis state to “cover” the entire state space with sufficient amplitude.

Unfortunately, this argument is incorrect: in fact, the orbit may remain quite localized around the initial basis state. This happens with dramatic effect to the Grover walk on the complete graph $G = K_N$, which mixes in time $T' = \Theta(N)$ (cf [13]) even though the classical random walk on K_N mixes in a single time step. It also happens to the Grover walk on \mathbb{Z}_n^d for $d=2$, albeit less dramatically [27–29]. [41] The probability distribution p_t induced on the vertices of \mathbb{Z}_n^2 at time $t \leq n/2$ is primarily localized at the initial basis state [28] but has substantial secondary spikes which propagate across the lattice in orthogonal directions [27]. In particular, the standard deviation of p_t appears to grow linearly with t [27] and the

mixing time of the Grover walk on \mathbb{Z}_n^2 with measurement $\omega = \bar{v}$ appears to be $T = O(n)$. In the high-dimensional regime, Moore and Russell [11] proved that the Grover walk on \mathbb{Z}_n^d with measurement $\omega = \delta$ mixes almost perfectly in time $T = O(d)$. It seems plausible that the decoherent Grover walk on \mathbb{Z}_n^d with measurement $\omega = \bar{v}$ mixes as fast asymptotically as the continuous-time walk.

V. CONCLUSION AND OPEN PROBLEMS

We have shown that decoherent quantum walks have the potential to speed up a large class of classical MCMC mixing processes. Exactly how large this class is, and whether a generic quantum mixing speedup to $O(\sqrt{\delta^{-1}} \log 1/\pi_*)$ is possible, remain important open questions. Since it seems that quantum walks can outperform classical random walks in low-dimensional examples and underperform in very high-dimensional examples (such as the complete graph), a hybrid method may work best for generic Markov chains consisting of both low- and high-dimensional substructures. Also worth investigating is whether by randomizing the “coin” used in discrete-time quantum walks we can improve mixing on the complete graph and other adversarial examples.

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